

STAT 512 HW2 Due Jan 29, 2020

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$$1. Y|\theta \sim N(\theta, \sigma^2)$$

$$p(Y|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{1}{2}\frac{(Y-\theta)^2}{\sigma^2}\right\}$$

$$\theta \sim N(\mu, \tau^2)$$

$$p(\theta) = \frac{1}{\sqrt{2\pi}\tau} \exp\left\{-\frac{1}{2}\frac{(\theta-\mu)^2}{\tau^2}\right\}$$

$$p(Y, \theta) = p(Y|\theta)p(\theta)$$

$$= \frac{1}{2\pi\sigma\tau} \exp\left\{-\frac{1}{2}\left[\left(\frac{Y-\theta}{\sigma}\right)^2 + \left(\frac{\theta-\mu}{\tau}\right)^2\right]\right\}$$

$$p(Y) = \int_{-\infty}^{+\infty} p(Y, \theta) d\theta$$

$$= \frac{1}{2\pi\sigma\tau} \int \exp\left\{-\frac{1}{2}\left[\left(\frac{Y^2}{\sigma^2} + \frac{\mu^2}{\tau^2}\right) + (\theta^2(\frac{1}{\sigma^2} + \frac{1}{\tau^2}) - 2\theta(\frac{Y}{\sigma} + \frac{\mu}{\tau}))\right]\right\} d\theta$$

$$= \frac{1}{2\pi\sigma\tau} \exp\left\{-\frac{1}{2}\left(\frac{Y^2}{\sigma^2} + \frac{\mu^2}{\tau^2}\right)\right\} \int \exp\left\{-\frac{1}{2}\left[\theta^2\left(\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right) - 2\theta\left(\frac{Y}{\sigma} + \frac{\mu}{\tau}\right)\right]\right\} d\theta$$

We can use kernel, let $a = \frac{1}{\sigma^2} + \frac{1}{\tau^2}$, $b = \frac{Y}{\sigma} + \frac{\mu}{\tau}$

We have,

$$\int \frac{1}{\sqrt{2\pi/a}} \exp\left\{-\frac{1}{2}\frac{(\theta - b/a)^2}{1/a}\right\} d\theta = 1$$

$$\int \frac{\sqrt{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}}{\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left[(\frac{1}{\sigma^2} + \frac{1}{\tau^2})\theta^2 - 2\theta(\frac{Y}{\sigma} + \frac{\mu}{\tau}) + \frac{(\frac{Y}{\sigma} + \frac{\mu}{\tau})^2}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}\right]\right\} d\theta = 1$$

$$\int \exp\left\{-\frac{1}{2}\left[(\frac{1}{\sigma^2} + \frac{1}{\tau^2})\theta^2 - 2\theta(\frac{Y}{\sigma} + \frac{\mu}{\tau})\right]\right\} d\theta = \sqrt{\frac{2\pi}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}} \exp\left\{\frac{1}{2}\frac{(\frac{Y}{\sigma} + \frac{\mu}{\tau})^2}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}\right\}$$

$$\text{Then, } p(Y) = \frac{1}{\sqrt{2\pi(\frac{1}{\sigma^2} + \frac{1}{\tau^2})} \sigma\tau} \exp\left\{-\frac{1}{2}\left(\frac{Y^2}{\sigma^2} + \frac{\mu^2}{\tau^2} - \frac{(\frac{Y}{\sigma} + \frac{\mu}{\tau})^2}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}\right)\right\}$$

$$\propto N(\mu, \tau^2 + \sigma^2)$$

$$p(\theta|Y) = \frac{p(Y, \theta)}{p(Y)} = \frac{\frac{1}{\sqrt{2\pi}\gamma} \exp\left\{-\frac{1}{2}\left[\frac{(Y-\theta)^2}{\sigma^2} + \frac{(\theta-\mu)^2}{\gamma^2}\right]\right\}}{\frac{1}{\sqrt{2\pi(\gamma^2+\sigma^2)}} \exp\left\{-\frac{1}{2}\frac{(Y-\mu)^2}{\gamma^2+\sigma^2}\right\}}$$

2. if r.v. X and Y are independent, $x, y \in \mathbb{R}$

$$\begin{aligned}
 p_{X,Y}(x,y) &= \Pr(X=x, Y=y) \\
 &= \Pr(X=x) \Pr(Y=y) \\
 &= p_X(x) p_Y(y)
 \end{aligned}$$

if $p_{X,Y}(x,y) = p_X(x) p_Y(y)$, $x, y \in \mathbb{R}$
 We take $A \subset \mathbb{R}$, $B \subset \mathbb{R}$

for continuous case:

$$\begin{aligned}
 \Pr(X \in A, Y \in B) &= \int_B \int_A p_{X,Y}(x,y) dx dy \\
 &= \int_B \int_A p_X(x) p_Y(y) dx dy \\
 &= \int_A p_X(x) dx \int_B p_Y(y) dy \\
 &= \Pr(X \in A) \Pr(Y \in B)
 \end{aligned}$$

