ProbLimI: Problem Set VI

Youngduck Choi CIMS New York University yc1104@nyu.edu

Abstract

This work contains solutions to the exercises of the problem set V. The chosen problems are 1,2, and 3.

Question 1.

- 1. Let $\{A_n\}$ be pairwise independent events with $\sum_{n=1}^{\infty} \mathbb{P}(A_n) = \infty$, and let $S_n = \sum_{k=1}^n \mathbf{1}_{A_k}$.
 - (a) Show that $Var(S_n) \leq \mathbb{E}S_n$ and deduce that $S_n/\mathbb{E}S_n \stackrel{p}{\to} 1$.
 - (b) Show that if $n_k = \inf\{n : \mathbb{E}S_n \ge k^2\}$ then $S_{n_k}/\mathbb{E}S_{n_k} \stackrel{a.s.}{\to} 1$. (Hint: use Borel-Cantelli I.)
 - (c) Prove that $\mathbb{E}[S_{n_{k+1}}]/\mathbb{E}[S_{n_k}] \to 1$ and deduce that $S_n/\mathbb{E}S_n \stackrel{a.s.}{\to} 1$.

Solution.

Observe that

$$\sum_{k=1}^{n} \mathbb{P}(A_k) = \mathbb{E}[S_n]$$

for any $n \in \mathbb{N}$. As the LHS tends to ∞ as $n \to \infty$, we can choose N large enough such that $\mathbb{E}[S_n] > 0$ for all $n \ge N$. We relabel the indices to start from N so that the random variables $\{\frac{S_n}{\mathbb{E}[S_n]}\}$ are well-defined for the problem.

(i) By independence,

$$Var(S_n) = \sum_{k=1}^{n} Var(1_{A_k}) = \sum_{k=1}^{n} \mathbb{E}[1_{A_k}^2] - \mathbb{E}[1_{A_k}]^2 = \sum_{k=1}^{n} \mathbb{P}(1_{A_k}) - \mathbb{P}(1_{A_k})^2$$

$$\leq \sum_{k=1}^{n} \mathbb{P}(1_{A_k}) = \mathbb{E}[S_n]$$

for each $n \ge 1$. Now, we prove the claimed convergence in probability. Let $\epsilon > 0$. By Chebyshev's inequality and the above result,

$$\mathbb{P}(\left|\frac{S_n}{\mathbb{E}[S_n]} - 1\right| > \epsilon) = \mathbb{P}(\left|S_n - \mathbb{E}[S_n]\right| > \epsilon \mathbb{E}[S_n])$$

$$\leq \frac{\operatorname{Var}(S_n)}{\epsilon^2 \mathbb{E}[S_n]^2} \leq \frac{1}{\epsilon^2 \mathbb{E}[S_n]}$$

for any $n \in \mathbb{N}$. Therefore, taking $n \to \infty$,

$$\lim_{n\to\infty} \mathbb{P}(\left|\frac{S_n}{\mathrm{E}[S_n]} - 1\right| > \epsilon) = 0.$$

Since $\epsilon > 0$ was arbitrary, $\frac{S_n}{\mathrm{E}[S_n]} \to 1$ in probability.

(ii) As $\mathbb{E}[S_n]$ tends to ∞ as $n \to \infty$, we can find a subsequence with the given property. Let $\epsilon > 0$. By the same argument as above, and the property of the chosen subsequence,

$$\mathbb{P}(\left|\frac{S_n}{\mathbb{E}[S_n]} - 1\right| > \epsilon) \leq \frac{1}{\epsilon^2 \mathbb{E}[S_n]} \leq \frac{1}{\epsilon^2 k^2}$$

for all $k \in \mathbb{N}$, which implies

$$\sum_{k=1}^{\infty} \mathbb{P}(\left|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1\right| > \epsilon) < \infty.$$

By Borel-Cantelli I,

$$\mathbb{P}(\left|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1\right| > \epsilon \quad \text{i.o.}) = 0$$

for any $\epsilon > 0$. Now, by definition of pointwise convergence,

$$\mathbb{P}\big(\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} \to 1\big) \quad = \quad \mathbb{P}\big(\bigcap_{\epsilon>0}\{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| \le \epsilon \text{ a.a.}\}\big) = 1 - \mathbb{P}\big(\bigcup_{\epsilon>0}\{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} - 1| > \epsilon \text{ i.o.}\}\big)$$

