# MH3510 Regression Analysis



## **Group Project Report**

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## **Modelling Linear Regression Model with R**

## 1. Load Required Libraries

```
library(dplyr) library(caret)
```

## 2. Data Preparation

```
# Load raw data
data_raw <- read.table('aadt.txt', header = FALSE)

df <- data.frame(y = data_raw$V1, x1 = data_raw$V2, x2 = data_raw$V3, x3 = data_raw$V4, x4 = data_raw$V5)

# Convert x4 values of 2 to 0

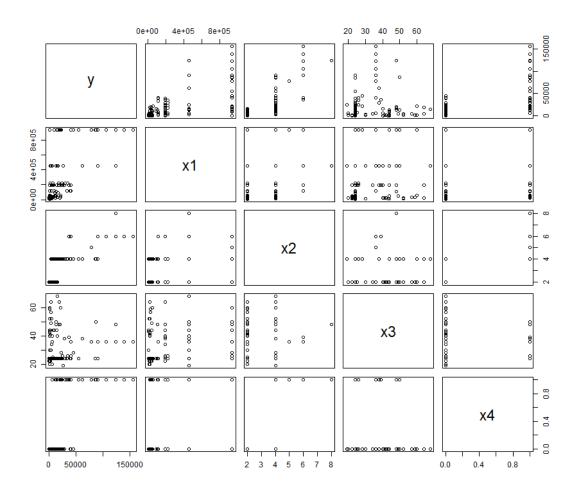
df$x4[df$x4 == 2] <- 0

# Scale the data (optional)

df_scaled <- df
```

## 3. Exploratory Data Analysis

# Graphical display of the observed data plot(df)



- y vs. x1: There seems to be some clustering or grouping in the data points. The relationship doesn't appear clearly linear.
- y vs. x2: The plot between y and x2 does not show a clear trend. The clustering observed here might imply the presence of some categorical elements or distinct groups within the dataset.
- y vs. x3: There appears to be minimal linear correlation between y and x3. The scatter is widespread, indicating that x3 might not be a significant predictor for y.
- y vs. x4: The points between y and x4 indicate a distinct grouping with many points overlapping at zero. This might indicate that x4 is binary or categorical.
- x1 vs. x2: There is no clear linear relationship between x1 and x2.
- x2 vs. x3: The scatter plot is widely spread, indicating that x2 and x3 are likely not correlated with each other.
- x3 vs. x4: There is significant correlation between x3 and x4.

## 4. Multiple Linear Regression Model

```
# Fit a multiple linear regression model
  mIr <- Im(y \sim x1 + x2 + x3 + x4, data = df)
  summary(mlr)
Call:
lm(formula = y \sim x1 + x2 + x3 + x4, data = df_scaled)
Residuals:
  Min
           10 Median
                          3Q
                                Max
-36263 -8501
                3493
                       6018 68317
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept) -2.604e+04 5.255e+03 -4.955 2.49e-06 ***
             3.303e-02 4.708e-03
                                     7.017 1.63e-10 ***
х1
             9.158e+03 1.531e+03
                                     5.983 2.49e-08 ***
x2
             1.003e+02 1.243e+02 0.807
                                              0.421
х3
             2.361e+04 4.520e+03 5.223 7.83e-07 ***
x4
                0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Signif. codes:
Residual standard error: 15290 on 116 degrees of freedom
                               Adjusted R-squared: 0.7442
Multiple R-squared: 0.7527,
F-statistic: 88.29 on 4 and 116 DF, p-value: < 2.2e-16
```

#### • Coefficients:

- $\circ$  Coefficient of **x3** has a p-value of 0.421 which is greater than the significance level of 0.05, thus H<sub>0</sub>:  $\beta_3$  = 0 is not rejected. Hence, **x3** is the least significant variable to the model in terms of t-test
- $\circ$  Coefficients of x1, x2 and x4 have p-values which are smaller than 0.05, thus H<sub>0</sub>:  $\beta_i = 0$  for i = 1, 2, 4 are rejected. Hence, x1, x2 and x4 are highly significant variables.

#### • R<sup>2</sup> Statistic:

- Multiple R-squared = 0.7527 indicates that approximately 75.27% of the variability in the response variable **y** is explained by the model.
- Adjusted R-squared = 0.7442 is slightly lower than the R-squared value, indicating that the model fits the data well, considering the number of predictors.

