

STA 512 HW9 Due 09/08/2020.

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1.

(a) Under H , $Y \sim N(0, 1)$

the acc region is $A(\alpha) = \{Z : z_{\alpha(1-\alpha)} < Z < z_{1-\alpha}\}$

thus, type I error rate is

$$\begin{aligned} \Pr(Y \notin A | H) &= 1 - \Pr(Y \in A | H) \\ &= 1 - \Pr(z_{\alpha(1-\alpha)} < Z < z_{1-\alpha}) \\ &= 1 - [\Pr(Z < z_{1-\alpha}) - \Pr(Z < z_{\alpha(1-\alpha)})] \\ &= 1 - [(1-\alpha) - \alpha(1-\alpha)] \\ &= \alpha \end{aligned}$$

(b) power is $\Pr(Y \in A_0 | \mu)$

$$\begin{aligned} &= 1 - \Pr(Y \notin A_0 | \mu) \\ &= 1 - \Pr(z_{\alpha(1-\alpha)} < Y < z_{1-\alpha} | \mu) \\ &= 1 - \Pr(z_{\alpha(1-\alpha)} - \mu < Y - \mu < z_{1-\alpha} - \mu | \mu) \\ \because Y &\sim N(\mu, 1) \rightarrow Y - \mu \sim N(0, 1) \quad \downarrow \\ &= 1 - [\Pr(Y - \mu < z_{1-\alpha} - \mu) - \Pr(Y - \mu < z_{\alpha(1-\alpha)} - \mu)] \\ &= 1 - [\Phi(z_{1-\alpha} - \mu) - \Phi(z_{\alpha(1-\alpha)} - \mu)] \end{aligned}$$

the plot is in appendix $\alpha = 0.05$

when I want to make a hyp test with a symmetric double tails, I would use $\alpha > 1/2$, otherwise I would use $1/4$

2.

$$(a) F_\theta = \int_0^y p_\theta dy = \int_0^y e^{-y/\theta} / \theta dy = 1 - e^{-y/\theta}$$

$$\begin{aligned} (b) \Pr(\text{reject } H | H) &= \Pr(Y > b | H) \\ &= 1 - \Pr(Y \leq b | H) \\ &= 1 - F_\theta(b | H) \\ &= 1 - (1 - e^{-b/\theta}) \\ &= e^{-b/\theta} \end{aligned}$$

the prob of rejection $\leq \alpha$ for each $\theta < \theta_0$

$$\Rightarrow \alpha \geq \sup_{\theta < \theta_0} \Pr(Y > b | \theta) = \sup_{\theta < \theta_0} e^{-b/\theta} = e^{-b/\theta_0}$$

$$b \geq -\theta_0 \log \alpha$$

$$\text{power} = e^{\theta_0/\theta \cdot \log \alpha}$$

the plot is in appendix ($\alpha = 0.05$, $\theta_0 = 1$)

$$\begin{aligned} (c) c(y) &= -\frac{y}{\log \alpha} \\ \Pr(\theta \in (c(y), \infty) | \theta) &= \Pr(\theta > -\frac{y}{\log \alpha} | \theta) \\ &= \Pr(Y < -\theta \log \alpha | \theta) \\ &= 1 - \alpha \end{aligned}$$

3.

(a) the usual two-sample t-statistic (with equal var) is

$$t(\underline{y}) = \frac{|\bar{y}_A - \bar{y}_B|}{s \sqrt{1/n_A + 1/n_B}}$$

The p-value is the probability under H_0 of observing a value of the test statistic the same as or more extreme than what was actually observed.

Under the continuous distribution, the p-value has a uniform(0,1) distribution.

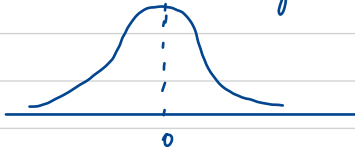
So the smallest p-value is 0.

(b) For permutation test,

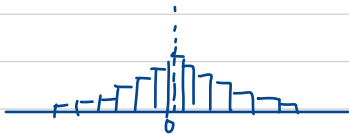
$$p\text{-value} = P_0(T > t_{\text{obs}}) = \frac{1}{N!} \sum_{j=1}^N \mathbb{I}(T_j > t_{\text{obs}})$$

Because t_{obs} is one of the permutations statistics, it will be among those found within permutation distribution. Therefore, the smallest p-value under the most extreme situation is $1/N!$

(c) Normal-theory test



Randomization Test



For normal theory-test, $p\text{-value} = 0.0016$.