By density of rationals and the above result,

$$\begin{split} \mathbb{P}\big(\bigcup_{\epsilon>0}\{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]}-1|>\epsilon \text{ i.o.}\}\big) &= \mathbb{P}\big(\bigcup_{\epsilon>0;\epsilon\in\mathbb{Q}}\{|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]}-1|>\epsilon \text{ i.o.}\}\big) \\ &\leq \sum_{\epsilon>0:\epsilon\in\mathbb{Q}}\mathbb{P}\big(|\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]}-1|>\epsilon \text{ i.o.}\big)=0 \end{split}$$

and hence

$$\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} \to 1 \text{ almost surely.}$$

(iii) Observe that

$$|\mathbb{E}[S_{n+1}] - \mathbb{E}[S_n]| = \mathbb{P}(A_{n+1}) \le 1$$

for all $n \ge 1$, which implies that $\{n_k\}$ chosen is strictly increasing as a function k and

$$\mathbb{E}[S_{n_k}] < (k+1)^2$$

for all $k \ge 1$. Therefore,

$$1 \le \frac{\mathbb{E}[S_{n_{k+1}}]}{\mathbb{E}[S_{n_k}]} \le \frac{(k+2)^2}{k^2}$$

for all $k \ge 1$, and hence, taking $k \to \infty$,

$$\lim_{k \to \infty} \frac{\mathbb{E}[S_{n_{k+1}}]}{\mathbb{E}[S_{n_k}]} = 1.$$

Now, let $w \in \Omega$ such that $\frac{S_{n_k(w)}}{\mathbb{E}[S_{n_k}]} \to 1$. Recall that

$$S_n(w) \le S_{n+1}(w)$$
 and $\mathbb{E}[S_n] \le \mathbb{E}[S_{n+1}]$

for all $n \ge 1$, and hence

$$\frac{\mathbb{E}[S_{n_k}]}{\mathbb{E}[S_{n_{k+1}}]} \frac{S_{n_k}(w)}{\mathbb{E}[S_{n_k}]} = \frac{S_{n_k}(w)}{\mathbb{E}[S_{n_{k+1}}]} \leq \frac{S_n(w)}{\mathbb{E}[S_n]} \leq \frac{S_{n_{k+1}}(w)}{\mathbb{E}[S_{n_k}]} = \frac{S_{n_{k+1}}(w)}{\mathbb{E}[S_{n_{k+1}}]} \frac{\mathbb{E}[S_{n_{k+1}}]}{\mathbb{E}[S_{n_k}]}$$

for any $k \in \mathbb{N}$ and $n_k \le n < n_{k+1}$. Set

$$L_n = \frac{\mathbb{E}[S_l]}{\mathbb{E}[S_u]} \frac{S_l(w)}{\mathbb{E}[S_l]} \quad \text{ and } \quad U_n = \frac{\mathbb{E}[S_u]}{\mathbb{E}[S_l]} \frac{S_u(w)}{\mathbb{E}[S_u]}$$

where $l = \sup\{n_k : n_k \le n; k \in \mathbb{N}\}$ and $u = \inf\{n_k : n_k > n; k \in \mathbb{N}\}$, for any $n \in \mathbb{N}$. Then,

$$1 = \lim_{n \to \infty} L_n \le \limsup_{n \to \infty} \frac{S_n(w)}{\mathbb{E}[S_n]} \le \liminf_{n \to \infty} \frac{S_n(w)}{\mathbb{E}[S_n]} \le \lim_{n \to \infty} U_n = 1$$

and hence

$$\big\{\frac{S_{n_k}}{\mathbb{E}[S_{n_k}]} \to 1\big\} \quad \subset \quad \big\{\frac{S_n}{\mathbb{E}[S_n]} \to 1\big\}$$

which implies

$$\frac{S_n}{\mathbb{E}[S_n]} \to 1 \text{ almost surely.}$$

Question 2.

