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# ProbLimI: Problem Set VI

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## Abstract

This work contains solutions to the exercises of the problem set V. The chosen problems are 1, 2, and 3.

### Question 1.

1. Let  $\{A_n\}$  be pairwise independent events with  $\sum_{n=1}^{\infty} \mathbb{P}(A_n) = \infty$ , and let  $S_n = \sum_{k=1}^n 1_{A_k}$ .
  - (a) Show that  $\text{Var}(S_n) \leq \mathbb{E}S_n$  and deduce that  $S_n/\mathbb{E}S_n \xrightarrow{P} 1$ .
  - (b) Show that if  $n_k = \inf\{n : \mathbb{E}S_n \geq k^2\}$  then  $S_{n_k}/\mathbb{E}S_{n_k} \xrightarrow{a.s.} 1$ . (Hint: use Borel-Cantelli I.)
  - (c) Prove that  $\mathbb{E}[S_{n_{k+1}}]/\mathbb{E}[S_{n_k}] \rightarrow 1$  and deduce that  $S_n/\mathbb{E}S_n \xrightarrow{a.s.} 1$ .

### Solution.

Observe that

$$\sum_{k=1}^n \mathbb{P}(A_k) = \mathbb{E}[S_n]$$

for any  $n \in \mathbb{N}$ . As the LHS tends to  $\infty$  as  $n \rightarrow \infty$ , we can choose  $N$  large enough such that  $\mathbb{E}[S_n] > 0$  for all  $n \geq N$ . We relabel the indices to start from  $N$  so that the random variables  $\{\frac{S_n}{\mathbb{E}[S_n]}\}$  are well-defined for the problem.

(i) By independence,

$$\begin{aligned} \text{Var}(S_n) &= \sum_{k=1}^n \text{Var}(1_{A_k}) = \sum_{k=1}^n \mathbb{E}[1_{A_k}^2] - \mathbb{E}[1_{A_k}]^2 = \sum_{k=1}^n \mathbb{P}(1_{A_k}) - \mathbb{P}(1_{A_k})^2 \\ &\leq \sum_{k=1}^n \mathbb{P}(1_{A_k}) = \mathbb{E}[S_n] \end{aligned}$$

for each  $n \geq 1$ . Now, we prove the claimed convergence in probability. Let  $\epsilon > 0$ . By Chebyshev's inequality and the above result,

$$\begin{aligned} \mathbb{P}\left(\left|\frac{S_n}{\mathbb{E}[S_n]} - 1\right| > \epsilon\right) &= \mathbb{P}(|S_n - \mathbb{E}[S_n]| > \epsilon \mathbb{E}[S_n]) \\ &\leq \frac{\text{Var}(S_n)}{\epsilon^2 \mathbb{E}[S_n]^2} \leq \frac{1}{\epsilon^2 \mathbb{E}[S_n]} \end{aligned}$$

for any  $n \in \mathbb{N}$ . Therefore, taking  $n \rightarrow \infty$ ,

$$\lim_{n \rightarrow \infty} \mathbb{P}\left(\left|\frac{S_n}{\mathbb{E}[S_n]} - 1\right| > \epsilon\right) = 0.$$

Since  $\epsilon > 0$  was arbitrary,  $\frac{S_n}{\mathbb{E}[S_n]} \rightarrow 1$  in probability.

(ii) As  $\mathbb{E}[S_n]$  tends to  $\infty$  as  $n \rightarrow \infty$ , we can find a subsequence with the given property. Let  $\epsilon > 0$ . By the same argument as above, and the property of the chosen subsequence,

$$\mathbb{P}(|\frac{S_n}{\mathbb{E}[S_n]} - 1| > \epsilon) \leq \frac{1}{\epsilon^2 \mathbb{E}[S_n]} \leq \frac{1}{\epsilon^2 k^2}$$

for all  $k \in \mathbb{N}$ , which implies

$$\sum_{k=1}^{\infty} \mathbb{P}(|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon) < \infty.$$

By Borel-Cantelli I,

$$\mathbb{P}(|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon \text{ i.o.}) = 0$$

for any  $\epsilon > 0$ . Now, by definition of pointwise convergence,

$$\mathbb{P}(\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} \rightarrow 1) = \mathbb{P}(\bigcap_{\epsilon > 0} \{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| \leq \epsilon \text{ a.a.}\}) = 1 - \mathbb{P}(\bigcup_{\epsilon > 0} \{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon \text{ i.o.}\})$$

