

$$X \sim \text{MVN}(\mu, \Sigma)$$

$$\mu \in \mathbb{R}^d$$

$$\Sigma \in \mathbb{R}^{d \times d} \quad \text{positive-definite sym matrix}$$

$$E[X] = \mu$$

$$\text{Var}(X) = \Sigma$$

$$X \sim \text{MVN}(0, I)$$

X is a vector of
iid $N(0, 1)$
r.v.s.

$$X \sim \text{MVN}(\mu, \Sigma)$$

$$a + AX \sim \text{MVN}(a + A\mu, A\Sigma A^T)$$

$$\Sigma = LL^T$$

$$X \sim \text{MVN}(0, I)$$

$$\mu + L X \sim \text{MVN}(\mu, \Sigma)$$

Copula is the joint cdf of

$$(U_1, \dots, U_d)$$

$$U_i \sim \text{Unif}(0, 1)$$

but not necessarily
ind.

$$X = (X_1, \dots, X_d) \quad F_1, \dots, F_d$$

copula of X is of the joint cdf

$$(F_1(X_1), \dots, F_d(X_d))$$

Gaussian copula correlation matrix

$$\Sigma \Rightarrow \Sigma_{ii} = 1$$

copula of $\text{MVN}(0, \Sigma)$

$$X = (X_1, \dots, X_d)$$

$$Y = (\Phi(X_1), \dots, \Phi(X_d))$$

sample from a random vector X
 w/ marginals given by F_1, \dots, F_d
 Gaussian copula w/ Σ (unit diagonal)

$$\textcircled{1} Y \sim \text{MVN}(0, \Sigma)$$

$$\textcircled{2} U = (\Phi(Y_1), \dots, \Phi(Y_d))$$

$$\textcircled{3} X = (F_1^{-1}(\Phi(Y_1)), \dots, F_d^{-1}(\Phi(Y_d)))$$

$$X_1 = (X_{11}, \dots, X_{1d})$$

\vdots

$$X_n = (X_{n1}, \dots, X_{nd})$$

fit models for marginals

$$\hat{F}_1, \dots, \hat{F}_d$$

$$U_i = (\hat{F}_1(X_{i1}), \dots, \hat{F}_d(X_{id}))$$

$$Y_i = (\Phi^{-1}(\hat{F}_1(X_{i1})), \dots,$$

$$\Phi^{-1}(\hat{F}(X_{id})))$$

t-copula can be used to
capture more extreme tail
dependencies

$$Z \sim N(0,1) \quad T = \frac{Z}{\sqrt{Y/v}} \sim t_v$$

$$Y \sim \chi^2_v$$

$$X \sim \text{MVN}(0, \Sigma)$$

$$Y \sim \chi^2_v$$

$$T = \frac{X}{\sqrt{Y/v}} \sim T_{\Sigma, v}$$

$$\textcircled{1} \quad X \sim \text{MVN}(0, \Sigma) \quad \Sigma_{ii} = 1$$

$$Y \sim \chi_v^2$$

$$T = \frac{X}{\sqrt{Y/v}}$$

$F_{T,v}$ = cdf of
a t-dist
with v

$\textcircled{2}$

$$U = (F_{T,v}(T_1), \dots, F_{T,v}(T_d))$$

$$\textcircled{3} \quad G_1, \dots, G_d$$

$$B = (G_1^{-1}(F_{T,v}(T_1)), \dots, \\ G_d^{-1}(F_{T,v}(T_d)))$$

Markov Chain Monte Carlo

Def Time homogeneous Markov Chain is a collection of random variable

$$X_1, \dots, X_t, \dots$$

that have a common support S

where conditional distribution

$$\begin{aligned} p(x_t | x_{t-1}, \dots, x_1) & \quad \text{for all } t \\ &= \underbrace{p(x_t | x_{t-1})}_{p(x|y)} \end{aligned}$$

Given a Markov Chain

$$p(x_t | x_{t-1})$$

P_t ← marginal dist'n of X_t
↳ this depends only on P_1

$$P_t(x_t) = \int p(x_t | x_{t-1}) p_{t-1}(x_{t-1}) dx_{t-1}$$

$$= \int \dots \int p(x_t | x_{t-1}) p(x_{t-1} | x_{t-2}) \dots p(x_2 | x_1) P_1(x_1) dx_1 dx_2 \dots dx_t$$

Def a distribution P^* is stationary wrt to a transition kernel $p(x|y)$ if

$$p^*(x) = \int p(x|y) p^*(y) dy$$

\Rightarrow If I have Markov Chain
 w/ $p(x|y)$, and I initialize
 w/ p^* that is stationary wrt
 $p(x|y)$,

then the marginal of X_t
 is p^* for every t .

$$\underline{P_1 = p^*}$$

$$\begin{aligned} P_2(x) &= \int p(x|y) P_1(y) dy \\ &= \int p(x|y) p^*(y) dy \\ &= p^*(x) \end{aligned}$$

Under some technical conditions,
given $p(x|y)$, \exists unique P^*
stationary.

If P^* exists, regardless of
what we use as P_1

$$\lim_{t \rightarrow \infty} P_t = P^*$$

Idea: inverse of the above
want to sample from f
find a Markov Chain whose
stationary distribution is
 f

sample X_1 however I want
generate a large # of transitions,

$$X_t \sim f \text{ for } t \text{ large}$$

Metropolis-Hastings

A sufficient condition
for f to be stationary w.r.t
 $p(x|y)$

$$\forall x, y \quad p(x|y) f(y) = p(y|x) f(x) \quad]$$

$$\begin{aligned} P_2(x) &= \int p(x|y) f(y) dy \\ &= \int p(y|x) \underbrace{f(x)} dy \\ &= f(x) \int p(y|x) dy \\ &= f(x) \end{aligned}$$

$$\frac{p(x|y)}{p(y|x)} = \frac{f(x)}{f(y)}$$

Split into 2 steps:

① generate $X_t = y$ proposal from

$$\theta \sim h(x|y)$$

② given θ , accept or reject

$$w/ a(\theta = x | y)$$

$$U \sim \text{Unif}(0,1)$$

$$\text{If } U \leq a(\theta | y)$$

$$X_{t+1} = \theta$$

O/w

$$X_{t+1} = y$$

$$p(x|y) = h(x|y) a(x|y) \quad x \neq y$$

$$p(x|x) = \int h(y|x) (1 - a(y|x)) dy$$

can re-write this:

$$\frac{p(x|y)}{p(y|x)} = \frac{f(x)}{f(y)}$$

$$\frac{h(x|y) a(x|y)}{h(y|x) a(y|x)} = \frac{f(x)}{f(y)}$$

assume we have a proposal kernel

$$\frac{a(x|y)}{a(y|x)} = \frac{f(x) h(y|x)}{f(y) h(x|y)}$$

$$a(x|y) = \min \left(1, \frac{f(x) h(y|x)}{f(y) h(x|y)} \right)$$

$$f(x) h(y|x) < f(y) h(x|y)$$

$$\textcircled{1} \quad X_1 \sim \text{whites} \quad t=1$$

$$\textcircled{2} \quad \theta_{t+1} \sim h(\cdot | X_t)$$

$$U \sim \text{Unif}(0,1)$$

$$\text{if } U \leq a(\theta_{t+1} | X_t)$$

$$\text{then } X_{t+1} = \theta_{t+1}$$

$$\text{olw } X_{t+1} = X_t$$

$$t = t + 1$$

repeat until I have generated
enough transitions