

Assign.6_YM

May 28, 2018

@author: Yiming

```
In [1]: import pandas as pd
import numpy as np
import statsmodels.api as sm
from patsy import dmatrices
import matplotlib.pyplot as plt
%matplotlib inline
```

```
/Users/yimingcai/anaconda/lib/python3.6/site-packages/statsmodels/compat/pandas.py:56: FutureWarning
from pandas.core import datetools
```

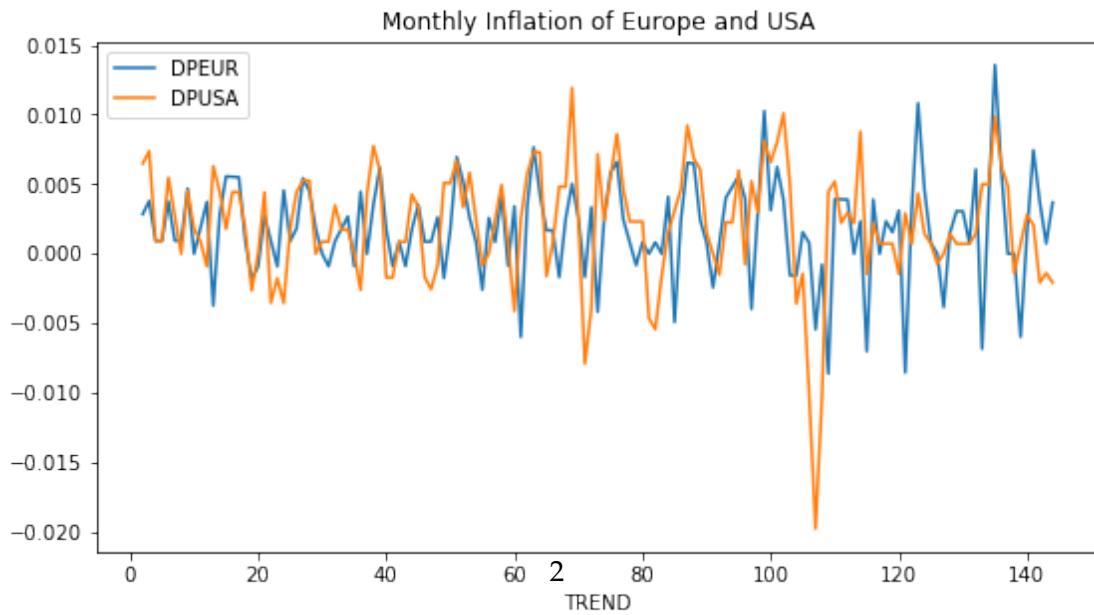
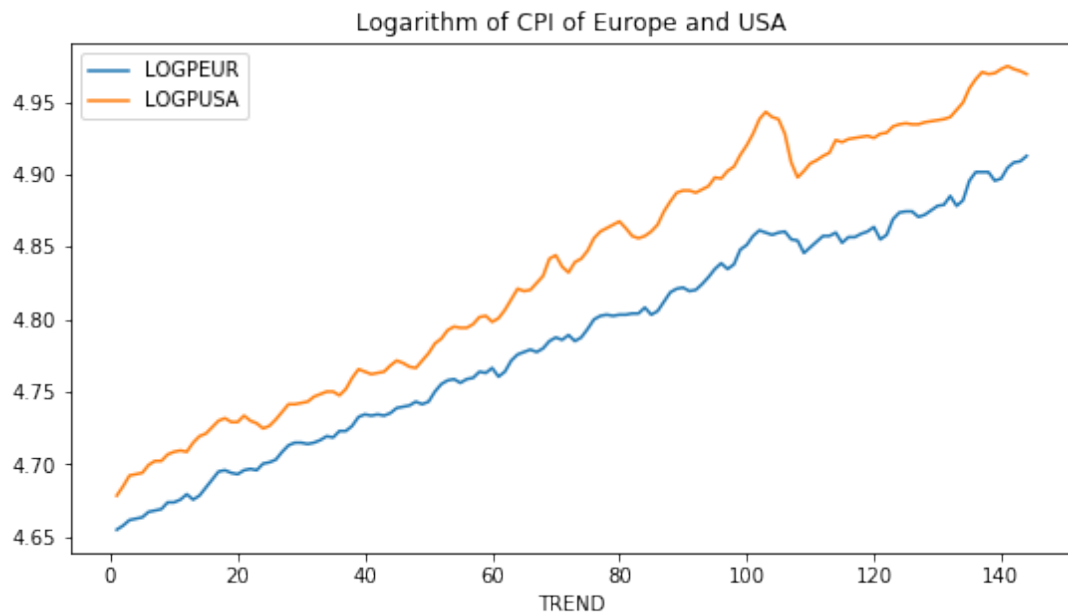
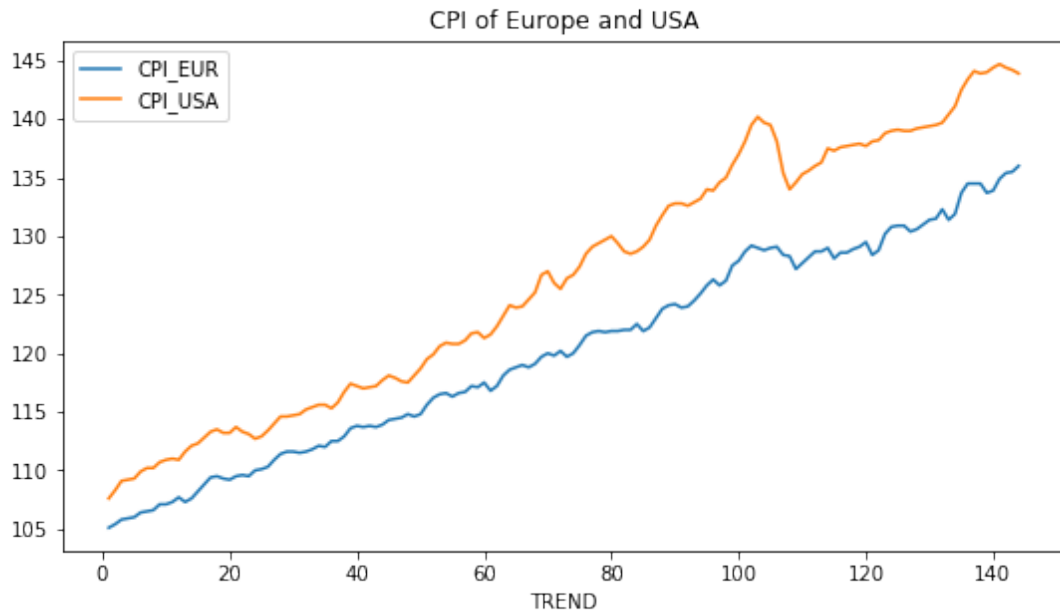
```
In [2]: df = pd.read_excel("Test6_data.xlsx")
```