By density of rationals and the above result,

$$\begin{aligned} \mathbb{P}(\bigcup_{\epsilon > 0} \{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon \text{ i.o.}\}) &= \mathbb{P}(\bigcup_{\epsilon > 0; \epsilon \in \mathbb{Q}} \{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon \text{ i.o.}\}) \\ &\leq \sum_{\epsilon > 0; \epsilon \in \mathbb{Q}} \mathbb{P}(|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon \text{ i.o.}) = 0 \end{aligned}$$

and hence

$$\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} \rightarrow 1 \text{ almost surely.}$$

(iii) Observe that

$$|\mathbb{E}[S_{n+1}] - \mathbb{E}[S_n]| = \mathbb{P}(A_{n+1}) \leq 1$$

for all  $n \geq 1$ , which implies that  $\{n_k\}$  chosen is strictly increasing as a function  $k$  and

$$\mathbb{E}[S_{n_k}] < (k+1)^2$$

for all  $k \geq 1$ . Therefore,

$$1 \leq \frac{\mathbb{E}[S_{n_{k+1}}]}{\mathbb{E}[S_{n_k}]} \leq \frac{(k+2)^2}{k^2}$$

for all  $k \geq 1$ , and hence, taking  $k \rightarrow \infty$ ,

$$\lim_{k \rightarrow \infty} \frac{\mathbb{E}[S_{n_{k+1}}]}{\mathbb{E}[S_{n_k}]} = 1.$$

Now, let  $w \in \Omega$  such that  $\frac{S_{n_k}(w)}{\mathbb{E}[S_{n_k}]} \rightarrow 1$ . Recall that

$$S_n(w) \leq S_{n+1}(w) \quad \text{and} \quad \mathbb{E}[S_n] \leq \mathbb{E}[S_{n+1}]$$

for all  $n \geq 1$ , and hence

$$\frac{\mathbb{E}[S_{n_k}]}{\mathbb{E}[S_{n_{k+1}}]} \frac{S_{n_k}(w)}{\mathbb{E}[S_{n_k}]} = \frac{S_{n_k}(w)}{\mathbb{E}[S_{n_{k+1}}]} \leq \frac{S_n(w)}{\mathbb{E}[S_n]} \leq \frac{S_{n_{k+1}}(w)}{\mathbb{E}[S_{n_k}]} = \frac{S_{n_{k+1}}(w)}{\mathbb{E}[S_{n_{k+1}}]} \frac{\mathbb{E}[S_{n_{k+1}}]}{\mathbb{E}[S_{n_k}]}$$

for any  $k \in \mathbb{N}$  and  $n_k \leq n < n_{k+1}$ . Set

$$L_n = \frac{\mathbb{E}[S_l]}{\mathbb{E}[S_u]} \frac{S_l(w)}{\mathbb{E}[S_l]} \quad \text{and} \quad U_n = \frac{\mathbb{E}[S_u]}{\mathbb{E}[S_l]} \frac{S_u(w)}{\mathbb{E}[S_u]}$$

where  $l = \sup\{n_k : n_k \leq n; k \in \mathbb{N}\}$  and  $u = \inf\{n_k : n_k > n; k \in \mathbb{N}\}$ , for any  $n \in \mathbb{N}$ . Then,

$$1 = \lim_{n \rightarrow \infty} L_n \leq \limsup_{n \rightarrow \infty} \frac{S_n(w)}{\mathbb{E}[S_n]} \leq \liminf_{n \rightarrow \infty} \frac{S_n(w)}{\mathbb{E}[S_n]} \leq \lim_{n \rightarrow \infty} U_n = 1$$

and hence

$$\left\{ \frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} \rightarrow 1 \right\} \subset \left\{ \frac{S_n}{\mathbb{E}[S_n]} \rightarrow 1 \right\}$$