2. (a) Let X be a nonnegative random variable. Show that $Y = \lfloor X \rfloor$ satisfies $Y = \sum_{n=1}^{\infty} 1_{\{X \geq n\}}$, and deduce that $\mathbb{E}X - 1 \leq \sum_{n=1}^{\infty} \mathbb{P}(X \geq n) \leq \mathbb{E}X$.

(b) Let X_1,\dots,X_n,\dots be i.i.d. r.v.'s with $\mathbb{E}|X_1|^\alpha=\infty$ for $\alpha>0$. Show that for every $\beta>0$ one has $\sum_{n=1}^\infty \mathbb{P}\left(|X_n|>\beta n^{1/\alpha}\right)=\infty$, and deduce that $\limsup_{n\to\infty} n^{-1/\alpha}|X_n|=\infty$, a.s.

(c) Conclude that $S_n := \sum_{k=1}^n X_k$ satisfies $\limsup_{n\to\infty} n^{-1/\alpha} |S_n| = \infty$, a.s.

Solution.

(a) As X is non-negative real-valued RV and $1_{\{X \ge n\}}(w) = 0$ for each $n > \max\{k \in \mathbb{N} : k \le X(w)\}$.

$$[X(w)] = \max\{k \in \mathbb{N} : k \le X(w)\} = \sum_{n=1}^{\max\{k \in \mathbb{N} : k \le X(w)\}} 1_{\{X \ge n\}}(w) = \sum_{n=1}^{\infty} 1_{\{X \ge n\}}(w)$$

for any $w \in \Omega$, and hence

$$\lfloor X \rfloor = \sum_{n=1}^{\infty} 1_{\{X \ge n\}} = Y.$$

Observe that $\{\sum_{n=1}^k 1_{\{X \geq n\}}\}_k$ is a pointwise non-decreasing and non-negative sequence of RVs, which converges pointwise everywhere to $Y = \lfloor X \rfloor$. Hence, by MCT,

$$\mathbb{E}Y = \sum_{n=1}^{\infty} \mathbb{E}1_{\{X \ge n\}} = \sum_{n=1}^{\infty} \mathbb{P}(X \ge n).$$

Since $X - 1 \le \lfloor X \rfloor \le X$, if X is integrable, by monotonicity of integration,

$$\mathbb{E}[X] - 1 \le \sum_{n=1}^{\infty} \mathbb{P}(X \ge n) \le \mathbb{E}[X].$$

If $\mathbb{E}[X] = \mathbb{E}[X] - 1 = \infty$, then X - 1 is not integrable, as otherwise it will contradict the non-integrability of X by linearity. Therefore, $\mathbb{E}[Y] = \mathbb{E}[X - 1] = \infty$, so the inequality holds trivially.

(b) Let $\beta > 0$. Observe that

$$\sum_{n=0}^{\infty} 1_{\{\beta^{-\alpha}|X_1|^{\alpha} > n\}} = \lceil \beta^{-\alpha}|X_1|^{\alpha} \rceil.$$

Similar to (a), by MCT,

$$\sum_{n=0}^{\infty} \mathbb{P}(\beta^{-\alpha}|X_1|^{\alpha} > n) = \mathbb{E}[\beta^{-\alpha}|X_1|^{\alpha}].$$

We now have the following pointwise estimate:

$$\beta^{-\alpha}|X_1|^{\alpha} \le [\beta^{-\alpha}|X_1|^{\alpha}].$$

As $\mathbb{E}|X_1|^{\alpha} = \infty$, we see $\mathbb{E}[\beta^{-\alpha}|X_1|^{\alpha}] = \infty$ and combined with the above estimate

$$\sum_{n=0}^{\infty} \mathbb{P}(|X_n| > \beta n^{\frac{1}{\alpha}}) = \sum_{n=0}^{\infty} \mathbb{P}(\beta^{-\alpha} |X_1|^{\alpha} > n) = \infty$$

which implies

$$\sum_{n=1}^{\infty} \mathbb{P}(|X_n| > \beta n^{\frac{1}{\alpha}}) = \infty.$$

Since $\beta > 0$ was arbitrary, we have the result for all $\beta > 0$.

