

Causal Inference

Week 5 Randomized Experiments and Identification

<https://www.youtube.com/watch?v=z91LnTDyhtI&list=PLoazKTcS0Rzb6bb9L508cyJ1z-U9iWkA0&index=37>

전은주

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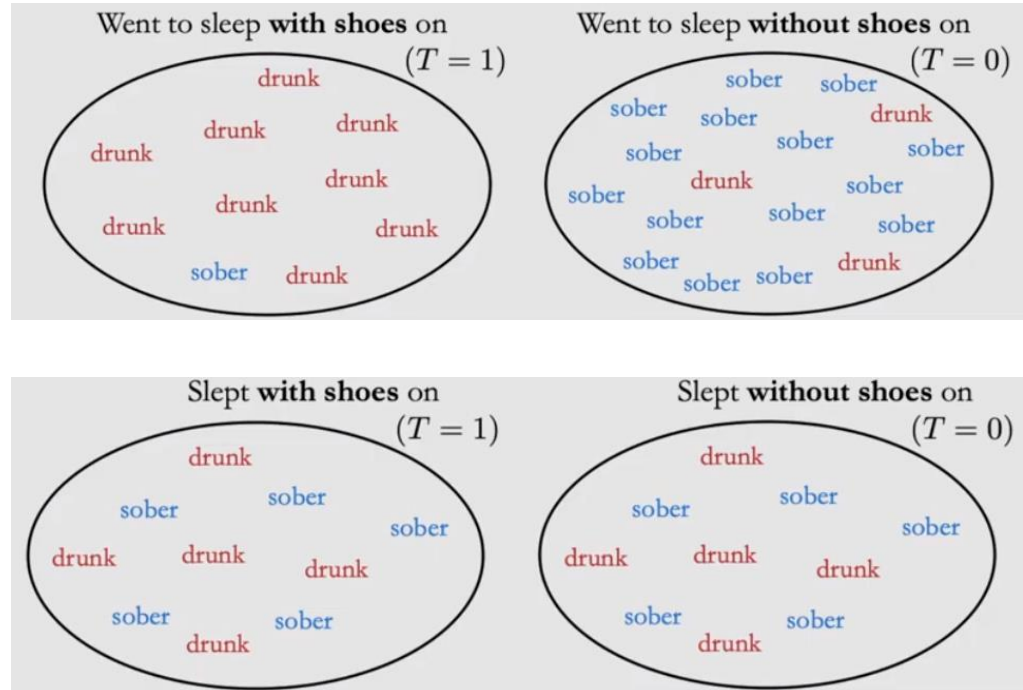
이전시간까지 Identification하는 방법으로
Backdoor adjustment를 활용하는 것을 배웠다.
이번 시간에는 그 외의 Causal inference를 계산
할 수 있는 방법들에 대해 배운다.

- 5.1 Randomized Experiments and Identification (Intro and Outline)
- 5.2 The Magic of Randomized Experiments
- 5.3 The Frontdoor Adjustment
- 5.4 Pearl's do-calculus
- 5.5 Determining Identifiability from the Graph

5.2 The Magic of Randomized Experiments

Confound factor나 Backdoor criteria에 만족하는지 정확히 알 수 없을 때,
RCT(randomize control trial)를 적용하면

No unobserved confounding



관찰만 했을 때: Drinking \rightarrow confounder



Coin 던져서 (random) $T=1$, $T=0$ 를 선택해서 배분
하면, drunk 비율은 두 그룹에서 같아진다

1) Comparability and covariate balance: intuition

Treatment외의 모든 변수에 대한 covariate가 Treatment그룹과 control그룹 같다.
두 그룹간 차이가 있다면 'Treatment'때문이다 (Causation)

Covariate balance definition

We have covariate balance if the distribution of covariates X is the same across treatment groups. More formally,

X : Treatment외의 x 변수들, d = dimension

$$P(X | T = 1) \stackrel{d}{=} P(X | T = 0)$$

Randomization implies covariate balance

Because T is not at all determined by X (solely by a coin flip), $T \perp\!\!\!\perp X$

T 는 X 에 의해서 선택되는 것이 아니므로, T 와 X 는 independent

$$P(X | T = 1) \stackrel{d}{=} P(X)$$

$$P(X | T = 0) \stackrel{d}{=} P(X)$$

따라서, T 가 conditioning되어도 $P(X)$ 와 같다.
결국 $P(X|T=1)$ 과 $P(X|T=0)$ 은 $P(X)$ 로 서로 같다.

$$P(X | T = 1) \stackrel{d}{=} P(X | T = 0)$$

Covariate balance implies association is causation

RCT일 때, T와 X가 독립적이고 모든 변수x가 covariate same

$$P(X | T = 1) \stackrel{d}{=} P(X | T = 0) \implies T \perp\!\!\!\perp X \quad \text{Let X be a sufficient adjustment set}$$

X가 sufficient adjustment set이면 backdoor criteria 사용

분자, 분모에 $P(t|x)$ 둘 다 곱하고
Chain rule에 의해서 조건부 확률을 joint distribution로 변환
 $P(y|x)P(t|x)P(x) \rightarrow P(y,t,x)$

