

Introduction to Regression

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Professor

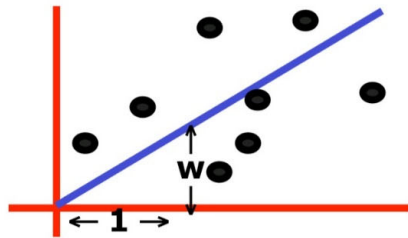
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Linear Regression

- 지도학습의 한 분야로 연속적인 숫자(실수)를 예측하는 것
 - 어떤 사람의 교육수준, 나이, 주거지를 바탕으로 연간소득 예측하는 문제
 - 측정된 점들의 열로부터 가장 근사한 방정식을 구하는 문제



inputs	outputs
$x_1 = 1$	$y_1 = 1$
$x_2 = 3$	$y_2 = 2.2$
$x_3 = 2$	$y_3 = 2$
$x_4 = 1.5$	$y_4 = 1.9$
$x_5 = 4$	$y_5 = 3.1$

- Linear regression assumes that the expected value of the output given an input, $E[Y|X]$, is linear.
i.e., $E[Y|X] = \alpha + \beta X$ or $E[Y|X] = \alpha + \beta_1 X_1 + \dots + \beta_p X_p$ where $E[Y|X] = \int y f(y/x) dy$
 - Simplest case: $Out(x) = wx$ for some unknown w .
 - Given the data, we can estimate w .

2-parameter linear regression

Observable dataset : $\mathbf{d}_1(x_1, y_1), \mathbf{d}_2(x_2, y_2) \dots \mathbf{d}_n(x_n, y_n)$

Model : $y = wx + b$

Compute mean squared error of the model on the dataset

$$MSE = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \frac{1}{n} \sum_{i=1}^n (y_i - wx_i - b)^2$$

To minimize MSE

$$\begin{cases} \frac{\partial}{\partial w} MSE = \frac{\partial}{\partial w} \frac{1}{n} \sum_{i=1}^n (y_i - wx_i - b)^2 = 0 \rightarrow \sum_{i=1}^n x_i (y_i - wx_i - b) = 0 \\ \frac{\partial}{\partial b} MSE = \frac{\partial}{\partial b} \frac{1}{n} \sum_{i=1}^n (y_i - wx_i - b)^2 = 0 \rightarrow \sum_{i=1}^n (y_i - wx_i - b) = 0 \end{cases}$$

$$\Rightarrow \begin{cases} w \sum_{i=1}^n x_i^2 + b \sum_{i=1}^n x_i = \sum_{i=1}^n x_i y_i \\ w \sum_{i=1}^n x_i + nb = \sum_{i=1}^n y_i \end{cases} \quad \text{if } b = 0, \quad w = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2} \quad (\text{즉, } y = wx \text{ 로 모델링 하면})$$

Bayesian linear regression

- Assume that the data is formed by $y_i = wx_i + noise_i$
 - the noise signals are independent
 - the noise has a normal distribution with mean 0 and unknown variance σ^2
 - $p(y|w,x)$ has a normal distribution with mean wx and variance σ^2
- $y \sim N(wx, \sigma^2)$
- We have a set of data $\mathbf{d}_1(x_1, y_1), \mathbf{d}_2(x_2, y_2) \dots \mathbf{d}_n(x_n, y_n)$.
- We want to infer w from the data.

$$P(w|\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_n) = P(w|\mathbf{D})$$

- We can use BAYES rule to work out a posterior distribution for w given the data.
- Or, we could do Maximum Likelihood Estimation.

Maximum likelihood estimation of w

- Choose the parameter w that maximizes the probability of the data, given that parameter.
- MLE asks: “For which value of w is this data most likely to have happened?”

For what w , is $P(\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_n | w)$ maximized?

\equiv For what w , is $\prod_{i=1}^n P(\mathbf{d}_i | w)$ maximized?

\equiv For what w , is $\prod_{i=1}^n \exp\left(-\frac{1}{2}\left(\frac{y_i - wx_i}{\sigma}\right)^2\right)$ maximized?