Set

$$A_k = \{n^{-\frac{1}{\alpha}}|X_n| > k \text{ i.o.}\}$$

for each $k \in \mathbb{N}$. By Borel-Cantelli II, combined with the above result,

$$\mathbb{P}(A_k) = 1$$

for each $k \in \mathbb{N}$. Since $\{A_k\}$ is descending, by continuity of probability,

$$\mathbb{P}(\bigcap_{k=1}^{\infty} A_k) = 1.$$

Suppose $w \in \bigcap_{k=1}^{\infty} A_k$. By induction, we construct a subsequence, which diverges to ∞ . Choose n_1 such that

$$(n_1)^{-\frac{1}{\alpha}}|X_{n_1}(w)| > 1.$$

Given $\{n_i\}_{i=1}^l$, choose n_{l+1} larger than all previous indices such that

$$(n_{l+1})^{-\frac{1}{\alpha}}|X_{n_{l+1}}(w)| > l+1.$$

By induction, we have constructed a subsequence $\{n_l\}$ such that

$$(n_l)^{-\frac{1}{\alpha}}|X_{n_l}(w)| > l$$

for each $l \in \mathbb{N}$, and hence

$$\bigcap_{k=1}^{\infty} A_k \subset \{\limsup_{n \to \infty} n^{-\frac{1}{\alpha}} |X_n| = \infty \}.$$

Therefore,

$$\limsup_{n\to\infty} n^{-\frac{1}{\alpha}} |X_n| = \infty \quad \text{a.s.}$$

(c) Firstly, by reverse triangle inequality,

$$|n^{-\frac{1}{\alpha}}|S_{n-1}| - n^{-\frac{1}{\alpha}}|X_n|| \le n^{-\frac{1}{\alpha}}|S_n|$$

for all $n \ge 2$, and hence, by elementary properties of \limsup

$$\limsup_{n \to \infty} n^{-\frac{1}{\alpha}} |X_{n}| \leq \limsup_{n \to \infty} (n^{-\frac{1}{\alpha}} |X_{n}| - n^{-\frac{1}{\alpha}} |S_{n-1}|) + \limsup_{n \to \infty} n^{-\frac{1}{\alpha}} |S_{n-1}| \\
= \limsup_{n \to \infty} (n^{-\frac{1}{\alpha}} |X_{n}| - n^{-\frac{1}{\alpha}} |S_{n-1}|) + \limsup_{n \to \infty} (n - 1)^{-\frac{1}{\alpha}} |S_{n-1}| \lim \sup_{n \to \infty} (\frac{n}{n-1})^{-\frac{1}{\alpha}} \\
= \limsup_{n \to \infty} (n^{-\frac{1}{\alpha}} |X_{n}| - n^{-\frac{1}{\alpha}} |S_{n-1}|) + \limsup_{n \to \infty} n^{-\frac{1}{\alpha}} |S_{n}| \\
\leq 2 \limsup_{n \to \infty} n^{-\frac{1}{\alpha}} |S_{n}|.$$

By the above estimate,

$$\{\limsup n^{-\frac{1}{\alpha}}|X_n|=\infty\} \quad \subset \quad \{\limsup n^{-\frac{1}{\alpha}}|S_n|=\infty\}$$

and hence

$$\limsup n^{-\frac{1}{\alpha}} |S_n| = \infty \qquad \text{a.s.}$$

Question 3.

3. Let (X_k) be i.i.d. r.v.'s taking values in $\overline{\mathbb{R}}$ and let $M_n = \max_{k \le n} X_k$.

- (a) Show that $\mathbb{P}(\{|X_n| > n\} \text{ i.o.}) = 0$ if and only if $\mathbb{E}|X_1| < \infty$.
- (b) Show that $n^{-1}X_n \stackrel{a.s.}{\to} 0$ if and only if $\mathbb{E}|X_1| < \infty$.
- (c) Show that $n^{-1}M_n\stackrel{a.s.}{\to}0$ if and only if $\mathbb{E}(X_1)_+<\infty$ and $\mathbb{P}(X_1>-\infty)>0$. Further show that $n^{-1}M_n\stackrel{P}{\to}0$ if and only if $n\mathbb{P}(X_1>n)\to 0$ and $\mathbb{P}(X_1>-\infty)>0$.
- (d) Show that $n^{-1}X_n \stackrel{p}{\to} 0$ if and only if $\mathbb{P}(|X_1| < \infty) = 1$.

