

Assignment 9

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April 28, 2019

1 Problem 9.1

1.1 Part a: Loading and plotting data

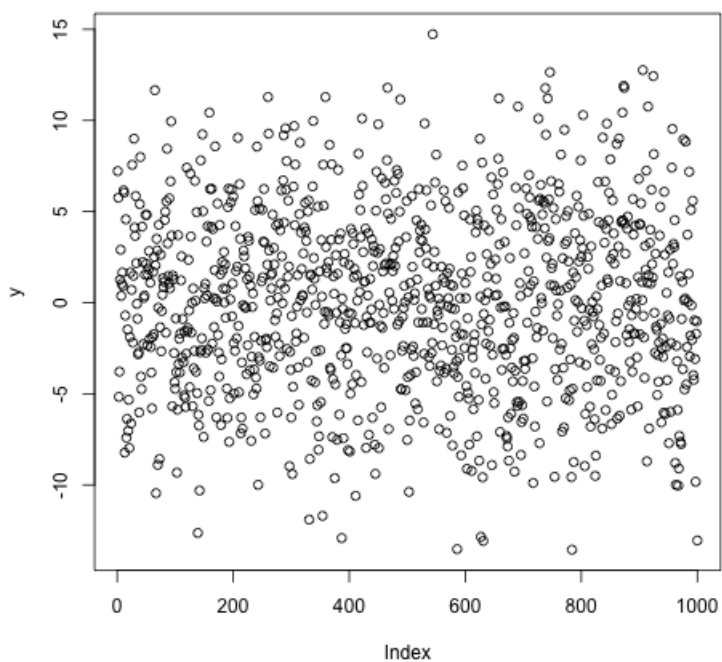


Figure 1: Simulated $MA(q)$ process with gaussian noise

1.2 Part b: simulate MA(k)

The AIC (Akaike's Information Criteria) and BIC (Bayesian Information Criteria) are penalised-likelihood information criteria [1, 2] used for model selection. A model with too few parameters can have a large bias while a model with too many parameters can promote overfitting by not generalising enough. Furthermore, the mean square error of the forecasts model depends on the white noise variance of the fitted model and on the error arising from the estimation of the parameters of the model [1]. Therefore the mean square error of the forecasts model for a model with many parameters, will be large [2]. To mitigate this error a penalty factor is introduced to discourage the fitting of a model with many parameters [2], giving rise to AIC and BIC which are unified log-likelihood function with penalties [1].

The AIC is an estimation of the difference between the unknown true likelihood function of the data and the fitted likelihood function of the model [1]. Therefore the best model will have a lower AIC [1]. The BIC is an estimate of a function of the posterior probability of a model being true, under a certain Bayesian setup, [1]. A lower BIC means the fitted model is likely to be the true model [1]. The optimum model is the one with order q that minimises the AIC and BIC respectively.

Now we simulate the data for k between 1 and 20

```
maq <- read.delim("ma_q.txt")
png("../note/maq.png")
y <- maq$x
K <- seq(1,20)
AIC = c()
BIC = c()

for (k in K)
{
  print(k)
  model <- arima(x=y, order=c(0,0,k), method="ML")
  aic <- AIC(model)
  bic <- AIC(model,k = log(length(y)))
  AIC = append(AIC, aic)
  BIC = append(BIC, bic)
}

min_bic = min(BIC)
index_min_bic = which(BIC == min_bic)
#max_bic = max(BIC)

min_aic = min(AIC)
index_min_aic = which(AIC == min_aic)
#max_aic = max(AIC)
```

From the simulation we find that for a model with order 5 both the AIC and BIC are minimised, therefore a model with order 5 is the optimum model.

1.3 Part c

Comparing the sample ACF and the PACF with the one of the from model selection using BIC and AIC.