\equiv For what w , is $\sum_{i=1}^n (y_i - wx_i)^2$ minimized?

where $P(\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_n | w)$ is called the **Likelihood**, and

$$P(\mathbf{d}_i | w) = P(y_i | w, x_i) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{y_i - wx_i}{\sigma}\right)^2\right).$$

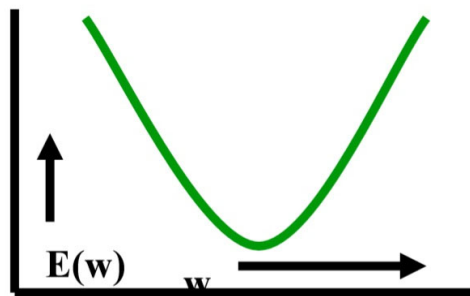
First result

- MLE with Gaussian noise is the same as minimizing the L_2 error

$$\operatorname{argmin}_w \sum_{i=1}^n (y_i - wx_i)^2$$

The maximum likelihood w is the one that minimizes sum-of-squares of residuals

$$\begin{aligned} E &= \sum_{i=1}^n (y_i - wx_i)^2 \\ &= \sum_{i=1}^n y_i^2 - (2 \sum x_i y_i)w + (\sum x_i^2)w^2 \end{aligned}$$



We want to minimize a quadratic function of w .

Linear regression

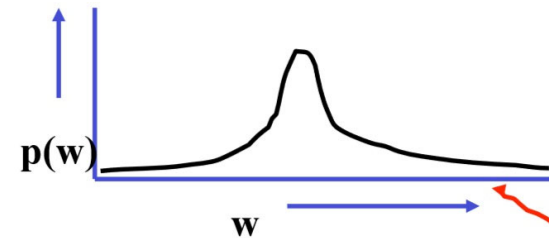
- Easy to show the sum of squares is minimized when

$$w = \frac{\sum x_i y_i}{\sum x_i^2}$$

The maximum likelihood model is:

$$Out(x) = wx$$

We can use it for prediction.



Note: In Bayesian stats you'd have ended up with a prob distribution of w

And predictions would have given a prob distribution of expected output

Often useful to know your confidence.
Max likelihood can give some kinds of confidence too.

Maximum a Posteriori estimation of w

■ MAP

- Choose w that maximizes the posteriori probability of w .
- Posterior probability of w is given by the Bayes Rule:

$$P(w|\mathbf{D}) = \frac{P(w)P(\mathbf{D}|w)}{P(\mathbf{D})}$$

where $P(w)$: Prior probability of w assumed as $w \sim N(0, \gamma^2)$

$P(\mathbf{D})$: Probability of data (independent of w)

$$P(\mathbf{D}) = \int P(w)P(\mathbf{D}|w)dw$$

Maximum a Posteriori estimation - cont'd

- MAP

$$\begin{aligned}\hat{w}_{MAP} &= \operatorname{argmax}_w P(w|\mathbf{D}) \\ &= \operatorname{argmax}_w \frac{P(w)P(\mathbf{D}|w)}{P(\mathbf{D})} \\ &\cong \operatorname{argmax}_w P(w)P(\mathbf{D}|w) \\ &= \operatorname{argmax}_w \prod_{i=1}^n P(\mathbf{d}_i|w) P(w) \\ &= \operatorname{argmax}_w \sum_{i=1}^n \log P(\mathbf{d}_i|w) + \log P(w)\end{aligned}$$

$$\begin{aligned}(cf: \hat{w}_{MLE} &= \operatorname{argmax}_w P(\mathbf{D}|w) \\ &= \operatorname{argmax}_w \prod_{i=1}^n P(\mathbf{d}_i|w))\end{aligned}$$

Maximum a Posteriori estimation - cont'd

For what w , is $\prod_{i=1}^n P(\mathbf{d}_i | w) P(w)$ maximized?

≡

For what w , is $\prod_{i=1}^n \exp\left(-\frac{1}{2}\left(\frac{y_i - wx_i}{\sigma}\right)^2\right) \exp\left(-\frac{1}{2}\left(\frac{w}{\gamma}\right)^2\right)$ maximized?

