

# Bayesian Synthetic Control with a Soft Simplex Constraint

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# Overview

- The Synthetic Control Method (SCM) has at least two limitations;
  - High-dimensional problem: the number of control units needs to be much smaller than the total observed time periods before the intervention;
  - Debate over the simplex constraint: all coefficients are restricted to be non-negative and sum-to-one.
- We propose a novel extension of SCM, using Bayesian variable selection (spike-and-slab prior) with a soft simplex constraint, BVS-SS;
  - Contribution 1: BVS-SS solves the high dimension problem while still preserves the simplex constraint;
  - Contribution 2: BVS-SS introduces a notion of “soft” simplex constraint, asking data’s advice on whether the constraint is proper.

# SCM Setup

- $T_0$  is # of pre-treatment periods;  $T_1$  is # of post-treatment periods;
- $X_{1,t}$  is the outcome of the treated unit;  $X_{0,t}$  is the  $N \times 1$  outcome vector of the control units;

$$\hat{w} = \underset{w \in \Delta^{N-1}}{\operatorname{argmin}} \frac{1}{T_0} \sum_{t=1}^{T_0} (X_{1,t} - X_{0,t}^\top w)^2, \quad (1)$$

$$\Delta^{N-1} = \left\{ u \in \mathbb{R}^N : u_i \geq 0 \text{ for each } i, \text{ and } \sum_{i=1}^N u_i = 1 \right\}$$

where  $\Delta^{N-1}$  denotes the  $(N-1)$ -simplex.

- The average treatment on the treated:

$$\widehat{\text{ATT}} = \frac{1}{T_1} \sum_{t=T_0+1}^T (X_{1,t} - X_{0,t}^\top \hat{w})$$

# Challenges

- How to solve the optimization (1) when  $X_0$  is high-dimensional ( $N \gg T_0$ )?
  - Panel data approach: Elastic net ([Doudchenko and Imbens, 2016](#)), Lasso ([Carvalho et al., 2018](#); [Hollingsworth and Wing, 2020](#)), forward selection ([Shi and Huang, 2023](#)), spike-and-slab ([Kim et al., 2020](#)) etc.;
  - Trade-off: dropping the simplex constraint;

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  - Trade-off: dropping the simplex constraint;
- Is the simplex constraint necessary?
  - Pros: no hedging (e.g.,  $w_1 = 300\%$  and  $w_2 = -250\%$ ); the counterfactual has economic interpretation;
  - Cons: too rigid (may have bad fitting); high-dimensional problem.

# Motivations

We want to propose a novel SCM framework that

- ① deals with the high-dimensional control group, while still keeping the simplex constraint (**NOT TRIVIAL!**);
- ② learns the attitude of data on the constraint case by case.

An intuitive way is to make a “**soft**” constraint for the selected control units, that is, the expectation of the selected weights satisfy the simplex constraint.

# Problem Setup

- Pre: outcome of the treated unit  $X_1 \in \mathbb{R}^{T_0}$ ; Outcome of the control units  $X_0 \in \mathbb{R}^{T_0 \times N}$  ( $N \gg T_0$ );
- Post:  $\tilde{X}_1 \in \mathbb{R}^{T_1}$ ;  $\tilde{X}_0 \in \mathbb{R}^{T_0 \times N}$ ;
- Treatment:  $\delta$ ;

$$X_1 = X_0 w + \epsilon;$$

$$\tilde{X}_1 = \tilde{X}_0 w + \delta + \tilde{\epsilon}.$$

# Problem Setup

Suppose the true model of the counterfactual is,

$$X_1 = X_{0,\gamma^*} w_{\gamma^*} + \epsilon.$$

- $\gamma \in \{0, 1\}^N$ :  $\gamma_i = 1$  if and only if  $w_i \neq 0$ ;
- $w_\gamma$ : the subvector of  $w$  with entries indexed by  $\{i: \gamma_i = 1\}$ ;

