Regression Discontinuity Designs

Hansen (2022, Chapter 21)

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https://yasu0704xx.github.io

Introduction

- Regression discontinuity designs (RDDs) are quasi-experimental designs which allow researchers to identify the causal effect of endogenous treatment on an outcome based on discontinuous policy rules.
- Local randomization is a key idea.
  - Consider a certain discontinuous rule under which treatment (e.g. college scholarship) is determined by whether a continuous covariate (e.g. admission score) is greater than a known threshold.
  - If all factors determined prior to the treatment are balanced just above and just below the threshold, the average causal effect can be estimated by comparing the mean outcome just above the threshold with that just below the threshold.

Literature

- Here we review Chapter 21 of Hansen (2022) [24].
- Excellent reviews/textbooks on regression discontinuity designs include Abadie and Cattaneo (2018) [1], and Cattaneo, Idrobo and Titiunik (2021, 2024) [13] [14].
- The common software package is `rdrobust` by Calonico, Cattaneo, Farrell and Titiunik.
- 日本語の文献：
  - 川口・澤田（2024）[44]
  - 末石（2024）[45]
  - 高野（2025）[46]

Contents

Identification

Setup

- $Y_d \in \mathbb{R}$, $d \in \{0, 1\}$: potential outcome
- Y : observed outcome
- $D \in \{0, 1\}$: treatment, which may be endogenous in that some unobserved factors may affect both D and Y.
- $\theta = Y_1 - Y_0$: treatment effect for an individual
  $\Rightarrow$ We cannot identify θ without restrictive assumptions, because either Y_1 or Y_0 is unobservable.
- Instead, we are interested in the conditional average treatment effect (conditional ATE)

$$\theta(x) = \mathbb{E}[\theta|X = x] = \mathbb{E}[Y_1 - Y_0|X = x],$$

where X is an observable covariate.

Sharp RD

- Suppose that treatment is determined by

$$D = 1(X \geq c),$$

  where the cut-off c is determined by policy or rule and common to all individuals.
- The covariate X is called the score, forcing variable, running variable, assignment variable, etc.
- In a standard RD setting, X is assumed to be continuously distributed on a subset of $\mathbb{R}$.

Identification in Sharp RD

- Set $\bar{\theta} = \theta(c) \equiv \mathbb{E}[\theta|X = c]$, which is the parameter of interest.
- Let $m(x) = \mathbb{E}[Y|X = x]$ and $m_d(x) = \mathbb{E}[Y_d|X = x]$ for each $d \in \{0, 1\}$. Note that $\theta(x) = m_1(x) - m_0(x)$.
- Set $m(x+) = \lim_{z \downarrow x} m(z)$ and $m(x-) = \lim_{z \uparrow x} m(z)$.

Theorem 21.1

Assume that treatent is assigned as $D = 1(X \geq c)$. *Suppose that* $m_0(x)$ *and* $m_1(x)$ *are continuous at* $x = c$. *Then,*

$$\bar{\theta} = m(c+) - m(c-).$$

Proof By construction,

$$Y = Y_0 \cdot 1(X < c) + Y_1 \cdot 1(X \geq c).$$

Taking expectations conditional on $X = x$, we obtain

$$m(x) = m_0(x)1(x < c) + m_1(x)1(x \geq c).$$

Since $m_0(x)$ and $m_1(x)$ are continuous at $x = c$,

$$m(c+) = m_1(c), \text{ and } m(c-) = m_0(c),$$

which completes the proof.

- The conditions for Theorem 21.1 mean that the conditional expectation of the untreated and treated outcome are continuously affected by the running variable.
- It is implied that the distributions of confounding factors, including observable covariates determined prior to the treatment, are balanced near the cutoff.
- In particular, there should be no policy/legal/experimental changes at the cutoff, except for the treatment assignment.
- Counterfactual: The continuity of $m_0(x)$ and $m_1(x)$ cannot be directly examined, since Y_0 and Y_1 are unobservable under $X \geq c$ and $X < c$, respectively.

Fuzzy RD

- In the fuzzy RD, D is partially determined by whether X is no less than a known fixed cutoff c, such that

$$\lim_{x \downarrow c} \mathbb{E}[D|X = x] \neq \lim_{x \uparrow c} \mathbb{E}[D|X = x],$$

  where $\lim_{x \downarrow c}$ and $\lim_{x \uparrow c}$ denote the right and left limits at $x = c$, respectively.
- Notice that $\mathbb{E}[D|X] = \mathbb{P}(D = 1|X)$.