T와 x는 서로 independent하므로 $P(t|x) \rightarrow P(t)$

BayesRule: $P(y,t,x)/P(t) \rightarrow P(y,x|t)$

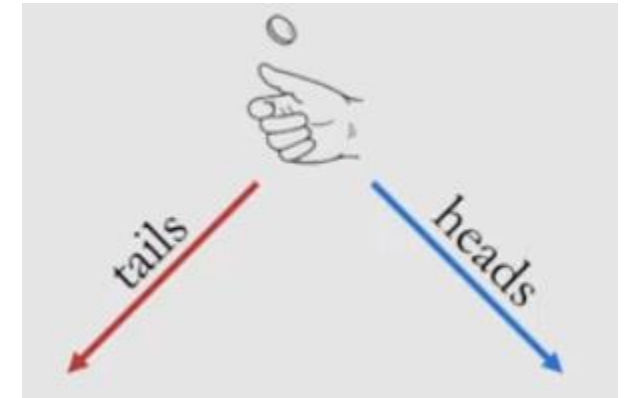
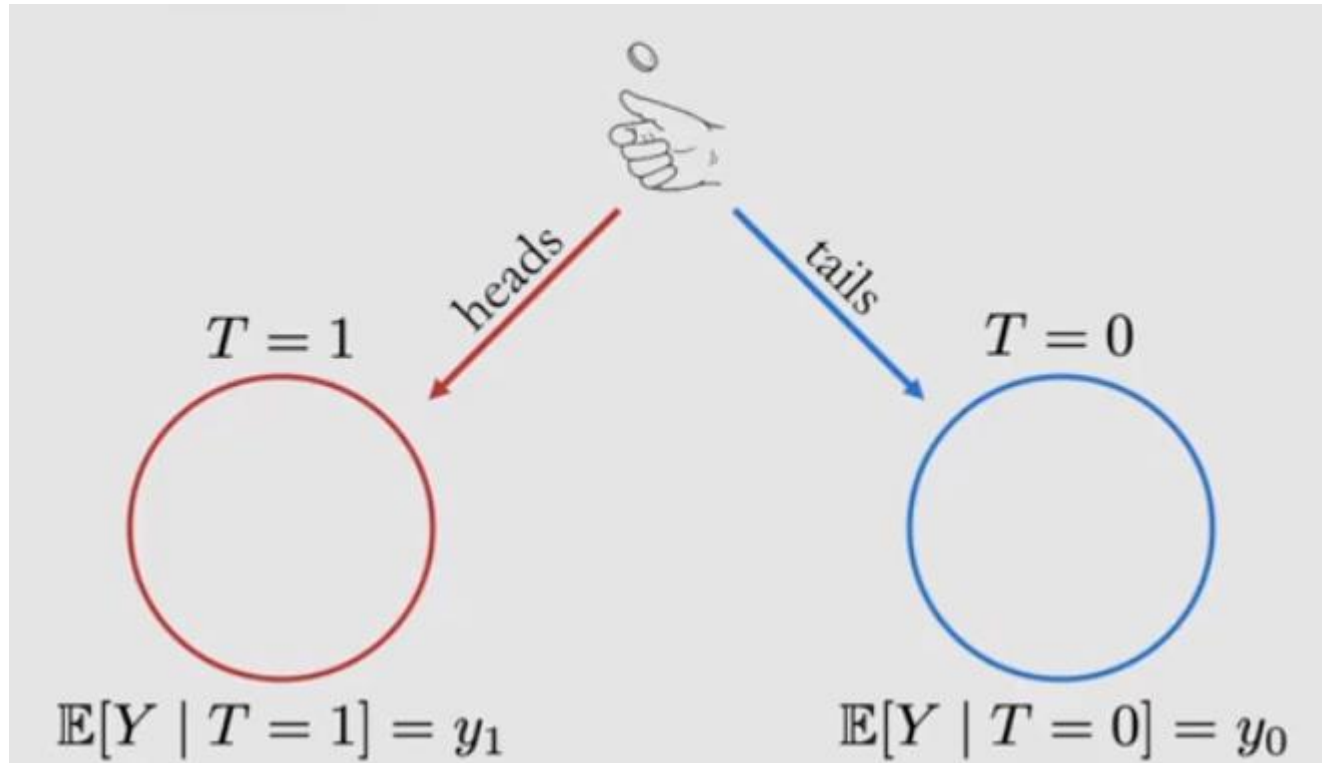
마지막으로 Sum x이므로 $P(y|t)$ 로 marginalize

$$\begin{aligned} P(y | do(t)) &= \sum_x P(y | t, x)P(x) \\ &= \sum_x \frac{P(y | t, x)P(t | x)P(x)}{P(t | x)} \\ &= \sum_x \frac{P(y, t, x)}{P(t | x)} \\ &= \sum_x \frac{P(y, t, x)}{P(t)} \\ &= \sum_x P(y, x | t) \\ &= P(y | t) \end{aligned}$$