(a) Make time series plots of the CPI of the Euro area and the USA, and also of their logarithm $\log(\text{CPI})$ and of the two monthly inflation series $\text{DP} = \log(\text{CPI})$. What conclusions do you draw from these plots?

```
In [3]: fig, ax = plt.subplots(3, 1, figsize = (9, 16))
df_plot = df.set_index("TREND").copy()
df_plot[["CPI_EUR", "CPI_USA"]].plot(ax = ax[0])
df_plot[["LOGPEUR", "LOGPUSA"]].plot(ax = ax[1])
df_plot[["DPEUR", "DPUSA"]].plot(ax = ax[2])

ax[0].set_title("CPI of Europe and USA")
ax[1].set_title("Logarithm of CPI of Europe and USA")
ax[2].set_title("Monthly Inflation of Europe and USA")
```

```
Out[3]: <matplotlib.text.Text at 0x11dff2668>
```



Conclusion:

1. There exists a deterministic trend for CPI in both Europe and USA, so as their logarithm terms.
2. The CPI for both Europe and USA are non-stationary as their means varies with time.
3. The inflation rate, however, seems stationary and thus can be used for time-series analysis.
4. The inflation rate of Europe and USA seem correlated.

(b) Perform the Augmented Dickey-Fuller (ADF) test for the two $\log(\text{CPI})$ series. In the ADF test equation, include a constant (α), a deterministic trend term (β), three lags of $\text{DP} = \log(\text{CPI})$ and, of course, the variable of interest $\log(\text{CPI})_t$. Report the coefficient of $\log(\text{CPI})_t$ and its standard error and t-value, and draw your conclusion.

```
In [4]: df_lagged1_term = df[["LOGPEUR", "LOGPUSA", "DPEUR", "DPUSA"]].shift(1).rename(columns=

df_lagged2_term = df[["DPEUR", "DPUSA"]].shift(2).rename(columns= {"DPEUR": "DPEUR_L2",
                                                                    "DPUSA": "DPUSA_L2"})

df_lagged3_term = df[["DPEUR", "DPUSA"]].shift(3).rename(columns = {"DPEUR": "DPEUR_L3",
                                                                    "DPUSA": "DPUSA_L3"})

df_lagged = pd.concat([df, df_lagged1_term, df_lagged2_term, df_lagged3_term], axis =1)

In [5]: y_eur, X_eur = dmatrices("DPEUR ~ TREND+LOGPEUR_L1 +DPEUR_L1+DPEUR_L2+DPEUR_L3", df_lag
y_usa, X_usa = dmatrices("DPUSA ~TREND+LOGPUSA_L1+DPUSA_L1+DPUSA_L2+DPUSA_L3" , df_lag

In [6]: eur_mod = sm.OLS(y_eur, X_eur).fit()
        usa_mod = sm.OLS(y_usa, X_usa).fit()

In [7]: print (eur_mod.summary())
```

```

                        OLS Regression Results
=====
Dep. Variable:          DPEUR      R-squared:            0.120
Model:                  OLS        Adj. R-squared:       0.087
Method:                 Least Squares    F-statistic:      3.662
Date:                   Mon, 28 May 2018    Prob (F-statistic): 0.00388
Time:                   11:49:13      Log-Likelihood:    601.82
No. Observations:       140          AIC:               -1192.
Df Residuals:           134          BIC:               -1174.
```

```

Df Model:                    5
Covariance Type:            nonrobust
=====
              coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept      0.6420      0.226      2.837      0.005      0.194      1.090
TREND           0.0002     8.5e-05      2.795      0.006     6.94e-05      0.000
LOGPEUR_L1     -0.1374      0.049     -2.826      0.005     -0.234     -0.041
DPEUR_L1        0.1442      0.087      1.665      0.098     -0.027      0.316
DPEUR_L2       -0.0902      0.085     -1.059      0.292     -0.259      0.078
DPEUR_L3       -0.1128      0.086     -1.317      0.190     -0.282      0.057
=====
Omnibus:                17.430   Durbin-Watson:                2.029
Prob(Omnibus):           0.000   Jarque-Bera (JB):         28.871
Skew:                    -0.606   Prob(JB):                 5.38e-07
Kurtosis:                 4.865   Cond. No.                 7.04e+04
=====

```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
[2] The condition number is large, 7.04e+04. This might indicate that there are strong multicollinearity or other numerical problems.

```
In [8]: print (usa_mod.summary())
```

```

                                OLS Regression Results
=====
Dep. Variable:                  DPUSA      R-squared:                0.326
Model:                          OLS      Adj. R-squared:            0.301
Method:                        Least Squares      F-statistic:              12.97
Date:                          Mon, 28 May 2018      Prob (F-statistic):      2.72e-10
Time:                          11:49:13      Log-Likelihood:          595.89
No. Observations:                140      AIC:                     -1180.
Df Residuals:                    134      BIC:                     -1162.
Df Model:                        5
Covariance Type:                nonrobust
=====
              coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept      0.3494      0.127      2.747      0.007      0.098      0.601
TREND           0.0002     5.72e-05      2.645      0.009     3.82e-05      0.000
LOGPUSA_L1     -0.0743      0.027     -2.734      0.007     -0.128     -0.021
DPUSA_L1        0.6091      0.084      7.248      0.000      0.443      0.775
DPUSA_L2       -0.1513      0.096     -1.568      0.119     -0.342      0.040
DPUSA_L3       -0.0064      0.086     -0.075      0.941     -0.177      0.164
=====
Omnibus:                6.073   Durbin-Watson:                1.993

```

Prob(Omnibus):	0.048	Jarque-Bera (JB):	7.828
Skew:	-0.228	Prob(JB):	0.0200
Kurtosis:	4.065	Cond. No.	3.82e+04

=====

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
[2] The condition number is large, 3.82e+04. This might indicate that there are strong multicollinearity or other numerical problems.

