

Bayesian Synthetic Control with a Soft Simplex Constraint

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Overview

- The Synthetic Control Method (SCM) has at least two limitations;
 - High-dimensional problem: the number of control units needs to be much smaller than the total observed time periods before the intervention;
 - Debate over the simplex constraint: all coefficients are restricted to be non-negative and sum-to-one.
- We propose a novel extension of SCM, using Bayesian variable selection (spike-and-slab prior) with a soft simplex constraint, BVS-SS;
 - Contribution 1: BVS-SS solves the high dimension problem while still preserves the simplex constraint;
 - Contribution 2: BVS-SS introduces a notion of “**soft**” simplex constraint, asking data’s advice on whether the constraint is proper.

SCM Setup

- T_0 is # of pre-treatment periods; T_1 is # of post-treatment periods;
- $X_{1,t}$ is the outcome of the treated unit; $X_{0,t}$ is the $N \times 1$ outcome vector of the control units;

$$\hat{w} = \operatorname{argmin}_{w \in \Delta^{N-1}} \frac{1}{T_0} \sum_{t=1}^{T_0} (X_{1,t} - X_{0,t}^\top w)^2, \quad (1)$$

$$\Delta^{N-1} = \left\{ u \in \mathbb{R}^N : u_i \geq 0 \text{ for each } i, \text{ and } \sum_{i=1}^N u_i = 1 \right\}$$

where Δ^{N-1} denotes the $(N-1)$ -simplex.

- The average treatment on the treated:

$$\widehat{\text{ATT}} = \frac{1}{T_1} \sum_{t=T_0+1}^T (X_{1,t} - X_{0,t}^\top \hat{w})$$

Challenges

- How to solve the optimization (1) when X_0 is high-dimensional ($N \gg T_0$)?
 - Panel data approach: Elastic net ([Doudchenko and Imbens, 2016](#)), Lasso ([Carvalho et al., 2018](#); [Hollingsworth and Wing, 2020](#)), forward selection ([Shi and Huang, 2023](#)), spike-and-slab ([Kim et al., 2020](#)) etc.;
 - Trade-off: dropping the simplex constraint;

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 - Trade-off: dropping the simplex constraint;
- Is the simplex constraint necessary?
 - Pros: no hedging (e.g., $w_1 = 300\%$ and $w_2 = -250\%$); the counterfactual has economic interpretation;
 - Cons: too rigid (may have bad fitting); high-dimensional problem.

Motivations

We want to propose a novel SCM framework that

- 1 deals with the high-dimensional control group, while still keeping the simplex constraint (**NOT TRIVIAL!**);
- 2 learns the attitude of data on the constraint case by case.

An intuitive way is to make a “**soft**” constraint for the selected control units, that is, the expectation of the selected weights satisfy the simplex constraint.

Problem Setup

- Pre: outcome of the treated unit $X_1 \in \mathbb{R}^{T_0}$; Outcome of the control units $X_0 \in \mathbb{R}^{T_0 \times N}$ ($N \gg T_0$);
- Post: $\tilde{X}_1 \in \mathbb{R}^{T_1}$; $\tilde{X}_0 \in \mathbb{R}^{T_0 \times N}$;
- Treatment: δ ;

$$X_1 = X_0 w + \epsilon;$$

$$\tilde{X}_1 = \tilde{X}_0 w + \delta + \tilde{\epsilon}.$$

Problem Setup

Suppose the true model of the counterfactual is,

$$X_1 = X_{0,\gamma^*} w_{\gamma^*} + \epsilon.$$

- $\gamma \in \{0, 1\}^N$: $\gamma_i = 1$ if and only if $w_i \neq 0$;
- w_γ : the subvector of w with entries indexed by $\{i: \gamma_i = 1\}$;

Two concerns:

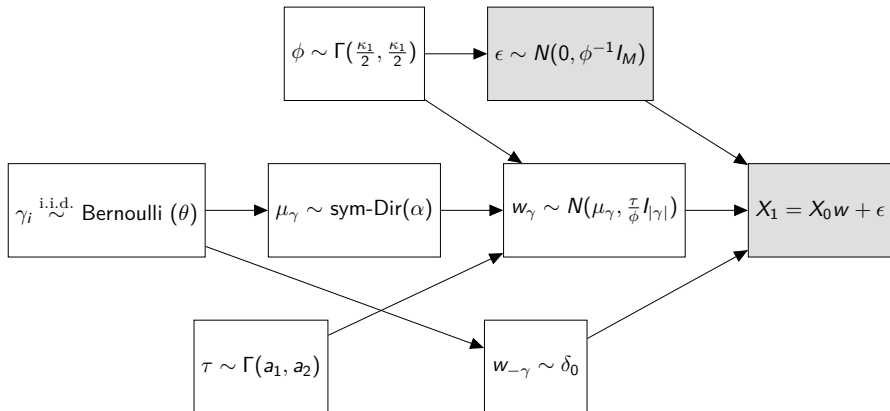
- **Variable selection:** estimate unknown $\gamma^* \in \{0, 1\}^N$;
- **Soft constraint:**
 - Hard: $w_\gamma \in \Delta^{|\gamma|-1}$;
 - Soft: $\mathbb{E}[w_\gamma] \in \Delta^{|\gamma|-1}$; w_γ is random.

Bayesian estimation produces not just a single “point estimate” but an entire probability distribution for the parameter.

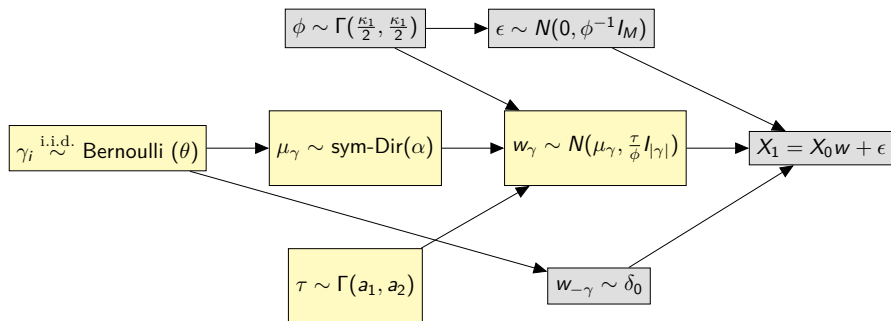
$$\underbrace{p(w_\gamma, \gamma, \phi \mid X_1)}_{\text{Posterior}} \propto \underbrace{f(X_1; w_\gamma, \gamma, \phi)}_{\text{Likelihood}} \underbrace{p(w_\gamma, \gamma, \phi)}_{\text{Prior}}$$

Model and Priors: BVS-SS

BVS-SS: Bayesian variable selection with a soft simplex constraint.



Model and Priors: BVS-SS



Remarks:

- A small θ penalizes the large models;
- Main diff from spike-and-slab: $\mu_\gamma \in \Delta^{|\gamma|-1}$;
- Variance τ : quantifying the deviation from simplex (mean μ_γ).

Hindrance

- Once we have samples from the posterior $p(\mu_\gamma, \gamma, \tau, \phi \mid X_1)$ [▶ Details](#), we can predict the counterfactual by

$$\mathbb{E}_p[\tilde{X}_0 w] = \mathbb{E}_p \left[\mathbb{E}_p(\tilde{X}_0 w \mid \mu_\gamma, \gamma, \tau, \phi) \right]; \widehat{ATT} = \frac{1}{\tilde{T}} \sum_{t=1}^{\tilde{T}} \left(\tilde{X}_{1,t} - \mathbb{E}_p[\tilde{X}_0 w]_t \right).$$

- How to generate the samples from the posterior distribution $p(\mu_\gamma, \gamma, \tau, \phi \mid X_1)$? The simplex constraint makes a Gibbs sampling non-trivial;
- What is the (non-asymptotic) property of the posterior distribution $p(\mu_\gamma, \gamma, \tau, \phi \mid X_1)$?

A Gibbs Scheme with the Simplex Constraint

How to generate the samples from the posterior distribution $p(\mu_\gamma, \gamma, \tau, \phi \mid X_1)$?