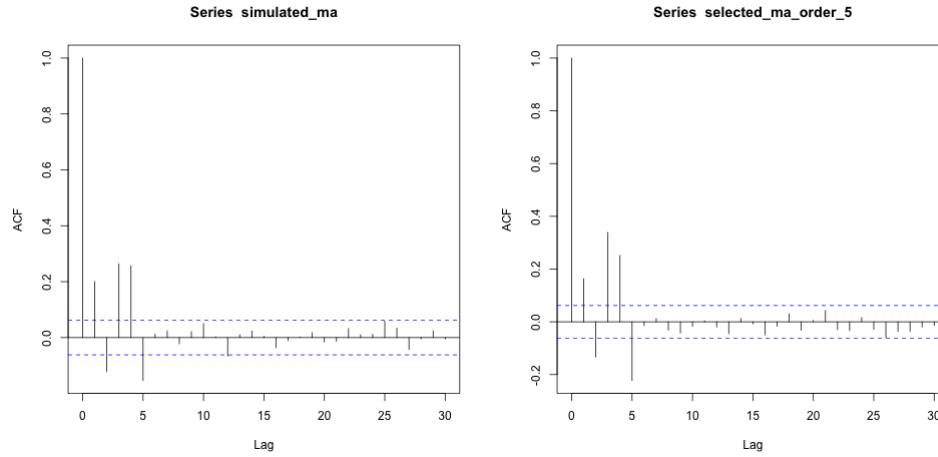


Figure 2: Comparison of ACF of simulated MA(q) (left) and selected MA(5) (right)

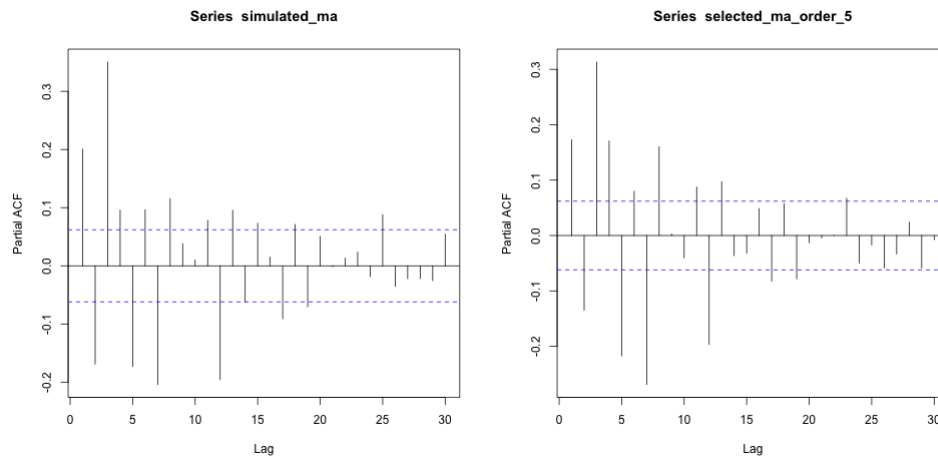


Figure 3: Comparison of PACF of simulated MA(q)(left) and selected MA(5) (right)

Figure 2 and 3 shows the ACF and the PACF of the simulated MA(q) and a MA(5) where $q = 5$ is selected from AIC and BIC penalty criteria. We can observe that the simulated MA(q) is some what consistent with the selected MA(5) model.

1.4 Part d

We simulate an ARMA(2,2) of length $n = 50$, with

$$\phi = (\phi_1 = -0.7, \phi_2 = 0.2) \quad (1)$$

$$\theta = (\theta_1 = 0.3, \theta_2 = -0.2) \quad (2)$$

with

$$N(0, 4^2)$$

The process is simulated in R by

```
arma2_2 <- arima.sim(n=50, list(ar=c(-0.7, 0.2), ma=c(0.3, -0.2)), sd=4)
```

which generate the following figure

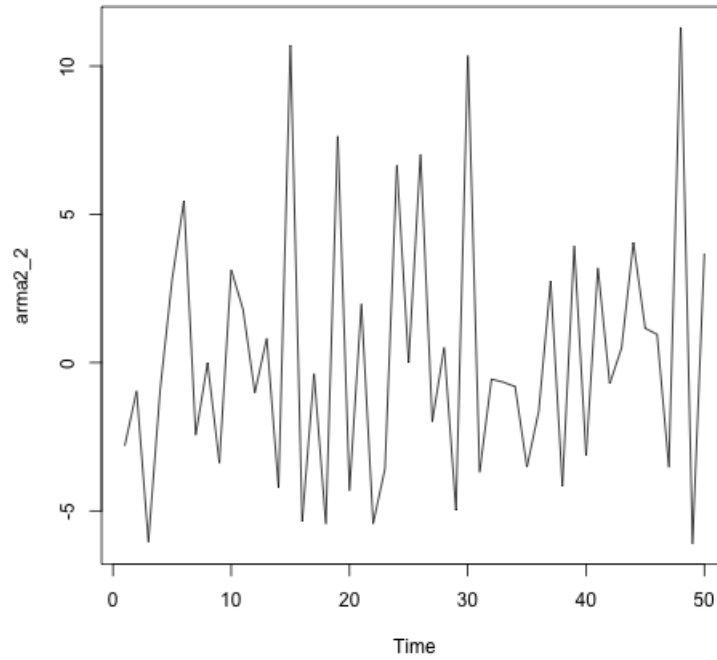


Figure 4: Simulated ARMA(2,2) with coefficients given by equation (1) and (2), with standard deviation 4

1.5 Part e

Now we fit an ARMA(2,2) and a MA(5) from the simulated data from part d and use BIC and AIC criteria to select the optimum model.