#### • F-statistic:

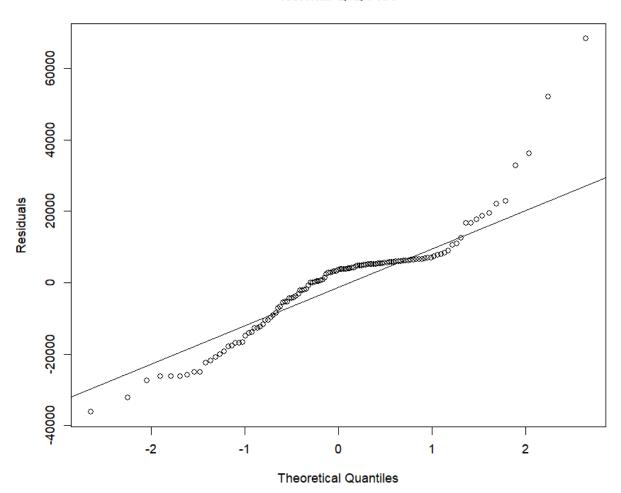
 88.29 on 4 and 116 degrees of freedom, with a p-value < 2.2e-16. This suggests that the model is significant overall, and at least one of the predictors is related to

У

## 5. Normality Check of Residuals

# Normality checking
qqnorm(residuals(mlr), ylab = 'Residuals')
qqline(residuals(mlr))

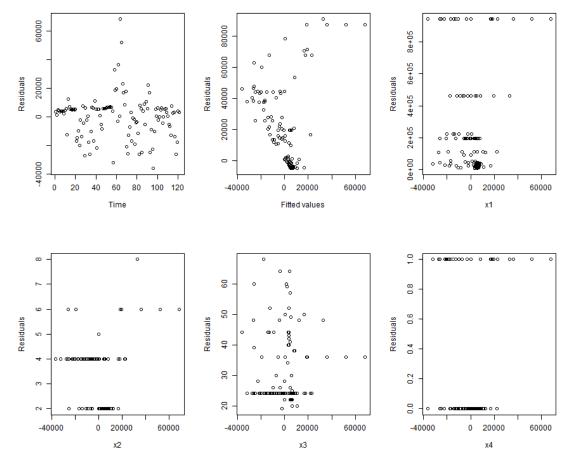
#### **Normal Q-Q Plot**



- Points generally follow the normal line on the QQ plot, except for some deviations in the upper tail.
- This suggests that the data mostly follows a normal distribution but has a few outliers in the upper tail.

## 6. Residual Analysis

```
# Draw some plots of residuals par(mfrow = c(2, 3)) plot(residuals(mlr), ylab = 'Residuals', xlab = 'Time') plot(residuals(mlr), fitted(mlr), ylab = 'Residuals', xlab = 'Fitted values') plot(residuals(mlr), df_scaled$x1, ylab = 'Residuals', xlab = 'x1') plot(residuals(mlr), df_scaled$x2, ylab = 'Residuals', xlab = 'x2') plot(residuals(mlr), df_scaled$x3, ylab = 'Residuals', xlab = 'x3') plot(residuals(mlr), df_scaled$x4, ylab = 'Residuals', xlab = 'x4') par(mfrow = c(1, 1))
```



- **Residuals vs. Time:** The points are somehow distributed around the zero line. There doesn't appear to be a strong trend over time, suggesting no clear temporal pattern.
- **Residuals vs. Fitted Values:** The plot shows a funnel-shaped pattern, which means that the variability of residuals is not constant across different levels of the fitted values.
- Residuals vs. Predictors (x1, x2, x3, x4): There are noticeable patterns and clusters, particularly for x2 and x3, which may indicate non-linearity. The residuals for x4 are clustered, suggesting x4 might be categorical

#### 7. Durbin-Watson Test

- DW = 1.3137 suggests positive autocorrelation of residuals, meaning that consecutive residuals are correlated in a positive manner
- The p-value is **3.101e-05**, which is very small. The null hypothesis for the Durbin-Watson test is that there is no autocorrelation in the residuals. Since the p-value is much smaller than typical significance levels, we reject the null hypothesis. This indicates that autocorrelation is present in the residuals.

