

STA 512 HW3 Due Wed 2/5

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1. (a) Let $f(y) = y^{p/q}$, $1 \leq q < p$

$$\frac{d^2 f(y)}{dy^2} = \frac{p}{q} \times \left(\frac{p}{q} - 1\right) \times y^{\frac{p}{q}-2} > 0 \text{ for } y \in (0, +\infty)$$

so $f(y)$ is strictly convex

Jensen: $E[f(Y)] \geq f[E(Y)]$ for convex $f(y)$

$$\Rightarrow E[(Y^q)^{p/q}] \geq [E(Y^q)]^{p/q}$$

$$E[Y^p] \geq [E(Y^q)]^{p/q}$$

$$E(Y^p)^{1/p} \geq E(Y^q)^{1/q}$$

(b) since $g(Y) = 1/Y$ is strictly convex in $Y \in (0, +\infty)$

$$E[1/Y] \geq 1/E[Y]$$

(c) since $h(Y) = \log Y$ is strictly concave in $Y \in (0, +\infty)$

$$E[\log Y] \leq \log E[Y]$$

2. (a) $(w_1, w_2, w_3) \sim \text{Dirichlet}(\alpha_1, \alpha_2, \alpha_3)$

The pdf for $\underline{w} = (w_1, w_2, w_3)$ is

$$p(\underline{w}) = \frac{1}{B(\alpha)} \prod_{i=1}^3 w_i^{\alpha_i-1},$$

$$\text{where } B(\alpha) = \frac{\prod_{i=1}^3 \Gamma(\alpha_i)}{\Gamma(\sum_{i=1}^3 \alpha_i)}$$

Let $\alpha_0 = \alpha_1 + \alpha_2 + \alpha_3$,

$$p(\underline{w}) = \frac{\Gamma(\alpha_0)}{\prod_{i=1}^3 \Gamma(\alpha_i)} \prod_{i=1}^3 w_i^{\alpha_i-1}$$

Mean: $E(w_1) = \iiint w_1 p(\underline{w}) dw_1 dw_2 dw_3$

$$= \iiint w_1 \frac{\Gamma(\alpha_0)}{\prod_{i=1}^3 \Gamma(\alpha_i)} w_1^{\alpha_1-1} w_2^{\alpha_2-1} w_3^{\alpha_3-1} dw_1 dw_2 dw_3$$

with kernel: $\iiint \frac{\Gamma(\alpha_0+1)}{\Gamma(\alpha_1+1)\Gamma(\alpha_2)\Gamma(\alpha_3)} w_1^{\alpha_1} w_2^{\alpha_2-1} w_3^{\alpha_3-1} dw_1 dw_2 dw_3 = 1$

we have:

$$E(w_1) = \frac{\Gamma(\alpha_0)}{\prod_{i=1}^3 \Gamma(\alpha_i)} \times \frac{\Gamma(\alpha_1+1)\Gamma(\alpha_2)\Gamma(\alpha_3)}{\Gamma(\alpha_0+1)}$$

$$= \frac{\alpha_1}{\alpha_0}$$

Therefore: $E(w_i) = \frac{\alpha_i}{\alpha_0}$ for $i \in \{1, 2, 3\}$

Variance: $\text{Var}(w_i) = E(w_i^2) - E(w_i)^2$

Similarly, $E(w_i^2) = \iiint w_i^2 p(\underline{w}) dw_1 dw_2 dw_3$

$$= \frac{\prod_{i=1}^3 \Gamma(\alpha_i)}{\prod_{i=1}^3 \Gamma(\alpha_i)} \times \frac{\Gamma(\alpha_1+2) \Gamma(\alpha_2) \Gamma(\alpha_3)}{\Gamma(\alpha_0+2)}$$

$$= \frac{(\alpha_1+1) \alpha_1}{(\alpha_0+1) \alpha_0}$$

Thus, $E(w_i^2) = \frac{(\alpha_i+1) \alpha_i}{(\alpha_0+1) \alpha_0}$

And, $\text{Var}(w_i) = \frac{(\alpha_i+1) \alpha_i}{(\alpha_0+1) \alpha_0} - \left(\frac{\alpha_i}{\alpha_0}\right)^2$