## Two concerns:

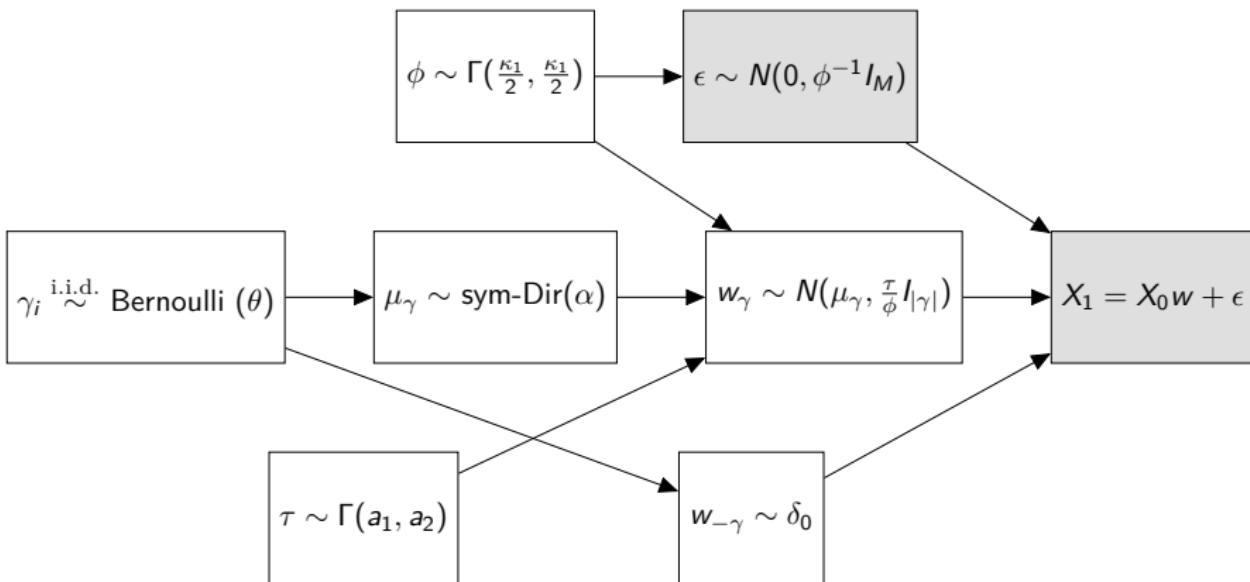
- **Variable selection:** estimate unknown  $\gamma^* \in \{0, 1\}^N$ ;
- **Soft constraint:**
  - Hard:  $w_\gamma \in \Delta^{|\gamma|-1}$ ;
  - Soft:  $\mathbb{E}[w_\gamma] \in \Delta^{|\gamma|-1}$ ;  $w_\gamma$  is random.

Bayesian estimation produces not just a single “point estimate” but an entire probability distribution for the parameter.

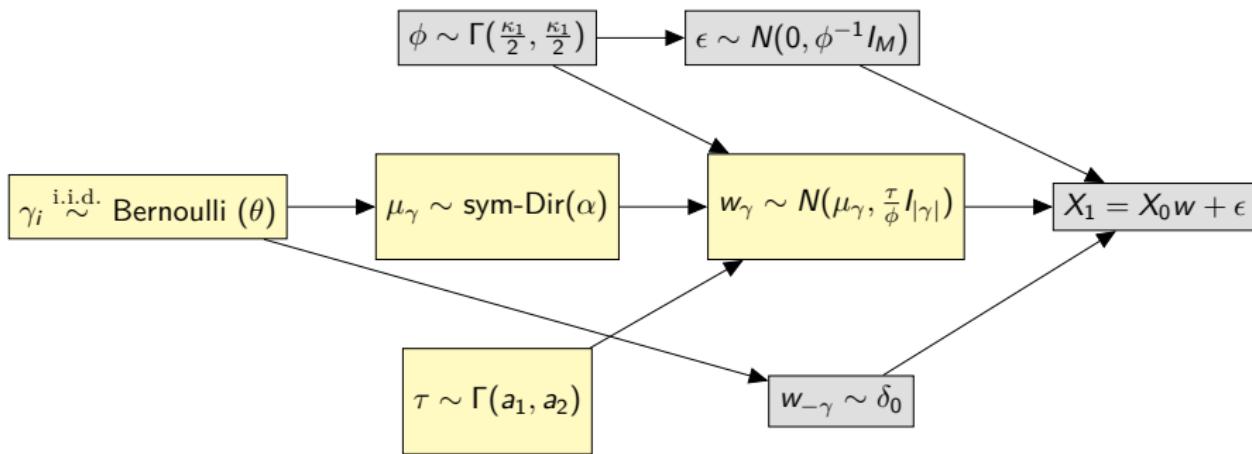
$$\underbrace{p(w_\gamma, \gamma, \phi | X_1)}_{\text{Posterior}} \propto \underbrace{f(X_1; w_\gamma, \gamma, \phi)}_{\text{Likelihood}} \underbrace{p(w_\gamma, \gamma, \phi)}_{\text{Prior}}$$

# Model and Priors: BVS-SS

BVS-SS: Bayesian variable selection with a soft simplex constraint.