which implies

$$\frac{S_n}{\mathbb{E}[S_n]} \rightarrow 1 \text{ almost surely.}$$

□

**Question 2.**

2. (a) Let  $X$  be a nonnegative random variable. Show that  $Y = \lfloor X \rfloor$  satisfies  $Y = \sum_{n=1}^{\infty} \mathbf{1}_{\{X \geq n\}}$ , and deduce that  $\mathbb{E}X - 1 \leq \sum_{n=1}^{\infty} \mathbb{P}(X \geq n) \leq \mathbb{E}X$ .
- (b) Let  $X_1, \dots, X_n, \dots$  be i.i.d. r.v.'s with  $\mathbb{E}|X_1|^\alpha = \infty$  for  $\alpha > 0$ . Show that for every  $\beta > 0$  one has  $\sum_{n=1}^{\infty} \mathbb{P}(|X_n| > \beta n^{1/\alpha}) = \infty$ , and deduce that  $\limsup_{n \rightarrow \infty} n^{-1/\alpha} |X_n| = \infty$ , a.s.
- (c) Conclude that  $S_n := \sum_{k=1}^n X_k$  satisfies  $\limsup_{n \rightarrow \infty} n^{-1/\alpha} |S_n| = \infty$ , a.s.

**Solution.**

(a) As  $X$  is non-negative real-valued RV and  $\mathbf{1}_{\{X \geq n\}}(w) = 0$  for each  $n > \max\{k \in \mathbb{N} : k \leq X(w)\}$ .

$$\lfloor X(w) \rfloor = \max\{k \in \mathbb{N} : k \leq X(w)\} = \sum_{n=1}^{\max\{k \in \mathbb{N} : k \leq X(w)\}} \mathbf{1}_{\{X \geq n\}}(w) = \sum_{n=1}^{\infty} \mathbf{1}_{\{X \geq n\}}(w)$$

for any  $w \in \Omega$ , and hence

$$\lfloor X \rfloor = \sum_{n=1}^{\infty} \mathbf{1}_{\{X \geq n\}} = Y.$$

Observe that  $\{\sum_{n=1}^k \mathbf{1}_{\{X \geq n\}}\}_k$  is a pointwise non-decreasing and non-negative sequence of RVs, which converges pointwise everywhere to  $Y = \lfloor X \rfloor$ . Hence, by MCT,

$$\mathbb{E}Y = \sum_{n=1}^{\infty} \mathbb{E} \mathbf{1}_{\{X \geq n\}} = \sum_{n=1}^{\infty} \mathbb{P}(X \geq n).$$

Since  $X - 1 \leq \lfloor X \rfloor \leq X$ , if  $X$  is integrable, by monotonicity of integration,

$$\mathbb{E}[X] - 1 \leq \sum_{n=1}^{\infty} \mathbb{P}(X \geq n) \leq \mathbb{E}[X].$$

If  $\mathbb{E}[X] = \mathbb{E}\lfloor X \rfloor - 1 = \infty$ , then  $X - 1$  is not integrable, as otherwise it will contradict the non-integrability of  $X$  by linearity. Therefore,  $\mathbb{E}[Y] = \mathbb{E}[X - 1] = \infty$ , so the inequality holds trivially.

(b) Let  $\beta > 0$ . Observe that

$$\sum_{n=0}^{\infty} \mathbf{1}_{\{\beta^{-\alpha} |X_1|^\alpha > n\}} = \lceil \beta^{-\alpha} |X_1|^\alpha \rceil.$$

Similar to (a), by MCT,

$$\sum_{n=0}^{\infty} \mathbb{P}(\beta^{-\alpha} |X_1|^\alpha > n) = \mathbb{E} \lceil \beta^{-\alpha} |X_1|^\alpha \rceil.$$

We now have the following pointwise estimate:

$$\beta^{-\alpha} |X_1|^\alpha \leq \lceil \beta^{-\alpha} |X_1|^\alpha \rceil.$$

As  $\mathbb{E}|X_1|^\alpha = \infty$ , we see  $\mathbb{E} \lceil \beta^{-\alpha} |X_1|^\alpha \rceil = \infty$  and combined with the above estimate

$$\sum_{n=0}^{\infty} \mathbb{P}(|X_n| > \beta n^{\frac{1}{\alpha}}) = \sum_{n=0}^{\infty} \mathbb{P}(\beta^{-\alpha} |X_1|^\alpha > n) = \infty$$

which implies

$$\sum_{n=1}^{\infty} \mathbb{P}(|X_n| > \beta n^{\frac{1}{\alpha}}) = \infty.$$

Since  $\beta > 0$  was arbitrary, we have the result for all  $\beta > 0$ .

