Lab 2: Exercises

ML701: Machine Learning MBZUAI

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Problem 1. Suppose $\mathbb{P}[A]$, $\mathbb{P}[A']$, $\mathbb{P}[B|A]$, and $\mathbb{P}[B|A']$ are known, where A' is a complement of A, i.e., $A' = \mathcal{X} \setminus A$. Find an expression for $\mathbb{P}[A|B]$ in terms of these four probabilities.

Solution. Using Bayes' rule with the fact that $\mathbb{P}[A' \cup A] = 1$, i.e., it is the certain event, we have

$$\mathbb{P}\left[A|B\right] = \frac{\mathbb{P}\left[A \cup B\right]}{\mathbb{P}\left[B\right]} = \frac{\mathbb{P}\left[A\right] \mathbb{P}\left[B|A\right]}{\mathbb{P}\left[A\right] \mathbb{P}\left[B|A\right] + \mathbb{P}\left[A'\right] \mathbb{P}\left[B|A'\right]}.$$

Problem 2. Assume the probability of having tuberculosis (TB) is 0.0005, and a test for TB is 99% accurate. What is the probability one has TB if one tests positive for the disease?

Solution. We apply the solution to Problem 1 with A correspond to having tubercolosis, and B refers to tubercolosis test being positive. Then, we can rewrite the problem formulation as

$$\mathbb{P}[A] = 0.005, \quad \mathbb{P}[A'] = 0.995, \quad \mathbb{P}[B|A] = 0.99, \quad \text{and} \quad \mathbb{P}[B|A'] = 0.01.$$

Therefore, plugging this to the derived formula

$$\begin{split} \mathbb{P}\left[\text{Having Tuberculosis}|\text{Positive Test}\right] &= \mathbb{P}\left[A|B\right] = \frac{\mathbb{P}\left[A\right]\mathbb{P}\left[B|A\right]}{\mathbb{P}\left[A\right]\mathbb{P}\left[B|A\right] + \mathbb{P}\left[A'\right]\mathbb{P}\left[B|A'\right]} \\ &= \frac{0.005*0.99}{0.005*0.99 + 0.995*0.01} \approx 0.0472. \end{split}$$

Problem 3. [Coupon Collector's Problem]

Imagine you're collecting trading cards from cereal boxes. There are n distinct cards to collect, and each cereal box contains one card, randomly chosen with equal probability. The question is: on average, how many cereal boxes do you need to buy in order to collect all n distinct cards?

Hint: Try to use linearity of expectation.

Solution. For card i, let X_i be the number of boxes you have to buy after collecting the i-1 distinct card until you get the i^{th} distinct card. The goal is to find $\mathbb{E}[X]$ where $X = X_1 + X_2 + ... + X_n$. Notice:

- 1. $\mathbb{E}[X_1] = 1$ because you are guaranteed to get a new card in the first box.
- 2. After obtaining the first card, the probability that the next box contains a new card is $\frac{n-1}{n}$. Therefore, $\mathbb{E}[X_2] = \frac{n}{n-1}$.
- 3. After obtaining the first two cards, the probability that the next box contains a new card is $\frac{n-2}{n}$. Therefore, $\mathbb{E}[X_3] = \frac{n}{n-2}$.

4. And so on...

Therefore, the expected number of boxes to get all n distinct cards is:

$$\mathbb{E}[X] = \mathbb{E}[X_1 + X_2 + \dots + X_n]$$

$$= \mathbb{E}[X_1] + \mathbb{E}[X_2] + \dots + \mathbb{E}[X_n]$$

$$= 1 + \frac{n}{n-1} + \frac{n}{n-2} + \dots + n$$

$$= n\left(1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n}\right).$$

The sum in the parentheses is the n^{th} Harmonic number, which is approximately $\log(n) + \gamma$ where γ is the Euler-Mascheroni constant.

Problem 4. [Decision Rule Based on Posterior Probabilities and Joint Densities]

Consider a binary prediction problem where the outcome Y can take values in the set $\{0,1\}$. Suppose we're making decisions based on the ratio:

$$\frac{\mathbb{P}\left[Y=1|X=x\right]}{\mathbb{P}\left[Y=0|X=x\right]}$$

Prove that any decision rule using the above ratio is equivalent to considering the ratio of the joint densities:

$$\frac{p(x, Y = 1)}{p(x, Y = 0)}$$

Solution. Using Bayes' rule, we have

$$\frac{\mathbb{P}[Y=1|X=x]}{\mathbb{P}[Y=0|X=x]} = \frac{\frac{p(x,Y=1)}{p(x)}}{\frac{p(x,Y=0)}{p(x)}} = \frac{p(x,Y=1)}{p(x,Y=0)}.$$

Problem 5. [Optimal Binary Prediction with Symmetric Loss and MAP Rule]

Consider a binary prediction problem, where the outcome Y can take values in the set $\{0,1\}$. Suppose we have a symmetric loss function, i.e.,

$$loss(y, y') = loss(y', y)$$

for all y, y' in $\{0, 1\}$. Furthermore, assume loss(1, 1) = loss(0, 0) = 0.

- 1. Show that for every such symmetric loss, the optimal binary predictor is equivalent to the Maximum A Posteriori (MAP) rule.
- 2. Furthermore, if $\mathbb{P}[Y=1] = \mathbb{P}[Y=0] = \frac{1}{2}$, demonstrate that the likelihood estimator yields the same result as the MAP rule.

Solution. For the first part, using Lemma 16, we have

$$\begin{split} \hat{Y}^{\star}(x) &= \mathbbm{1}\left\{\mathbb{P}\left[Y = 1 \mid X = x\right] \geq \frac{loss(1,0) - loss(0,0)}{loss(0,1) - loss(1,1)} \cdot \mathbb{P}\left[Y = 0 \mid X = x\right]\right\} \\ & \stackrel{loss(1,1) = loss(0,0) = 0}{=} \mathbbm{1}\left\{\mathbb{P}\left[Y = 1 \mid X = x\right] \geq \frac{loss(1,0)}{loss(0,1)} \cdot \mathbb{P}\left[Y = 0 \mid X = x\right]\right\} \\ & \stackrel{loss(1,0) = loss(0,1)}{=} \mathbbm{1}\left\{\mathbb{P}\left[Y = 1 \mid X = x\right] \geq \cdot \mathbb{P}\left[Y = 0 \mid X = x\right]\right\} \\ &= \arg\max_{y \in \{0,1\}} \mathbb{P}\left[Y = y \mid X = x\right]. \end{split}$$

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For the second part,

$$\begin{split} \hat{Y}^{\star}(x) &= \arg\max_{y \in \{0,1\}} \mathbb{P}\left[Y = y \mid X = x\right] \\ &= \arg\max_{y \in \{0,1\}} \frac{p(x,Y = y)}{p(x)} \\ &\stackrel{p(x) \perp y \text{ appears in both }}{=} \arg\max_{y \in \{0,1\}} p(x,Y = y) \\ &\stackrel{\text{Bayes' rule }}{=} \arg\max_{y \in \{0,1\}} \mathbb{P}\left[Y = y\right] p(x|Y = y) \\ &\stackrel{\mathbb{P}[Y = y] = 1/2, \ \forall y \in \{1,2\}}{=} \arg\max_{y \in \{0,1\}} p(x|Y = y). \end{split}$$