Conditional Expected Shortfall

Nonparametric Estimation

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- Risk Measures
- Nonparametric Estimation
- Simulation
- Oata Analysis

Reviewed Paper

[1] Z. Cai and X. Wang. "Nonparametric estimation of conditional VaR and expected shortfall". In: *Journal of Econometrics* 147.1 (2008), pp. 120-130. ISSN: 0304-4076. DOI: 10.1016/j.jeconom.2008.09.005.

Keywords

Boundary effects, Empirical likelihood, Expected shortfall, Local linear estimation, Nonparametric smoothing, Value-at-risk, Weighted double kernel

Section 1

Risk Measures

Value at Risk

Given time horizon,

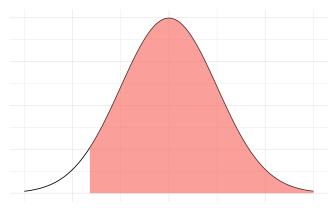


Figure 1: Loss Distribution - Can the financial institution still be in business after a catastrophic event?

Two Viewpoints

Tsay (2010) says that

- Financial institution: Maximal loss of a financial position during a given time period for a given probability
 - Measure of loss under normal market conditions
- Regulatory committe: Minimal loss under extraordinary market circumstances

Definition

- p: **Right** tail probability
- I: Time horizon
- L(I): loss function of the asset from t to t+I
- F_I : CDF of L(I)

$$p = P[L(I) \ge VaR]$$

i.e. VaR can be computed by finding the p-th quantile.

Quantile Loss

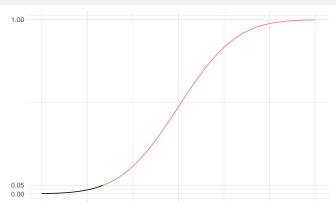


Figure 2: CDF of Loss

$$VaR = \inf \{x: F_I(x) \ge 1 - p\}$$

Loss function and log Returns

Consider dollar $\{P_t\}$.

$$L(I) = P_{t+1} - P_t = R_{t+1} + \ldots + R_{t+1}$$

where $R_t = P_t - P_{t-1}$ is the return series.

- Loss occurs when the return R_t are negative
- Log returns $Y_t = \ln R_t$ corresponds approximately to percentage changes
- Use negative log returns
- VaR by Upper quantile of the distribution of log return

VaR using log Return

$$\text{percentage change} \approx \ln \frac{P_{t+1}}{P_t} = Y_{t+1}$$

From t to t + 1,

$$VaR = P_t \times VaR(Y_{t+1})$$

Subadditivity

- When two portfolios are merged, the risk measure should not be greater than the sum of each.
- VaR underestimates the actual loss.
- Thus, Expected shortfall

VaR and Expected Loss

- VaR: Quantile loss of the loss distribution
- ES: Expected value of loss function if the VaR is exceeded

$$ES:=E[L(I) \mid L(I) \geq VaR]$$

Conditional information

- $\{X_t: t = 1, \dots n\}$
- Exogenous variable: economic, market variables
- or past observed returns e.g. $\{Y_{t-1}\}$

Conditional VaR

- Stationary log-return $\{Y_t: t=1,\ldots n\}$
- Conditional information $\{X_t: t = 1, \dots n\}$
- Conditional VaR (CVaR)

$$\nu_p(x) = S^{-1}(p \mid x)$$

where

- $S(y \mid x) := 1 F(y \mid x)$
- F: conditional CDF of Y_t given $X_t = x$.

Conditional Expected Shortfall

$$\mu_p(x) = E[Y_t \mid Y_t \ge \nu_p(x), X_t = x]$$

Formulating CES

Let
$$B \equiv \{\omega \colon Y_t(\omega) \mid X_t = x \ge \nu_p(x)\} \in \mathcal{B}$$
. Then

$$\mu_{p}(x) = E[Y_{t} \mid Y_{t} \geq \nu_{p}(x), X_{t} = x]$$

$$= \frac{1}{P(B)} \int_{B} Y_{t} dP$$

$$= \frac{1}{P(Y_{t} \geq \nu_{p}(x) \mid X_{t} = x)} \int_{\nu_{p}(x)}^{\infty} yf(y \mid x) dy$$

$$= \frac{1}{p} \int_{\nu_{p}(x)}^{\infty} yf(y \mid x) dy$$

Section 2

Nonparametric Estimation

Goal

Risk measures

- CVaR: $\hat{\nu}_p(x)$
- CES: $\hat{\mu}_p(x)$

Workflow of Estimation

Plugging-in Method

$$\hat{\mu}_p(x) = \frac{1}{p} \int_{\hat{\nu}_p(x)}^{\infty} y \hat{f}(y \mid x) dy$$