따라서, RCT하에서 $P(y|do(t)) = P(y|t)$ causation 과 같다

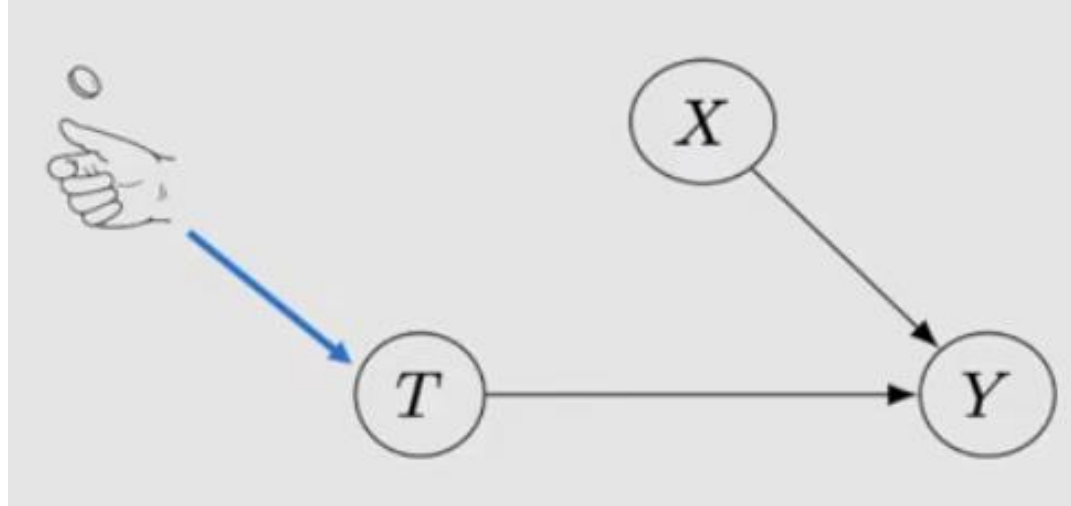
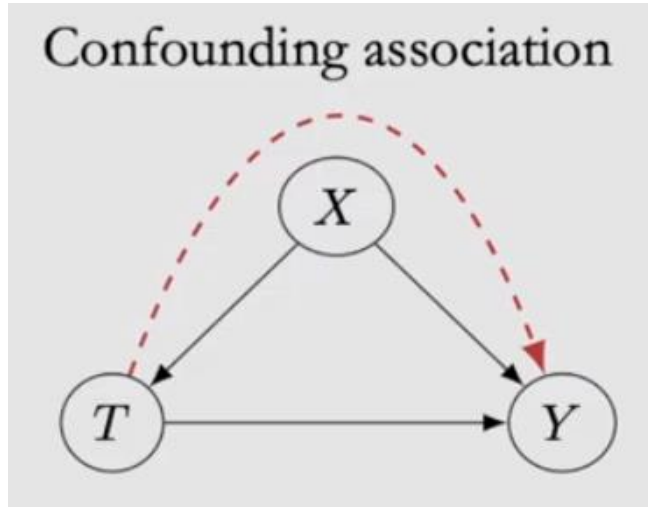
2) Exchangeability

Coin toss에 의해서 $T=1$ 과 $T=0$ 을 바꿔도 두 $T=0, T=1$ 간의 차이 (Average Outcome) 은 같다.
 T 에 의한 결과일 뿐 다른 영향은 없다



3) No backdoor paths

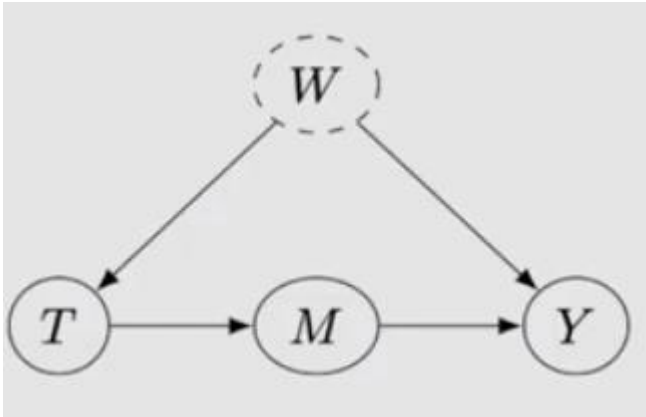
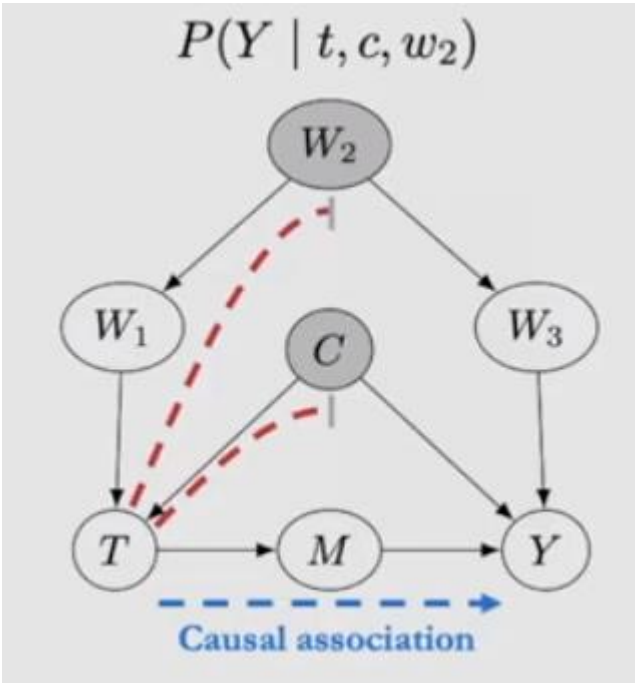
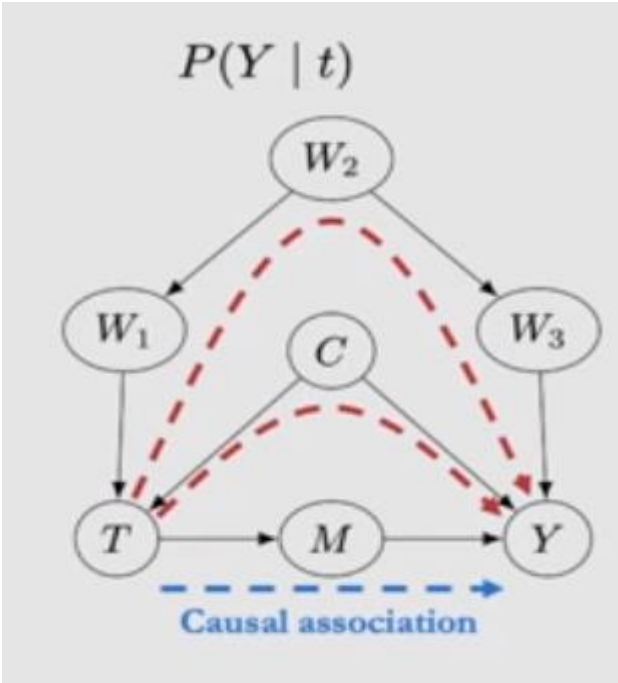
T가 coin toss에만 영향을 받기 때문에 (X와 T는 독립)
관찰 못한 X (교란변수, confounder)가 측정되더라도 T→Y에 영향을 미치지 않는다



