

Feature Selection for Discovering Distributional Treatment Effect Modifiers


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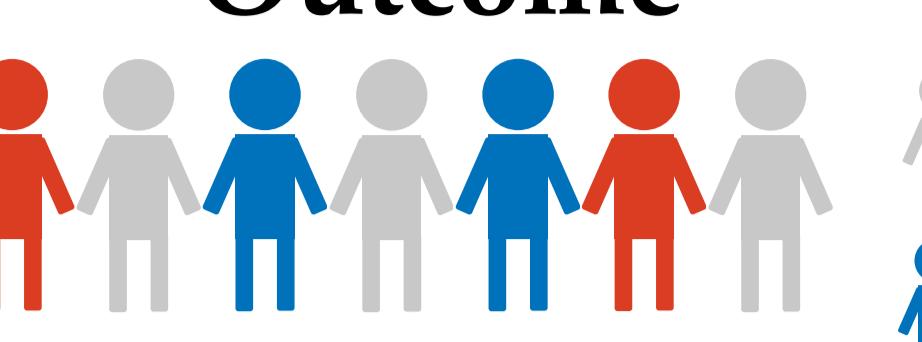
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Motivation: Elucidate why treatment effects are different

Treatment



Outcome



: Harmed

: No effect

: Benefited

e.g., vaccination, education program

e.g., immunity, grades

Many existing methods use a complex ML model to accurately estimate heterogeneous treatment effects across individuals.

However, they offer no answer to the following question:

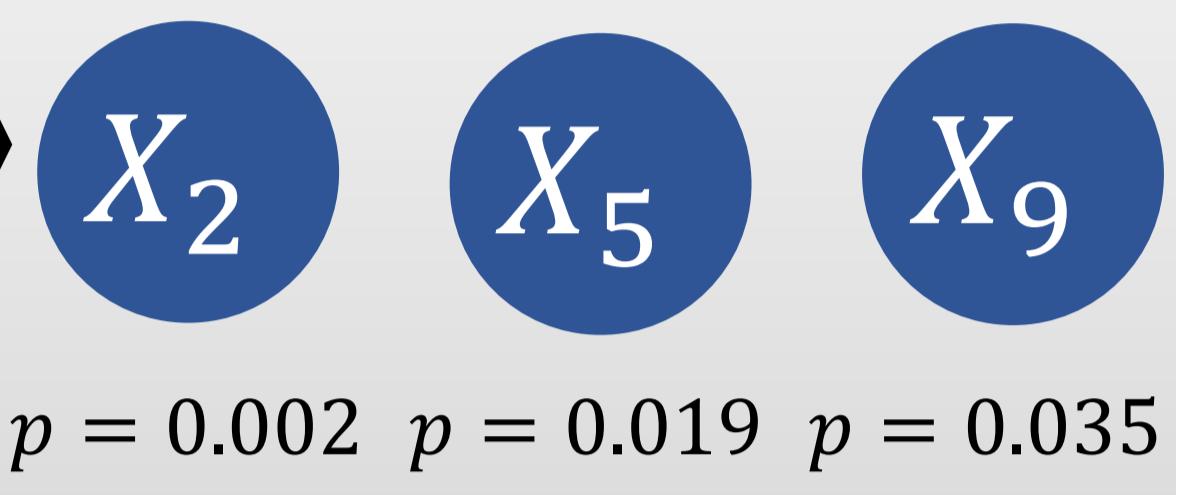
Different individuals have different treatment effects. Why?

We answer this question by solving the feature selection problem:

Input: Observations of features X , treatment A , and outcome Y

Output: Features related to treatment effect heterogeneity

X_1	...	X_d	A	Y	Y^1	Y^0	$Y^1 - Y^0$
Male	15 y.o.	0	82	?	82	?	
Male	80 y.o.	0	174	?	174	?	
Female	64 y.o.	1	135	135	?	?	
Female	32 y.o.	1	110	110	?	?	



