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## Pretest with Caution: Event-Study Estimates after Testing for Parallel Trends

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*This paper discusses two important limitations of the common practice of testing for preexisting differences in trends (“pre-trends”) when using difference-in-differences and related methods. First, conventional pre-trends tests may have low power. Second, conditioning the analysis on the result of a pretest can distort estimation and inference, potentially exacerbating the bias of point estimates and under-coverage of confidence intervals. I analyze these issues both in theory and in simulations calibrated to a survey of recent papers in leading economics journals, which suggest that these limitations are important in practice. I conclude with practical recommendations for mitigating these issues. (JEL A14, C23, C51)*

This slide is available on

<https://github.com/yasu0704xx/ArticleReview>.

# Testing for Pre-Trends

- When using DiD and related methods, researchers often test for pretreatment differences in trends (“pre-trends”) as a way of assessing the plausibility of the parallel trends assumption.
- These tests are remarkably common: 70 publications in *AER*, *AEJ Applied Economics*, *AEJ Economic Policy* between 2014 and June 2018 use an “event-study” plot to visually test for pre-trends.

## Limitations with Pretesting for Pre-Trends

- **Low Power:** Conventional pretests may have low power, meaning that preexisting trends that produce meaningful bias in the treatment effects estimates may not be detected with substantial probability.
- **Distortion in Estimation:** Conditioning the analysis on the result of a pre-trends test induces distortions to estimation and inference from pretesting. In other words, the draws of the data that survive a pretest are a selected sample from the true DGP. Because of this selection, the bias caused by a violation of parallel trends can actually be worse conditional on passing the pretest.
- These results imply that pre-trends tests may be ineffective in avoiding biases from violations of parallel trends and can even exacerbate these biases.

## Related Literature: Pretesting, Model-Selection

- Issues arising with pretesting or model-selection setup:
  - Keynes (1939) [36], Friedman (1940) [26]
  - Giles and Giles (1993) [29], Leeb and Potscher (2005) [42], Lee et al. (2016) [41], Andrews (2018) [2]
- Publication bias matters:
  - Rothstein, Sutton and Borenstein (2005) [50], Christensen and Miguel (2016) [14], Snyder and Zhuo (2018) [52], Andrews and Kasy (2019) [3]
- Low power of the pre-trends tests to detect meaningful violations of parallel trends:
  - Freyaldenhoven, Hansen and Shapiro (2019) [24], Kahn-Lang and Lang (2020) [35], Bilinski and Hatfield (2018) [8]
  - This paper also provides theoretical and empirical evidence on additional statistical distortions from pretesting.
  - The alternative approach can be found in Bilinski and Hatfield (2018) [8] and Freyaldenhoven, Hansen and Shapiro (2019).

- Failure of standard TWFE models:
  - Borusyak and Jaravel (2016) [9], Sun and Abraham (2021) [53], de Chaisemartin and D'Haultfoeuille (2020) [16], Goodman-Bacon (2021) [31], Callaway and Sant'Anna (2021) [12], Athey and Imbens (2022) [5]
  - This paper points out that conventional pretests may do a poor job detecting violations of the relevant parallel trends assumption.

Survey of Recent Papers

Theoretical Analysis

Practical Recommendations

Appendix

References

# Survey of Recent Papers

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- Candidates: 70 empirical papers with the phrase “event study” in *AER*, *AEJ Applied Economics*, *AEJ Economic Policy* between 2014 and June 2018.
- The final sample of 12 papers:
  - Bailey and Goodman-Bacon (2015) [6], Bosch and Campos-Vazquez (2014) [10], Deryugina (2017) [17], Deschenes et al. (2017) [19], Fitzpatrick and Lovenheim (2014) [22], Gallagher (2014) [27], He and Wang (2017) [33], Kuziemko et al. (2018) [37], Lafortune et al. (2018) [39], Markevich and Zhuravskaya (2018) [45], Tewari (2014) [54], Ujhelyi (2014) [56]
- The above 12 studies show an event plot with pointwise CIs that allows ..... (続きここから)