# Model and Priors: BVS-SS



## Remarks:

- A small  $\theta$  penalizes the large models;
- Main diff from spike-and-slab:  $\mu_\gamma \in \Delta^{|\gamma|-1}$ ;
- Variance  $\tau$ : quantifying the deviation from simplex (mean  $\mu_\gamma$ ).

# Hindrance

- Once we have samples from the posterior  $p(\mu_\gamma, \gamma, \tau, \phi | X_1)$  ► Details, we can predict the counterfactual by

$$\mathbb{E}_p[\tilde{X}_0 w] = \mathbb{E}_p \left[ \mathbb{E}_p(\tilde{X}_0 w | \mu_\gamma, \gamma, \tau, \phi) \right]; \widehat{\text{ATT}} = \frac{1}{\tilde{T}} \sum_{t=1}^{\tilde{T}} \left( \tilde{X}_{1,t} - \mathbb{E}_p[\tilde{X}_0 w]_t \right).$$

- How to generate the samples from the posterior distribution  $p(\mu_\gamma, \gamma, \tau, \phi | X_1)$ ? The simplex constraint makes a Gibbs sampling non-trivial;
- What is the (non-asymptotic) property of the posterior distribution  $p(\mu_\gamma, \gamma, \tau, \phi | X_1)$ ?

# A Gibbs Scheme with the Simplex Constraint

How to generate the samples from the posterior distribution  $p(\mu_\gamma, \gamma, \tau, \phi | X_1)$ ?

- Gibbs sampling: Fix  $\mu_{-j}$ , update  $\mu_j$  according to  $p(\mu_j | Y, \mu_{-j}, \tau, \phi)$ ;
- $\mu_j = 1 - \sum_{i \neq j} \mu_i$  almost surely;  $p(\mu_j | Y, \mu_{-j}, \tau, \phi)$  is degenerate;
- Standard methods like [Geyer \(1992\)](#); [George and McCulloch \(1993\)](#); [Gilks et al. \(1995\)](#); [George and McCulloch \(1997\)](#); [Brooks \(1998\)](#); [Carlo \(2004\)](#) don't work.

Update  $\mu_i, \mu_j$  (with  $i \neq j$ ) simultaneously from the full conditional posterior  $p(\mu_i, \mu_j | X_1, \mu_{-(i,j)}, \tau, \phi)$ .

- ① Selection based on  $p(\gamma_i, \gamma_j | X_1, \mu_{-(i,j)}, \tau, \phi)$ ;
- ② Sampling based on  $p(\mu_i, \mu_j | X_1, \gamma_i, \gamma_j, \mu_{-(i,j)}, \tau, \phi)$

► Details →

# Non-asymptotic Bounds ( $\tau = 0$ )

Fix  $\tau = 0$ . The true model strictly satisfies the simplex constraint. We show the convergence rate of the posterior probability of selecting the true model  $\gamma^*$ .

## Theorem 1

Under mild conditions, it is true that

$$p(\gamma^* \mid Y) \geq 1 - c_3 N^{-1},$$

with probability at least  $1 - c_1 N^{-c_2 L}$ .

## Corollary 1

Under the setting of Theorem 1,

$$\mathbb{E} [\|\mu - \mu^*\|_2^2 \mid Y] = O_p \left( C_4 \frac{L \log N}{T_0} \right).$$

# More Results

- We also show the property when the simplex constraint is substantially violated ( $\tau \rightarrow \infty$ );
- Several simulation studies show the advantages of BVS-SS;
- Revisit an empirical example, with 87 control units but only 35 pre-intervention periods. [▶ Details →](#)