Fit an ARMA(2,2) to the model

```
arma2_2 <- arima.sim(n=50,list(ar=c(-0.7, 0.2),ma=c(0.3, -0.2)),sd=4)
fitted_ARMA_2_2_model <- arima(x=arma2_2,order=c(2,0,2), method="ML")
aic <- AIC(fitted_ARMA_2_2_model)
bic <- AIC(fitted_ARMA_2_2_model,k=log(length(fitted_ARMA_2_2_model)))
print(aic)
print(bic)

>295.4357
>299.2701
```

Fit a MA(5) process

```
arma2_2 <- arima.sim(n=50,list(ar=c(-0.7, 0.2),ma=c(0.3, -0.2)),sd=4)

fitted_MA5_model <- arima(x=arma2_2, order=c(0,0,5), method="ML")
aic <- AIC(fitted_MA5_model)
bic <- AIC(fitted_MA5_model,k=log(length(fitted_MA5_model)))
print(aic)
print(bic)

>286.4791
>290.9525
```

Since the BIC and AIC for the MA(5) process are the smallest we pick the MA(5) model

1.6 Part f

Now we repeat part e by increasing the sample size 500 times

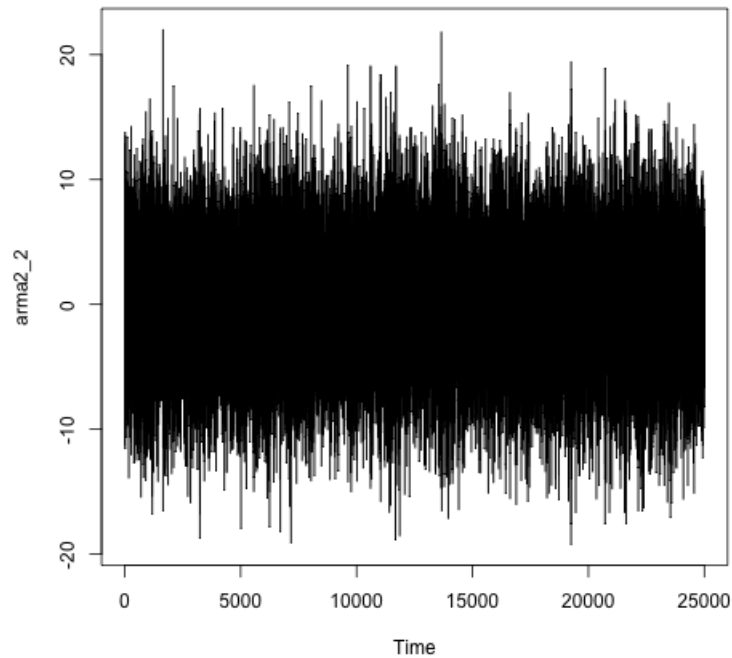


Figure 5: Simulated ARMA(2,2) with sample size 50×500

Fit an ARMA(2,2)

```
arma2_2 <- arima.sim(n=50,list(ar=c(-0.7, 0.2),ma=c(0.3, -0.2)),sd=4)
fitted_ARMA_2_2_model <- arima(x=arma2_2,order=c(2,0,2), method="ML")
aic <- AIC(fitted_ARMA_2_2_model)
bic <- AIC(fitted_ARMA_2_2_model,k=log(length(fitted_ARMA_2_2_model)))
print(aic)
print(bic)

>142228.2
>142232.7
```

Fit a MA(5) process

```
arma2_2 <- arima.sim(n=50,list(ar=c(-0.7, 0.2),ma=c(0.3, -0.2)),sd=4)
fitted_MA5_model <- arima(x=arma2_2, order=c(0,0,5), method="ML")
aic <- AIC(fitted_MA5_model)
bic <- AIC(fitted_MA5_model,k=log(length(fitted_MA5_model)))
```

```
print(aic)
print(bic)

>142130
>142134.5
```

Again the MA(5) process comes with the lowest BIC and AIC and is the preferred model.

