Baysian modeling of hurricane trajectories P8160 Group Project 3 Baysian modeling of hurricane trajectories

Jingchen Chai, Yi Huang, Zining Qi, Ziyi Wang, Ruihan Zhang

Columbia University

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Content

- Introduction
- EDA
- Method
- Results
- 6 Limitations and Conclusion

Introduction

- Hurricanes cause fatalities and property damage
- There is a growing need to accurately predict hurricane behavior, including location and speed (Taboga, 2021)
- This project aims to forecast wind speeds by modeling hurricane trajectories using a Hierarchical Bayesian Model.

Data

ID: ID of hurricanes

Year: In which year the hurricane occurred

Month: In which month the hurricane occurred

Nature: Nature of the hurricane

ET: Extra Tropical

DS: Disturbance

NR: Not Rated

SS: Sub Tropical

TS: Tropical Storm

Time: dates and time of the record

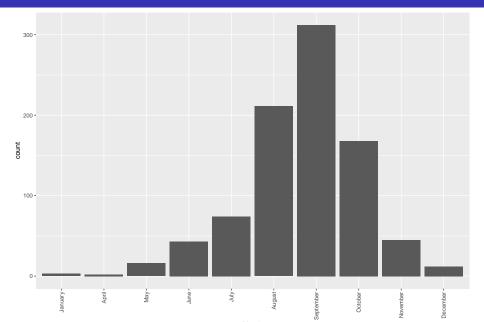
Latitude and Longitude: The location of a hurricane check point

Wind.kt: Maximum wind speed (in Knot) at each check point

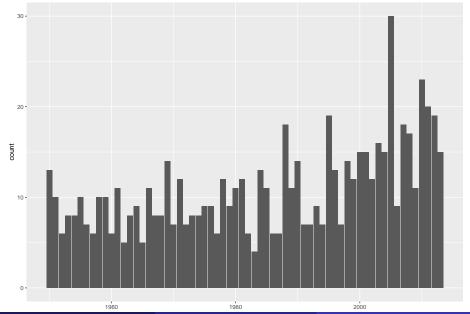
Data Pre-processing

- We have filtered observations that occurred on a 6-hour intervals. (e.g., hour 0, 6, 12, 18)
- Calculated the lag difference for latitude, longitude and wind speed.
- After data cleaning, we obtained 20293 observations and with 699 different hurricanes.

EDA-Count of Hurricanes in each Month

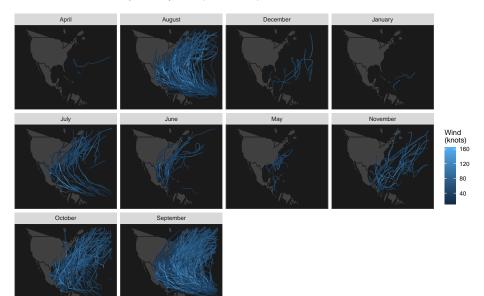


EDA-Count of Hurricanes in each Year



Show hurricance tracks by month

Atlantic named Windstorm Trajectories by Month (1950 - 2013)



Bayesian Model

The suggested Bayesian model is $Y_i(t+6)=\beta_{0,i}+\beta_{1,i}Y_i(t)+\beta_{2,i}\Delta_{i,1}(t)+\beta_{3,i}\Delta_{i,2}(t)+\beta_{4,i}\Delta_{i,3}(t)+X_i\gamma+\epsilon_i(t)$

- where $Y_i(t)$ the wind speed at time t (i.e. 6 hours earlier), $\Delta_{i,1}(t)$, $\Delta_{i,2}(t)$ and $\Delta_{i,3}(t)$ are the changes of latitude, longitude and wind speed between t and t-6, and $\epsilon_{i,t}$ follows a normal distributions with mean zero and variance σ^2 , independent across t.
- $X_i=(x_{i,1},x_{i,2},x_{i,3})$ are covariates with fixed effect γ , where $x_{i,1}$ be the month of year when the i-th hurricane started, $x_{i,2}$ be the calendar year of the i hurricane, and $x_{i,3}$ be the type of the i-th hurricane.
- $\boldsymbol{\beta}_i = (\beta_{0,i}, \beta_{1,i}, ..., \beta_{5,i})$, we assume that $\boldsymbol{\beta}_i \sim N(\mu, \Sigma)$.

