## Semiparametric Single Index Models

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#### Introduction

A semiparametric single index model is given by

$$Y = g(X^T \beta_0) + u,$$

where

 $Y\in\mathbb{R}$ : a dependent variable,  $X\in\mathbb{R}^q: \text{a }q\times 1 \text{ explanatory vector,}$   $\beta_0\in\mathbb{R}^q: \text{a }q\times 1 \text{ vector of unknown parameters,}$   $u\in\mathbb{R}: \text{an error term which satisfies }\mathbb{E}(u\mid X)=0,$   $g(\cdot): \text{an unknown distribution function.}$ 

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#### Introduction

- Even though x is a  $q \times 1$  vector,  $x^T \beta_0$  is a scalar of a single linear combination, which is called a single index.
- By the form of the single index model, we obtain

$$\mathbb{E}(Y \mid X) = g(X^T \beta_0),$$

which means that the conditional expectation of Y only depends on the vector X through a single index  $X^T\beta_0$ .

- The model is semiparametric when  $\beta \in \mathbb{R}^q$  is estimated with the parametric methods and  $g(\cdot)$  with the nonparametric methods.
- Some of the parametric single index models are really familiar with us.

#### **Examples of Parametric Single Index Model**

• If  $g(\cdot)$  is the identity function, then the model turns out to be a linear regression model:

$$Y = g(X^T \beta_0) + u = X^T \beta_0 + u.$$

- If  $g(\cdot)$  is the CDF of Normal(0,1), then the model turns out to be a probit model.
  - See the textbook for further discussions on a probit model.
- If  $g(\cdot)$  is the CDF of logistic distribution, then the model turns out to be a logistic regression model.

#### **Identification Conditions**

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Proposition 8.1 (Identification of a Single Index Model) -

For the semiparametric single index model  $Y=g(x^T\beta_0)+u$ , identification of  $\beta_0$  and  $g(\cdot)$  requires that

- (i) x should not contain a constant/intercept, and must contain at least one continuous variable. Moreover,  $\|\beta_0\|=1$ .
- (ii)  $g(\cdot)$  is differentiable and is not a constant function on the support of  $x^T\beta_0$ .
- (iii) For the discrete components of x, varying the values of the discrete variables will not divide the support of  $x^T\beta_0$  into disjoint subsets.

#### Identification Condition (i)

- Note that the location and the scale of  $\beta_0$  are not identified.
- The vector x cannot include an intercept because the function  $g(\cdot)$  (which is to be estimated in nonparametric manners) includes any location and level shift.
  - That is,  $\beta_0$  cannot contain a location parameter.
- Some normalization criterion (scale restrictions) for  $\beta_0$  are needed.
  - One approach is to set  $\|\beta_0\|$ .
  - The second approach is to set one component of  $\beta_0$  to equal one. This approach requires that the variable corresponding to the component set to equal one is continuously distributed and has a non-zero coefficient.
  - Then, x must be dimension 2 or larger. If x is one-dimensional, then  $\beta_0 \in \mathbb{R}^1$  is simply normalized to 1, and the model is the one-dimensional nonparametric regression  $E(Y \mid x) = g(x)$  with no semiparametric component.

## Identification Conditions (ii) and (iii)

- The function  $g(\cdot)$  cannot be a constant function and must be differentiable on the support of  $x^T \beta_0$ .
- x must contain at least one continuously distributed variable and this continuous variable must have non-zero coefficient.
  - If not,  $x^T\beta_0$  only takes a discrete set of values and it would be impossible to identify a continuous function  $g(\cdot)$  on this discrete support.

Method

Estimation: Ichimura (1993)'s

- Suppose that the functional form of  $g(\cdot)$  were known.
- Then we could estimate  $\beta_0$  by minimizing the least-squares criterion:

$$\sum_{i=1} \left[ Y_i - g(X_i^T \beta) \right]^2$$

with respect to  $\beta$ .

- We could think about replacing  $g(\cdot)$  with a nonparametric estimator  $\hat{g}(\cdot)$ .
- However, since g(z) is the conditional mean of  $Y_i$  given  $X_i^T\beta_0=z,\ g(\cdot)$  depends on unknown  $\beta_0$ , so we cannot estimate  $g(\cdot)$  here.

• Nevertheless, for a fixed value of  $\beta$ , we can estimate

$$G(X_i^T \beta) := \mathbb{E}(Y_i \mid X_i^T \beta) = \mathbb{E}(g(X_i^T \beta) \mid X_i^T \beta).$$

- In general  $G(X_i^T \beta) \neq g(X_i^T \beta)$ .
- When  $\beta = \beta_0^{-1}$ , it holds that  $G(X_i^T \beta_0) = g(X_i^T \beta_0)$ .

<sup>&</sup>lt;sup>1</sup>Recall that  $\beta_0$  is the true value of  $\beta$ .

• First, we estimate  $G(X_i^T\beta)$  with the leave-one-out NW estimator:

$$\hat{G}_{-i}(X_i^T \beta) := \hat{\mathbb{E}}_{-i}(Y_i \mid X_i^T \beta)$$

$$= \frac{\sum_{j \neq i} Y_j K\left(\frac{X_j^T \beta - X_i^T \beta}{h}\right)}{\sum_{j \neq i} K\left(\frac{X_j^T \beta - X_i^T \beta}{h}\right)}.$$

• Second, using the leave-one-out NW estimator  $\hat{G}_{-i}(X_i^T\beta)$ , we estimate  $\beta$  with

$$\hat{\beta} := \arg\min_{\beta} \sum_{i=1}^{n} \left[ Y_i - \hat{G}_{-i}(X_i^T \beta) \right]^2 w(X_i) \mathbf{1}(X_i \in A_n) := S_n(\beta),$$

which is called Ichimura (1993)'s estimator.

- $w(X_i)$  is a nonnegative weight function.
- $\mathbf{1}(X_i \in A_n)$  is a trimming function to trim out small values of  $\frac{1}{nh} \sum_{j \neq i} K\left(\frac{X_j^T \beta X_i^T \beta}{h}\right)$ , so that we do not suffer the random dominator problem.

### Asymptotic Distribution of Ichimura (1993)'s Estimator

• See Theorem 8.1. for the statement on the asymptotic distribution of Ichimura's estimator.

## Direct Semiparametric Estimators for $\beta$

**Bandwidth Selection** 

# Klein and Spady (1993)

Lewbel (2000)

Manski's (1975) Maximum Score

**Estimator** 

Horowitz's (1992) Smoothed

**Maximum Score Estimator** 

Han's (1987) Maximum Rank

**Estimator** 

**Multinomial Discrete Choice Models** 

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Maximum Likelihood Approach

Ai's (1997) Semiparametric

### References

## References (1)

- Li, Q. and J. S. Racine, (2007). *Nonparametric Econometrics:* Theory and Practice, Princeton University Press.
- 末石直也 (2024) 『データ駆動型回帰分析:計量経済学と機械学習の融合』日本評論社.
- 西山慶彦, 人見光太郎 (2023) 『ノン・セミパラメトリック 統計解析 (理論統計学教程: 数理統計の枠組み)』共立出版.

## References (2)

Useful references also include some lecture notes of the following topic courses:

- ECON 718 NonParametric Econometrics (Bruce Hansen, Spring 2009, University of Wisconsin-Madison),
- セミノンパラメトリック計量分析(末石直也, 2014 年度後期, 京都大学).