For Randomization test, $p\text{-value} \rightarrow 0$

the distribution of Randomization test would be more flat than that of normal-theory test, since in the former one we have limited candidates values.

7.

(a) We know that $S^2 \stackrel{d}{=} \sigma^2 \chi^2 / (n-1)$ with d.f. = $n-1$

thus $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1}$

let $t(Y) = \frac{(n-1)S^2}{\sigma_0^2}$

and Acc region = $\{y: \chi^2_{1-\alpha/2} < t(Y) < \chi^2_{\alpha/2}\}$

this is a level- α hypothesis test for $H: \sigma^2 = \sigma_0^2$

(b) $\chi^2_{1-\alpha/2} < \frac{(n-1)S^2}{\sigma_0^2} < \chi^2_{\alpha/2}$

the $1-\alpha$ CI for σ^2 is $\left(\frac{(n-1)S^2}{\chi^2_{\alpha/2}}, \frac{(n-1)S^2}{\chi^2_{1-\alpha/2}} \right)$

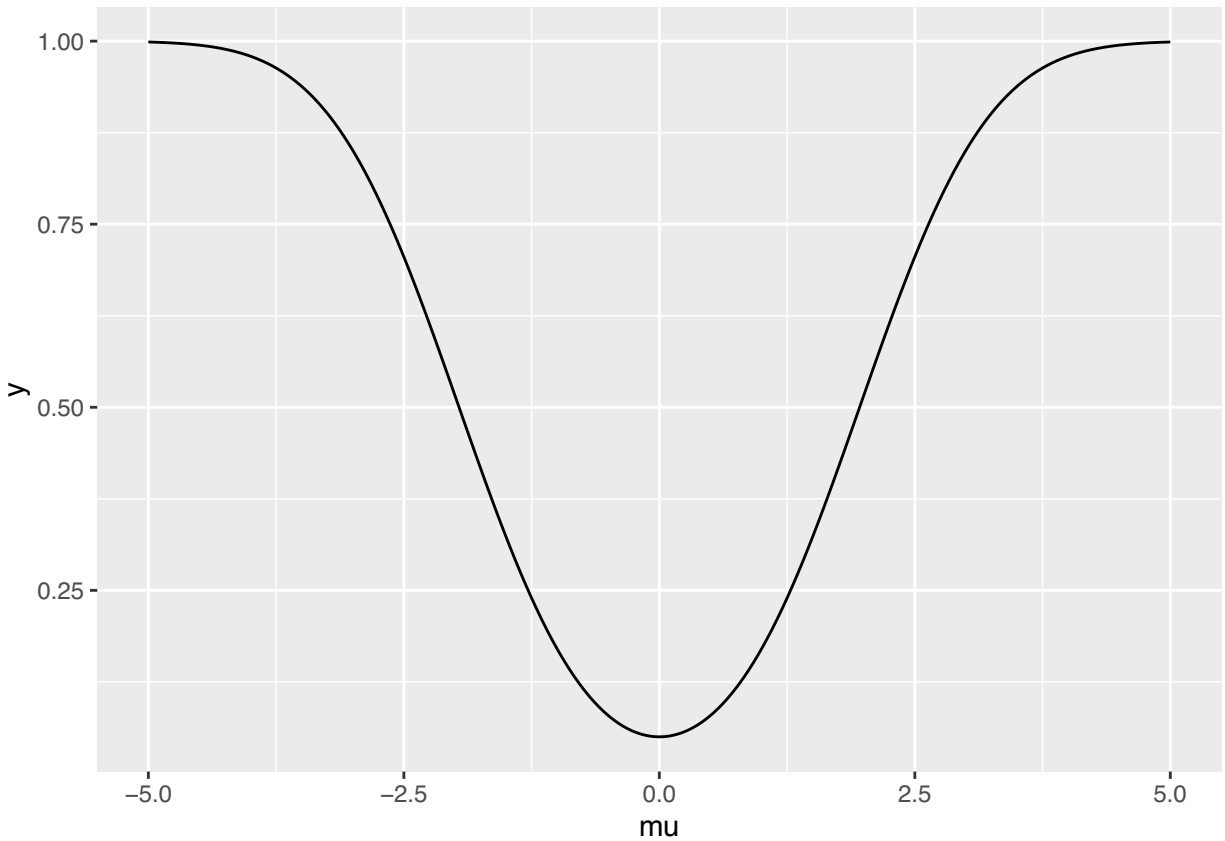
STA 532 Homework 9 - Appendix

Yi Mi

1.(b)