Identification in Fuzzy RD

- Define $Z = 1(X \geq c)$.
- Let $D_z, z \in \{0, 1\}$ be the potential treatment status when $Z = z$. By construction, $D = ZD_1 + (1 - Z)D_0$.
- Consider the following causal parameter for the "compliers," sometimes called the local Wald estimand:

$$\tau_{\text{FRD}} \equiv \mathbb{E}[Y_1 - Y_0 \mid D_1 > D_0, X \in \{c - \epsilon, c + \epsilon\}].$$

- Under several assumptions, τ_{FRD} can be identified by

$$\tau_{\text{FRD}} = \frac{\lim_{x \downarrow c} \mathbb{E}[Y|X = x] - \lim_{x \uparrow c} \mathbb{E}[Y|X = x]}{\lim_{x \downarrow c} \mathbb{E}[D|X = x] - \lim_{x \uparrow c} \mathbb{E}[D|X = x]}.$$

- The arguments are quite similar to the identification of LATE parameter in the IV estimations (so skiped in the class).
- See Hahn, Todd and van der Klaauw (2001) [23] and Hansen (2022, Sections 21.10-11) [24] for details.

Estimation

Notes

LLR or LPR with $p = 2$

Bandwidth: Imbens and Kalyanaraman (2012) [25]

`rdrobust`

LPR ($p > 2$) should be avoided: Gelman and Imbens (2019) [21]

Regression Specification

Inference

Bandwidth

A naive solution: undersmoothing

Robust bias-corrected:
Calonico, Cattaneo and Titiunik (2014) [9] & Calonico, Cattaneo and Farrell (2020) [8]

`rdrobust`

Unifromly honest CI: Kolesar and Rothe (2018) [30]

`rdhonest`

Covariates

Notes

Calonico et al. (2019) [10]

Kreiss and Rothe (2023) [31]

Arai, Otsu and Seo (2024) [5]

Chernozhukov et al. (2025) [16]

Noack, Olma and Rothe (2025) [36]

Falsification Test

Identification Assumptions

Theorem 21.1 (re)

Assume that treatent is assigned as $D = 1(X \geq c)$. *Suppose that* $m_0(x)$ *and* $m_1(x)$ *are continuous at* $x = c$. *Then,*

$$\bar{\theta} = m(c+) - m(c-).$$

- Counterfactual: The continuity of $m_0(x)$ and $m_1(x)$ cannot be directly examined, since Y_0 and Y_1 are unobservable under $X \geq c$ and $X < c$, respectively.
- Instead, researchers often examine certain necessary conditions.
- No manipulation: Under the above assumptions, it is necessary for the density of running variable X to be continuous at the cutoff point.

Manipulation Test

- Testing the continuity of the density of the assignment variable:
  - McCrary (2008) [34]
  - Otsu, Xu and Matsushita (2013) [38]
  - Cattaneo, Jansson and Ma (2020) [15]
- Testing the continuity of the conditional distribution of covariates:
  - Lee (2008) [32]
  - Canay and Kamat (2018) [11]
  - Fusejima, Ishihara and Sawada (2025) [20]

Placebo Test

- Take several placebo cutoff points.
- At a placebo cutoff point $c' < c$, researchers can observe whether the dendity of untreated potential outcome Y_0 is continuous or not.
- Similarly, at $c' > c$, they can observe whether the density of treated potential outcome Y_1 is continuous or not.
- If the continuity is observed, then there arises some plausibility of the identification assumptions in Theorem 21.1.
- Howevere, it is just a plausibility. Note that the continuity at placebo cutoff points is neither necessary nor sufficient for the identification assumptions.

Caveats on Pretesting

- Recent works argue that, in general, researchers should not implement such pretesting.
  - Roth (2022) [39] : Pre-trend test
  - Sueishi (2023) [42] : Hausman test
- See Section 5.2 of Fusejima, Ishihara and Sawada (2025) [20] for pretesting analysis in the RD setting.

Empirical Application

"Waiting for Life"

- The first study relying on RD can be found in Thistlehwaite and Cambell (1960) [43].
- As Cook (2008) [17] says that RD was "waiting for life," RD was not popular until 1999, the year when Angrist and Lavy (1999) [4] and Black (1999) [7] were published.
- Hahn, Todd, and Klaauw (2001) [23] formalize general RDDs and establish identification results for treatment effects.