Our Contributions

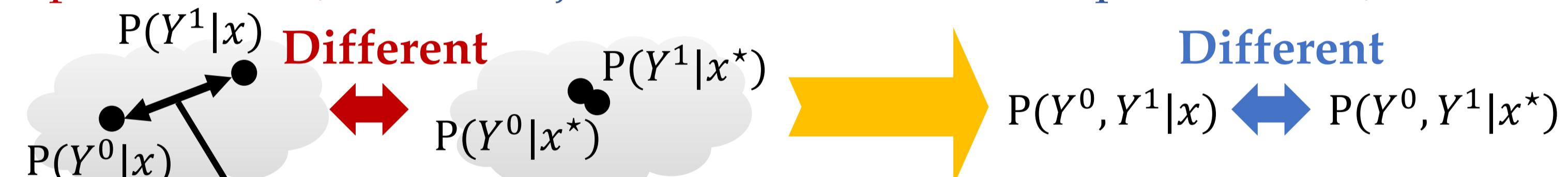
1. Novel feature importance measure
2. Its computationally efficient estimator
3. Selection algorithm that controls Type I error

Proposed method

Our Goal: Detect features whose values affect the **functionals** of joint distribution $P(Y^0, Y^1 | X_m = x)$ (e.g., treatment effect variance)

1. Detecting distributional treatment effect modifiers

Idea: If the discrepancy between $P(Y^0 | X_m = x)$ and $P(Y^1 | X_m = x)$ depends on $X_m = x$, then joint distribution also depends on $X_m = x$.



Measured by kernel MMD [2]: $D_m^2(x) := \text{MMD}^2(P(Y^0 | X_m = x), P(Y^1 | X_m = x)) = \mathbb{E}_{Y^0, Y^0' | X_m = x} [k_Y(Y^0, Y^0')] + \mathbb{E}_{Y^1, Y^1' | X_m = x} [k_Y(Y^1, Y^1')] - 2 \mathbb{E}_{Y^0, Y^1 | X_m = x} [k_Y(Y^0, Y^1)]$

To detect the **MMD value variation**, we formulate feature importance measure as the variance of squared MMD:

$$I_m := \text{Var}[D_m^2(X_m)].$$

2. Estimating importance measure with IPW and RFFs

Using **inverse probability weighting** (IPW), we reformulate $D_m^2(x)$ as

$$\begin{aligned} & \text{WCMM}^2_{X_m=x} \\ & := \mathbb{E}_{A, A', X_{-m}, Y, Y' | X_m = x} [w^0(A, X) w^0(A', X') k_Y(Y, Y')] \\ & + \mathbb{E}_{A, A', X_{-m}, Y, Y' | X_m = x} [w^1(A, X) w^1(A', X') k_Y(Y, Y')] \\ & - 2 \mathbb{E}_{A, A', X_{-m}, Y, Y' | X_m = x} [w^0(A, X) w^1(A', X') k_Y(Y, Y')]. \end{aligned}$$

$$\begin{aligned} w^0(A, X) &= \frac{\mathbf{I}(A = 0)}{1 - e(X)} \\ w^1(A, X) &= \frac{\mathbf{I}(A = 1)}{e(X)} \\ e(X) &:= P(A = 1 | X) \end{aligned}$$

Empirical estimator: $\widehat{D}_m^2(x) := \sum_{i=1}^n \sum_{j=1}^n (\omega_i^{0,x} \omega_j^{0,x} + \omega_i^{1,x} \omega_j^{1,x}) k_Y(y_i, y_j) - 2 \sum_{i=1}^n \sum_{j=1}^n \omega_i^{0,x} \omega_j^{1,x} k_Y(y_i, y_j)$

If X_m is discrete, $\omega_i^{a,x} = \frac{\mathbf{I}(x_{m,i} = x)}{\sum_{l=1}^n \mathbf{I}(x_{m,l} = x)} w^a(a_i, x_i)$; otherwise, $\omega_i^{a,x} = \frac{1}{\sum_{l=1}^n k_{X_m}(x_{m,l}, x)} w^a(a_i, x_i)$

To reduce the computation time, we approximate k_Y with RFFs [3]:

$$k_Y(y_i, y_j) \approx \tilde{k}_Y(y_i, y_j) = \langle z(y_i), z(y_j) \rangle_{\mathbb{R}^r}, \quad z(y) = \begin{bmatrix} \sqrt{2} \cos(\lambda_1 y_1 + \zeta_1) \\ \vdots \\ \sqrt{2} \cos(\lambda_r y_r + \zeta_r) \end{bmatrix}$$

which yields

$$\widehat{D}_m^2(x) := \langle \tilde{\mu}_{Y^0|X}, \tilde{\mu}_{Y^0|X} \rangle_{\mathbb{R}^r} + \langle \tilde{\mu}_{Y^1|X}, \tilde{\mu}_{Y^1|X} \rangle_{\mathbb{R}^r} - 2 \langle \tilde{\mu}_{Y^0|X}, \tilde{\mu}_{Y^1|X} \rangle_{\mathbb{R}^r}.$$

Estimated feature importance:

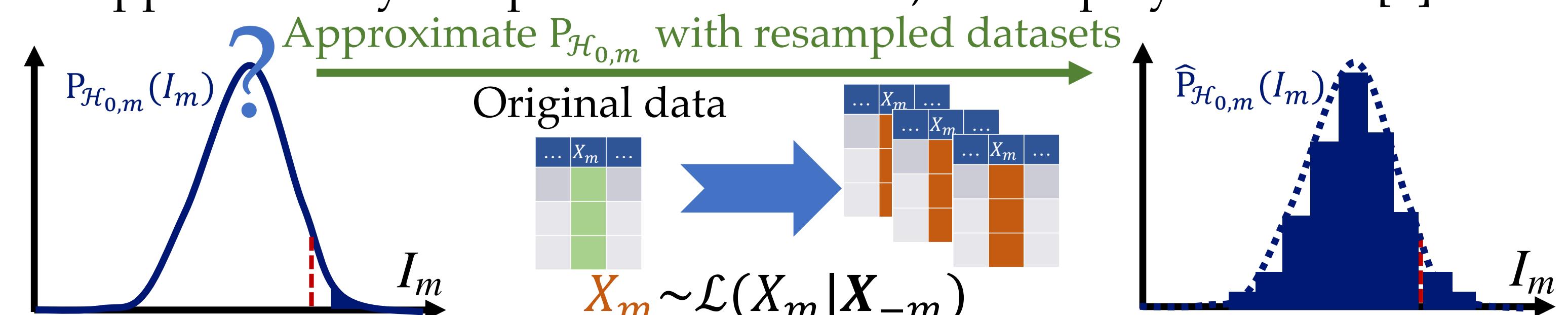
$$\widetilde{I}_m = \frac{1}{n-1} \sum_{i=1}^n \left(\widehat{D}_m^2(x_{m,i}) - \frac{1}{n} \sum_{s=1}^n \widehat{D}_m^2(x_{m,s}) \right)^2$$

3. Multiple tests with conditional randomization test (CRT)

We select features by performing multiple hypothesis tests:

$$\mathcal{H}_{0,m}: I_m = 0 \quad \text{and} \quad \mathcal{H}_{1,m}: I_m > 0. \quad (m=1, \dots, d)$$

To approximately compute the **threshold**, we employ the CRT [4]:



Traditional mean-based approaches

Using the CATE conditioned on a single feature (i.e., the average treatment effect across individuals with identical attribute $X_m = x$):

$$\begin{aligned} T_m(x) &:= \mathbb{E}[Y^1 - Y^0 | X_m = x] \\ &= \mathbb{E}[Y^1 | X_m = x] - \mathbb{E}[Y^0 | X_m = x], \end{aligned} \quad (1)$$

the existing methods (e.g., [1]) seek **treatment effect modifiers**:

Definition 1 (Rothman et al. [2008]). Feature X_m is said to be a treatment effect modifier if there are at least two values of X_m , x_m and x_m^* ($x_m \neq x_m^*$), such that CATE T_m in (1) takes different values, i.e., $T_m(x_m) \neq T_m(x_m^*)$.

Weakness: Mean-based methods may overlook important features

Example:

$P(Y^0, Y^1 X = 0)$			$P(Y^0, Y^1 X = 1)$						
$Y^1 \backslash Y^0$	-1	0	1	Total	$Y^1 \backslash Y^0$	-1	0	1	Total
-1	0	0	0	0	-1	0	0	0	0
0	0.5	0	0.5	1.0	0	0	1.0	0	1.0
1	0	0	0	0	1	0	0	0	0
Total	0.5	0	0.5	1.0	Total	0	1.0	0	1.0

Individuals with $X=0$ Individuals with $X=1$

$$\begin{aligned} Y^1 - Y^0: & \quad -1 + 1 + 1 - 1 \\ \mathbb{E}[Y^1 - Y^0 | X]: & \quad 0 \\ \text{Var}[Y^1 - Y^0 | X]: & \quad 1 \end{aligned} = \begin{aligned} 0 \\ 0 \end{aligned} \quad \begin{aligned} \text{Identical means} \\ \text{Different variances} \end{aligned}$$