$$= \frac{(\alpha_0 - \alpha_i) \alpha_i}{(\alpha_0+1) \alpha_0^2}$$

(b). $\text{Cov}(w_1, w_2) = E[(w_1 - E(w_1))(w_2 - E(w_2))]$

$$= E(w_1 w_2) - E(w_1)E(w_2) - E(w_1)E(w_2) + E(w_1)E(w_2)$$

$$= E(w_1 w_2) - E(w_1)E(w_2)$$

Similarly, $E(w_1 w_2) = \iiint w_1 w_2 p(\underline{w}) dw_1 dw_2 dw_3$

$$= \frac{\prod_{i=1}^3 \Gamma(\alpha_i)}{\prod_{i=1}^3 \Gamma(\alpha_i)} \times \frac{\Gamma(\alpha_1+1) \Gamma(\alpha_2+1) \Gamma(\alpha_3)}{\Gamma(\alpha_0+2)}$$

$$= \frac{\alpha_1 \alpha_2}{(\alpha_0+1) \alpha_0}$$

Thus, $E(w_i w_j) = \frac{\alpha_i \alpha_j}{(\alpha_0+1) \alpha_0}$

$$\text{Cov}(w_i, w_j) = \frac{\alpha_i \alpha_j}{(\alpha_0+1) \alpha_0} - \frac{\alpha_i}{\alpha_0} \frac{\alpha_j}{\alpha_0} = -\frac{\alpha_i \alpha_j}{\alpha_0^2 (\alpha_0+1)} \quad i \neq j$$

explain the sign: $\sum_{i=1}^n w_i = 1$, when w_i increases, w_j may decrease, so their relationship is negative.

$$\begin{aligned}
 (c) \text{Var}(w_1 + w_2) &= E(w_1 + w_2)^2 - [E(w_1 + w_2)]^2 \\
 &= E(w_1^2 + 2w_1w_2 + w_2^2) - [E(w_1)^2 + 2E(w_1)E(w_2) + E(w_2)^2] \\
 &= \frac{(\alpha_1 + 1)\alpha_1}{(\alpha_0 + 1)\alpha_0} + 2\frac{\alpha_1\alpha_2}{(\alpha_0 + 1)\alpha_0} + \frac{(\alpha_2 + 1)\alpha_2}{(\alpha_0 + 1)\alpha_0} - \left[\frac{\alpha_1^2}{\alpha_0^2} + 2\frac{\alpha_1\alpha_2}{\alpha_0^2} + \frac{\alpha_2^2}{\alpha_0^2} \right] \\
 &= \frac{(\alpha_1 + \alpha_2)^2 + \alpha_1 + \alpha_2}{(\alpha_0 + 1)\alpha_0} - \frac{(\alpha_1 + \alpha_2)^2}{\alpha_0^2}
 \end{aligned}$$

$$\begin{aligned}
 \text{Cov}(w_1, w_1 + w_2) &= E[(w_1 - E(w_1))(w_1 + w_2 - E(w_1 + w_2))] \\
 &= E[w_1(w_1 + w_2)] - E(w_1 + w_2)E(w_1) - E(w_1)E(w_1 + w_2) + E(w_1)E(w_1 + w_2) \\
 &= E(w_1^2) + E(w_1w_2) - E(w_1)[E(w_1) + E(w_2)] \\
 &= \frac{(\alpha_1 + 1)\alpha_1}{(\alpha_0 + 1)\alpha_0} + \frac{\alpha_1\alpha_2}{(\alpha_0 + 1)\alpha_0} - \frac{\alpha_1}{\alpha_0} \left(\frac{\alpha_1}{\alpha_0} + \frac{\alpha_2}{\alpha_0} \right) \\
 &= \frac{\alpha_1(\alpha_0 + \alpha_1 - \alpha_2)}{(\alpha_0 + 1)\alpha_0^2}
 \end{aligned}$$

$$(d) \quad p(\underline{w}) = \frac{\Gamma(\alpha_0)}{\prod_{i=1}^3 \Gamma(\alpha_i)} \prod_{i=1}^3 w_i^{\alpha_i-1}$$

$$\text{let } w = \frac{w_2}{1-w_1}$$

$$\text{we have } w_1 = 1 - w_1 - w_2 = (1-w_1)(1-w)$$

$$\text{thus } p(w_1, w) = \frac{\Gamma(\alpha_0)}{\prod_{i=1}^3 \Gamma(\alpha_i)} w_1^{\alpha_1-1} [w(1-w_1)]^{\alpha_2-1} [(1-w_1)(1-w)]^{\alpha_3-1}$$