Set

$$A_k = \{n^{-\frac{1}{\alpha}} |X_n| > k \text{ i.o.}\}$$

for each  $k \in \mathbb{N}$ . By Borel-Cantelli II, combined with the above result,

$$\mathbb{P}(A_k) = 1$$

for each  $k \in \mathbb{N}$ . Since  $\{A_k\}$  is descending, by continuity of probability,

$$\mathbb{P}\left(\bigcap_{k=1}^{\infty} A_k\right) = 1.$$

Suppose  $w \in \bigcap_{k=1}^{\infty} A_k$ . By induction, we construct a subsequence, which diverges to  $\infty$ . Choose  $n_1$  such that

$$(n_1)^{-\frac{1}{\alpha}} |X_{n_1}(w)| > 1.$$

Given  $\{n_i\}_{i=1}^l$ , choose  $n_{l+1}$  larger than all previous indices such that

$$(n_{l+1})^{-\frac{1}{\alpha}} |X_{n_{l+1}}(w)| > l + 1.$$

By induction, we have constructed a subsequence  $\{n_l\}$  such that

$$(n_l)^{-\frac{1}{\alpha}} |X_{n_l}(w)| > l$$

for each  $l \in \mathbb{N}$ , and hence

$$\bigcap_{k=1}^{\infty} A_k \subset \{\limsup_{n \rightarrow \infty} n^{-\frac{1}{\alpha}} |X_n| = \infty\}.$$

Therefore,

$$\limsup_{n \rightarrow \infty} n^{-\frac{1}{\alpha}} |X_n| = \infty \quad \text{a.s.}$$

(c) Firstly, by reverse triangle inequality,

$$|n^{-\frac{1}{\alpha}} |S_{n-1}| - n^{-\frac{1}{\alpha}} |X_n|| \leq n^{-\frac{1}{\alpha}} |S_n|$$

for all  $n \geq 2$ , and hence, by elementary properties of  $\limsup$

$$\begin{aligned} \limsup n^{-\frac{1}{\alpha}} |X_n| &\leq \limsup (n^{-\frac{1}{\alpha}} |X_n| - n^{-\frac{1}{\alpha}} |S_{n-1}|) + \limsup n^{-\frac{1}{\alpha}} |S_{n-1}| \\ &= \limsup (n^{-\frac{1}{\alpha}} |X_n| - n^{-\frac{1}{\alpha}} |S_{n-1}|) + \limsup (n-1)^{-\frac{1}{\alpha}} |S_{n-1}| \limsup \left(\frac{n}{n-1}\right)^{-\frac{1}{\alpha}} \\ &= \limsup (n^{-\frac{1}{\alpha}} |X_n| - n^{-\frac{1}{\alpha}} |S_{n-1}|) + \limsup n^{-\frac{1}{\alpha}} |S_n| \\ &\leq 2 \limsup n^{-\frac{1}{\alpha}} |S_n|. \end{aligned}$$

By the above estimate,

$$\{\limsup n^{-\frac{1}{\alpha}} |X_n| = \infty\} \subset \{\limsup n^{-\frac{1}{\alpha}} |S_n| = \infty\}$$

and hence

$$\limsup n^{-\frac{1}{\alpha}} |S_n| = \infty \quad \text{a.s.}$$

### Question 3.

3. Let  $(X_k)$  be i.i.d. r.v.'s taking values in  $\overline{\mathbb{R}}$  and let  $M_n = \max_{k \leq n} X_k$ .
- (a) Show that  $\mathbb{P}(\{|X_n| > n\} \text{ i.o.}) = 0$  if and only if  $\mathbb{E}|X_1| < \infty$ .
  - (b) Show that  $n^{-1}X_n \xrightarrow{\text{a.s.}} 0$  if and only if  $\mathbb{E}|X_1| < \infty$ .
  - (c) Show that  $n^{-1}M_n \xrightarrow{\text{a.s.}} 0$  if and only if  $\mathbb{E}(X_1)_+ < \infty$  and  $\mathbb{P}(X_1 > -\infty) > 0$ . Further show that  $n^{-1}M_n \xrightarrow{p} 0$  if and only if  $n\mathbb{P}(X_1 > n) \rightarrow 0$  and  $\mathbb{P}(X_1 > -\infty) > 0$ .
  - (d) Show that  $n^{-1}X_n \xrightarrow{p} 0$  if and only if  $\mathbb{P}(|X_1| < \infty) = 1$ .