≡

For what w , is $\sum_{i=1}^n -\frac{1}{2}\left(\frac{y_i - wx_i}{\sigma}\right)^2 - \frac{1}{2}\left(\frac{w}{\gamma}\right)^2$ maximized?

≡

For what w , is $\sum_{i=1}^n (y_i - wx_i)^2 + \left(\frac{\sigma w}{\gamma}\right)^2$ minimized?

Second result

- MAP with a Gaussian prior on w is the same as minimizing the L_2 error plus an L_2 penalty on w

$$\operatorname{argmin}_w \sum_{i=1}^n (y_i - wx_i)^2 + \rho w^2$$

$$\rho = \frac{\sigma}{\gamma}$$

- MLE estimation of a parameter leads to unregularized solutions.
- MAP estimation of a parameter leads to regularized solutions.
- The prior distribution $P(w)$ acts as a regularizer in MAP estimation.

Multivariate regression

- What if the inputs are vectors?

Write matrix \mathbf{X} and \mathbf{y} :

$$\mathbf{X} = \begin{bmatrix} x_{11} & \cdots & x_{1p} \\ \vdots & \ddots & \vdots \\ x_{n1} & \cdots & x_{np} \end{bmatrix} \quad \mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}, \quad \mathbf{w} = \begin{bmatrix} w_1 \\ \vdots \\ w_p \end{bmatrix}$$

where $\mathbf{x}_1 = (x_{11}, \dots, x_{1p})$, $\mathbf{x}_2 = (x_{21}, \dots, x_{1p})$, \dots
 $\mathbf{x}_n = (x_{n1}, \dots, x_{np})$

- Assume that the data is formed by $y_i = \mathbf{w}^T \mathbf{x}_i + \text{noise}_i$

$$y \sim N(\mathbf{w}^T \mathbf{x}, \sigma^2)$$

Multivariate regression - cont'd

- Probability of each response variable

$$P(\mathbf{d}_i|\mathbf{w}) = P(y_i|\mathbf{w}, \mathbf{x}_i) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma}\right)^2\right).$$

- Given data $\mathbf{D} = \{\mathbf{d}_1(\mathbf{x}_1, y_1), \dots, \mathbf{d}_n(\mathbf{x}_n, y_n)\}$, we want to estimate the weight vector \mathbf{w} .

Likelihood:

$$\begin{aligned} L(\mathbf{w}) &= P(\mathbf{D}|\mathbf{w}) = P(y|\mathbf{w}, \mathbf{X}) = \prod_{i=1}^n P(\mathbf{d}_i|\mathbf{w}) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma}\right)^2\right) \end{aligned}$$

Log-likelihood:

$$\log L(\mathbf{w}) = \sum_{i=1}^n \left\{ -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2} \left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma} \right)^2 \right\}$$

Multivariate regression - cont'd

- Maximum likelihood solution:

$$\begin{aligned}\hat{\mathbf{w}}_{MLE} &= \operatorname{argmax}_{\mathbf{w}} \prod_{i=1}^n P(\mathbf{d}_i | \mathbf{w}) \\ &= \operatorname{argmax}_{\mathbf{w}} \sum_{i=1}^n -\frac{1}{2} \left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma} \right)^2 \\ &= \operatorname{argmin}_{\mathbf{w}} \sum_{i=1}^n \left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma} \right)^2 \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \\ &\quad \left(\text{from } \frac{d}{d\mathbf{w}} (\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w}) = \mathbf{0} \right)\end{aligned}$$

Multivariate regression - cont'd

- Maximum-a-Posteriori Solution:

- Assume a Gaussian prior distribution over the weight vector \mathbf{w} .

$$P(\mathbf{w}) \sim N(0, \lambda^{-1} \mathbf{I}) = \frac{1}{(2\pi)^{p/2}} \exp\left(-\frac{\lambda}{2} \mathbf{w}^T \mathbf{w}\right)$$

- Posteriori probability:

$$P(\mathbf{w}|\mathbf{D}) = \frac{P(\mathbf{w})P(\mathbf{D}|\mathbf{w})}{P(\mathbf{D})}$$