2 Problem 9.2: Plotting basis function

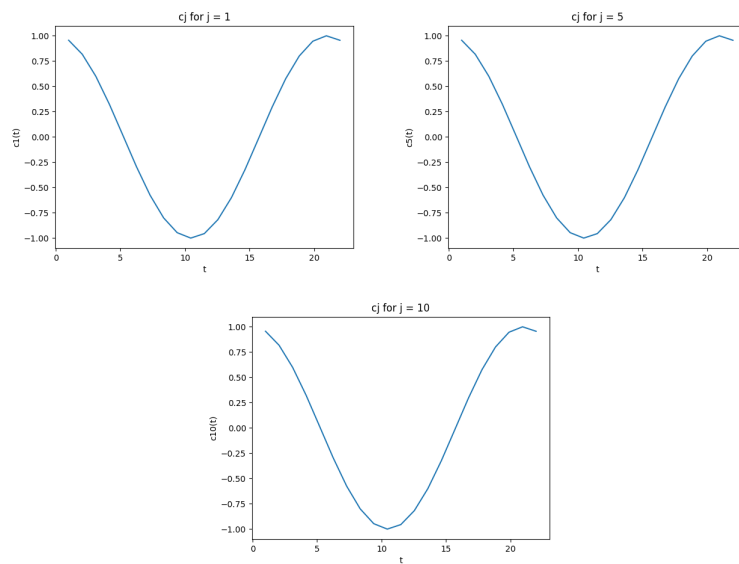


Figure 6: Plot of $c_1(t)$, $c_5(t)$, $c_{10}(t)$ on $t = [1, 21]$

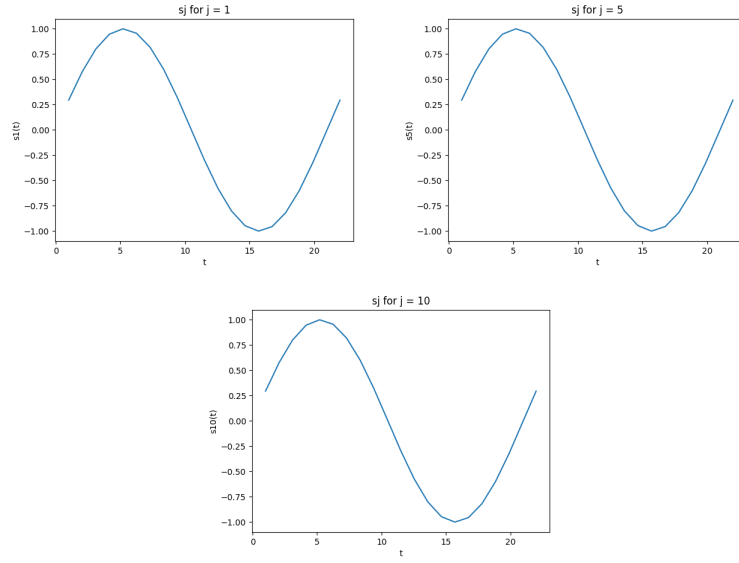


Figure 7: Plot of $s_1(t)$, $s_5(t)$, $s_{10}(t)$ on $t = [1, 21]$

3 Problem 9.3: White noise

A white noise is a sequence of uncorrelated random variable $\{X_t\}$, each with zero mean and finite variance σ^2 [2]. The white noise $\{X_t\}$ is a stationary process with mean zero and covariance function given by [2]:

$$\gamma_X(t+h, t) = \begin{cases} \sigma^2, & \text{if } h = 0 \\ 0, & \text{if } h \neq 0 \end{cases} \quad (3)$$

If $\{X_t\}$ is a white noise we write

$$\{X_t\} \sim \mathbf{WN}(0, \sigma^2). \quad (4)$$

4 Problem 9.4

4.1 Part a: simulation of an AR(2) process

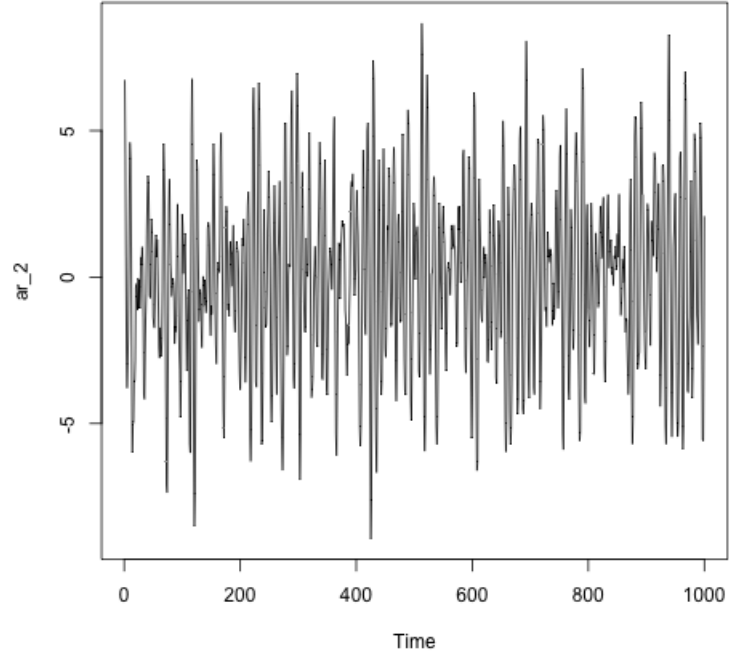


Figure 8: Simulation of the AR(2) process with $n = 1000$, $\phi = (1.4, -0.8)$ and $\sigma^2 = 1$

