


3.

$$\frac{E_{\text{out}}(h)}{E} = \frac{E_{\text{out}}(h)}{\sum_{t=1}^M E_{\text{out}}(g_t)} = \frac{\frac{1}{N} E_{\text{out}}(h)}{\frac{1}{N} \sum_{t=1}^M E_{\text{out}}(g_t)}$$

$$= \frac{P[h(x_n) \neq y_n]}{\sum_{n=1}^M P[g(x_n) \neq y_n]} = \frac{P\left[\sum_{t=1}^M ([g_t(x_n) \neq y_n]) \geq 1\right]}{\sum_{n=1}^M P[g(x_n) \neq y_n]}$$

根據 Markov's inequality: $P(x \geq \alpha) \leq \frac{E(x)}{\alpha}$

$$\leq \frac{1}{1} \frac{E\left[\sum_{t=1}^M ([g_t(x_n) \neq y_n])\right]}{\sum_{t=1}^M P[g(x_n) \neq y_n]}$$

$$= \frac{\frac{1}{N} \sum_{n=1}^N \sum_{t=1}^M ([g_t(x_n) \neq y_n])}{\sum_{t=1}^M P[g(x_n) \neq y_n]}$$

$$= \frac{1}{\sum_{t=1}^m P[g(x_n) \neq y_n]} \quad |$$

$$\therefore \frac{E_{out}(h)}{E} \leq \frac{1}{q}, \quad \text{即 } \leq \frac{1}{q}$$

4.

抽一次一个 sample 没被抽到的概率

$$\left(1 - \frac{1}{N}\right)$$

抽 $\frac{3}{4}N$ 次一个 sample 没被抽到的概率

$$\left(1 - \frac{1}{N}\right)^{\frac{3}{4}N}$$

$$= \frac{1}{\left(\frac{N}{N-1}\right)^{\frac{3}{4}N}} = \left(\frac{1}{\left(1 + \frac{1}{N-1}\right)^N}\right)^{\frac{3}{4}}$$

$$\approx \left(\frac{1}{e}\right)^{\frac{3}{4}}$$

5.

obtain iteration 1 (1st weight) 需要

需要 $\sum_{n=1}^N y_n$ 和 $\sum_{n=1}^N u_n^{(1)}$

$$E_1 = \frac{\sum_{n=1}^N M_n^{(1)} [(y_n + 9.1x_n)]}{\sum_{n=1}^N M_n^{(1)}} = 2\%$$

$$\diamond_1 = \sqrt{\frac{48\%}{2\%}} = \sqrt{49\%} = 70\%$$

$$u_n^{(2)} = u_n^{(1)} \times 70\%, \text{ if } y_n > 0$$

$$u_n^{(2)} = u_n^{(1)} / 70\%, \text{ if } y_n < 0$$

$$\frac{\sum_{n: y_n > 0} M_n^{(2)}}{\sum_{n: u_n < 0} M_n^{(2)}} = \frac{\cancel{48\%} \times \cancel{70\%} + \cancel{2\%} \times M_n^{(1)}}{\cancel{2\%} \div \cancel{70\%} \times M_n^{(1)}} = 24.0$$

6. 使用 Jensen 不等式证明 $\frac{V_{t+1}}{V_t} = 2\sqrt{\epsilon_t(1-\epsilon_t)}$

consider $t=1$, $V_1 = 1$,

$$V_2 = \sum_{n=1}^N \left[\epsilon_1 * \frac{1}{N} + (1-\epsilon_1) * \frac{1}{N} \right]$$

$$= \epsilon_1 * \sqrt{\frac{1-\epsilon_1}{\epsilon_1}} + (1-\epsilon_1) * \sqrt{\frac{\epsilon_1}{1-\epsilon_1}}$$

$$= 2\sqrt{\epsilon_1(1-\epsilon_1)}, \quad \frac{V_2}{V_1} = 2\sqrt{\epsilon_1(1-\epsilon_1)} \text{ 成立}$$

consider $t=N$, 对于的随机变量分别以 ϵ_N 及 $1-\epsilon_N$ 的概率

分布 $\epsilon_N \neq \frac{1}{N}$ 及 $\frac{1}{N}$

$$V_{N+2} = V_{N+1} \left[\epsilon_{N+1} \frac{1}{N+1} + (1-\epsilon_{N+1}) \frac{1}{N+1} \right]$$

$$= V_{N+1} \left[\epsilon_{N+1} \sqrt{\frac{1-\epsilon_{N+1}}{\epsilon_{N+1}}} + (1-\epsilon_{N+1}) \sqrt{\frac{\epsilon_{N+1}}{1-\epsilon_{N+1}}} \right]$$

$$= V_{N+1} \sqrt{\varepsilon_{N+1} (1 - \varepsilon_{N+1})}$$

$$\frac{V_{N+2}}{V_{N+1}} = 2 \sqrt{\varepsilon_{N+1} (1 - \varepsilon_{N+1})} \quad \text{得 2 个 } \varepsilon$$

$$\therefore \frac{V_{T+1}}{V_1} = \frac{V_2}{V_1} \cdots \frac{V_T}{V_{T-1}} \frac{V_{T+1}}{V_T}$$

$$= \prod_{t=1}^T 2 \sqrt{\varepsilon_t (1 - \varepsilon_t)}$$

$$8. \quad s_i^{(l-1)} \xRightarrow{\tanh} x_i^{(l-1)} \xRightarrow{W_{ij}} s_j^{(l)} \xRightarrow{\tanh}$$

$$\delta_j^{(l)} = \frac{\partial e_n}{\partial s_j^{(l)}} = \sum_{k=1}^{d^{(l+1)}} \frac{\partial e_n}{\partial s_k^{(l+1)}} \frac{\partial s_k^{(l+1)}}{\partial x_j^{(l)}} \frac{\partial x_j^{(l)}}{\partial s_j^{(l)}}$$

$$= \sum_k \delta_k^{(l+1)} w_{jk}^{(l+1)} \tanh'(s_j^{(l)})^{(l)}$$

$$\delta_j^{(l-1)} = \sum_k \delta_k^{(l)} w_{jk}^{(l)} \tanh'(s_j^{(l-1)})^{(l-1)}$$