Prior Distribution

$$P(\mu) = \frac{1}{\sqrt{2\pi}|V|^{\frac{1}{2}}} \exp\{-\frac{1}{2}\mu^\top V^{-1}\mu\} \propto |V|^{-\frac{1}{2}} \exp\{-\frac{1}{2}\mu^\top V^{-1}\mu\}$$

where V is a variance-covariance matrix

$$P(\Sigma) \propto |\Sigma|^{-\frac{(\nu+d+1)}{2}} \exp(-\frac{1}{2} tr(S\Sigma^{-1}))$$

where d is the dimension of $\boldsymbol{\beta}_i.$

$$P(\gamma) \propto exp(-\frac{\gamma^2}{2*(0.05)^2}) = e^{-200\gamma^2}$$

$$P(\sigma) = \frac{2\alpha}{\pi + \alpha^2} \propto \frac{1}{\sigma^2 + \alpha^2} = \frac{1}{\sigma^2 + 100}$$

Posterior

Let $\mathbf{B} = (\beta_1^\top, ..., \beta_n^\top)^\top$, derive the posterior distribution of the parameters $\Theta = (\mathbf{B}^\top, \mu^\top, \sigma^2, \Sigma, \gamma)$.

$$\begin{split} \operatorname{Let} & Z_i(t) \boldsymbol{\beta}_i^\top = \\ & \beta_{0,i} + \beta_{1,i} Y_i(t) + \beta_{2,i} \Delta_{i,1}(t) + \beta_{3,i} \Delta_{i,2}(t) + \beta_{4,i} \Delta_{i,3}(t) + X_i \boldsymbol{\gamma} + \epsilon_i(t) \\ \operatorname{We can find that} & Y_i \sim MVN(\boldsymbol{Z}_i \boldsymbol{\beta}_i, \sigma^2 \boldsymbol{I}) \end{split}$$

The likelihood for Y is

$$\begin{split} &f(Y\mid B,\mu,\sigma^2,\Sigma,\gamma) = \prod_{i=1}^n f(Y_i|B,\mu,\Sigma,\sigma^2) = \\ &\prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\{-\frac{1}{2}(y_i - Z_i\beta_i - X_i\gamma_i)^\top (\sigma^2I)^{-1}(y_i - Z_i\beta_i - X_i\gamma_i)\} \\ &\propto (2\pi\sigma^2)^{-\frac{n}{2}} \prod_{i=1}^n \exp\big\{-\frac{1}{2}(Y_i - Z_i\beta_i - X_i\gamma_i)^\top (\sigma^2I)^{-1}(Y_i - Z_i\beta_i - X_i\gamma_i)\big\} \end{split}$$

Joint Posterior

$$\begin{split} \pi(\Theta|Y) &= P(B,\mu,\sigma^2,\Sigma,\gamma|Y) \propto \underbrace{L(Y|B,\sigma^2)}_{\text{likelihood of }Y} \underbrace{L(B|\mu,\Sigma)}_{\text{distribution of B}} \underbrace{p(\mu)p(\sigma)p(\Sigma)p(\gamma)}_{\text{priors}} \\ &\propto \frac{1}{\sigma^N(\sigma^2+10^2)} \prod_{i=1}^n \exp\big\{-\frac{1}{2}(Y_i-Z_i\beta_i-X_i\gamma_i)^\top(\sigma^2I)^{-1}(Y_i-Z_i\beta_i-X_i\gamma_i)\big\} \\ &\times \exp\{-\frac{1}{2}\sum_i^n(\beta_i-\mu)^\top\Sigma^{-1}(\beta_i-\mu)\}|\Sigma^{-1}|^{\frac{N+d+v+1}{2}} \exp\{-\frac{1}{2}tr(S\Sigma^{-1})\}|V|^{-\frac{1}{2}} \\ &\times \exp\{-\frac{1}{2}\mu^\top V^{-1}\mu\} \\ &\times \exp\{-200\gamma^2\} \end{split}$$

where V is a variance-covariance matrix, N is the total number of hurricanes and d is the dimension of β , and v is the degree of freedom.