As the results indicate, for both variables, the ADF statistic is greater than the critical value of 3.5. Therefore, the non-stationarity hypothesis is not rejected.

(c) As the two series of log(CPI) are not cointegrated (you need not check this), we continue by modelling the monthly inflation series $DPEUR = \log(CPIEUR)$ for the Euro area. Determine the sample autocorrelations and the sample partial autocorrelations of this series to motivate the use of the following AR model: $DPEUR_t = +1DPEUR_{t6} + 2DPEUR_{t12} + t$. Estimate the parameters of this model (sample Jan 2000 - Dec 2010).

```
In [9]: dpeur_l6 = df[["DPEUR"]].shift(6).rename(columns = {"DPEUR": "DPEUR_L6"})
        dpeur_l12 = df[["DPEUR"]].shift(12).rename(columns = {"DPEUR": "DPEUR_L12"})
        df_lagged_c = pd.concat([df_lagged, dpeur_l6, dpeur_l12], axis= 1)
```

```
In [10]: from statsmodels.tsa.stattools import acf, pacf
```

```
In [11]: dpeur = df_lagged_c[df_lagged_c.TREND <= 132].DPEUR.dropna().values
```

```
In [12]: acf(dpeur, nlags= 12)[1:]
```

```
Out[12]: array([ 0.08325178, -0.10916313, -0.1990793 , -0.15896745, -0.08844387,
                0.40291713, -0.0350489 , -0.17326768, -0.16204276, -0.11141802,
                0.01458082,  0.55447498])
```

```
In [13]: pacf(dpeur, nlags= 12)[1:]
```

```
Out[13]: array([ 0.08389217, -0.11872908, -0.18732104, -0.15337675, -0.12507951,
                0.39304979, -0.20987763, -0.18067838, -0.07363718, -0.08214183,
                0.05004201,  0.45590301])
```

The lags with largest ACF and PACF were found at lag6 and lag12

Estimation:

```
In [14]: y_c, X_c = dmatrices("DPEUR ~ DPEUR_L6+DPEUR_L12", data= df_lagged_c[df_lagged_c.TREND
```

```
In [15]: mod_c = sm.OLS(y_c, X_c).fit()
        print (mod_c.summary())
```

OLS Regression Results

```
=====
Dep. Variable:          DPEUR    R-squared:          0.423
Model:                  OLS      Adj. R-squared:       0.413
Method:                 Least Squares    F-statistic:       42.55
Date:                  Mon, 28 May 2018    Prob (F-statistic): 1.38e-14
Time:                  11:49:13    Log-Likelihood:     542.43
No. Observations:      119    AIC:               -1079.
Df Residuals:          116    BIC:               -1071.
Df Model:               2
Covariance Type:       nonrobust
=====
```

```
=====
              coef    std err          t      P>|t|      [0.025    0.975]
-----
Intercept    0.0004      0.000      1.365      0.175     -0.000     0.001
DPEUR_L6     0.1887      0.077      2.442      0.016      0.036     0.342
DPEUR_L12    0.5980      0.084      7.157      0.000      0.432     0.763
=====
```

```
=====
Omnibus:            10.597    Durbin-Watson:       1.626
Prob(Omnibus):      0.005    Jarque-Bera (JB):    19.695
Skew:               -0.321    Prob(JB):            5.29e-05
Kurtosis:           4.887    Cond. No.            406.
=====
```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

(d) Extend the AR model of part (c) by adding lagged values of monthly inflation in the USA at lags 1, 6, and 12. Check that the coefficient at lag 6 is not significant, and estimate the ADL model $DPEUR_t = + 1DPEUR_{t6} + 2DPEUR_{t12} + 1DPUSA_{t1} + 2DPUSA_{t12} + t$ (sample Jan 2000 - Dec 2010).

```
In [16]: dpusa_l6 = df[["DPUSA"]].shift(6).rename(columns = {"DPUSA": "DPUSA_L6"})
dpusa_l12 = df[["DPUSA"]].shift(12).rename(columns = {"DPUSA": "DPUSA_L12"})
df_lagged_d = pd.concat([df_lagged_c, dpusa_l6, dpusa_l12], axis= 1)
```

```
In [17]: y_d, X_d = dmatrices("DPEUR~DPEUR_L6+DPEUR_L12+DPUSA_L1+DPUSA_L6+DPUSA_L12", df_lagged_d)
```

```
In [18]: mod_d = sm.OLS(y_d, X_d).fit()
print (mod_d.summary())
```