5.3 Frontdoor adjustment

Backdoor adjustment

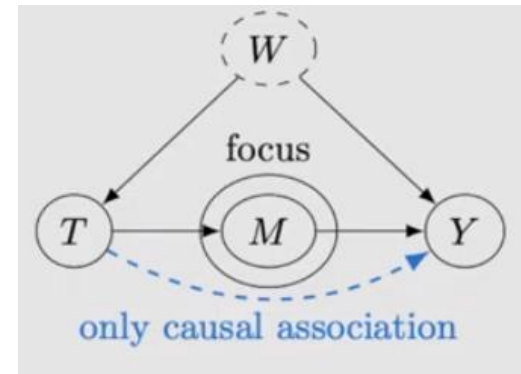
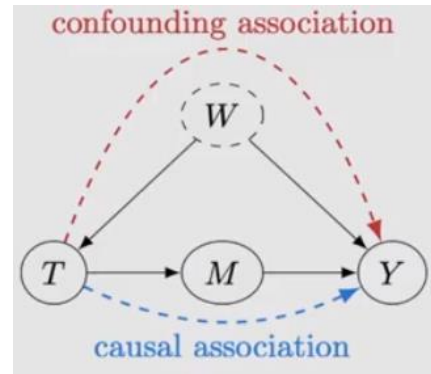
다음과 같이 non-causal association이 흐를 때, W_2, C 에 conditioning하여 흐름을 막는다



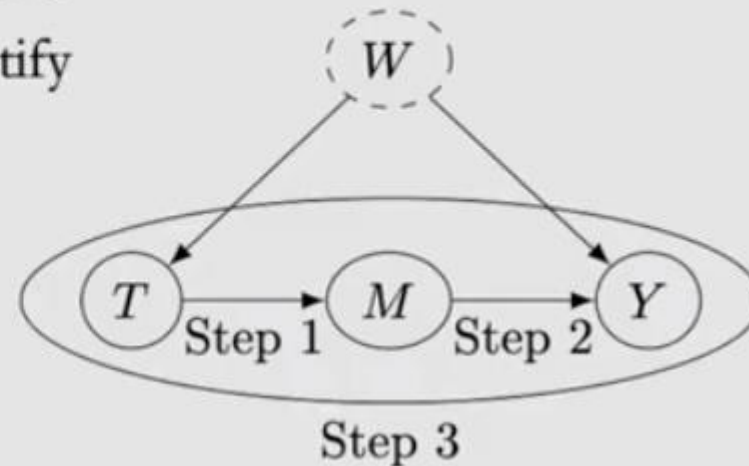
근데 만약에 W 를 관찰을 못한다면?

5.3 Frontdoor adjustment: big picture

W를 알 수 없을 때, M에 집중해서 막아보자



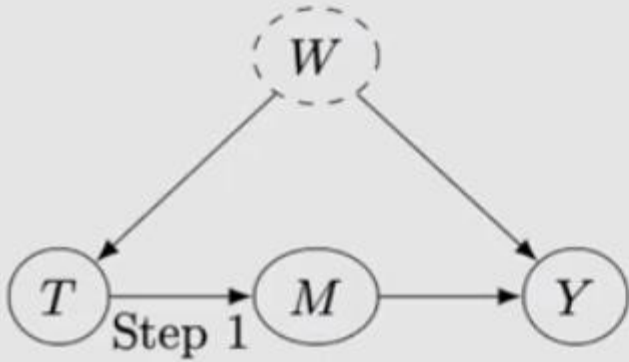
1. Identify the causal effect of T on M
2. Identify the causal effect of M on Y
3. Combine the above steps to identify the causal effect of T on Y