# Concluding Remarks

- We construct a novel Bayesian synthetic control framework to solve the high dimension issue (BVS-SS);
  - The posterior of the weight shows an empirical attitude on the simplex constraint;
  - [Li \(2020\)](#); [Chernozhukov et al. \(2021\)](#); [Martinez and Vives-i-Bastida \(2022\)](#); [Goh and Yu \(2022\)](#);
- We propose a novel Markov chain Monte Carlo (MCMC) sampler for our BVS-SS posterior;
  - to select the “proper” control units and estimate the weights and counterfactuals through our posterior;
  - [Geyer \(1992\)](#); [George and McCulloch \(1993\)](#); [Gilks et al. \(1995\)](#); [George and McCulloch \(1997\)](#); [Brooks \(1998\)](#); [Carlo \(2004\)](#);
- We derive non-asymptotic bounds for our Bayesian estimations;
  - on control units selection and posterior mean of coeffs/weights;
  - [George and McCulloch \(1993, 1997\)](#); [Brown et al. \(1998\)](#); [Li and Li \(2008\)](#); [Casella et al. \(2009\)](#); [Choi et al. \(2010\)](#); [Farcomeni \(2010\)](#); [Guan and Stephens \(2011\)](#); [Yang et al. \(2016\)](#); [Zhou and Guan \(2019\)](#).

*Thank You!*

# References I

- S. Brooks. Markov chain monte carlo method and its application. *Journal of the royal statistical society: series D (the Statistician)*, 47(1):69–100, 1998.
- P. J. Brown, M. Vannucci, and T. Fearn. Multivariate Bayesian variable selection and prediction. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 60(3):627–641, 1998.
- C. M. Carlo. Markov chain monte carlo and gibbs sampling. *Lecture notes for EEB*, 581(540):3, 2004.
- C. Carvalho, R. Masini, and M. C. Medeiros. Arco: An artificial counterfactual approach for high-dimensional panel time-series data. *Journal of econometrics*, 207(2):352–380, 2018.
- G. Casella, F. J. Girón, M. L. Martínez, and E. Moreno. Consistency of Bayesian procedures for variable selection. *The Annals of Statistics*, 37(3):1207–1228, 2009.
- V. Chernozhukov, K. Wüthrich, and Y. Zhu. An exact and robust conformal inference method for counterfactual and synthetic controls. *Journal of the American Statistical Association*, 116(536):1849–1864, 2021.

## References II

- N. H. Choi, W. Li, and J. Zhu. Variable selection with the strong heredity constraint and its oracle property. *Journal of the American Statistical Association*, 105(489):354–364, 2010.
- N. Doudchenko and G. W. Imbens. Balancing, regression, difference-in-differences and synthetic control methods: A synthesis. Technical report, National Bureau of Economic Research, 2016.
- A. Farcomeni. Bayesian constrained variable selection. *Statistica Sinica*, 20(3):1043–1062, 2010. URL <https://www.jstor.org/stable/24309479>.
- E. I. George and R. E. McCulloch. Variable selection via Gibbs sampling. *Journal of the American Statistical Association*, 88(423):881–889, 1993.
- E. I. George and R. E. McCulloch. Approaches for Bayesian variable selection. *Statistica sinica*, pages 339–373, 1997.
- C. J. Geyer. Practical markov chain monte carlo. *Statistical science*, pages 473–483, 1992.
- W. R. Gilks, S. Richardson, and D. Spiegelhalter. *Markov chain Monte Carlo in practice*. CRC press, 1995.

## References III

- G. Goh and J. Yu. Synthetic control method with convex hull restrictions: a bayesian maximum a posteriori approach. *The Econometrics Journal*, 25(1): 215–232, 2022.
- Y. Guan and M. Stephens. Bayesian variable selection regression for genome-wide association studies and other large-scale problems. *The Annals of Applied Statistics*, 5(3):1780, 2011.
- A. Hollingsworth and C. Wing. Tactics for design and inference in synthetic control studies: An applied example using high-dimensional data. Available at SSRN 3592088, 2020.
- S. Kim, C. Lee, and S. Gupta. Bayesian synthetic control methods. *Journal of Marketing Research*, 57(5):831–852, 2020.
- C. Li and H. Li. Network-constrained regularization and variable selection for analysis of genomic data. *Bioinformatics*, 24(9):1175–1182, 2008. doi: 10.1093/bioinformatics/btn081.