What to estimate

- Conditional PDF: $\hat{f}(y \mid x)$
- CVaR: $\hat{\nu}_p(x) = \hat{S}^{-1}(p \mid x)$ by inverting the conditional CDF

Section 3

Simulation

AR(1)-GARCH(1, 0)

$$\begin{cases} X_t = Y_{t-1} \\ Y_t = 0.01 + 0.62X_t + \sigma_t \epsilon_t \\ \sigma_t^2 = 0.15 + 0.65\sigma_{t-1}^2 \\ \epsilon_t \sim N(0, 1) \end{cases}$$

True conditional distribution

Since $\epsilon_t \sim N(0,1)$,

$$\sigma_t \epsilon_t \sim N(0, \sigma_t^2)$$

$$Y_t \mid X_t \sim N\left(0.01 + 0.62X_t, \sigma_t^2\right)$$

True CES

- For each X_t , pnorm(x, mean, sd) gives the conditional cdf value.
- Inverting $S(y \mid x) = 1 F(y \mid x)$ gives $\nu_p(x)$.

$$\nu_p(x) = S^{-1}(p \mid x)$$

• Plugging-in method gives $\mu_p(x)$.

$$\mu_p(x) = \frac{1}{p} \int_{\nu_p(x)}^{\infty} y f(y \mid x) dy$$

Goal of MC Simulation

- Compute the error between the true $\mu_p(x)$ and $\hat{\mu}_p(x)$
- Is the estimator of Cai and Wang (2008) good?

Process

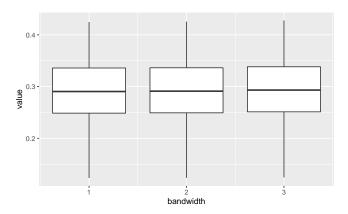
Monte Carlo Samples:

- For fixed x_t (pre-determined grid points)
- ② Generate GARCH(1, 0): (σ_t, ϵ_t)
- **3** Generate Y_t using AR(1) for each $X_t = Y_{t-1}$
- **4** AR(1): $Y_t = 0.01 + 0.62 Y_{t-1} + \sigma_t \epsilon_t$

Monte Carlo Samples

```
#>
         sigma
               garch
                         mc
                               хt
                                      уt
                         s1 0.005 -0.397
#>
      1: 0.655 -0.41011
#>
     2: 0.655
              0.12022
                         s1 0.010
                                   0.136
     3: 0.655 -0.54705
                         s1 0.015 -0.528
#>
#>
     4: 0.655
              1.04436
                         s1 0.020
                                   1.067
#>
     5: 0.655 0.21571
                         s1 0.025
                                   0.241
#>
        0.655 - 0.00527 \text{ s} 20 \ 0.480
                                   0.302
  1997:
         0.655 0.67677 s20 0.485
                                   0.987
  1998: 0.655 -0.52315 s20 0.490 -0.209
  1999: 0.655 0.65742 s20 0.495
                                   0.974
#> 2000: 0.655 -0.20423 s20 0.500
                                   0.116
```

MC

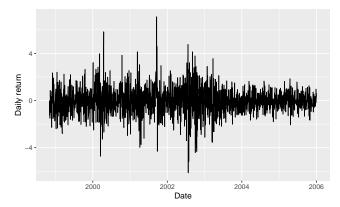


Section 4

Data Analysis

Dow Jones Industrial Index

Cai and Wang (2008) used Dow Jones index with daily return defined by y_t : = $-100 \ln \frac{P_t}{P_{t-1}}$



References

Cai, Zongwu, and Xian Wang. 2008. "Nonparametric estimation of conditional VaR and expected shortfall." *Journal of Econometrics* 147 (1): 120–30. https://doi.org/10.1016/j.jeconom.2008.09.005.

Tsay, Ruey S. 2010. Analysis of Financial Time Series. John Wiley & Sons.