5.3 Frontdoor adjustment: step1

Identify the causal effect of T on M

$$P(m \mid do(t)) = P(m \mid t)$$



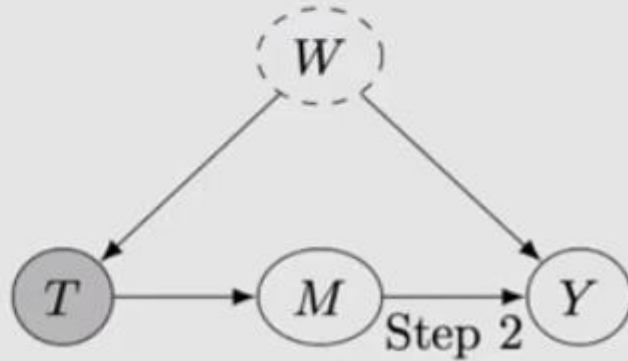
T->W->Y->M으로 가는 path는 Y가 collider로 막아버리니까

T에서 M으로 가는 backdoor path가 없으므로 Do(t)가 |t로 변화

5.3 Frontdoor adjustment: step 2

Identify the causal effect of M on Y

$$P(y \mid do(m)) = \sum_t P(y \mid m, t) P(t)$$



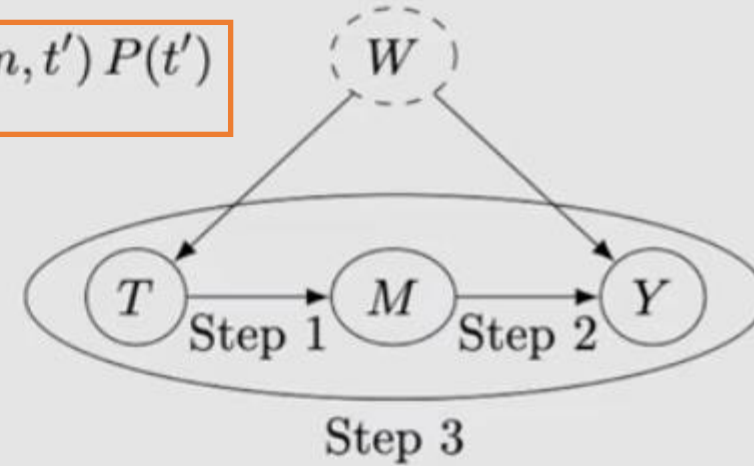
M-→T-→W-→Y 로의 path는 T를 conditioning함으로써 막을 수 있음. 따라서

5.3 Frontdoor adjustment: step 3

Combine steps 1 and 2 to identify the causal effect of T on Y

Goal

$$P(y \mid do(t)) = \sum_m \overset{\text{STEP1}}{P(m \mid do(t))} \overset{\text{STEP2}}{P(y \mid do(m))}$$
$$= \sum_m \overset{\text{STEP1}}{P(m \mid t)} \overset{\text{STEP2}}{\sum_{t'} P(y \mid m, t') P(t')}$$



5.3 Frontdoor adjustment and criterion

If (T, M, Y) satisfy the frontdoor criterion, and we have positivity, then

$$P(y \mid do(t)) = \sum_m P(m \mid t) \sum_{t'} P(y \mid m, t') P(t')$$

A set of variables M satisfies the **frontdoor criterion** relative to T and Y if the following are true:

1. M completely mediates the effect of T on Y (i.e. all causal paths from T to Y go through M).
2. There is no unblocked backdoor path from T to M .
3. All backdoor paths from M to Y are blocked by T .

조건1. M 은 full mediator여야 한다 ($T \rightarrow M \rightarrow Y$ 100프로)

조건2. $T \rightarrow M$ 으로의 backdoor는 막혀야 함 (step 1)

조건3. $M \rightarrow Y$ 로의 path는 T conditioning으로 막혀야 함 (step 2)

5.4 Pearl's do-calculus

Backdoor 혹은 frontdoor criterion 둘 다 만족 못할 때도, causal effect를 identify할 수 있냐? **Yes, do-calculus**

Pearl's do-calculus로 어떤 causal quantity등 identify할 수 있다.