## References IV

- K. T. Li. Statistical inference for average treatment effects estimated by synthetic control methods. *Journal of the American Statistical Association*, 115(532):2068–2083, 2020.
- I. Martinez and J. Vives-i-Bastida. Bayesian and frequentist inference for synthetic controls. *arXiv preprint arXiv:2206.01779*, 2022. URL <https://arxiv.org/abs/2206.01779>.
- Z. Shi and J. Huang. Forward-selected panel data approach for program evaluation. *Journal of Econometrics*, 234(2):512–535, 2023.
- R. Tibshirani. Regression shrinkage and selection via the lasso. *Journal of the Royal Statistical Society Series B: Statistical Methodology*, 58(1):267–288, 1996.
- Y. Yang, M. J. Wainwright, and M. I. Jordan. On the computational complexity of high-dimensional Bayesian variable selection. *The Annals of Statistics*, 44(6):2497–2532, 2016.
- Q. Zhou and Y. Guan. Fast model-fitting of Bayesian variable selection regression using the iterative complex factorization algorithm. *Bayesian analysis*, 14(2):573, 2019.

## Motivation 1: High Dimensional $X$ (Contd)

To see why no frequentist implementations with simplex, consider the standard Lasso regression ([Tibshirani, 1996](#)).

$$\hat{w} = \underset{w}{\operatorname{argmin}} \frac{1}{T} \sum_{t=1}^T (X_{1,t} - X_{0,t}^\top w)^2 \text{ such that } \|w\|_1 \leq \lambda$$

where  $\lambda$  determines the degree of regularization.

The constraint in Lasso contradicts the simplex constraint  $\Delta^{N-1}$ !

▶ Back ←

# Posterior and ATT

A routine calculation using Woodbury identity yields the marginal likelihood of  $(\mu, \tau, \phi)$ :

$$p(X_1 | \mu, \tau, \phi) \propto \frac{\phi^{T/2}}{\tau^{|\gamma|/2} \det(V_{\gamma, \tau})^{1/2}} \exp \left\{ -\frac{\phi}{2} (X_1 - X_{0, \gamma} \mu_\gamma)^\top \Sigma_{\gamma, \tau} (X_1 - X_{0, \gamma} \mu_\gamma) \right\},$$

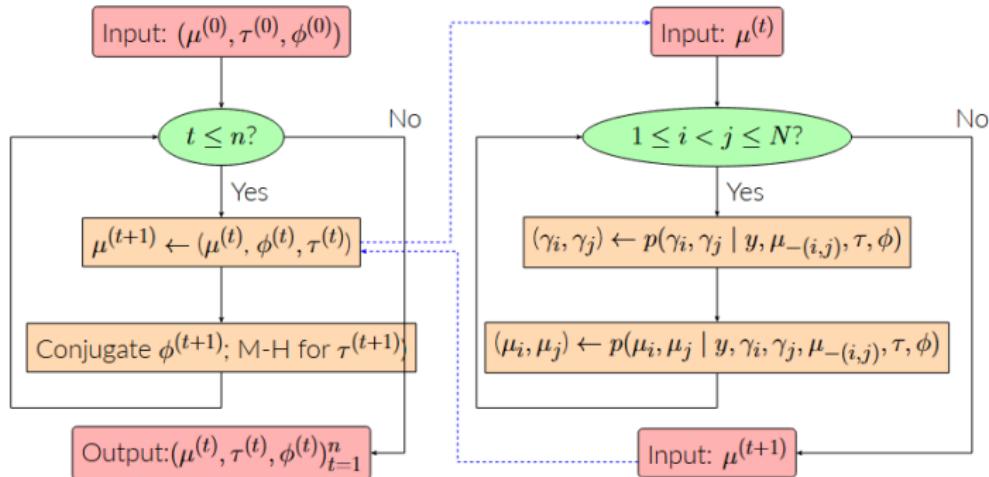
where

$$V_{\gamma, \tau} = X_{0, \gamma}^\top X_{0, \gamma} + \tau^{-1} I, \quad \Sigma_{\gamma, \tau} = I - X_{0, \gamma} V_{\gamma, \tau}^{-1} X_{0, \gamma}^\top.$$

w's conditional posterior:

$$w | X_1, \mu, \tau, \phi \sim N \left( V_{\gamma, \tau}^{-1} (X_{0, \gamma}^\top X_1 + \tau^{-1} \mu_\gamma), \phi^{-1} V_{\gamma, \tau}^{-1} \right).$$

## Metropolis-within-Gibbs



- When  $\alpha = 1$ ,  $\gamma$ 's conditional posterior can be expressed by CDF of standard normals; given  $\gamma_i = \gamma_j = 1$ ,  $\mu_i$  and  $\mu_j$ 's marginal posteriors degenerate to a **truncated Gaussian** distribution.