$\Rightarrow X$ and Y are independent

for discrete case:

$$\begin{aligned}
 \Pr(X \in A, Y \in B) &= \sum_B \sum_A p_{X,Y}(x,y) \\
 &= \sum_B \sum_A p_X(x) p_Y(y) \\
 &= \sum_A p_X(x) \sum_B p_Y(y) \\
 &= \Pr(X \in A) \Pr(Y \in B)
 \end{aligned}$$

$\Rightarrow X$ and Y are independent.

3. if r.v. X, Y are independent, $X, Y \in \mathbb{R}$

$$\begin{aligned}
 p_{X|Y}(x|y) &= \frac{p_{X,Y}(x,y)}{p_Y(y)} = \frac{p_X(x) p_Y(y)}{p_Y(y)} \\
 &= \frac{p_X(x) p_Y(y)}{p_Y(y)} \\
 &= p_X(x) = p_X(x)
 \end{aligned}$$

if $p_{X|Y}(x|y) = p_X(x)$, $X, Y \in \mathbb{R}$,

We take $A \subset \mathbb{R}$, $B \subset \mathbb{R}$

in continuous case:

$$\begin{aligned}
 P(X \in A, Y \in B) &= \int_A \int_B p_{X,Y}(x,y) dx dy \\
 &= \int_A \int_B p_{X|Y}(x|y) p_Y(y) dx dy \\
 &= \int_A \int_B p_X(x) p_Y(y) dx dy \\
 &= P(X \in A) P(Y \in B) \\
 \Rightarrow X \text{ and } Y \text{ are independent}
 \end{aligned}$$

in discrete case:

$$\begin{aligned}
 P(X \in A, Y \in B) &= \sum_A \sum_B p_{X,Y}(x,y) \\
 &= \sum_A \sum_B p_{X|Y}(x,y) p_Y(y) \\
 &= \sum_A \sum_B p_X(x) p_Y(y) \\
 &= \sum_A p_X(x) \sum_B p_Y(y) \\
 &= P(X \in A) P(Y \in B)
 \end{aligned}$$