Will allow us to identify any causal quantity that is identifiable

$$P(Y \mid do(T = t), X = x)$$

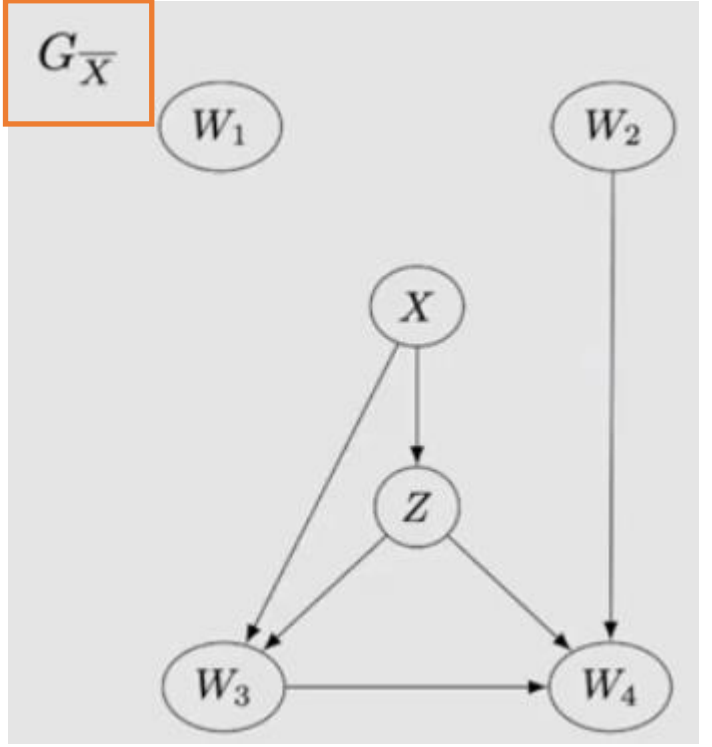
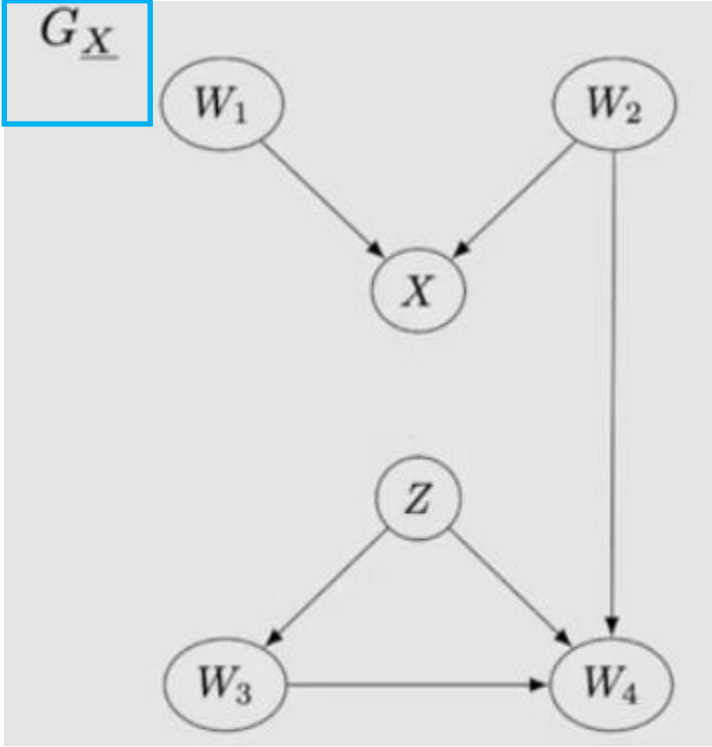
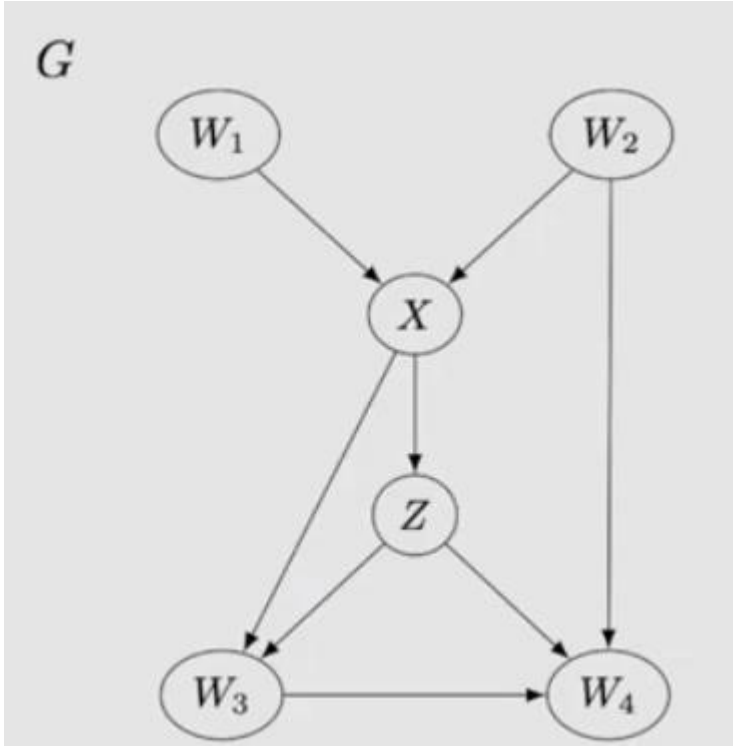
where Y , T , and X are arbitrary sets

Multiple treatments and/or multiple outcomes

Notation for Pearl's *do*-calculus

X로부터 나가는 모든 edge 모두 삭제

X로부터 들어오는 모든 edge 삭제



Rules for Pearl's *do*-calculus

Graph에서 T의 incoming edge를 삭제했을 때 (T는 coin toss니까)
Y가 T, W가 conditioning 되었을 때 Z랑 de-separate면 (흐르지 않으면)

Rule 1 of *do*-calculus

$$P(y \mid \cancel{do(t)}, z, w) = P(y \mid \cancel{do(t)}, w) \quad \text{if } Y \perp\!\!\!\perp_{G_{\cancel{T}}} Z \mid T, W$$

Question: What concept does this remind you of?