# China's anti-corruption campaign

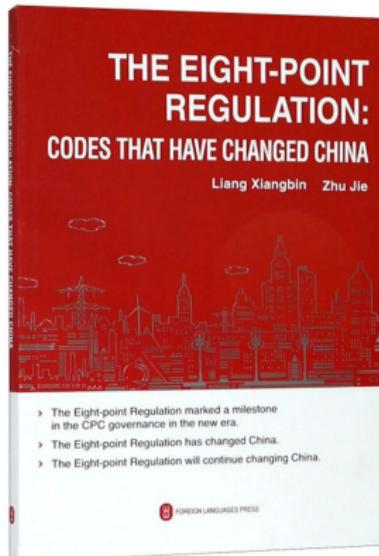


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Forward-selected panel data approach for program evaluation

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- Shi and Huang (2023): China anti-corruption campaign in 2012;
- $X_1$ : monthly growth rate of luxury watch imports;
- Category: “watches with case of, or clad with, precious metal”;
- $X_0$ : monthly growth rate of other 87 commodity categories;
- Only  $T_0 = 35$  pre-treatment periods.

# China's anti-corruption campaign

Table: Anti-corruption Campaign's Impact on Luxury Watches Importation

|                       | ATT              | $\tau$     | $\phi$         | $ \gamma $ |
|-----------------------|------------------|------------|----------------|------------|
| Mean                  | -0.021           | 0.069      | 20.86          | 5.09       |
| 95% credible interval | (-0.032, -0.008) | (0, 0.641) | (12.22, 32.76) | (1, 20)    |

**Remark:** The ratio  $\tau/\phi$  is very small → empirical support on the validity of SCM type estimator.

# Control Units Selection

- Shi and Huang (2023):

- ① knitted or crocheted fabric;
- ② cork and articles of cork;
- ③ salt, sulfur, earth, stone, plaster, lime and cement.

- BVS-SS:

- ① knitted or crocheted fabric (51.4%);
- ② special woven fabrics (lace, tapestries, trimmings, embroidery) (16.6%);
- ③ man-made filaments (14.8%);
- ④ wool, fine or coarse animal hair (12%);
- ⑤ other fabric/textile;
- ⑥ optical, photographic, cinematographic, measuring, checking, medical or surgical instruments and apparatus; parts and accessories" (10.2%).

**Remarks:** seems a more reasonable selection.

# Counterfactual Prediction

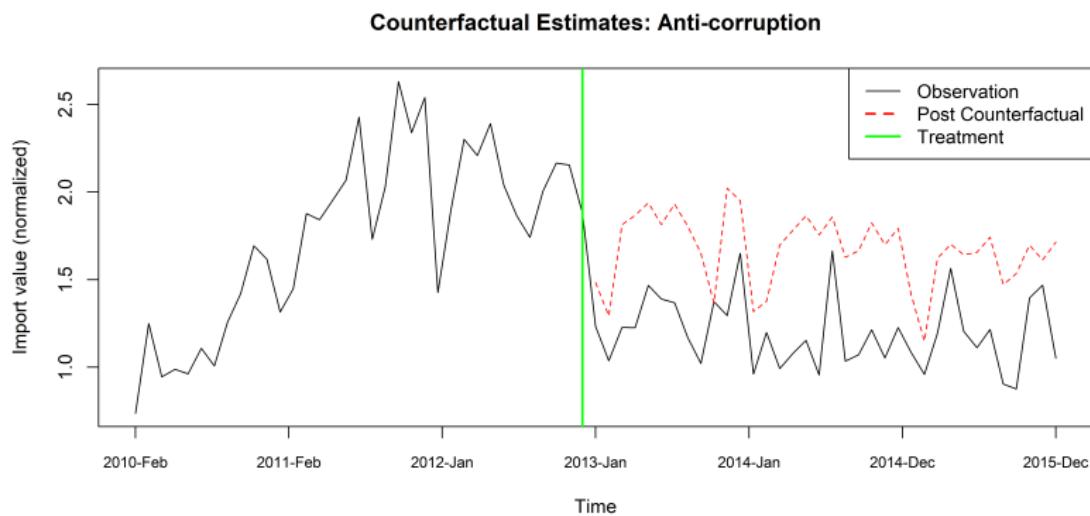


Figure: Significant impact right after the Eight-point regulation