MCMC for Hierarchical Bayesian Model: Method

Conditional Distribution of each parameter:

- $\begin{aligned} \bullet \ \beta_i \sim MVN(N^{-1}M,N^{-1}), \ \text{where} \ N &= \frac{Z_i^\top Z_i}{\sigma^2} + \Sigma^{-1}, \\ M &= \frac{Z_i^\top Y_i Z_i^\top X_i \gamma}{\sigma^2} + \mu \Sigma^{-1} \end{aligned}$
- $\mu \sim MVN(H^{-1}M,N^{-1})$, where $H=N\Sigma^{-1}-\frac{1}{V}$, $M=\sum_i^n \beta_i \Sigma^{-1}$
- $\bullet \ \Sigma \sim W^{-1}(S + \textstyle\sum_{i}^{n}(\beta_{i} \mu)(\beta_{i} \mu)^{\top}, n + v)$
- $$\begin{split} \bullet \ \, \gamma \sim MVN(N^{-1}M,N^{-1}) \text{, where } N = \frac{\sum_i^n X_i^\top X_i}{\sigma^2} + 400I \text{,} \\ \mathsf{M} = & \frac{\sum_i^n (X_i^\top Y_i X_i^\top Z_i \beta_i)}{\sigma^2} \end{split}$$

•

$$\begin{split} \pi(\sigma|\boldsymbol{Y}, \mathbf{B}^{\top}, \boldsymbol{\mu}^{\top}, \boldsymbol{\Sigma}, \boldsymbol{\gamma}) &\propto \frac{1}{\sigma^{N}(\sigma^{2} + 10^{2})} \\ &\times \prod_{i=1}^{n} \exp\big\{ -\frac{1}{2(\sigma^{2}I)} (\boldsymbol{Y}_{i} - \boldsymbol{Z}_{i}\boldsymbol{\beta}_{i} - \boldsymbol{X}_{i}\boldsymbol{\gamma}_{i})^{\top} (\boldsymbol{Y}_{i} - \boldsymbol{Z}_{i}\boldsymbol{\beta}_{i} - \boldsymbol{X}_{i}\boldsymbol{\gamma}_{i}) \big\} \end{split}$$

MCMC Algorithm - Metropolis-Hastings

Target distribution is

$$\begin{split} \pi(\sigma|Y,\mathbf{B}^\top,\mu^\top,\Sigma,\gamma) &\propto \frac{1}{\sigma^N(\sigma^2+10^2)} \\ &\times \prod_{i=1}^n \exp\big\{-\frac{1}{2\sigma^2}(\boldsymbol{Y}_i-\boldsymbol{Z}_i\boldsymbol{\beta}_i-\boldsymbol{X}_i\boldsymbol{\gamma}_i)^\top(\boldsymbol{Y}_i-\boldsymbol{Z}_i\boldsymbol{\beta}_i-\boldsymbol{X}_i\boldsymbol{\gamma}_i)\big\} \end{split}$$

- Choose a random walk with step size distributed as a uniform random variable
- \bullet The conditional density is $q(x|y) = \frac{1}{2a} 1_{[y-a,y+a]}(x)$
- \bullet Proposed q is symmetric, thus the acceptance rate is only depend on $P(\sigma|B,\mu,A,\gamma,Y)$

MCMC Algorithm - Metropolis-Hastings

- The acceptance rate $\alpha_{XY} = \min(1, \frac{P(X|B,\mu,A,\gamma,Y)}{P(Y|B,\mu,A,\gamma,Y)})$
- $\bullet \ \, \text{Accept X if} \,\, U < \alpha_{XY}$
- Iterate over 1000 times
- New σ is the mean of last 200 values in the chain

MCMC Algorithm - Gibbs Sampling

We apply a MCMC algorithm consisting of Gibb Samping and Metropolis-Hastings steps.