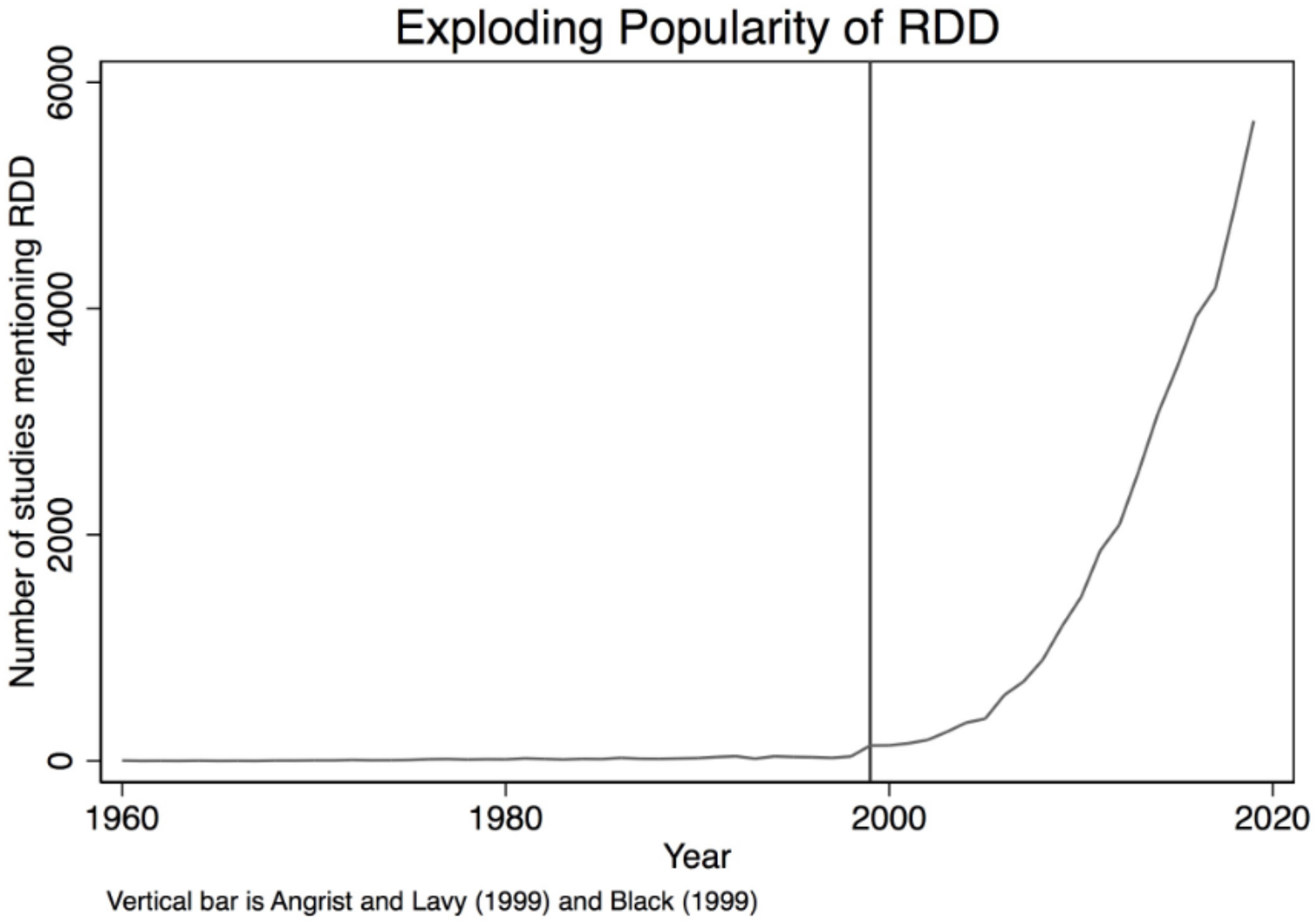

Figure 6.1 of Cunningham (2021) [18]

Recent Empirical RD Studies

Ludwig and Miller (2007) [33],
Lee (2008) [32],
Matsudaira (2008) [35],
Battistin et al. (2009) [6],
Carpenter, Christopher and Carlos Dobkin (2009) [12],
Greenstone, Hornbeck and Moretti (2010) [22],
Abdulkadiroglu, Angrist and Pathak (2014) [2],
Ito (2014) [26],
Kleven et al. (2014) [29],
Shigeoka (2014) [40],
Shigeoka (2016) [41],
Ito and Sallee (2018) [27],
Oizer (2018) [37],
Kawai et al. (2023) [28], etc.