### Solution.

Verbatim repeat the argument given in the problem 2 to have

$$\mathbb{E}|X| < \infty \iff \sum_n \mathbb{P}(|X| > n) \text{ converges}$$

and

$$\mathbb{E}X_+ < \infty \iff \sum_n \mathbb{P}(X > n) \text{ converges.}$$

(a) As  $(X_n)$  are identically distributed,

$$\sum_n \mathbb{P}(|X_1| > n) = \sum_n \mathbb{P}(|X_n| > n).$$

Therefore, by Borel-Cantelli I,II and the above equivalences

$$\mathbb{P}(|X_n| > n \text{ i.o.}) = 0 \iff \mathbb{E}|X_1| < \infty.$$

(b) Suppose  $\mathbb{E}|X_1| < \infty$ . Then,

$$\mathbb{E}\left|\frac{X_1}{\alpha}\right| < \infty$$

so, by (a)

$$\mathbb{P}(n^{-1}|X_n| > \alpha \text{ i.o.}) = 0$$

for any  $\alpha > 0$ . By the same argument given in 1-b,

$$n^{-1}|X_n| \rightarrow 0 \text{ a.s.}$$

Conversely, suppose  $n^{-1}X_n \rightarrow 1$  a.s. Observe that

$$\{n^{-1}X_n \rightarrow 0\} = \bigcap_{\alpha > 0} \{n^{-1}X_n \leq \alpha \text{ a.s.}\}.$$

With  $\alpha = 1$ ,

$$\mathbb{P}(n^{-1}|X_n| \leq 1 \text{ a.s.}) = 1$$

so

$$\mathbb{P}(|X_1| > n \text{ i.o.}) = 0.$$

which by (a) implies  $\mathbb{E}|X_1| < \infty$ .

(c) We first show the forward direction. Assume  $n^{-1}M_n \rightarrow 0$  a.s. Suppose for sake of contradiction that  $\mathbb{P}(X_1 > -\infty) = 0$ . Then,  $M_n = -\infty$  a.s. for all  $n \geq 1$ . Therefore,  $n^{-1}M_n \rightarrow -\infty$  a.s, which is a contradiction, and  $\mathbb{P}(X_1 > -\infty) > 0$ . Now, suppose again for sake of contradiction that  $\mathbb{E}(X_1)_+ = \infty$ , then by the equivalence established before, and Borel Cantelli II, we have  $\mathbb{P}(\{X_n > n \text{ i.o.}\}) = 1$ . Then, it follows that for a.s.  $w \in \Omega$ , there exists a subsequence  $X_{n_k}(w) > 1$  for all  $k$ , and

$$\limsup M_n(w) \geq \limsup X_n(w) \geq 1 \text{ a.s.}$$

which is a contradiction. Now, conversely, suppose there exists a set  $A$  with positive probability, where  $n^{-1}M_n \not\rightarrow 0$ , and  $\mathbb{E}(X_1)_+ < \infty$ . Then, we wish to show that  $X_1 = -\infty$  a.s. Since negative

values do not affect the limit behavior of  $M_n$ , we can further assume that  $X$ s take values that are non-negative or  $-\infty$ . Therefore,

$$\mathbb{P}(X_1 \geq 0) = 0.$$

Since  $\mathbb{E}(X_1)_+ < \infty$ . by Borel Cantelli I,

$$\mathbb{P}(X_n > n \text{ i.o.}) = 0$$

Hence, there almost everywhere  $-w$ , we have  $X_n(w) \leq n$  for all sufficiently large  $n$ , and

$$n^{-1} \limsup_n M_n(w) \leq 1 \text{ a.s.}$$

Assume  $n^{-1}M_n \rightarrow_p 0$ . If  $\mathbb{P}(X_1 > -\infty) = 0$ , then  $n^{-1}M_n \rightarrow -\infty$  a.s. so  $n^{-1}M_n \rightarrow_p -\infty$ , a contradiction. Hence,  $\mathbb{P}(X_1 > -\infty) > 0$ . Now, by iid assumption,

$$\begin{aligned} \mathbb{P}(n^{-1}M_n > \epsilon) &= 1 - \mathbb{P}(n^{-1}M_n \leq \epsilon) = 1 - \mathbb{P}(|n^{-1}X_1| \leq \epsilon)^n \\ &= 1 - (1 - \mathbb{P}(|n^{-1}X_1| > \epsilon))^n = 1 - e^{-n\mathbb{P}(|X_1| > n\epsilon)} \end{aligned}$$

for any  $\epsilon > 0$  and  $n \geq 1$ . Hence, by convergence in probability, with  $\epsilon = 1$ ,

