# CSCI567 Machine Learning (Spring 2021)

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March 19, 2021

### Outline

Review of last lecture

Question mixture models

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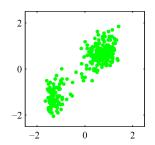
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- Question mixture models

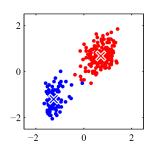
# Clustering: formal definition

**Given**: data points  $x_1, \ldots, x_N \in \mathbb{R}^{\mathsf{D}}$  and #clusters K we want

**Output**: group the data into K clusters, which means

- find assignment  $\gamma_{nk} \in \{0,1\}$  for each data point  $n \in [N]$  and  $k \in [K]$  s.t.  $\sum_{k \in [K]} \gamma_{nk} = 1$  for any fixed n
- find the cluster centers  $\mu_1, \ldots, \mu_K \in \mathbb{R}^{\mathsf{D}}$





# Alternating minimization

Instead, use a heuristic that alternatingly minimizes over  $\{\gamma_{nk}\}$  and  $\{\mu_k\}$ :

Initialize 
$$\{oldsymbol{\mu}_k^{(1)}\}$$

For 
$$t = 1, 2, ...$$

find

$$\{\gamma_{nk}^{(t+1)}\} = \operatorname*{argmin}_{\{\gamma_{nk}\}} F\left(\{\gamma_{nk}\}, \{\boldsymbol{\mu}_k^{(t)}\}\right)$$

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# The K-means algorithm

**Step 0** Initialize  $\mu_1, \dots, \mu_K$ 

**Step 1** Fix the centers  $\mu_1, \ldots, \mu_K$ , assign each point to the closest center:

$$\gamma_{nk} = \mathbb{I}\left[k = \operatorname*{argmin}_{c} \|oldsymbol{x}_n - oldsymbol{\mu}_c\|_2^2
ight]$$

**Step 2** Fix the assignment  $\{\gamma_{nk}\}$ , update the centers

$$oldsymbol{\mu}_k = rac{\sum_n \gamma_{nk} oldsymbol{x}_n}{\sum_n \gamma_{nk}}$$

Step 3 Return to Step 1 if not converged

### Outline

- Review of last lecture
- Gaussian mixture models
  - Motivation and Model
  - EM algorithm
  - EM applied to GMMs

### Gaussian mixture models

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To solve GMM, we will introduce a powerful method for learning probabilistic model: **Expectation–Maximization (EM) algorithm** 

For classification, we discussed the sigmoid model to "explain" how the labels are generated.

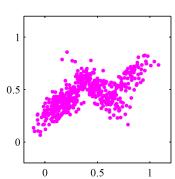
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That is, each point is an independent sample of  $x \sim p$ .

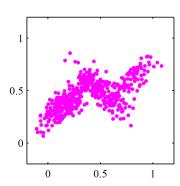


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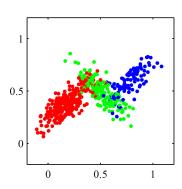
That is, each point is an independent sample of  $\boldsymbol{x} \sim p$ .

What probabilistic model generates data like this?



GMM is a natural model to explain such data

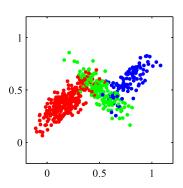
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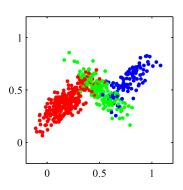
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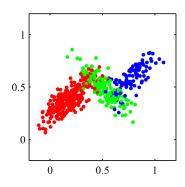
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Hence the name "Gaussian mixture model".

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$$p(\boldsymbol{x}) = \sum_{k=1}^K \omega_k N(\boldsymbol{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

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- ullet N: the density function for a Gaussian

By introducing a **latent variable**  $z \in [K]$ , which indicates cluster membership, we can see p as a **marginal distribution** 

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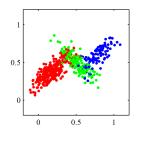
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 $oldsymbol{x}$  and z are both random variables drawn from the model

- x is observed
- z is unobserved/latent

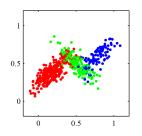
# An example



The conditional distributions are

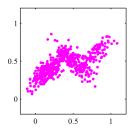
$$\begin{split} p(\boldsymbol{x} \mid z = \text{red}) &= N(\boldsymbol{x} \mid \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) \\ p(\boldsymbol{x} \mid z = \text{blue}) &= N(\boldsymbol{x} \mid \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2) \\ p(\boldsymbol{x} \mid z = \text{green}) &= N(\boldsymbol{x} \mid \boldsymbol{\mu}_3, \boldsymbol{\Sigma}_3) \end{split}$$