T가 empty set이면, 아래와 같이 그냥 일반적인 d-separation 식

Rule 1 with *do*(t) removed:

$$P(y \mid z, w) = P(y \mid w) \quad \text{if } Y \perp\!\!\!\perp_G Z \mid W$$

Generalization of d-separation to interventional distributions

Rules for Pearl's *do*-calculus

Rule 2 of *do*-calculus

$$P(y \mid \cancel{do(t)}, do(z), w) = P(y \mid \cancel{do(t)}, z, w) \quad \text{if } Y \perp\!\!\!\perp_{G_{\cancel{T}, Z}} Z \mid T, W$$

Question: What concept does this remind you of?

Rule 2 with $do(t)$ removed:

$$P(y \mid do(z), w) = P(y \mid z, w) \quad \text{if } Y \perp\!\!\!\perp_{G_Z} Z \mid W$$

Backdoor adjustment 의 확장

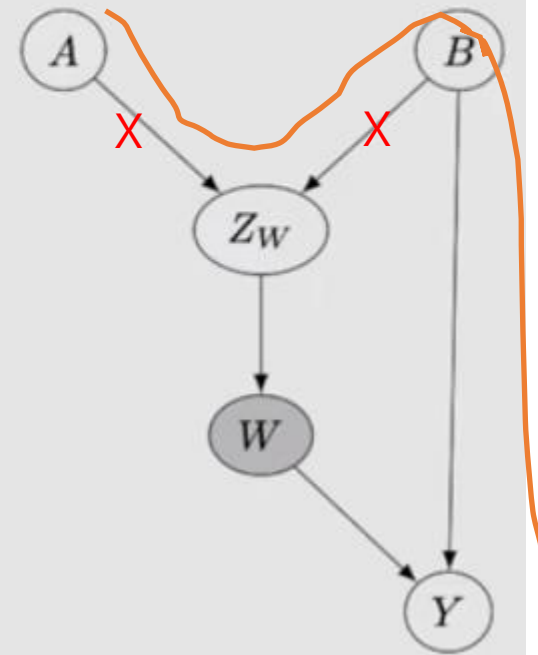
Generalization of backdoor adjustment/criterion

Rules for Pearl's *do*-calculus

Rule 3 of *do*-calculus

$$P(y \mid do(t), do(z), w) = P(y \mid do(t), w) \quad \text{if } Y \perp\!\!\!\perp_{G_{\overline{T}, \overline{Z(W)}}} Z \mid T, W$$

where $Z(W)$ denotes the set of nodes of Z that aren't ancestors of any node of W in $G_{\overline{T}}$



Rule 3 with $do(t)$ removed:

$$P(y \mid do(z), w) = P(y \mid w) \quad \text{if } Y \perp\!\!\!\perp_{G_{\overline{Z(W)}}} Z \mid W$$

$$P(y \mid do(z), w) = P(y \mid w) \quad \text{if } Y \perp\!\!\!\perp_{G_{\overline{Z}}} Z \mid W$$

Do(z)를 없애고 싶다 -> Y랑 Z랑 graph상에서 연결이 없으면 된다.

그럼, Z로 들어오는 incoming edge만 삭제하면 될까? 아니

W의 부모인 Z의 incoming edge를 삭제 해야 한다 Z(W)

W가 conditioning되면 Z_W 가 collider가 되어서 $A \rightarrow Z(W) \rightarrow B \rightarrow Y$ 로 흐르게 됨으로. $Z(W)$ 의 incoming edge삭제

Rules for Pearl's *do*-calculus

The rules of *do*-calculus

Rule 1: $P(y \mid do(t), z, w) = P(y \mid do(t), w)$ if $Y \perp\!\!\!\perp_{G_{\overline{T}}} Z \mid T, W$

Rule 2: $P(y \mid do(t), do(z), w) = P(y \mid do(t), z, w)$ if $Y \perp\!\!\!\perp_{G_{\overline{T}, Z}} Z \mid T, W$

Rule 3: $P(y \mid do(t), do(z), w) = P(y \mid do(t), w)$ if $Y \perp\!\!\!\perp_{G_{\overline{T}, \overline{Z(W)}}} Z \mid T, W$

Completeness of *do*-calculus

Completeness of *do*-calculus

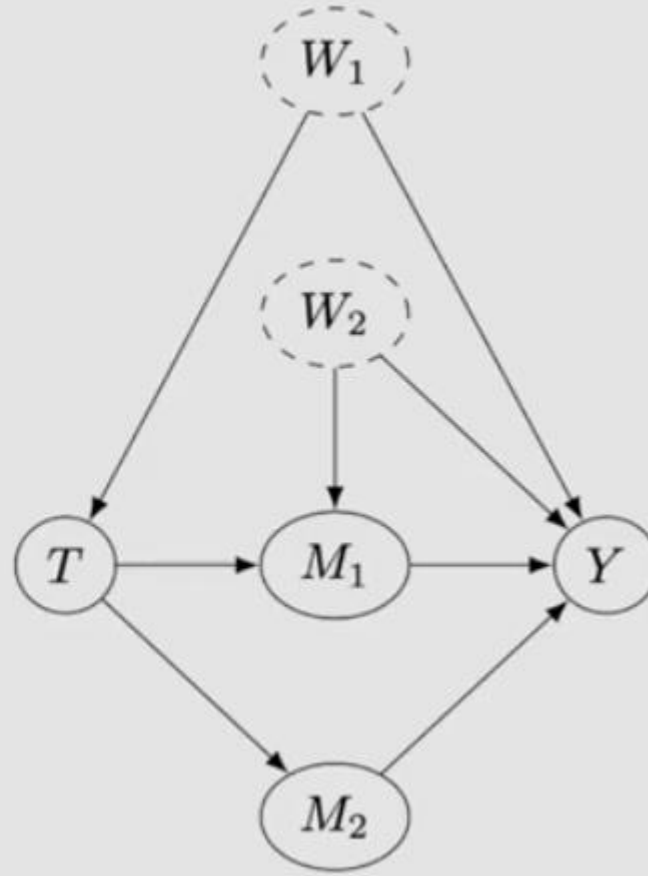
Maybe there are some identifiable causal estimands that can't be identified using the rules of *do*-calculus