$\Rightarrow X$ and Y are independent

7. $U \in \mathcal{F}, V \in \mathcal{G}$, g and h are invertible

$$\Pr(U=u, V=v)$$

$$= \Pr(g(X)=u, h(Y)=v)$$

$$= \Pr(X=g^{-1}(u), Y=h^{-1}(v))$$

because X and Y are independent

$$= \Pr(X=g^{-1}(u)) \Pr(Y=h^{-1}(v))$$

$$= \Pr(g(X)=u) \Pr(h(Y)=v)$$

$$= \Pr(U=u) \Pr(V=v)$$

$\Rightarrow U$ and V are independent

5. Y_1, \dots, Y_n iid $p(y)$, cdf is $F_Y(y)$

$$\textcircled{1} Y_{(n)} = \min \{Y_1, \dots, Y_n\}$$

$$F_{Y_{(n)}}(y) = \Pr(Y_{(n)} \leq y) = 1 - \Pr(Y_{(n)} > y)$$

$$= 1 - \prod_{i=1}^n \Pr(Y_i > y)$$

$$= 1 - \prod_{i=1}^n (1 - \Pr(Y_i \leq y))$$

$$= 1 - \prod_{i=1}^n (1 - F_Y(y))$$

$$= 1 - (1 - F_Y(y))^n$$

$$p_{Y_{(n)}}(y) = \frac{dF_{Y_{(n)}}(y)}{dy} = n(1 - F_Y(y))^{n-1} p(y)$$

$$\textcircled{2} Y_{(n)} = \max \{Y_1, \dots, Y_n\}$$

$$F_{Y_{(n)}}(y) = \Pr(Y_{(n)} \leq y)$$

$$= \prod_{i=1}^n \Pr(Y_i \leq y)$$

$$= (F_Y(y))^n$$

$$p_{Y_{(n)}}(y) = \frac{dF_{Y_{(n)}}(y)}{dy} = n F_Y(y)^{n-1} p(y)$$

$$b. (a) F_{Y^{(n)}}(y) = 1 - [1 - F(y)]^n$$

$$p_{Y^{(n)}}(y) = n(1 - F(y))^{n-1} p(y)$$

$$p(y) = \lambda e^{-\lambda y}, \quad F_Y(y) = 1 - e^{-\lambda y}$$

$$F_{Y^{(n)}}(y) = 1 - e^{-n\lambda y}$$

$$p_{Y^{(n)}}(y) = n\lambda e^{-n\lambda y}$$

$$T_{(n)} \sim \exp(1/\lambda n)$$

$$c. b) Y \in [0, 1]$$

$$p(y) = 1, \quad F_Y(y) = y$$

$$F_{Y^{(n)}}(y) = 1 - (1 - y)^n$$

$$p_{Y^{(n)}}(y) = n(1 - y)^{n-1}$$

(c) Y follows a discrete distribution with p.m.f

$$p(y) = \theta(1 - \theta)^{y-1} \text{ for } Y \in \{1, 2, \dots\}$$

$$F_Y(y) = \sum_{y'=1}^y p(y') = \sum_{y'=1}^y \theta(1 - \theta)^{y'-1} = 1 - (1 - \theta)^y$$

$$F_{Y^{(n)}}(y) = 1 - (1 - \theta)^{ny}$$

$$p_{Y^{(n)}}(y) = \Pr(T_{(n)} = y)$$

$$= \Pr(T_{(n)} \leq y) - \Pr(T_{(n)} \leq y-1)$$

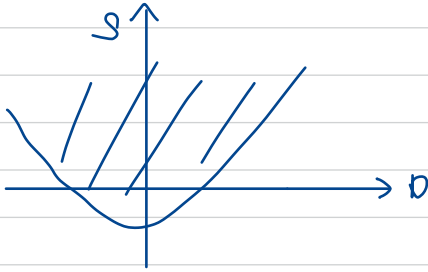
$$= (1 - (1 - \theta)^{ny}) - (1 - (1 - \theta)^{n(y-1)})$$

$$= (1 - \theta)^{ny} ((1 - \theta)^{-n} - 1)$$

$$T \sim (0, \infty)$$

$$\begin{aligned} \Pr(T \leq y) &= \int_0^y p(t) dt \\ &= \int_0^y \frac{1}{\lambda} e^{-t/\lambda} dt \\ &= 1 - e^{-y/\lambda} \end{aligned}$$

$$\begin{aligned} (b) \quad S + D &= 2T_1 & S \in (-D, +\infty) \\ S - D &= 2T_2 & S \in (-D, +\infty) \end{aligned}$$



$$(c) \quad T_1, \dots, T_n \stackrel{iid}{\sim} p(y)$$

$$p_{T_1, T_2}(y_1, y_2) = p_{T_1}(y_1) p_{T_2}(y_2) = \left(\frac{1}{\lambda}\right)^2 e^{-\frac{y_1 + y_2}{\lambda}}$$

$$\text{Let } X = \begin{pmatrix} T_1 + T_2 \\ T_1 - T_2 \end{pmatrix} = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}$$

$$T' = \begin{pmatrix} T_1 \\ T_2 \end{pmatrix} = \begin{pmatrix} \frac{X_1 + X_2}{2} \\ \frac{X_1 - X_2}{2} \end{pmatrix}$$

$$p_X(x) = p_{T'}(y') \left| \frac{\partial T'}{\partial x} \right| = \frac{1}{2\lambda^2} e^{-\frac{y_1 + y_2}{\lambda}} = \frac{1}{2\lambda^2} e^{-\frac{x_1}{\lambda}}$$

$$\Rightarrow p_{s,0}(s,d) = \frac{1}{2\lambda} e^{-s/\lambda} \quad |d| < s$$

$$\begin{aligned} p_s(s) &= \int_{-s}^s p_{s,0}(s,d) dd \\ &= \frac{s}{\lambda^2} e^{-s/\lambda} \\ 0 < s < \infty \end{aligned}$$

$$p_{D|S}(d|s) = \frac{p_{s,0}(s,d)}{p_s(s)} = \frac{1}{2s}$$

$$(d) \quad p_D(d) = \int_0^{\infty} p_{s,0}(s,d) ds$$

$$\because s > |d|$$

$$\begin{aligned} \text{if } d > 0, \quad s \in (d, \infty), \quad p_D(d) &= \int_d^{\infty} p_{s,0}(s,d) ds \\ &= \frac{1}{2\lambda} e^{-d/\lambda} \end{aligned}$$

$$\begin{aligned} \text{if } d < 0, \quad s \in (-d, \infty), \quad p_D(d) &= \int_{-d}^{\infty} p_{s,0}(s,d) ds \\ &= \frac{1}{2\lambda} e^{d/\lambda} \end{aligned}$$