Parameters are updated component-wise for each k=1,...,N,N=5000:

- Generate $\beta_{ij}, j=0,1,2,3,4$ for i^{th} hurricane from $\pi(\mathbf{B}|Y,\mu_{k-1}^\intercal,\sigma_{k-1},\Sigma_{k-1},\gamma_{k-1})$
- Generate $\mu_j, j=0,1,2,3,4$ from $\pi(\mu|Y,\mathbf{B}_k,\sigma_{k-1},\Sigma_{k-1},\gamma_{k-1})$
- ullet Generate σ_k from the Metropolis-Hastings steps
- \bullet Generate Σ_k from $\pi(\Sigma|Y,\mathbf{B}_k,\boldsymbol{\mu}_k,\boldsymbol{\sigma}_k,\boldsymbol{\gamma}_{k-1})$
- \bullet Generate γ_k from $\pi(\gamma|Y,\mathbf{B}_k,\boldsymbol{\mu}_k,\boldsymbol{\sigma}_k,\boldsymbol{\Sigma}_k)$

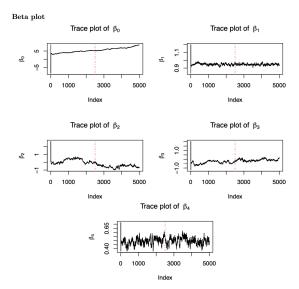
MCMC Algorithm - Initial Values

We first fit a Generalize Linear Mixed Models(GLMM)

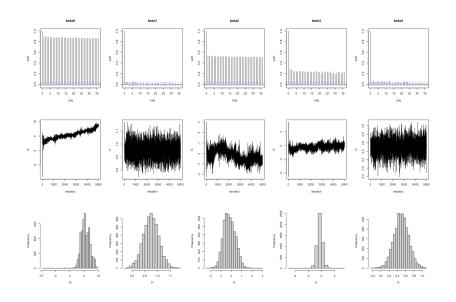
- ullet $eta_i^{(0)}$: The random effect for i^{th} hurricane from GLMM as start values
- ullet $\mu^{(0)}$: Average over $eta_i^{(0)}$
- ullet $\sigma^{(0)}$: Residuals from the GLMM
- ullet $\Sigma^{(0)}$: Variance-Covariance matrix of $eta_i^{(0)}$
- \bullet $\gamma^{(0)}$: Fixed effects from the GLMM

MCMC Results - Convergence Plots of B

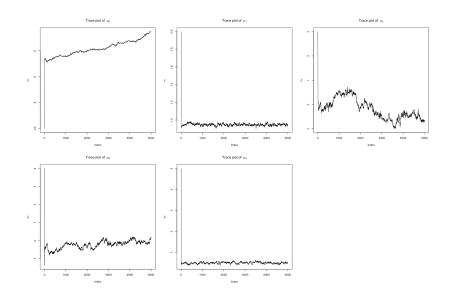
• Trace plots based on 5000 MCMC sample.



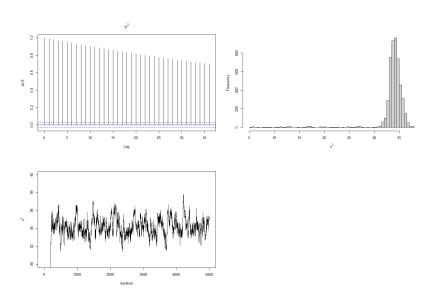
Selected B: Hurricane GEORGE.1951



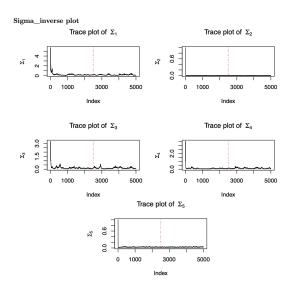
MCMC Results - Convergence Plots of μ



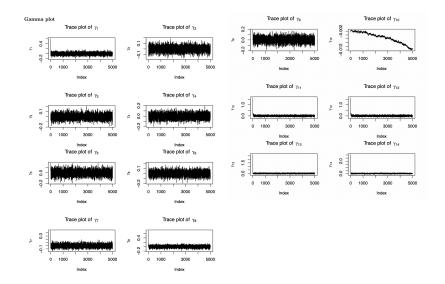
MCMC Results - Convergence and Distribution of σ^2