4.2 Part b

Let compute the matrix G given by

$$G = \begin{bmatrix} | & & | \\ e_{-\lceil \frac{n-1}{2} \rceil} & \cdots & e_{\lceil \frac{n}{2} \rceil} \\ | & & | \end{bmatrix}, \quad (5)$$

where

$$e_k = \frac{1}{\sqrt{n}} \begin{bmatrix} e^{i\frac{2\pi k}{n}} \\ e^{i\frac{4\pi k}{n}} \\ \vdots \\ e^{i2\pi k} \end{bmatrix}, \quad (6)$$

So that

$$\begin{aligned} G &= \frac{1}{\sqrt{1000}} \begin{bmatrix} | & & | \\ e^{-[\frac{n-1}{2}]} & \cdots & e^{[\frac{n}{2}]} \\ | & & | \end{bmatrix} \\ &= \frac{1}{\sqrt{1000}} \begin{bmatrix} e^{-i\frac{2\pi \times 499}{1000}} & \cdots & e^{i\pi} \\ e^{-i\frac{4\pi \times 499}{1000}} & \cdots & e^{i2\pi} \\ \vdots & \cdots & \vdots \\ e^{-i499} & \cdots & e^{i5.10^5} \end{bmatrix} \end{aligned} \quad (7)$$

And a_k is given by

$$a_k = \frac{1}{\sqrt{1000}} \sum_{t=1}^{1000} x_t e^{-it\frac{2\pi}{n}k}, \quad k = -499, \dots, 500 \quad (8)$$

5 Problem 9.5

We have

$$g(\lambda) = K|a(e^{-i\lambda})|^2, \quad K > 0, \quad (9)$$

where

$$a(z) = 1 + \sum_{i=1}^q a_i z^i \quad (10)$$

and g is strictly positive in $[-\pi, \pi]$

5.1 Part a

Suppose that the polynomial a has root on the unit circle. Then $z = \pm 1, \pm i$ are roots, meaning that

$$a(1) = 1 + \sum_{i=1}^q a_i = 0 \Rightarrow \sum_{i=1}^q a_i = -1 \quad (11)$$

$$a(-1) = 1 + \sum_{i=1}^q a_i (-1)^i = 0 \Rightarrow \sum_{i=1}^q a_i (-1)^i = -1 \quad (12)$$

$$a(i) = 1 + \sum_{i=1}^q a_i i^p = 0 \Rightarrow \sum_{i=1}^q a_i i^q = -1 \quad (13)$$

$$a(-i) = 1 + \sum_{i=1}^q a_i (-i)^p = 0 \Rightarrow \sum_{i=1}^q a_i (-i)^p = -1 \quad (14)$$

So we have

$$\begin{aligned} \sum_{i=1}^q a_i &= \sum_{i=1}^q a_i i^i \\ \sum_{i=1}^q a_i - \sum_{i=1}^q a_i i^i &= 0 \\ \sum_{i=1}^q a_i (1 - i^q) &= 0 \\ a_1(1 - i) + \dots + a_q(1 - i^q) &= 0 \end{aligned} \quad (15)$$

So that every coefficient with index of the sort $q = 2^r$ where r is even will be zeros. If we repeat the same with they other equations, we see that all coefficient $a_i = 0$ for $i = 1, \dots, p$. This mean that

$$a(z) = 1 \quad (16)$$

But if this is the case our assumption that a has root on the unit circle is not true, therefore a can not take root at the unit circle

5.2 Part b

By definition of invertible process, the polynomial $a(B)$ does not have roots on the unit circle. We have shown that in part a.

References

- [1] John J. Dziak, Donna L. Coffman, Stephanie T. Lanza, Runze Li *Sensitivity and specificity of information criteria*. Addison-Wesley, Reading, Massachusetts, 1993.
- [2] Petter J. Brockewell. Richard A. Davis *Introduction to time series and forecasting*. Springer Text In Statistics, second edition