## 8. Model Comparison with F-tests

```
I. # Model without x3
    mlr1 <- lm(y ~ x1 + x2 + x4, data = df)
    anova(mlr1, mlr)

    Analysis of Variance Table

    Model 1: y ~ x1 + x2 + x4
    Model 2: y ~ x1 + x2 + x3 + x4
    Res.Df    RSS Df Sum of Sq    F Pr(>F)
    1    117    2.7281e+10
    2    116    2.7128e+10    1    152302593    0.6512    0.4213
```

• The p-value is 0.4213. Since the p-value is greater than common significance levels, we fail to reject the null hypothesis, meaning that adding **x3** does not significantly improve the model.

```
II. # Model where \hat{\beta}_3 is constant using offset mlr3 <- lm(y ~ x1 + x2 + offset(100.3 * x3) + x4, data = df) summary(mlr3) anova(mlr3, mlr)

Analysis of Variance Table

Model 1: y ~ x1 + x2 + offset(100.3 * x3) + x4 Model 2: y ~ x1 + x2 + x3 + x4 Res.Df RSS Df Sum of Sq F Pr(>F) 1 117 2.7128e+10 2 116 2.7128e+10 1 1.831 0 0.9999
```

- From the results of the MLR, we get  $\hat{\beta}_3$  = 100.3
- To determine if setting a constant coefficient for x3 improves model fitting, we specified 100.3 as an offset for  $\beta_3$
- The p-value obtained is 0.9999. Since p-value is extremely high, we fail to reject the null hypothesis. It would be appropriate to treat **x3** as an offset or consider removing it altogether, as it does not add value in terms of improving model fit.

### 9. Prediction

- fit: The predicted value for the given input is 9106.94.
- lwr (Lower Bound): 1045.888 is the lower bound of the 95% confidence interval.
- upr (Upper Bound): 17167.99 is the upper bound of the 95% confidence interval.
- Interpretation: We are 95% confident that the mean response will fall between 1045.888 and 17167.99.

```
# Prediction interval prediction
predict(mlr, con, interval = 'prediction', level = 0.95)

fit | lwr | upr
1 9106.94 -22236.34 40450.22
```

- fit: The predicted value for the given input is 9106.94.
- lwr (Lower Bound): -22236.34 is the lower bound of the 95% prediction interval.
- upr (Upper Bound): 40450.22 is the upper bound of the 95% prediction interval.
- Interpretation: The prediction interval is very wide, ranging from -22236.34 to 40450.22, which suggests a lot of uncertainty in predicting individual responses.

### 10. Consideration of normalised data

```
# Normalised data
  process <- preProcess(df, method=c("range"))</pre>
  df scaled <- predict(process, df)</pre>
  mlr_scaled <- lm(y \sim x1+x2+x3+x4, data=df_scaled)
  summary(mlr scaled)
Call:
lm(formula = y \sim x1 + x2 + x3 + x4, data = df_scaled)
Residuals:
                    Median
     Min
               1Q
                                  3Q
                                          Max
-0.23343 -0.05472 0.02249 0.03874 0.43978
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.03708
                         0.01497 -2.477 0.0147 *
                                  7.017 1.63e-10 ***
x1
             0.19854
                         0.02830
x2
             0.35371
                         0.05912
                                  5.983 2.49e-08 ***
                                  0.807 0.4213
x3
             0.03163
                         0.03920
             0.15199
                         0.02910 5.223 7.83e-07 ***
x4
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '. ' 0.1 ' ' 1
Residual standard error: 0.09844 on 116 degrees of freedom
Multiple R-squared: 0.7527,
                                Adjusted R-squared: 0.7442
F-statistic: 88.29 on 4 and 116 DF, p-value: < 2.2e-16
```

• The result shows that the normalised data does not improve the model fitting. Therefore, normalised data is not necessary.