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#### The marginal distribution is

$$\begin{split} p(\boldsymbol{x}) &= p(\text{red}) N(\boldsymbol{x} \mid \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) + p(\text{blue}) N(\boldsymbol{x} \mid \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2) \\ &+ p(\text{green}) N(\boldsymbol{x} \mid \boldsymbol{\mu}_3, \boldsymbol{\Sigma}_3) \end{split}$$

Learning a GMM means finding all the parameters  $\boldsymbol{\theta} = \{\omega_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k\}_{k=1}^K$ .

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- ullet both learn the cluster centers  $oldsymbol{\mu}_k$ 's
- ullet in addition, GMM learns cluster weight  $\omega_k$  and covariance  $oldsymbol{\Sigma}_k$ , thus
  - we can predict probability of seeing a new point
  - we can generate synthetic data

### How to learn these parameters?

An obvious attempt is maximum-likelihood estimation (MLE): find

$$\underset{\boldsymbol{\theta}}{\operatorname{argmax}} \ln \prod_{n=1}^{N} p(\boldsymbol{x}_{n}; \boldsymbol{\theta}) = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \sum_{n=1}^{N} \ln p(\boldsymbol{x}_{n}; \boldsymbol{\theta}) \triangleq \underset{\boldsymbol{\theta}}{\operatorname{argmax}} P(\boldsymbol{\theta})$$

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One solution is to still apply GD/SGD, but a much more effective approach is the **Expectation–Maximization (EM) algorithm**.

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Step 1 (E-Step) update the "soft assignment" (fixing parameters)

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$$\mathbf{\Sigma}_k = \frac{1}{\sum_n \gamma_{nk}} \sum_n \gamma_{nk} (\mathbf{x}_n - \mathbf{\mu}_k) (\mathbf{x}_n - \mathbf{\mu}_k)^{\mathrm{T}}$$

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We will see how this is a special case of EM.

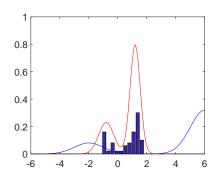
Generate 50 data points from a mixture of 2 Gaussians with

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### histogram represents the data



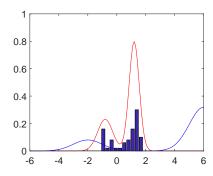
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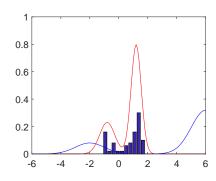
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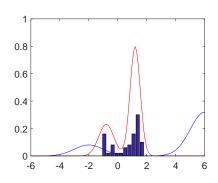
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EM\_demo.pdf shows how the blue curve moves towards red curve quickly via EM

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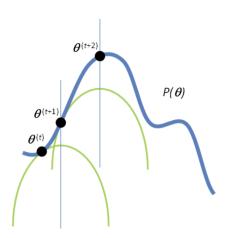
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Again, directly solving the objective is intractable.

## High level idea

Keep maximizing a lower bound of P that is more manageable



$$\ln p(\boldsymbol{x};\boldsymbol{\theta}) = \ln \int_{z} p(\boldsymbol{x}, z; \boldsymbol{\theta}) dz$$

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$$= \mathbb{E}_{z \sim q} \left[ \ln p(\boldsymbol{x},z\;;\boldsymbol{\theta}) \right] + H(q)$$

#### Finding the lower bound of P:

$$\begin{split} \ln p(\boldsymbol{x}\;;\boldsymbol{\theta}) &= \ln \int_{z} p(\boldsymbol{x},z\;;\boldsymbol{\theta}) dz \\ &= \ln \int_{z} q(z) \frac{p(\boldsymbol{x},z\;;\boldsymbol{\theta})}{q(z)} dz \qquad \qquad \text{(true for any dist. } q\text{)} \\ &= \ln \mathbb{E}_{z \sim q} \left[ \frac{p(\boldsymbol{x},z\;;\boldsymbol{\theta})}{q(z)} \right] \\ &\geq \mathbb{E}_{z \sim q} \left[ \ln \frac{p(\boldsymbol{x},z\;;\boldsymbol{\theta})}{q(z)} \right] \qquad \qquad \text{(Jensen's inequality)} \\ &= \mathbb{E}_{z \sim q} \left[ \ln p(\boldsymbol{x},z\;;\boldsymbol{\theta}) \right] + H(q) \end{split}$$

where,  $H(q) = -\mathbb{E}_{z \sim q} [\ln q(z)]$  is the Entropy.