TABLE 1—SUMMARY OF PRE-PERIOD EVENT-STUDY COEFFICIENTS

Paper	# pre-periods	# significant	Max  t	Joint <i>p</i> -value	t  for slope
Bailey and Goodman-Bacon (2015)	5	0	1.674	0.540	0.381
Bosch and Campos-Vazquez (2014)	11	2	2.357	0.137	0.446
Deryugina (2017)	4	0	1.090	0.451	1.559
Deschenes et al. (2017)	5	1	2.238	0.014	0.239
Fitzpatrick and Lovenheim (2014)	3	0	0.785	0.705	0.977
Gallagher (2014)	10	0	1.542	0.166	0.855
He and Wang (2017)	3	0	0.884	0.808	0.720
Kuziemko et al. (2018)	2	0	0.474	0.825	0.474
Lafortune et al. (2017)	5	0	1.382	0.522	1.390
Markevich and Zhuravskaya (2018)	3	0	0.850	0.591	0.676
Tewari (2014)	10	0	1.061	0.948	0.198
Ujhelyi (2014)	4	1	2.371	0.003	1.954

*Notes:* This table provides information about the pre-period event-study coefficients in the papers reviewed. The table shows the number of pre-period coefficients in the event study, the number of the pre-period coefficients that are significant at the 95 percent level, the maximum t-stat among those coefficients, the *p*-value for a chi-squared test of joint significance, and the t-stat for the slope of the linear trend through the pre-period coefficients. See Section I for more detail on the sample of papers reviewed.



# Theoretical Analysis

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# Practical Recommendations

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# Appendix

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## Canonical Two-Period DiD

- Suppose that we have panel data  $\{\{(Y_{it}, D_{it})\}_{t=0}^T\}_{i=1}^n$ , where  $i$  and  $t$  index units and time periods, respectively.  $Y_{it} \in \mathbb{R}$  and  $D_{it} \in \{0, 1\}$  denote the outcome and treatment, respectively, for unit  $i$  in period  $t$ .
- Consider the simplest case where  $T = 1$ , no units are treated at  $t = 0$  (i.e.,  $D_{i0} = 0$  for any  $i$ ), and some but not all units become treated at  $t = 1$ .
- Then, only  $D_{i1}$  is relevant, and we simply write  $D_i = D_{i1}$ .
- Let  $Y_{it}(d)$  denote the potential outcome given  $D_i = d$ .

# Parallel Trends in Canonical DiD

- **Parallel trends (PT)** is a key assumption for point identification of ATT in DiD.<sup>1</sup>
- In a canonical two-period DiD model, PT means

$$\mathbb{E}[Y_{i1}(0) - Y_{i0}(0)|D_i = 1] = \mathbb{E}[Y_{i1}(0) - Y_{i0}(0)|D_i = 0]$$

In words, the untreated potential outcome is required to have the same average path over time between the treated and untreated groups.

- Under assumptions of no anticipation and PT, ATT is identified as

$$\begin{aligned}\tau_{\text{ATT}} &= \mathbb{E}[Y_{i1}(1) - Y_{i1}(0)|D_i = 1] \\ &= \mathbb{E}[Y_{i1} - Y_{i0}|D_i = 1] - \mathbb{E}[Y_{i1} - Y_{i0}|D_i = 0]\end{aligned}$$

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<sup>1</sup>Note that identification of ATT also requires another assumption, namely, “no anticipation,” or “no anticipatory effects of treatment,” which we do not focus on here.








## Pre-Trends Tests

- Since  $Y_{i1}(0)|D_i = 1$  is counterfactual (i.e., cannot be observed), PT cannot be directly examined.
- Instead, researchers implement **pre-trends tests**, testing for preexisting differences in trends, so that they can assess the plausibility of the PT assumption.
- Note that this “plausibility” is neither a necessary nor sufficient condition for PT.

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