- Log Posteriori probability:

$$\log P(\mathbf{w}|\mathbf{D}) = \log \frac{P(\mathbf{w})P(\mathbf{D}|\mathbf{w})}{P(\mathbf{D})}$$

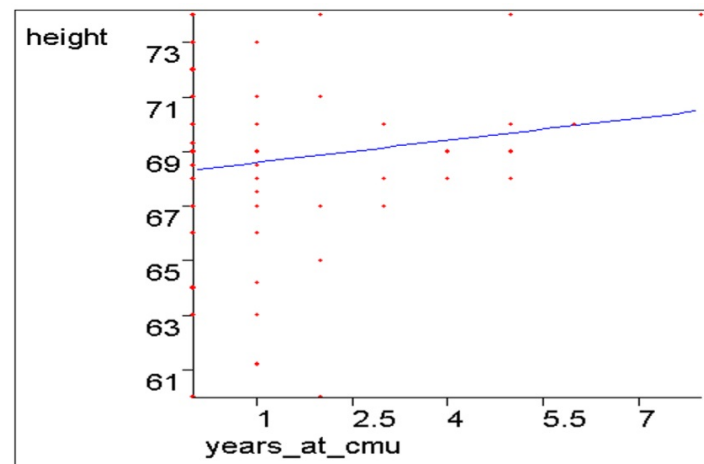
Multivariate regression - cont'd

- Maximum-a-Posteriori Solution:

$$\begin{aligned}\hat{\mathbf{w}}_{MAP} &= \operatorname{argmax}_{\mathbf{w}} \log P(\mathbf{w} | \mathbf{D}) \\ &= \operatorname{argmax}_{\mathbf{w}} \{ \log P(\mathbf{D} | \mathbf{w}) + \log P(\mathbf{w}) \} \\ &= \operatorname{argmax}_{\mathbf{w}} \{ \log P(\mathbf{w}) + \sum_{i=1}^n \log P(\mathbf{d}_i | \mathbf{w}) \} \\ &= \operatorname{argmax}_{\mathbf{w}} \left\{ -\frac{\lambda}{2} \mathbf{w}^T \mathbf{w} - \sum_{i=1}^n \frac{1}{2} \left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma} \right)^2 \right\} \\ &= \operatorname{argmin}_{\mathbf{w}} \left\{ \frac{\lambda}{2} \mathbf{w}^T \mathbf{w} + \sum_{i=1}^n \frac{1}{2} \left(\frac{y_i - \mathbf{w}^T \mathbf{x}_i}{\sigma} \right)^2 \right\} \\ &= \operatorname{argmin}_{\mathbf{w}} \left\{ \frac{\lambda}{2} \mathbf{w}^T \mathbf{w} + \frac{1}{2\sigma^2} (\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w}) \right\} \\ &= \left(\mathbf{X}^T \mathbf{X} + \frac{\sigma^2}{2} \lambda \mathbf{I} \right)^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$

Constant term in linear regression

- We may expect linear data that does not go through the origin.
- Statisticians and Neural Net Folks all agree on a simple obvious hack. Can you guess??



The constant term

- The trick is to create a fake input “ X_0 ” that always takes the value 1 .

X_1	X_2	Y
2	4	16
3	4	17
5	5	20

Before:

$$Y = w_1 X_1 + w_2 X_2$$

“ Poor model “

X_0	X_1	X_2	Y
1	2	4	16
1	3	4	17
1	5	5	20

After:

$$Y = w_0 X_0 + w_1 X_1 + w_2 X_2$$

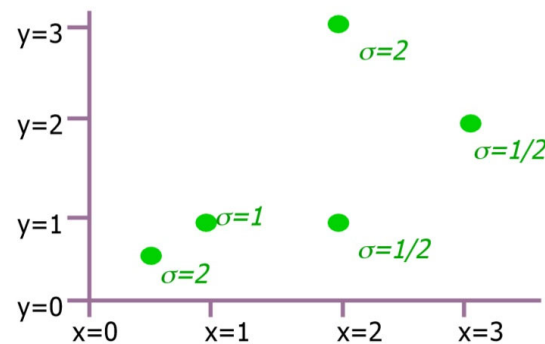
“ has a fine constant term “

you Should be able to see the
MLE w_0, w_1, w_2 by inspection.