## 11. Annex A (source code)

```
library(dplyr)
library(caret)
# Graphical display of the observed data.
data raw <- read.table('aadt.txt', header=FALSE)</pre>
df <-
data.frame(y=data raw$V1,x1=data raw$V2,x2=data raw$V3,x3=data raw$V4,x4=data raw$V5)
df$x4[df$x4 == 2]<-0
process <- preProcess(df, method=c("range"))</pre>
df scaled <- predict(process, df)</pre>
plot(df)
plot(df scaled)
# Fit a multiple linear regression model for both scaled and non-scaled data.
mlr < -lm(y \sim x1+x2+x3+x4, data=df)
summary(mlr)
mlr scaled <- lm(y \sim x1+x2+x3+x4, data=df scaled)
summary(mlr scaled)
# Normality checking.
qqnorm(residuals(mlr),ylab='Residuals')
qqline(residuals(mlr))
qqnorm(residuals(mlr scaled),ylab='Residuals')
qqline(residuals(mlr scaled))
# Draw some plots of residuals.
par(mfrow=c(2,3))
plot(residuals(mlr),ylab='Residuals',xlab='Time')
plot(residuals(mlr), fitted(mlr), ylab='Residuals', xlab='Fitted values')
plot(residuals(mlr), df$x1, ylab='Residuals', xlab='x1')
plot(residuals(mlr),df$x2,ylab='Residuals',xlab='x2')
plot(residuals(mlr), df$x3, ylab='Residuals', xlab='x3')
plot(residuals(mlr),df$x4,ylab='Residuals',xlab='x4')
par(mfrow=c(1,1))
# scaled data
par(mfrow=c(2,3))
plot(residuals(mlr scaled),ylab='Residuals',xlab='Time')
plot(residuals(mlr scaled),fitted(mlr),ylab='Residuals',xlab='Fitted values')
plot(residuals(mlr scaled),df scaled$x1,ylab='Residuals',xlab='x1')
plot(residuals(mlr scaled),df scaled$x2,ylab='Residuals',xlab='x2')
plot(residuals(mlr scaled),df scaled$x3,ylab='Residuals',xlab='x3')
plot(residuals(mlr scaled),df scaled$x4,ylab='Residuals',xlab='x4')
par(mfrow=c(1,1))
# Durbin-Watson tests.
# install.packages( "lmtest" )
library(lmtest)
dwtest(y \sim x1+x2+x3+x4, data=df)
dwtest(y \sim x1+x2+x3+x4, data=df scaled)
```

```
# Some F-tests.
mlr1 < -lm(y \sim x1+x2+x4, data=df) #remove x3 as insignificant from above
anova(mlr1,mlr)
mlr1 scaled <- lm(y \sim x1+x2+x4, data=df scaled) #remove x3 as insignificant from above
anova(mlr1_scaled,mlr scaled)
mlr2 <- lm(y \sim x1+x2+offset(100.3*x3)+x4,data=df) # from MLR results, B3^=100.3
summary(mlr2)
anova(mlr2,mlr)
mlr2 scaled <- lm(y \sim x1+x2+offset(0.03*x3)+x4,data=df scaled) # from MLR(scaled)
results, B3^=0.03163
summary(mlr2 scaled)
anova(mlr2 scaled, mlr scaled)
# Predicting non-scaled inputs
con <- data.frame(x1=50000, x2=3, x3=60, x4=0)
predict(mlr,con,interval='confidence',level=0.95)
predict(mlr,con,interval='prediction',level=0.95)
# Predicting scaled inputs
# scaling new input
new input <- data.frame(y=-1,x1=50000,x2=3,x3=60,x4=0)
df input <- rbind(new input,df)</pre>
process <- preProcess(df input[c('x1','x2','x3','x4')], method=c("range"))</pre>
df input scaled <- predict(process, df input)</pre>
## scaled input: x1=0.045286737 ,x2=0.1666667 ,x3=0.83673469 ,x4=0
to predict <- data.frame(x1=0.045286737 ,x2=0.1666667 ,x3=0.83673469 ,x4=0)
predict(mlr scaled, to predict, interval='confidence', level=0.95)
predict(mlr scaled, to predict, interval='prediction', level=0.95)
predicted val <- (0.05732971*(max(df['y'])-min(df['y'])))+min(df['y'])
predicted lwr conf <- (0.00543876*(max(df['y'])-min(df['y'])))+min(df['y'])
predicted upr conf <- (0.1092207*(max(df['y'])-min(df['y'])))+min(df['y'])
predicted lwr pred <- (-0.1444346 * (max(df['y'])-min(df['y'])))+min(df['y'])
predicted upr pred <- (0.259094*(max(df['y'])-min(df['y'])))+min(df['y'])</pre>
conf<-data.frame(fit=predicted val,lwr=predicted lwr conf,upr=predicted upr conf)</pre>
pred<-data.frame(fit=predicted val,lwr=predicted lwr pred,upr=predicted upr pred)
conf
pred
```