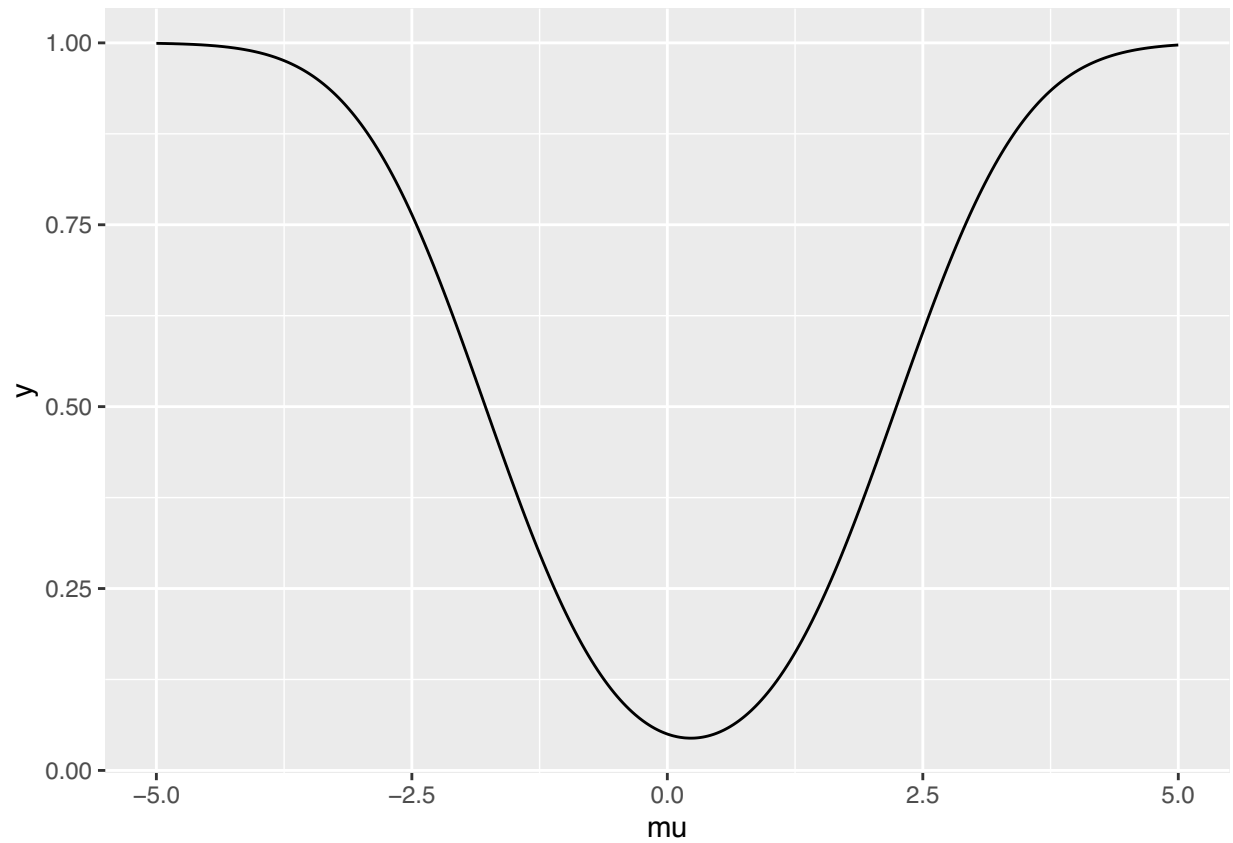
Let $\alpha = 0.05$. And $w = 1/2$.

```
mu = seq(-5, 5, 0.01)
#fa = qnorm(1 - a/2)*sqrt(a)
w = 0.5
y = 1-(pnorm(qnorm(1-0.05*w)-mu) - pnorm(qnorm(0.05 - 0.05*w)-mu))
df = data.frame(y,mu)
ggplot(df, aes(x=mu,y=y))+geom_line()
```



Let $\alpha = 0.05$. And $w = 1/4$.

```
mu = seq(-5, 5, 0.01)
#fa = qnorm(1 - a/2)*sqrt(a)
w = 0.25
y = 1-(pnorm(qnorm(1-0.05*w)-mu) - pnorm(qnorm(0.05 - 0.05*w)-mu))
df = data.frame(y,mu)
ggplot(df, aes(x=mu,y=y))+geom_line()
```



2.(b)

Let $\alpha = 0.05$ and $\theta_0 = 1$.

```
theta = seq(0, 10, 0.01)
p = exp(log(0.05)/theta)
df = data.frame(p, theta)
ggplot(df, aes(x=theta, y=p)) + geom_line()
```