Linear regression with varying noise

- Suppose you know the variance of the noise that was added to each data point.

x_i	y_i	σ_i^2
$1/2$	$1/2$	4
1	1	1
2	1	$1/4$
2	3	4
3	2	$1/4$



Assume $y_i \sim N(wx_i, \sigma_i^2)$
What is the MLE estimate of w ?

MLE estimation with varying noise

$$\operatorname{argmax}_w \log P(\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_n | w, \sigma_1^2, \dots, \sigma_n^2)$$

$$= \operatorname{argmin}_w \sum_{i=1}^n \frac{(y_i - wx_i)^2}{\sigma_i^2}$$

$$\rightarrow w = \frac{\sum_{i=1}^n \frac{x_i y_i}{\sigma_i^2}}{\sum_{i=1}^n \frac{x_i^2}{\sigma_i^2}}$$

Nonlinear regression

- Suppose you know that y is related to a function of x in such a way that the predicted values have a non-linear dependence on w , e.g. :

Assume $y_i \sim N(\sqrt{w} + x_i, \sigma^2)$

What is the MLE estimate of w ?

Nonlinear regression - cont'd

$$\operatorname{argmax}_w \log P(\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_n | w, \sigma_1^2, \dots, \sigma_n^2)$$

$$= \operatorname{argmin}_w \sum_{i=1}^n \frac{(y_i - \sqrt{w + x_i})^2}{\sigma_i^2}$$

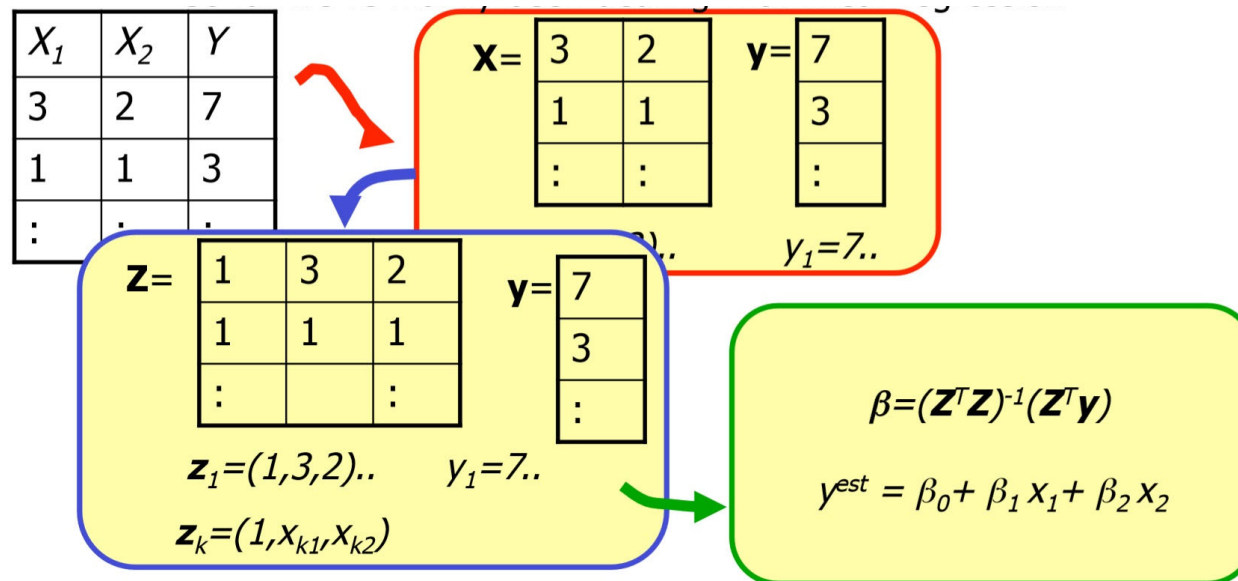
$$\rightarrow w \text{ such that } \sum_{i=1}^n \frac{y_i - \sqrt{w + x_i}}{\sigma_i^2 \sqrt{w + x_i}} = 0$$

Nonlinear regression - cont'd

- Common (but not only) approach:
- Numerical Solutions:
 - Line Search
 - Simulated Annealing
 - Gradient Descent
 - Conjugate Gradient
 - Levenberg Marquart
 - Newton's Method
 - Also, special purpose statistical-optimization-specific tricks such as E.M.