∵ 當 s 為 0 時 \tanh' 趨向無限大

∴ 2~L 層的 gradient 皆非 0

而第一層的 gradient 為

$$\delta_j^{(1)} = \sum_k \delta_k^{(2)} w_{jk}^{(2)} = 0$$

∴ 除了第一層 gradient 為 0, 其他層皆非 0.

7.

we know $s_n^{t+1} = s_n^t + \alpha_t g_t(x_n)$

$$\alpha_t = \operatorname{argmin} \frac{1}{N} \sum_{n=1}^N \left((y_n - s_n^t) - \eta g_t(x_n) \right)^2 - \textcircled{D}$$

$$\frac{\partial \textcircled{D}}{\partial \eta} = 0$$

$$= -\frac{1}{N} \sum_{n=1}^N \left(2(y_n - s_n^t) g_t(x_n) - 2\eta g_t(x_n)^2 \right)$$

$$= -\frac{1}{N} \sum_{n=1}^N 2 \left(y_n - \underbrace{s_n^t - \eta g_t(x_n)}_{= s_n^t - \alpha_t g_t(x_n)} \right) g_t(x_n)$$

$$= -\frac{1}{N} \sum_{n=1}^N 2 \left(y_n - s_n^{t+1} \right) g_t(x_n)$$

$$1. \quad K_{ds}(x, x') = (\phi_{ds}(x))^T \phi_{ds}(x')$$

$$= \sum_{s, i, \theta} g_{s, i, \theta}(x) g_{s, i, \theta}(x')$$

而 x 與 x' 有 2 種情況 ① $x = x'$ ② $x \neq x'$

case ① \Rightarrow 因為每個 decision stump 的結果相同,

$$g_{s, i, \theta}(x) g_{s, i, \theta}(x') = 1 \quad \text{對每個}$$

$$2 \times d \times (R-L)$$

case ② \Rightarrow 若值在 x_i 與 x'_i 間所組成的區間不同, 此時

$$\begin{cases} g_{+, i, \theta}(x) g_{+, i, \theta}(x') = -1 \\ g_{-, i, \theta}(x) g_{-, i, \theta}(x') = -1 \end{cases}$$

$$\text{因此 case ② 的 } \sum_{i=1}^d |x_i - x'_i| = 2 \|x - x'\|_1 \quad \therefore$$

$$\therefore \text{total } \frac{1}{2} 2 \times d \times (R-L) - 2 \times 2 \|X - X'\|$$

$$\Rightarrow K_{as}(X, X') = 2 \times d \times (R-L) - 4 \|X - X'\|$$

2. 找到 (α, b) 變 $(\tilde{\alpha}, \tilde{\gamma})$ 的關係式 以證明
新的 $\tilde{\alpha}, \tilde{\gamma}$ 與原問題等價

assume $\tilde{\alpha}_n = \frac{\alpha_n}{n}$

$$\min_{\alpha} \frac{1}{2} \sum_n^N \sum_m^N \tilde{\alpha}_n \tilde{\alpha}_m \gamma_n \gamma_m \tilde{K}(x_n, x_m) - \sum_n^N \tilde{\alpha}_n$$

$$\begin{aligned} &= \min_{\alpha} \frac{1}{2} \sum_n^N \sum_m^N \frac{\alpha_n}{n} \frac{\alpha_m}{m} \gamma_n \gamma_m (n K(x_n, x_m) + V) \\ &\quad - \sum_n^N \frac{\alpha_n}{n} \end{aligned}$$

$$= \min_{\alpha} \frac{1}{2} \sum_n^N \sum_m^N \alpha_n \alpha_m \gamma_n \gamma_m \left(K(x_n, x_m) + \frac{V}{n} \right) - \sum_n^N \alpha_n$$

$$= \min_{\alpha} \frac{1}{2} \sum_n^N \sum_m^N \alpha_n \alpha_m \gamma_n \gamma_m K(x_n, x_m) - \sum_n^N \alpha_n$$

$$\tilde{g}_{svm}(x) = \text{sign} \left(\sum_{sv} \tilde{\alpha}_n \gamma_n \tilde{k}(x_n, x) + \tilde{b} \right)$$

$$= \text{sign} \left(\sum_{sv} \frac{\alpha_n}{\mu} \gamma_n (\mu k(x_n, x) + v) + \tilde{b} \right)$$

$$= \text{sign} \left(\sum_{sv} \alpha_n \gamma_n \left(k(x_n, x) + \frac{v}{\mu} \right) + \tilde{b} \right)$$

$$= \text{sign} \left(\sum_{sv} \alpha_n \gamma_n k(x_n, x) + b \right)$$

$$\therefore \begin{cases} \tilde{\alpha}_n = \frac{\alpha_n}{\mu} \\ b = \tilde{b} + \alpha_n \gamma_n \frac{v}{\mu} \Rightarrow \tilde{b} = b - \alpha_n \gamma_n \frac{v}{\mu} \end{cases}$$

以上關係式說明 \tilde{k}, \tilde{b} 與原問題

的關係。