OLS Regression Results

```
=====
Dep. Variable:          DPEUR    R-squared:          0.560
Model:                  OLS      Adj. R-squared:       0.541
Method:                 Least Squares    F-statistic:       28.79
Date:                  Mon, 28 May 2018    Prob (F-statistic): 9.84e-19
Time:                  11:49:13    Log-Likelihood:     558.57
=====
```

```

No. Observations:      119    AIC:      -1105.
Df Residuals:          113    BIC:      -1088.
Df Model:              5
Covariance Type:      nonrobust

```

```

=====
              coef    std err          t      P>|t|      [0.025    0.975]
-----
Intercept      0.0004      0.000       1.545      0.125     -0.000     0.001
DPEUR_L6       0.2030      0.079       2.584      0.011      0.047     0.359
DPEUR_L12      0.6368      0.087       7.279      0.000      0.463     0.810
DPUSA_L1       0.2264      0.051       4.429      0.000      0.125     0.328
DPUSA_L6      -0.0560      0.055      -1.023      0.308     -0.165     0.052
DPUSA_L12     -0.2301      0.054      -4.247      0.000     -0.337    -0.123
=====

```

```

Omnibus:          10.600    Durbin-Watson:      2.011
Prob(Omnibus):    0.005    Jarque-Bera (JB):    15.286
Skew:             0.443    Prob(JB):            0.000479
Kurtosis:         4.516    Cond. No.            512.
=====

```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

The p-value for DPUSA at lag 6 ("DPUSA_L6") is 0.308, which is not significant at 95% confidence level. Therefore we keep variable "DPUSA_L6" out of the model, and new estimation process goes as follows:

```

In [19]: y_d2, X_d2 = dmatrices("DPEUR~DPEUR_L6+DPEUR_L12+DPUSA_L1+DPUSA_L12", df_lagged_d[df_
mod_d2 = sm.OLS(y_d2, X_d2).fit()
print (mod_d2.summary())

```

OLS Regression Results

```

=====
Dep. Variable:      DPEUR    R-squared:      0.556
Model:              OLS      Adj. R-squared:    0.541
Method:             Least Squares    F-statistic:      35.71
Date:               Mon, 28 May 2018    Prob (F-statistic): 2.55e-19
Time:               11:49:13    Log-Likelihood:    558.02
No. Observations:    119    AIC:      -1106.
Df Residuals:        114    BIC:      -1092.
Df Model:            4
Covariance Type:    nonrobust
=====

```

```

=====
              coef    std err          t      P>|t|      [0.025    0.975]
-----
Intercept      0.0003      0.000       1.267      0.208     -0.000     0.001
DPEUR_L6       0.1687      0.071       2.374      0.019      0.028     0.310
=====

```

DPEUR_L12	0.6552	0.086	7.651	0.000	0.486	0.825
DPUSA_L1	0.2326	0.051	4.582	0.000	0.132	0.333
DPUSA_L12	-0.2265	0.054	-4.189	0.000	-0.334	-0.119
=====						
Omnibus:		10.147	Durbin-Watson:			2.014
Prob(Omnibus):		0.006	Jarque-Bera (JB):			15.787
Skew:		0.386	Prob(JB):			0.000373
Kurtosis:		4.609	Cond. No.			481.
=====						

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

(e) Use the models of parts (c) and (d) to make two series of 12 monthly inflation forecasts for 2011. At each month, you should use the data that are then available, for example, to forecast inflation for September 2011 you can use the data up to and including August 2011. However, do not re-estimate the model and use the coefficients as obtained in parts (c) and (d). For each of the two forecast series, compute the values of the root mean squared error (RMSE), mean absolute error (MAE), and the sum of the forecast errors (SUM). Finally, give your interpretation of the outcomes.

```
In [20]: def predict_future(trend, mod = "C", data = df_lagged_d):
        if mod == "C":
            mod = mod_c
            exogs = ["const", "DPEUR_L6", "DPEUR_L12"]
        elif mod == "D":
            mod = mod_d2
            exogs = ["const", "DPEUR_L6", "DPEUR_L12", "DPUSA_L1", "DPUSA_L12"]
        else:
            raise Exception("Model does not exist")
        data = sm.add_constant(data)
        predicted_values = mod.predict(data[data.TREND == trend][exogs])
        return predicted_values.values[0]

In [21]: #predicted values
        trends = range(133, 145)
        mod_c_predicted = []
        mod_d_predicted = []
        for trend in trends:
            mod_c_predicted.append(predict_future(trend, mod = "C"))
            mod_d_predicted.append(predict_future(trend, mod = "D"))
        #actual values
        real_dpeur = df_lagged_d[df_lagged_d.TREND.isin(trends)].DPEUR.values

In [22]: fig, ax = plt.subplots(1, 2, figsize = (16, 6))
        xs = range(1, 13)
        ax[0].plot(xs, mod_c_predicted, label = "model C predicted", linestyle = "--", marker = "x")
        ax[0].plot(xs, mod_d_predicted, label = "model D predicted", marker = "o", c = "orange")
```