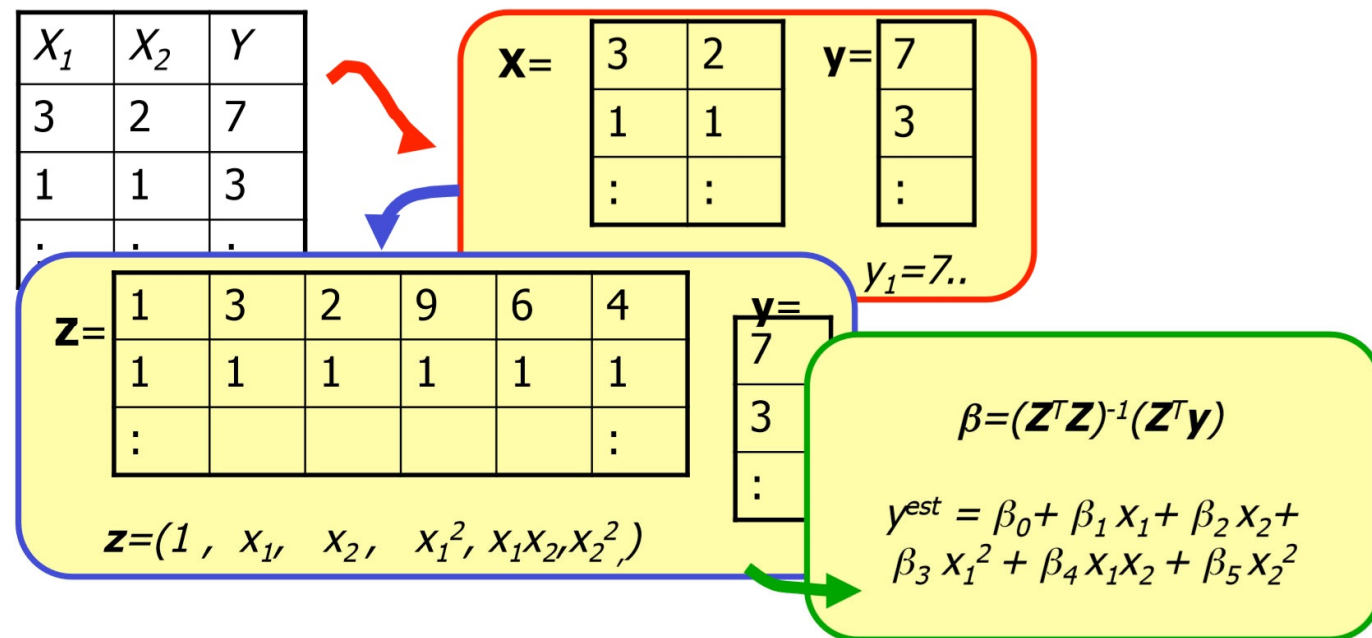
Polynomial regression

- So far we've mainly been dealing with linear regression



Quadratic regression

- It's trivial to do linear fits of fixed nonlinear basis functions.



Quadratic regression - cont'd

Each component of a \mathbf{z} vector is called a term.

Each column of the \mathbf{Z} matrix is called a term column

How many terms in a quadratic regression with m inputs?

- 1 constant term
- m linear terms
- $(m+1)\text{-choose-}2 = m(m+1)/2$ quadratic terms

$(m+2)\text{-choose-}2$ terms in total = $O(m^2)$

Note that solving $\beta = (\mathbf{Z}^T \mathbf{Z})^{-1} (\mathbf{Z}^T \mathbf{y})$ is thus $O(m^6)$

Q^{th} -degree polynomial regression

