

Autocorrelation in Time Series

KNNL – Chapter 12

Issues in Autocorrelated Data

One of the standard assumptions in the regression model is that the error terms ε_i and ε_j are uncorrelated for $i \neq j$. When error terms are correlated, problems occur when using OLS estimates:

- Regression Coefficients are still unbiased
- MSE *underestimates* σ^2
- Standard errors of regression coefficients based on OLS *underestimate* the true standard error
- Inflated t and F statistics and *artificially narrow* confidence intervals

Autocorrelated errors of 1st order:

$$\varepsilon_t = \rho \varepsilon_{t-1} + u_t$$

where u_t are uncorrelated disturbances assumed normally distributed.

First-Order Model - I

First-Order Autoregressive Model ($AR(1)$):

Simple Regression: $Y_t = \beta_0 + \beta_1 X_t + \varepsilon_t \quad t = 1, \dots, n \quad \varepsilon_t = \rho \varepsilon_{t-1} + u_t$

$\rho \equiv$ autoregression parameter with $|\rho| < 1$

$u_t \sim N(0, \sigma^2)$ and independent

Generalizes to Multiple Regression:

$Y_t = \beta_0 + \beta_1 X_{t1} + \dots + \beta_{p-1} X_{t,p-1} + \varepsilon_t \quad t = 1, \dots, n \quad \varepsilon_t = \rho \varepsilon_{t-1} + u_t$

Properties of Errors (assumption regarding ε_1 for model consistency):

$$\varepsilon_1 \sim N\left(0, \frac{\sigma^2}{1-\rho^2}\right)$$

$$\varepsilon_2 = \rho \varepsilon_1 + u_2 \Rightarrow E\{\varepsilon_2\} = \rho E\{\varepsilon_1\} + E\{u_2\} = 0 \quad \sigma^2\{\varepsilon_2\} = \rho^2 \sigma^2\{\varepsilon_1\} + \sigma^2\{u_2\} = \rho^2 \left(\frac{\sigma^2}{1-\rho^2}\right) + \sigma^2 = \frac{\sigma^2}{1-\rho^2}$$

$$\text{Covariance: } \sigma\{\varepsilon_2, \varepsilon_1\} = \sigma\{\rho \varepsilon_1 + u_2, \varepsilon_1\} = \rho \sigma^2\{\varepsilon_2\} + \sigma\{u_2, \varepsilon_1\} = \rho \sigma^2\{\varepsilon_2\} + 0 = \frac{\rho \sigma^2}{1-\rho^2}$$

$$\text{Correlation: } \rho\{\varepsilon_2, \varepsilon_1\} = \frac{\sigma\{\varepsilon_2, \varepsilon_1\}}{\sigma\{\varepsilon_2\} \sigma\{\varepsilon_1\}} = \frac{\frac{\rho \sigma^2}{1-\rho^2}}{\sqrt{\frac{\sigma^2}{1-\rho^2}} \sqrt{\frac{\sigma^2}{1-\rho^2}}} = \rho$$

First-Order Model - II

In General:

$$\varepsilon_t = \rho\varepsilon_{t-1} + u_t = \rho(\rho\varepsilon_{t-2} + u_{t-1}) + u_t = \rho^2\varepsilon_{t-2} + \rho u_{t-1} + u_t = \dots = \sum_{s=0}^{\infty} \rho^s u_{t-s}$$

$$E\{\varepsilon_t\} = 0 \quad \sigma^2\{\varepsilon_t\} = \sigma^2 \left\{ \sum_{s=0}^{\infty} \rho^s u_{t-s} \right\} = \sum_{s=0}^{\infty} \rho^{2s} \sigma^2\{u_{t-s}\} = \sigma^2 \sum_{s=0}^{\infty} \rho^{2s} = \frac{\sigma^2}{1-\rho^2}$$

$$\text{Covariance: } \sigma\{\varepsilon_t, \varepsilon_{t-s}\} = \frac{\rho^s \sigma^2}{1-\rho^2} \quad s \geq 0$$

$$\text{Correlation: } \rho\{\varepsilon_t, \varepsilon_{t-s}\} = \rho^s \quad s \geq 0$$

$$\sigma^2\{\varepsilon\} = \frac{\sigma^2}{1-\rho^2} \begin{bmatrix} 1 & \rho & \rho^2 & \dots & \rho^{n-1} \\ \rho & 1 & \rho & \dots & \rho^{n-2} \\ \rho^2 & \rho & 1 & \dots & \rho^{n-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \rho^{n-1} & \rho^{n-2} & \rho^{n-3} & \dots & 1 \end{bmatrix}$$

$AR(2)$: $\varepsilon_t = \rho_1\varepsilon_{t-1} + \rho_2\varepsilon_{t-2} + u_t$ Even Higher order models can be fit as well.

Test For Independence : Durbin-Watson Test

$$Y_t = \beta_0 + \beta_1 X_{t1} + \dots + \beta_{p-1} X_{t,p-1} + \varepsilon_t \quad \varepsilon_t = \rho \varepsilon_{t-1} + u_t \quad u_t \sim NID(0, \sigma^2) \quad |\rho| < 1$$

$H_0 : \rho = 0 \Rightarrow$ Errors are uncorrelated over time

$H_A : \rho > 0 \Rightarrow$ Positively correlated

1) Obtain Residuals from Regression

2) Compute Durbin-Watson Statistic (given below)

3) Obtain Critical Values from Table B.7, pp. 1330-1331 (R will provide a p-value)