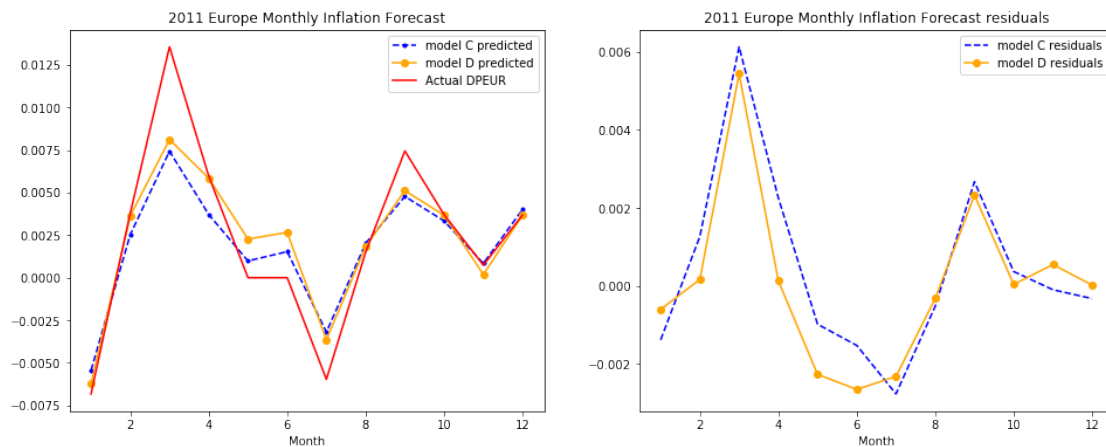
```

ax[0].plot(xs, real_dpeur, label = "Actual DPEUR", c = "red")
ax[0].legend()
ax[0].set_title("2011 Europe Monthly Inflation Forecast")
ax[0].set_xlabel("Month")

ax[1].plot(xs, real_dpeur- mod_c_predicted , label = "model C residuals", linestyle =
ax[1].plot(xs, real_dpeur- mod_d_predicted, label = "model D residuals", marker ="o"
ax[1].set_title("2011 Europe Monthly Inflation Forecast residuals")
ax[1].set_xlabel("Month")
ax[1].legend()

```

Out [22]: <matplotlib.legend.Legend at 0x11f13b278>



```

In [23]: RMSE_C = np.sqrt(sum([resid**2 for resid in (real_dpeur- mod_c_predicted)])/12)
RMSE_D = np.sqrt(sum([resid**2 for resid in (real_dpeur- mod_d_predicted)])/12)
MAE_C = sum([np.abs(resid) for resid in (real_dpeur- mod_c_predicted)])/12
MAE_D = sum([np.abs(resid) for resid in (real_dpeur- mod_d_predicted)])/12
SUM_C = sum([resid for resid in (real_dpeur- mod_c_predicted)])
SUM_D = sum([resid for resid in (real_dpeur- mod_d_predicted)])

```

In [24]: RMSE_C

Out [24]: 0.0023241954513586365

In [25]: RMSE_D

Out [25]: 0.0021105252005109189

In [26]: MAE_C

Out [26]: 0.0016924268523882621

In [27]: MAE_D

```
Out [27]: 0.0014036628892379894
```

```
In [28]: SUM_C
```

```
Out [28]: 0.0050653525975492527
```

```
In [29]: SUM_D
```

```
Out [29]: 0.00047846854944003504
```

Based on the statistics above, model D outperforms Model C.