MCMC Results - Convergece Plots of Σ



MCMC Results - Convergece Plots of γ



95% credible intervals of Gamma

95%CI	2.5%	97.5%
April	-0.09922543	0.09022636
May	-0.10038368	0.09254019
June	-0.09761796	0.09437084
July	-0.09928455	0.10249011
August	-0.1028288	0.0916931
September	-0.0931095	0.1077383
October	-0.09762039	0.09269326
November	-0.1097137	0.1001810
December	-0.09799494	0.09315765
Year	-0.003099822	-0.001501638
TS	-0.08982878	0.10641711
ET	-0.10332349	0.09418438
SS	-0.09844162	0.09384343
NR	-0.09398186	0.09398186

Are there seasonal differences in hurricane wind speeds?

- Summer month: June, July, August.
- Non-summer month: April, May, September, October, November, December
- $\bullet \ \, H_0\hbox{:}\,\, \gamma_{summer}=0\hbox{; vs } H_1\hbox{:}\,\, \gamma_{summer}\neq 0$

$\overline{95\%CI}$	2.5%	97.5%
	-0.1382282	-0.1369877

ullet Fail to reject H_1 . No evidence to support the claim that there are seasonal differences in hurricane wind speeds.

Are hrricane wind speeds increasing over the years?

• H_0 : $\gamma_{10} \le 0$; vs H_1 : $\gamma_{10} > 0$

$\overline{95\%CI}$	2.5%	97.5%
	-0.003099822	-0.001501638

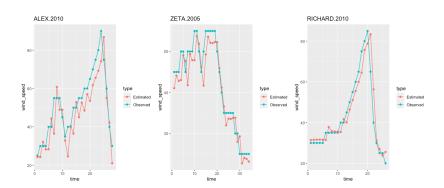
ullet Fail to reject $H_1.$ No evidence to support the claim that hurricane wind speeds have been increasing over the years.

Bayesian Model Performance

-The overall mean RMSE is 6.467.

	ID	r_square	rmse
1	SUBTROP:UNNAMED.1974	0.655	4.867
2	JEANNE.1980	0.921	5.437
3	FRANCES.2004	0.978	5.628
4	CHANTAL.1995	0.947	2.388
5	ETHEL.1960	0.473	27.218
6	PHILIPPE.2011	0.843	5.598
7	JOSEPHINE.1984	0.956	4.095
8	FRANCES.1976	0.895	6.114
9	BEULAH.1963	0.930	3.873
10	HOLLY.1969	0.873	5.670
11	ISAAC.2000	0.957	5.631
12	DAVID.1979	0.949	7.899
13	ALMA.1966	0.913	6.557
14	ERIN.1995	0.883	8.036
15	ANA.1997	0.880	2.156
16	DEBBIE.1969	0.851	8.869
17	HARVEY.2005	0.941	2.836
18	ALLISON.1995	0.768	4.339
19	LAURA.1971	0.967	2.112
20	EDNA.1968	0.957	2.006

Bayesian Model Performance



Estimated Wind Speed vs. Predicted Wind Speed

Limitations

- -Long running time for MCMC algorithm
- -Low performance on hurricanes without enough observations

Conclusion

- Our MCMC algorithm successfully estimates the high-dimensional parameters
 - All the parameters converges under a good initial values setting
 - ullet The overall R^2 is relatively large, and the overall RMSE is relatively small, so our model fits the data well
- There are no discernible variations between the months. The impact of the wind speed from six months ago on the current wind speed may gradually diminish over time.
- ullet When it comes to foretelling the harm and fatalities brought on by storms, the eta_i coefficients calculated from the Bayesian model are effective.

Reference

Taboga, Marco (2021). "Markov Chain Monte Carlo (MCMC) diagnostics", Lectures on probability theory and mathematical statistics. Kindle Direct Publishing. Online appendix.

Zhang, Z. (2021). A Note on Wishart and Inverse Wishart Priors for Covariance Matrix. Journal of Behavioral Data Science, 1(2), 119–126.

Q&A

Thank you for your attention. Any questions?