Anderson (2018, AER)

- Anderson (2018) [3] examines causal relationship between legal systems and female HIV infection rates in sub-Saharan Africa.
- RD, motiveted by Dell (2010) [19]:
  - As-if random borders can mitigate an endogeneity emerged within ethnicity level.
- Result 1 (HIV positive rates)
  - Female: common law countries $>$ civil law countries
  - Male: no significant difference
- Result 2 (Contraception use)
  - Female: common law countries $<$ civil law countries
  - Male: common law countries $<$ civil law countries
- Common Law $\Rightarrow$ Female bargaining power $\downarrow$
  $\Rightarrow$ Negotiation for safe sex practices $\times$ $\Rightarrow$ HIV prevalence $\uparrow$

Split Ethnic Groups with Different Legal Origins

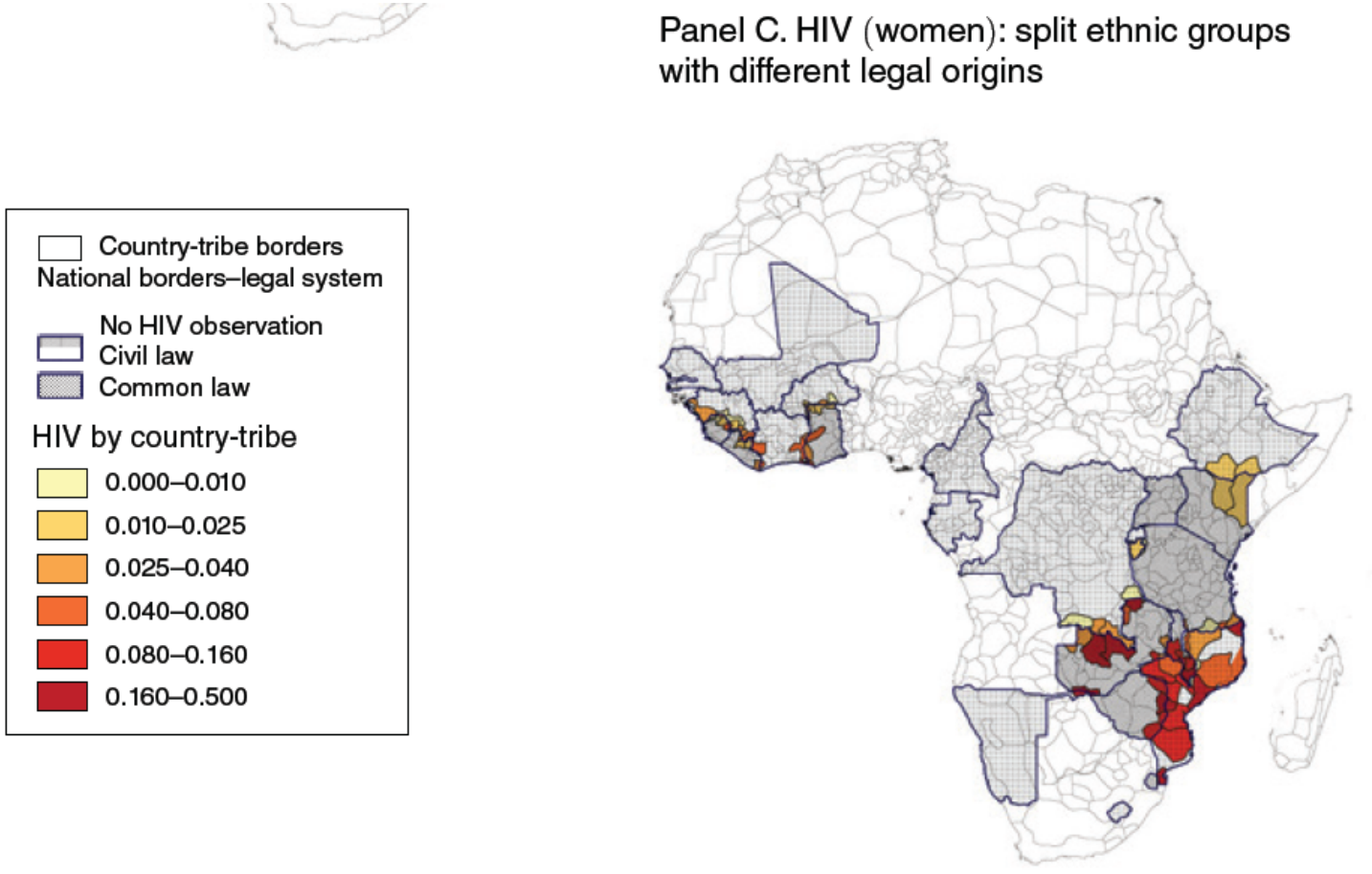

FIGURE 1. FEMALE HIV BY ETHNIC GROUP

Empirical Strategy & Main Result

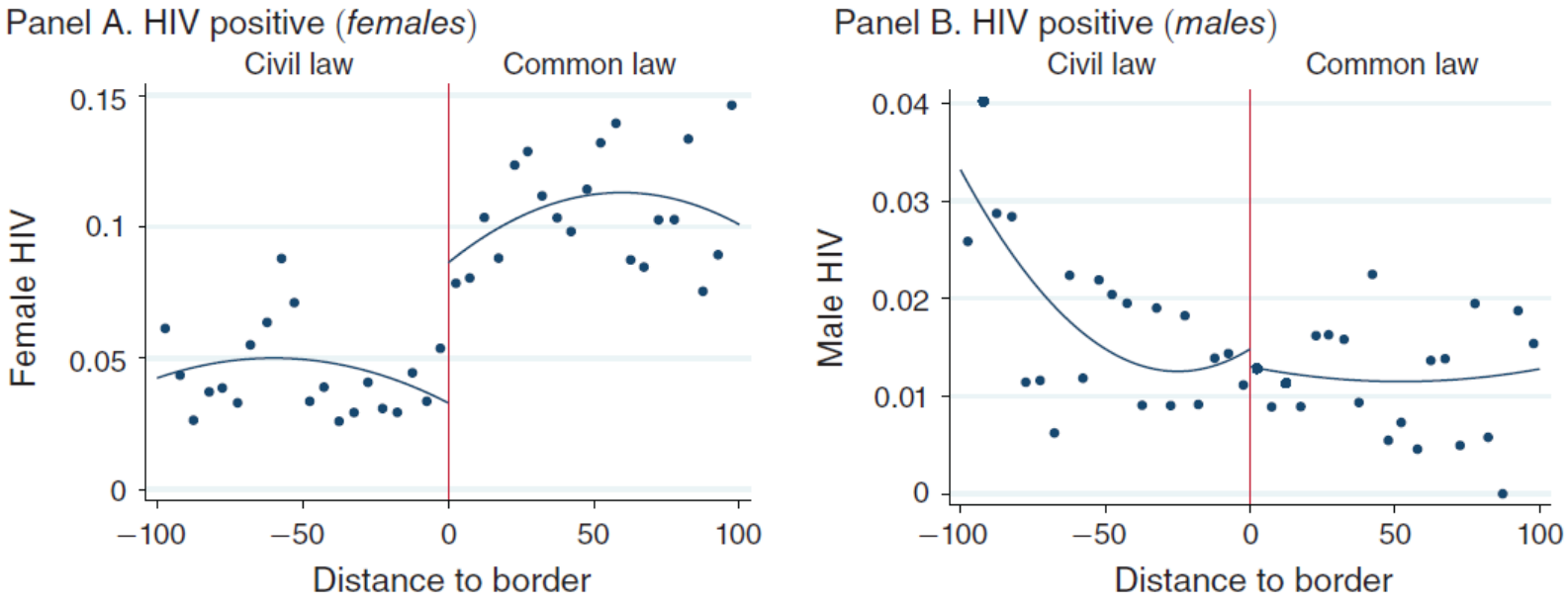

FIGURE 2. HIV POSITIVE

- $Y_i =$ HIV Positive Rate (Proxy for Females Bargaining Power)
- $D_i =$ Legal Origins (Civil Law vs Common Law)
- $X_i - c =$ Distance to Border
- As-if random borders of Scramble for Africa (1884)

  $\Longrightarrow$ Same cultures but different legal origins near a border!

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