Solution.

Verbatim repeat the argument given in the problem 2 to have

$$\mathbb{E}|X| < \infty \quad \Longleftrightarrow \quad \sum_{n} \mathbb{P}(|X| > n) \text{ converges}$$

and

$$\mathbb{E}X_+ < \infty \quad \Longleftrightarrow \quad \sum_n \mathbb{P}(X > n) \text{ converges.}$$

(a) As (X_n) are identically distributed,

$$\sum_{n} \mathbb{P}(|X_1| > n) = \sum_{n} \mathbb{P}(|X_n| > n).$$

Therefore, by Borel-Cantelli I,II and the above equivalences

$$\mathbb{P}(|X_n| > n \text{ i.o.}) = 0 \iff \mathbb{E}|X_1| < \infty.$$

(b) Suppose $\mathbb{E}|X_1| < \infty$. Then,

$$\mathbb{E}\left|\frac{X_1}{\alpha}\right| < \infty$$

so, by (a)

$$\mathbb{P}(n^{-1}|X_n| > \alpha \text{ i.o.}) = 0$$

for any $\alpha > 0$. By the same argument given in 1 - b,

$$n^{-1}|X_n| \to 0$$
 a.s.

Conversely, suppose $n^{-1}X_n \to 1$ a.s. Observe that

$$\{n^{-1}X_n \to 0\} = \bigcap_{\alpha>0} \{n^{-1}X_n \le \alpha \text{ a.a.}\}.$$

With $\alpha = 1$,

$$\mathbb{P}(n^{-1}|X_n| \le 1 \text{ a.a.}) = 1$$

so

$$\mathbb{P}(|X_1| > n \text{ i.o.}) = 0.$$

which by (a) implies $\mathbb{E}|X_1| < \infty$.

(c) We first show the forward direction. Assume $n^{-1}M_n \to 0$ a.s. Suppose for sake of contradiction that $\mathbb{P}(X_1 > -\infty) = 0$. Then, $M_n = -\infty$ a.s. for all $n \ge 1$. Therefore, $n^{-1}M_n \to -\infty$ a.s, which is a contradiction, and $\mathbb{P}(X_1 > -\infty) > 0$. Now, suppose again for sake of contradiction that $\mathbb{E}(X_1)_+ = \infty$, then by the equivalence established before, and Borel Cantelli II, we have $\mathbb{P}(\{X_n > n \text{ i.o.}\}) = 1$. Then, it follows that for a.s. $w \in \Omega$, there exists a subsequence $X_{n_k}(w) > 1$ for all k, and

$$\limsup M_n(w) \ge \limsup X_n(w) \ge 1$$
 a.s.

which is a contradiction. Now, conversely, suppose there exists a set A with positive probability, where $n^{-1}M_n \not \to 0$. There are two possible cases $\mathbb{P}(X_1 > \infty) = 0$ or $\mathbb{P}(X_1 > \infty) > 0$. It suffices to show that if $\mathbb{P}(X_1 > \infty)$

By iid assumption,

$$\mathbb{P}(n^{-1}M_n > \epsilon) = 1 - \mathbb{P}(n^{-1}M_n \le \epsilon) = 1 - \mathbb{P}(|n^{-1}X_1| \le \epsilon)^n = 1 - (1 - \mathbb{P}(|n^{-1}X_1| > \epsilon)^n)$$

for any $\epsilon > 0$ and $n \ge 1$. Hence

$$\lim_{n} \mathbb{P}(n^{-1}M_{n} > \epsilon) = 1 - \exp(-\mathbb{P}(|n^{-1}X_{1}| > \epsilon))$$

provided the limit exists.