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$$= \ln \mathbb{E}_{z \sim q} \left[ \frac{p(\boldsymbol{x},z\;;\boldsymbol{\theta})}{q(z)} \right]$$

$$\geq \mathbb{E}_{z \sim q} \left[ \ln \frac{p(\boldsymbol{x},z\;;\boldsymbol{\theta})}{q(z)} \right] \qquad \text{(Jensen's inequality)}$$

$$= \mathbb{E}_{z \sim q} \left[ \ln p(\boldsymbol{x},z\;;\boldsymbol{\theta}) \right] + H(q)$$

where,  $H(q) = -\mathbb{E}_{z \sim q} \left[ \ln q(z) \right]$  is the Entropy. Therefore, for an observation x we have

$$\ln p(\boldsymbol{x};\boldsymbol{\theta}) \geq \mathbb{E}_{z \sim q} \left[ \ln p(\boldsymbol{x}, z; \boldsymbol{\theta}) \right] + H(q)$$

Therefore, we obtain a lower bound for the log-likelihood function

$$P(\boldsymbol{\theta}) = \sum_{n=1}^{N} \ln p(\boldsymbol{x}_n ; \boldsymbol{\theta})$$

$$\geq \sum_{n=1}^{N} (\mathbb{E}_{z_n \sim q_n} [\ln p(\boldsymbol{x}_n, z_n ; \boldsymbol{\theta})] + H(q_n)) = F(\boldsymbol{\theta}, \{q_n\})$$

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Equivalently, this is the same as alternatingly maximizing F over  $\{q_n\}$  and  $\theta$  (similar to K-means).

Fix  $\theta^{(t)}$ , the solution to

$$\operatorname*{argmax}_{q_n} \mathbb{E}_{z_n \sim q_n} \left[ \ln p(\boldsymbol{x}_n, z_n ; \boldsymbol{\theta}^{(t)}) \right] + H(q_n)$$

is  $q_n^{(t)}$  s.t.

$$q_n^{(t)}(z_n) = p(z_n \mid \boldsymbol{x}_n ; \boldsymbol{\theta}^{(t)})$$

i.e., the *posterior distribution of*  $z_n$  given  $x_n$  and  $heta^{(t)}$ . (See MLaPP 11.4.7)

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$$\bullet \ F\left(\boldsymbol{\theta},\{q_n^{(t)}\}\right) \leq P(\boldsymbol{\theta}) \text{ for all } \boldsymbol{\theta}.$$

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- $F\left(\boldsymbol{\theta}, \{q_n^{(t)}\}\right) \leq P(\boldsymbol{\theta})$  for all  $\boldsymbol{\theta}$ .
- ullet  $F\left(m{ heta}^{(t)},\{q_n^{(t)}\}
  ight)=P(m{ heta}^{(t)})$  (verify using Slide 20 and MLaPP 11.4.7)

# Maximizing over heta

Fix  $\{q_n^{(t)}\}$ , maximize over  $\boldsymbol{\theta}$ :

$$\mathop{\mathrm{argmax}}_{\pmb{\theta}} F\left(\pmb{\theta}, \{q_n^{(t)}\}\right)$$

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$$\begin{split} & \underset{\boldsymbol{\theta}}{\operatorname{argmax}} F\left(\boldsymbol{\theta}, \{q_n^{(t)}\}\right) \\ &= \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \sum_{n=1}^N \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n \ ; \boldsymbol{\theta})\right] \quad \left(H(q_n^{(t)}) \text{ is independent of } \boldsymbol{\theta}\right) \\ &\triangleq \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \ Q(\boldsymbol{\theta} \ ; \boldsymbol{\theta}^{(t)}) & \left(\{q_n^{(t)}\} \text{ are computed via } \boldsymbol{\theta}^{(t)}\right) \end{split}$$

Q is the (expected) **complete likelihood** and is usually more tractable.