$$n\mathbb{P}(|X_1| > \epsilon) \rightarrow 0$$

so

$$n\mathbb{P}(|X_1| > \epsilon) \rightarrow 0$$

**(d)** Since

$$\{|X_1| < \infty\} = \left\{ \frac{|X_1|}{\epsilon} < \infty \right\} = \bigcup_{n=1}^{\infty} \left\{ \frac{|X_1|}{\epsilon} \leq n \right\}$$

by continuity of probability,

$$\mathbb{P}(|X_1| < \infty) = 1 - \lim_n \mathbb{P}(|n^{-1}X_n| > \epsilon)$$

for any  $\epsilon > 0$ . Therefore,

$$\mathbb{P}(|X_1| < \infty) = 1 \iff \lim_n \mathbb{P}(|n^{-1}X_n| > \epsilon) = 0 \quad \forall \epsilon > 0.$$

By definition  $n^{-1}X_n \rightarrow_p 0$  iff

$$\lim_n \mathbb{P}(|n^{-1}X_n| > \epsilon) = 0$$

for any  $\epsilon > 0$ , so we are done. □

**Question 4.**

4. Let  $(X_k)$  be integrable i.i.d. r.v.'s with  $\mathbb{E}X_k = 0$ .
- (a) Let  $\{a_n\}$  and  $\{b_n\}$  are to sequences of real numbers such that  $b_n > 0$  and  $b_n \uparrow \infty$ . Show that if  $\sum_n a_n/b_n$  converges then  $b_n^{-1} \sum_{k=1}^n a_k \rightarrow 0$ .
  - (b) Show that  $\sum_{k=1}^\infty k^{-2} \text{Var}(X_k \mathbf{1}_{\{|X_k| \leq k\}}) \leq 2\mathbb{E}|X_1|$ .
  - (c) Conclude from parts (a),(b) and Kolmogorov's One Series Theorem that if  $c_n$  is a bounded sequence of non-random constants, then  $n^{-1} \sum_{k=1}^n c_k X_k \xrightarrow{a.s.} 0$  as  $n \rightarrow \infty$ .

**Solution.**

(a) Let  $b_0, s_0 = 0$  and  $s_n = \sum_{k=1}^n \frac{a_n}{b_n}$ , so  $a_n = b_n(s_n - s_{n-1})$  for each  $n \in \mathbb{N}$ . Observe that

$$\frac{1}{b_n} \sum_{k=1}^n a_n = \frac{1}{b_n} \sum_{k=1}^n b_k(s_k - s_{k-1}) = s_n - \sum_{k=1}^n \left( \frac{b_k - b_{k-1}}{b_n} \right) s_{k-1}$$

for each  $n \in \mathbb{N}$ . Let  $s_\infty$  be the limit of  $\{s_n\}$ . It suffices to show that the right most term on the above formula converges to  $s_\infty$ . Let  $\epsilon > 0$ . By triangle inequality,

$$\begin{aligned} \left| \sum_{k=1}^n \left( \frac{b_k - b_{k-1}}{b_n} \right) s_{k-1} - s_\infty \right| &\leq \sum_{k=1}^n \left( \frac{b_k - b_{k-1}}{b_n} \right) |s_{k-1} - s_\infty| \\ &= \sum_{k=1}^m \left( \frac{b_k - b_{k-1}}{b_n} \right) |s_{k-1} - s_\infty| + \sum_{k=m+1}^n \left( \frac{b_k - b_{k-1}}{b_n} \right) |s_{k-1} - s_\infty| \end{aligned}$$

for each  $1 \leq m < n$ . Choose  $m_0$  such that  $|s_n - s_\infty| < \epsilon$  for each  $n \geq m_0$ . Then

$$\left| \sum_{k=1}^n \left( \frac{b_k - b_{k-1}}{b_n} \right) s_{k-1} - s_\infty \right| \leq \frac{1}{b_n} \sum_{k=1}^{m_0} (b_k - b_{k-1}) |s_{k-1} - s_\infty| + \frac{b_n - b_{m_0}}{b_n} \epsilon$$

for each  $n \geq m_0$ . Letting  $n \rightarrow \infty$ ,

$$\left| \sum_{k=1}^n \left( \frac{b_k - b_{k-1}}{b_n} \right) s_{k-1} - s_\infty \right| < \epsilon$$

as required.

(b)

(c)