Fortunately, not, as *do*-calculus is complete ([Shpitser & Pearl, 2006a](#); [Huang & Valtorta, 2006](#); [Shpitser & Pearl, 2006b](#))

5.5 Determining Identifiability from the Graph

Question:

In this graph, is the
backdoor criterion satisfied?



ator여야 한다 (T->M->Y 100프로)
backdoor는 막혀야 함 (step 1)
h는 T conditioning으로 막혀야 함 (step 2)

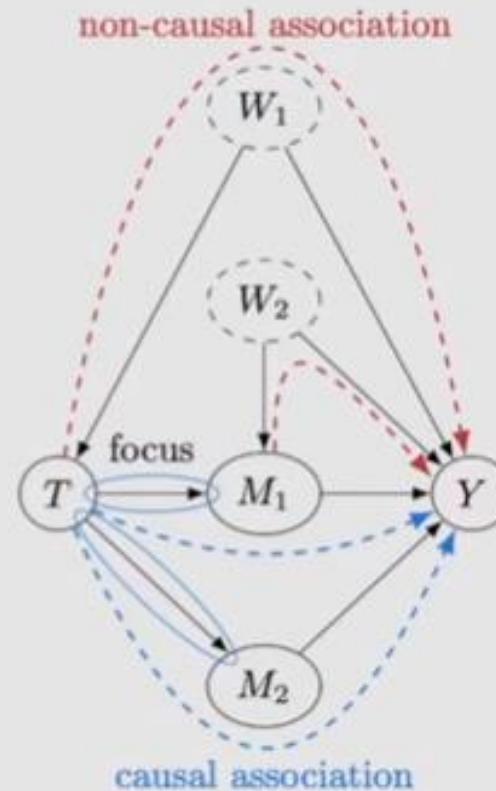
5.5 Determining Identifiability form the Graph

Unconfounded children criterion

This criterion is satisfied if it is possible to block all backdoor paths from the treatment variable T to all of its children that are ancestors of Y with a single conditioning set ([Tian & Pearl, 2002](#)).

T에서 나오는 모든 자식 노드이며 Y의 부모 노드인 것들을 막으면
Unconfounded children criterion

Sufficient condition for identifiability when T is a single variable



Backdoor: (W1 관찰 안 되서 block)
T->W1->Y

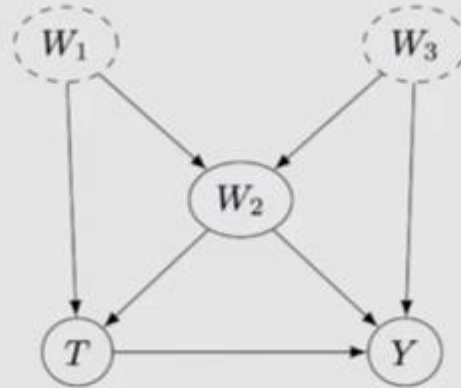
Frontdoor: (W2 관찰 안 되서 block)
M1->W2->Y

T의 자식노드이자 Y의 부모노드인 (T->M1, T->M2)만 남기고 다 막으면 된다.

(Backdoor, frontdoor보다 좀 더 general한 컨셉)

Necessary condition for identifiability

For each backdoor path from T to any child M of T that is an ancestor of Y , it is possible to block that path ([Pearl, 2009](#), p. 92).



T 의 자식이자, Y 의 부모노드인 M 에 영향을 미치는 path block 가능하다

Backdoor

1) $T \rightarrow W_1 \rightarrow W_2 \rightarrow W_3 \rightarrow Y$: blocked by conditioning W_2

2) $T \rightarrow W_2 \rightarrow Y$: blocked by conditioning W_2

근데, W_2 conditioning하면 두개를 동시에 막을 수 없음 (collider라서)

즉, single conditioning으로 모든 path를 막을 수 있는 상황은
Unconfounded children criterion: T 의 자식이며 Y 의 부모 노드에 대해서만 가능

지금 상황에서 W_2 는 T 의 자식이 아님. Y 의 부모는 맞지만

Necessary condition for identifiability

Recall: identification with the rules *do*-calculus is necessary and sufficient
([Shpitser & Pearl, 2006a](#); [Huang & Valtorta, 2006](#); [Shpitser & Pearl, 2006b](#))

For graphical criterion, see [Shpitser & Pearl, 2006a, 2006b](#): hedge criterion