If $DW < d_L(p-1, n) \Rightarrow$ Reject H_0 If $DW > d_U(p-1, n) \Rightarrow$ Conclude H_0 Otherwise Inconclusive

$$\text{Test Statistic: } DW = \frac{\sum_{t=2}^n (e_t - e_{t-1})^2}{\sum_{t=1}^n e_t^2}$$

$$E\{\varepsilon_t\} = 0 \quad E\{\varepsilon_t \varepsilon_{t-1}\} = \frac{\rho \sigma^2}{1 - \rho^2}$$

$$\Rightarrow \sum_{t=2}^n (e_t - e_{t-1})^2 = \sum_{t=2}^n e_t^2 + \sum_{t=2}^n e_{t-1}^2 - 2 \sum_{t=2}^n e_t e_{t-1} \approx 2 \sum_{t=1}^n e_t^2 - 2n \frac{\rho \sigma^2}{1 - \rho^2}$$

\Rightarrow Under H_0 , expect $DW \approx 2$

Autocorrelation - Remedial Measures

- Determine whether a missing predictor variable can explain the autocorrelation in the errors
- Include a linear (trend) term if the residuals show a consistent increasing or decreasing pattern
- Include seasonal dummy variables if data are quarterly or monthly and residuals show cyclic behavior
- Use transformed Variables that remove the (estimated) autocorrelation parameter (Cochrane-Orcutt and Hildreth-Lu Procedures)
- Use First Differences
- Estimated Generalized Least Squares

Transformed Variables

Suppose ρ is known: $Y_t = \beta_0 + \beta_1 X_t + \varepsilon_t$ $\varepsilon_t = \rho \varepsilon_{t-1} + u_t$

$$\text{Let } Y'_t = Y_t - \rho Y_{t-1} = (\beta_0 + \beta_1 X_t + \varepsilon_t) - \rho(\beta_0 + \beta_1 X_{t-1} + \varepsilon_{t-1}) = \\ \beta_0(1 - \rho) + \beta_1(X_t - \rho X_{t-1}) + (\varepsilon_t - \rho \varepsilon_{t-1}) = \beta_0(1 - \rho) + \beta_1(X_t - \rho X_{t-1}) + u_t$$

$\Rightarrow Y'_t = \beta'_0 + \beta'_1 X'_t + u_t$ (Standard Simple linear regression with independent errors)

where:

$$Y'_t = Y_t - \rho Y_{t-1} \quad X'_t = X_t - \rho X_{t-1} \quad \beta'_0 = \beta_0(1 - \rho) \quad \beta'_1 = \beta_1$$

In Practice, we need to estimate ρ with a sample based value r

$$Y'_t = Y_t - r Y_{t-1} \quad X'_t = X_t - r X_{t-1}$$

Fit: $\hat{Y}' = b'_0 + b'_1 X'$ and if errors are uncorrelated, back transform to:

$$\hat{Y} = b_0 + b_1 X \text{ where: } b_0 = \frac{b'_0}{1 - r} \quad s\{b_0\} = \frac{s\{b'_0\}}{1 - r} \quad b_1 = b'_1 \quad s\{b_1\} = s\{b'_1\}$$

Cochrane-Orcutt Method

- Start by estimating ρ in Model: $\varepsilon_t = \rho\varepsilon_{t-1} + u_t$ by regression through the origin for residuals (see below)
- Fit transformed regression model (previous slide)
- Check to see if new residuals are uncorrelated (Durbin-Watson test), based on the transformed model
- If uncorrelated, stop and keep current model
- If correlated, repeat process with new estimate r based on current regression residuals from the original (back transformed) model

$$r = \frac{\sum_{t=2}^n e_{t-1}e_t}{\sum_{t=2}^n e_{t-1}^2}$$

Hildreth-Lu and First Difference Methods

- Hildreth-Lu Method
 - Find value of r (between 0 and 1) that minimizes the SSE for the transformed model by grid search
 - Apply the transformed analysis based on the estimated r
- First Differences Method
 - Uses $\rho = 1$ in transformed model ($Y'_t = Y_t - Y_{t-1}$ $X'_t = X_t - X_{t-1}$)
 - Set $b'_0 = 0$ and fits regression through origin of Y' on X'
 - When back-transforming:

$$b_0 = \bar{Y} - b'_1 \bar{X} \quad b_1 = b'_1$$

Forecasting with Autocorrelated Errors

Makes use of any of the 3 estimation techniques (C-O, H-L, First Differences):

$$Y_t = \beta_0 + \beta_1 X_t + \varepsilon_t \quad \varepsilon_t = \rho \varepsilon_{t-1} + u_t$$

$$\Rightarrow Y_t = \beta_0 + \beta_1 X_t + \rho \varepsilon_{t-1} + u_t \Rightarrow Y_{n+1} = \beta_0 + \beta_1 X_{n+1} + \rho \varepsilon_n + u_{n+1}$$

3 Elements:

1. Expected Value: $\beta_0 + \beta_1 X_{n+1}$ Estimated as $\hat{Y}_{n+1} = b_0 + b_1 X_{n+1}$
2. Multiple of period n Error Term: $\rho \varepsilon_n$ Estimated as re_n
3. Current disturbance $u_{n+1} \sim N(0, \sigma^2)$

Forecast for period $n+1$ (note the notation is "Forecast", not F -distribution :

$$\hat{F}_{n+1} = \hat{Y}_{n+1} + re_n$$

Standard Error of the Prediction (based on transformed model):

$$s^2 \{\text{pred}\} = MSE' \left[1 + \frac{1}{n-1} + \frac{(X'_{n+1} - \overline{X'})^2}{\sum_{i=2}^n (X'_i - \overline{X'})^2} \right]$$

Approximate 95% PI: $\hat{F}_{n+1} \pm t(1 - (\alpha/2); n-3) s \{\text{pred}\}$ (First Differences has $n-2$ df)