**Step 0** Initialize  $\theta^{(1)}$ , t=1

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$$q_n^{(t)}(\cdot) = p(\cdot \mid \boldsymbol{x}_n ; \boldsymbol{\theta}^{(t)})$$

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and obtain Expectation of complete likelihood

$$Q(\boldsymbol{\theta} ; \boldsymbol{\theta}^{(t)}) = \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[ \ln p(\boldsymbol{x}_n, z_n ; \boldsymbol{\theta}) \right]$$

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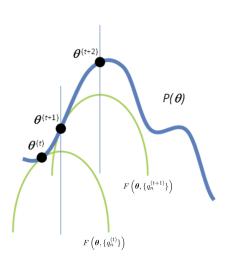
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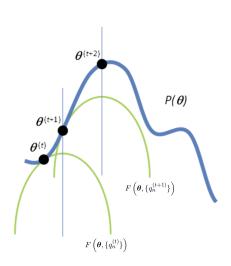
$$Q(\boldsymbol{\theta} ; \boldsymbol{\theta}^{(t)}) = \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[ \ln p(\boldsymbol{x}_n, z_n ; \boldsymbol{\theta}) \right]$$

Step 2 (M-Step) update the model parameter via Maximization

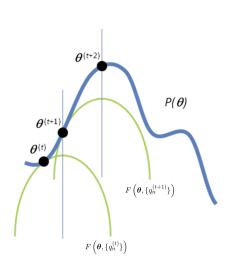
$$\boldsymbol{\theta}^{(t+1)} \leftarrow \operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta} \ ; \boldsymbol{\theta}^{(t)})$$

**Step 3**  $t \leftarrow t + 1$  and return to Step 1 if not converged

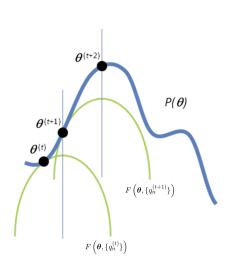




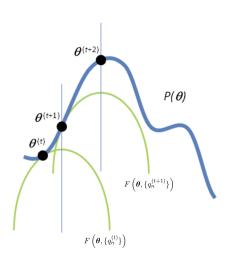
$$P(\boldsymbol{\theta}^{(\mathsf{t}+1)}) \ge F\left(\boldsymbol{\theta}^{(\mathsf{t}+1)}; \{q_n^{(t)}\}\right)$$



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$$= P(\boldsymbol{\theta}^{(t)})$$



 $P(\theta)$  is non-concave, but  $Q(\theta; \theta^{(t)})$  often is concave and easy to maximize.

$$P(\boldsymbol{\theta}^{(t+1)}) \ge F\left(\boldsymbol{\theta}^{(t+1)}; \{q_n^{(t)}\}\right)$$
$$\ge F\left(\boldsymbol{\theta}^{(t)}; \{q_n^{(t)}\}\right)$$
$$= P(\boldsymbol{\theta}^{(t)})$$

So EM always increases the objective value and will converge to some local maximum (similar to K-means).

#### E-Step:

$$q_n^{(t)}(z_n = k) = p\left(z_n = k \mid \boldsymbol{x}_n ; \boldsymbol{\theta}^{(t)}\right)$$
  
  $\propto p\left(\boldsymbol{x}_n, z_n = k ; \boldsymbol{\theta}^{(t)}\right)$ 

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This computes the "soft assignment"  $\gamma_{nk} = q_n^{(t)}(z_n = k)$ , i.e. conditional probability of  $x_n$  belonging to cluster k.

#### M-Step:

$$\operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) = \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_{n} \sim q_{n}^{(t)}} \left[ \ln p(\boldsymbol{x}_{n}, z_{n} ; \boldsymbol{\theta}) \right]$$

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To find  $\omega_1, \ldots, \omega_K$ , solve

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To find each  $\mu_k, \Sigma_k$ , solve

$$\underset{\boldsymbol{\omega}}{\operatorname{argmax}} \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma_{nk} \ln \omega_{k} \qquad \underset{\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}}{\operatorname{argmax}} \sum_{n=1}^{N} \gamma_{nk} \ln N(\boldsymbol{x}_{n} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

## M-Step (continued)

Solutions to previous two problems are very natural, for each  $\boldsymbol{k}$ 

$$\omega_k = \frac{\sum_n \gamma_{nk}}{N}$$

i.e. (weighted) fraction of examples belonging to cluster k

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EM for learning GMMs:

**Step 0** Initialize  $\omega_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k$  for each  $k \in [K]$ 

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GMM is a soft version of K-means and it provides a probabilistic interpretation of the data, which means we can predict and generate data after learning.