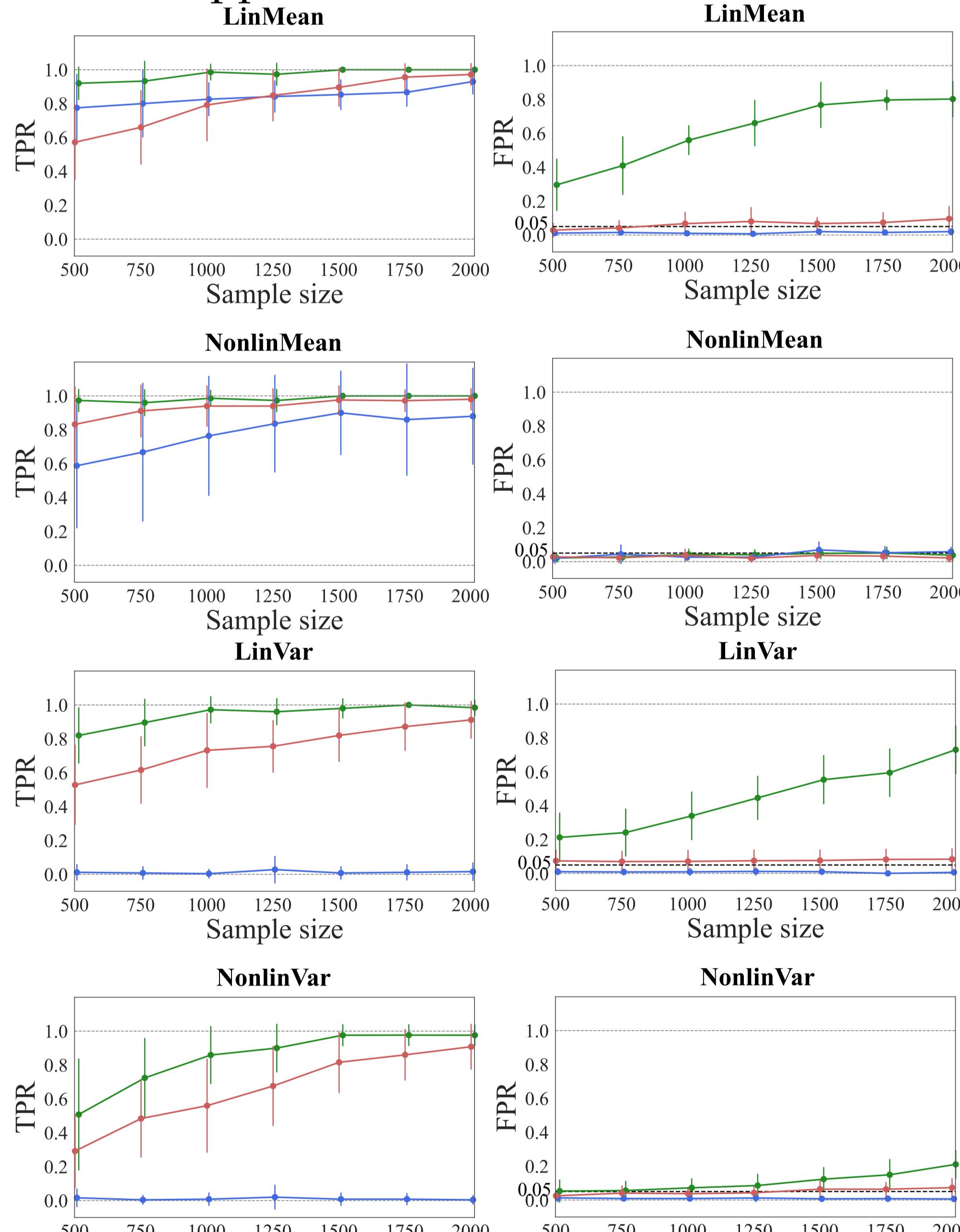
How can we detect distributional heterogeneity?

Experimental results

Synthetic data We compare our method with the two baselines:

1. **SI-EM** [1]: Mean-based approach

2. **Naive**: Approximate the null distribution via a naive bootstrap



Naive cannot control the FPR to $\alpha = 0.05$

SI-EM cannot detect the features related to treatment effect variance

Proposed achieves high TPR while controlling FPR

Real-world data We use health record dataset (from NHANES)

Treatment A : obesity Outcome Y : low-grade systemic inflammation
 Features X : e.g., age, gender, race, past medical history (e.g., asthma, stroke)

Feature	Adjusted p-value
Age	0.0075 ± 0.0305
Gender	0.0046 ± 0.0269
Number of cigarettes smoked	0.0 ± 0.0

Not detected by SI-EM

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[4] Emmanuel Candès, Yingying Fan, Lucas Janson, and Jinchi Lv. "Panning for gold: 'Model-X' knockoffs for high dimensional controlled variable selection". Journal of Royal Statistical Society: Series B (Statistical Methodology), 80 (3):551–577, 2018.