(d) Since

$$\big\{\big|X_1\big|<\infty\big\} \quad = \quad \big\{\frac{\big|X_1\big|}{\epsilon}<\infty\big\} = \bigcup_{n=1}^\infty \big\{\frac{\big|X_1\big|}{\epsilon} \leq n\big\}$$

by continuity of probability,

$$\mathbb{P}(|X_1| < \infty) = 1 - \lim_n \mathbb{P}(|n^{-1}X_n| > \epsilon)$$

for any $\epsilon > 0$. Therefore,

$$\mathbb{P}(|X_1|<\infty)=1\quad\Longleftrightarrow\quad \lim_n\mathbb{P}(|n^{-1}X_n|>\epsilon)=0\ \forall\epsilon>0.$$

By definition $n^{-1}X_n \to_p 0$ iff

$$\lim_{n} \mathbb{P}(|n^{-1}X_n| > \epsilon) = 0$$

for any $\epsilon > 0$, so we are done.

Question 4.

4. Let (X_k) be integrable i.i.d. r.v.'s with $\mathbb{E}X_k = 0$.

- (a) Let $\{a_n\}$ and $\{b_n\}$ are to sequences of real numbers such that $b_n > 0$ and $b_n \uparrow \infty$. Show that if $\sum_n a_n/b_n$ converges then $b_n^{-1} \sum_{k=1}^n a_k \to 0$.
- (b) Show that $\sum_{k=1}^\infty k^{-2} \operatorname{Var}(X_k \mathbf{1}_{\{|X_k| \leq k\}}) \leq 2\mathbb{E}|X_1|.$
- (c) Conclude from parts (a),(b) and Kolmogorov's One Series Theorem that if c_n is a bounded sequence of non-random constants, then $n^{-1}\sum_{k=1}^n c_k X_k \overset{a.s.}{\to} 0$ as $n \to \infty$.

Solution.

(a) Let $b_0, s_0 = 0$ and $s_n = \sum_{k=1}^n \frac{a_n}{b_n}$, so $a_n = b_n(s_n - s_{n-1})$ for each $n \in \mathbb{N}$. Observe that

$$\frac{1}{b_n} \sum_{k=1}^n a_n = \frac{1}{b_n} \sum_{k=1}^n b_k (s_k - s_{k-1}) = s_n - \sum_{k=1}^n (\frac{b_k - b_{k-1}}{b_n}) s_{k-1}$$

for each $n \in \mathbb{N}$. Let s_{∞} be the limit of $\{s_n\}$. It suffices to show that the right most term on the above formula converges to s_{∞} . Let $\epsilon > 0$. By triangle inequality,

$$\left| \sum_{k=1}^{n} \left(\frac{b_{k} - b_{k-1}}{b_{n}} \right) s_{k-1} - s_{\infty} \right| \leq \sum_{k=1}^{n} \left(\frac{b_{k} - b_{k-1}}{b_{n}} \right) \left| s_{k-1} - s_{\infty} \right|$$

$$= \sum_{k=1}^{m} \left(\frac{b_{k} - b_{k-1}}{b_{n}} \right) \left| s_{k-1} - s_{\infty} \right| + \sum_{k=m+1}^{n} \left(\frac{b_{k} - b_{k-1}}{b_{n}} \right) \left| s_{k-1} - s_{\infty} \right|$$

for each $1 \le m < n$. Choose m_0 such that $|s_n - s_\infty| < \epsilon$ for each $n \ge m_0$. Then

$$\left| \sum_{k=1}^{n} \left(\frac{b_k - b_{k-1}}{b_n} \right) s_{k-1} - s_{\infty} \right| \leq \frac{1}{b_n} \sum_{k=1}^{m_0} \left(b_k - b_{k-1} \right) \left| s_{k-1} - s_{\infty} \right| + \frac{b_n - b_{m_0}}{b_n} \epsilon$$

for each $n \ge m_0$. Letting $n \to \infty$,

$$\left|\sum_{k=1}^{n} \left(\frac{b_k - b_{k-1}}{b_n}\right) s_{k-1} - s_{\infty}\right| < \epsilon$$

as required.

- **(b)**
- (c)