- Gibbs sampling: Fix μ_{-j} , update μ_j according to $p(\mu_j \mid Y, \mu_{-j}, \tau, \phi)$;
- $\mu_j = 1 - \sum_{i \neq j} \mu_i$ almost surely; $p(\mu_j \mid Y, \mu_{-j}, \tau, \phi)$ is degenerate;
- Standard methods like [Geyer \(1992\)](#); [George and McCulloch \(1993\)](#); [Gilks et al. \(1995\)](#); [George and McCulloch \(1997\)](#); [Brooks \(1998\)](#); [Carlo \(2004\)](#) don't work.

Update μ_i, μ_j (with $i \neq j$) simultaneously from the full conditional posterior $p(\mu_i, \mu_j \mid X_1, \mu_{-(i,j)}, \tau, \phi)$.

- 1 Selection based on $p(\gamma_i, \gamma_j \mid X_1, \mu_{-(i,j)}, \tau, \phi)$;
- 2 Sampling based on $p(\mu_i, \mu_j \mid X_1, \gamma_i, \gamma_j, \mu_{-(i,j)}, \tau, \phi)$

► Details →

Non-asymptotic Bounds ($\tau = 0$)

Fix $\tau = 0$. The true model strictly satisfies the simplex constraint. We show the convergence rate of the posterior probability of selecting the true model γ^* .

Theorem 1

Under mild conditions, it is true that

$$p(\gamma^* \mid Y) \geq 1 - c_3 N^{-1},$$

with probability at least $1 - c_1 N^{-c_2 L}$.

Corollary 1

Under the setting of Theorem 1,

$$\mathbb{E} [\|\mu - \mu^*\|_2^2 \mid Y] = O_p \left(c_4 \frac{L \log N}{T_0} \right).$$

More Results

- We also show the property when the simplex constraint is substantially violated ($\tau \rightarrow \infty$);
- Several simulation studies show the advantages of BVS-SS;
- Revisit an empirical example, with 87 control units but only 35 pre-intervention periods. [▶ Details →](#)

Concluding Remarks

- We construct a novel Bayesian synthetic control framework to solve the high dimension issue (BVS-SS);
 - The posterior of the weight shows an empirical attitude on the simplex constraint;
 - [Li \(2020\)](#); [Chernozhukov et al. \(2021\)](#); [Martinez and Vives-i-Bastida \(2022\)](#); [Goh and Yu \(2022\)](#);
- We propose a novel Markov chain Monte Carlo (MCMC) sampler for our BVS-SS posterior;
 - to select the “proper” control units and estimate the weights and counterfactuals through our posterior;
 - [Geyer \(1992\)](#); [George and McCulloch \(1993\)](#); [Gilks et al. \(1995\)](#); [George and McCulloch \(1997\)](#); [Brooks \(1998\)](#); [Carlo \(2004\)](#);
- We derive non-asymptotic bounds for our Bayesian estimations;
 - on control units selection and posterior mean of coeffs/weights;
 - [George and McCulloch \(1993, 1997\)](#); [Brown et al. \(1998\)](#); [Li and Li \(2008\)](#); [Casella et al. \(2009\)](#); [Choi et al. \(2010\)](#); [Farcomeni \(2010\)](#); [Guan and Stephens \(2011\)](#); [Yang et al. \(2016\)](#); [Zhou and Guan \(2019\)](#).

Thank You!

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Motivation 1: High Dimensional X (Contd)

To see why no frequentist implementations with simplex, consider the standard Lasso regression ([Tibshirani, 1996](#)).

$$\hat{w} = \underset{w}{\operatorname{argmin}} \frac{1}{T} \sum_{t=1}^T (X_{1,t} - X_{0,t}^\top w)^2 \quad \text{such that } \|w\|_1 \leq \lambda$$

where λ determines the degree of regularization.

The constraint in Lasso contradicts the simplex constraint Δ^{N-1} !