$$\text{and } p(w_1) = \int p(w_1, w) dw$$

$$= \frac{\Gamma(\alpha_0)}{\prod_{i=1}^3 \Gamma(\alpha_i)} w_1^{\alpha_1-1} (1-w_1)^{\alpha_2+\alpha_3-2} \int w^{\alpha_2-1} (1-w)^{\alpha_3-1} dw$$

$$= w_1^{\alpha_1-1} (1-w_1)^{\alpha_2+\alpha_3-1} \frac{\Gamma(\alpha_0)}{\Gamma(\alpha_2+\alpha_3)\Gamma(\alpha_1)} \int \frac{\Gamma(\alpha_2+\alpha_3)}{\Gamma(\alpha_2)\Gamma(\alpha_3)} w^{\alpha_2-1} (1-w)^{\alpha_3-1} dw$$

$$\Rightarrow w_1 \sim \text{Beta}(\alpha_1, \alpha_2+\alpha_3)$$

$$\text{similarly, } w_2 \sim \text{Beta}(\alpha_3, \alpha_1+\alpha_2)$$

$$\text{and, } w_1 + w_2 = 1 - w_3$$

$$\text{thus, } p(w_1 + w_2) = P_{w_3}(1 - (w_1 + w_2)) \left| \frac{d(1-w_3)}{dw_3} \right|$$

$$= [1 - (w_1 + w_2)]^{\alpha_3-1} (w_1 + w_2)^{\alpha_1+\alpha_2-1} \frac{\Gamma(\alpha_0)}{\Gamma(\alpha_1+\alpha_2)\Gamma(\alpha_3)}$$

$$= \text{Beta}(\alpha_1+\alpha_2, \alpha_3)$$

3. (a) if X and Y are independent,

$$P(X, Y) = P(X) P(Y)$$

$$E(XY) = \iint XY P(X, Y) dx dy.$$

$$= \iint XY P(X) P(Y) dx dy$$

$$= \int X P(X) dx \int Y P(Y) dy$$

$$= E(X) E(Y)$$

$$\text{cov}(X, Y) = E[(X - E(X))(Y - E(Y))]$$

$$= E(XY) - E(X)E(Y)$$

$$= 0$$

$$(b) \text{Cov}(X, Y) = \text{cov}(a + bY, Y)$$

$$= E[(a + bY - E(a + bY))(Y - E(Y))]$$

$$= b E[(Y - E(Y))^2]$$

$$= b \text{Var}(Y)$$

$$\text{Var}(X) = E[a + bY - E(a + bY)]^2$$

$$= b^2 E[(Y - E(Y))^2]$$

$$= b^2 \text{Var}(Y)$$

$$\text{cor}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}} = \frac{b\text{Var}(Y)}{\sqrt{b^2 \text{Var}(Y)}} = \pm 1$$

$$\begin{aligned}
(c) i. \text{Cov}(a_1 + b_1 X_1, a_2 + b_2 X_2) &= E[(a_1 + b_1 X_1 - E(a_1 + b_1 X_1))(a_2 + b_2 X_2 - E(a_2 + b_2 X_2))] \\
&= b_1 b_2 E[(X_1 - E X_1)(X_2 - E X_2)] \\
&= b_1 b_2 \text{Cov}(X_1, X_2)
\end{aligned}$$

$$\begin{aligned}
ii. E(X_1 + X_2 + X_3) &= \iiint (X_1 + X_2 + X_3) p(X_1, X_2, X_3) dX_1 dX_2 dX_3 \\
&= \sum_{i=1}^3 \iiint X_i p(X_1, X_2, X_3) dX_1 dX_2 dX_3 \\
&= E(X_1) + E(X_2) + E(X_3)
\end{aligned}$$

$$\begin{aligned}
\text{Var}(X_1 + X_2 + X_3) &= E(X_1 + X_2 + X_3 - E(X_1 + X_2 + X_3))^2 \\
&= E(X_1 + X_2 + X_3)^2 - (E(X_1 + X_2 + X_3))^2 \\
&= E(X_1^2 + X_2^2 + X_3^2 + 2X_1X_2 + 2X_1X_3 + 2X_2X_3) \\
&\quad - (E^2 X_1 + E^2 X_2 + E^2 X_3 + 2E X_1 E X_2 + 2E X_1 E X_3 + 2E X_2 E X_3) \\
&= \sum_{i=1}^3 \text{Var}(X_i) + 2 \sum_{i < j} \text{Cov}(X_i, X_j) \\
&= \sum_i \sum_j \text{Cov}(X_i, X_j)
\end{aligned}$$

Yes, it matches the result from class.