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Posterior and ATT

A routine calculation using Woodbury identity yields the marginal likelihood of (μ, τ, ϕ) :

$$p(X_1 \mid \mu, \tau, \phi) \propto \frac{\phi^{T/2}}{\tau^{|\gamma|/2} \det(V_{\gamma, \tau})^{1/2}} \exp \left\{ -\frac{\phi}{2} (X_1 - X_{0, \gamma} \mu_\gamma)^\top \Sigma_{\gamma, \tau} (X_1 - X_{0, \gamma} \mu_\gamma) \right\},$$

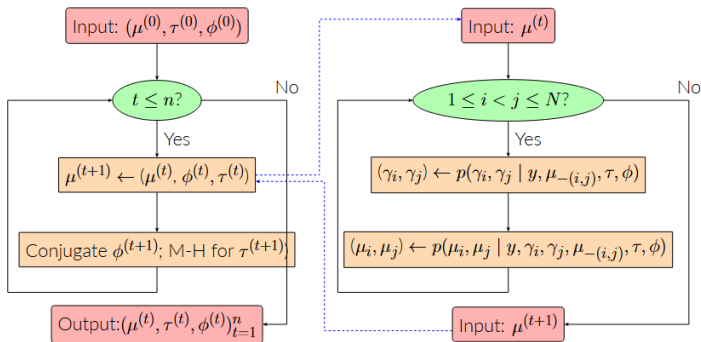
where

$$V_{\gamma,\tau} = X_{0,\gamma}^\top X_{0,\gamma} + \tau^{-1}I, \quad \Sigma_{\gamma,\tau} = I - X_{0,\gamma} V_{\gamma,\tau}^{-1} X_{0,\gamma}^\top.$$

w 's conditional posterior:

$$w \mid X_1, \mu, \tau, \phi \sim N(V_{\gamma, \tau}^{-1}(X_{0, \gamma}^\top X_1 + \tau^{-1} \mu_\gamma), \phi^{-1} V_{\gamma, \tau}^{-1}).$$

Metropolis-within-Gibbs



- When $\alpha = 1$, γ 's conditional posterior can be expressed by CDF of standard normals; given $\gamma_i = \gamma_j = 1$, μ_i and μ_j 's marginal posteriors degenerate to a **truncated Gaussian** distribution.

China's anti-corruption campaign

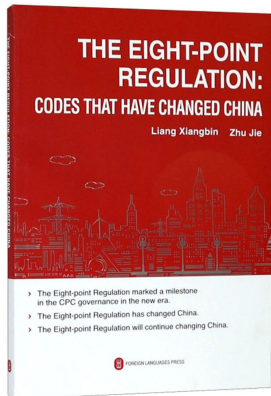


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Forward-selected panel data approach for program evaluation

Zhentao Shi ^a, Jingyi Huang ^b



- Shi and Huang (2023): China anti-corruption campaign in 2012;
- X_1 : monthly growth rate of luxury watch imports;
- Category: “watches with case of, or clad with, precious metal”;
- X_0 : monthly growth rate of other 87 commodity categories;
- Only $T_0 = 35$ pre-treatment periods.

China's anti-corruption campaign

Table: Anti-corruption Campaign's Impact on Luxury Watches Importation

	ATT	τ	ϕ	$ \gamma $
Mean	-0.021	0.069	20.86	5.09
95% credible interval	(-0.032, -0.008)	(0, 0.641)	(12.22, 32.76)	(1, 20)

Remark: The ratio τ/ϕ is very small \rightarrow empirical support on the validity of SCM type estimator.

Control Units Selection

- Shi and Huang (2023):

- ① knitted or crocheted fabric;
- ② cork and articles of cork;
- ③ salt, sulfur, earth, stone, plaster, lime and cement.

- BVS-SS:

- ① knitted or crocheted fabric (51.4%);
- ② special woven fabrics (lace, tapestries, trimmings, embroidery) (16.6%);
- ③ man-made filaments (14.8%);
- ④ wool, fine or coarse animal hair (12%);
- ⑤ other fabric/textile;
- ⑥ optical, photographic, cinematographic, measuring, checking, medical or surgical instruments and apparatus; parts and accessories" (10.2%).

Remarks: seems a more reasonable selection.

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Figure: Significant impact right after the Eight-point regulation