7. (a) Since $Z, X_1, X_2, X_3 \sim N(0, 1)$

$$E(Y_1) = E(Z + X_1) = 0$$

$$E(Y_2) = E(Z + X_2) = 0$$

$$E(Y_3) = E(Z^2 + X_3) = \text{Var}(Z) + (EZ)^2 + EX_3 = 1$$

$$\begin{aligned} \text{Var}(Y_1) &= EY_1^2 - (EY_1)^2 = EY_1^2 = E(Z^2 + 2ZX_1 + X_1^2) \\ &= EZ^2 + 2EZEX_1 + EX_1^2 \\ &= 2 \end{aligned}$$

Similarly, $\text{Var}(Y_2) = 2$

$$\begin{aligned} \text{Var}(Y_3) &= EY_3^2 - (EY_3)^2 = EY_3^2 - 1 = E(Z^4 + 2X_3Z^2 + X_3^2) - 1 \\ &= EZ^4 + 2EX_3EZ^2 + EX_3^2 - 1 \\ &= 3 + 1 - 1 \\ &= 3 \end{aligned}$$

$$(b) \text{Cov}(Y_i) = \begin{bmatrix} V(Y_1) & c(Y_1, Y_2) & c(Y_1, Y_3) \\ c(Y_2, Y_1) & V(Y_2) & c(Y_2, Y_3) \\ c(Y_3, Y_1) & c(Y_3, Y_2) & V(Y_3) \end{bmatrix}$$

$$\begin{aligned} c(Y_1, Y_2) &= EY_1Y_2 - EY_1EY_2 = E(Z + X_1)(Z + X_2) \\ &= EZ^2 + EZEX_1 + EZEX_2 + EX_1EX_2 \\ &= 1 \end{aligned}$$

$$\begin{aligned} c(Y_1, Y_3) &= EY_1Y_3 - EY_1EY_3 = E(Z + X_1)(Z^2 + X_3) \\ &= EZ^3 + EZEX_3 + EX_1Z^2 + EX_1EX_3 \end{aligned}$$

since $X_1 \perp Z, X_2 \perp Z$
 $EX_1Z = EX_1EZ$
 $EX_2Z = EX_2EZ$
 $EX_1X_2 = EX_1EX_2$

$$\begin{aligned}
 E X_1 Z^2 &= \iint x_1 z^2 p(x_1, z) dx dz \\
 &= \int x_1 p(x) dx \int z^2 p(z) dz \\
 &= E X_1 E Z^2
 \end{aligned}$$

$$\text{thus, } c(Y_1, Y_3) = E Z^3 = 0$$

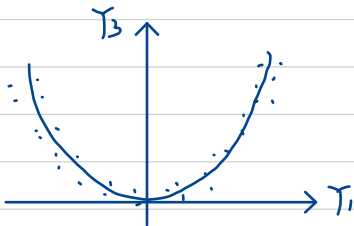
$$\text{Similarly, } c(Y_2, Y_3) = 0$$

$$\text{Cov}(\underline{Y}) = \begin{pmatrix} 2 & 1 & 0 \\ 1 & 2 & 0 \\ 0 & 0 & 3 \end{pmatrix}$$

(c) $c(Y_1, Y_3) = 0$, indicating that Y_1 and Y_3 have no linear relationship, however, it does not mean that Y_1 and Y_3 are independent.

$Y_1 = Z + X_1$, $Y_3 = Z^2 + X_3$, they both refer to the same random variable Z

We can see X_1 and X_3 as noise, and thus the sketch of the relationship between Y_1 and Y_3 is like the



randomly scattered points around the line $f(z) = z^2$, resulting from the noise X_1 and X_3 .

Therefore, Y_3 and Y_1 are not independent.

$$\begin{aligned}
5. (a) E[f(Y)] &= \int f(y) p(y) dy \\
&= \int_y f(y) \int_{\mathbb{R}} p(x, y) dx dy \\
&= \int_y f(y) \int_{\mathbb{R}} p(y|x) p(x) dx dy \\
&= \int_{\mathbb{R}} \int_y f(y) p(y|x) dy p(x) dx \\
&= \int_{\mathbb{R}} E(f(y)|x) p(x) dx \\
&= E[E(f(y)|x)]
\end{aligned}$$

$$\begin{aligned}
(b) E[f(x) g(x, Y) | x] &= \int f(x) g(x, Y) p(Y|x) dy \\
&= f(x) \int g(x, Y) p(Y|x) dy \\
&= f(x) E(g(x, Y) | x)
\end{aligned}$$