

# Expertise in Essay Scoring

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Achieving levels of reliability that allow large-scale essay assessments to be used to guide educational policy is a major hurdle for test developers. Previous studies have shown that one influential source of measurement error associated with essay scores is rater idiosyncracies [Engelhard 1994]. Although the literature dealing with scoring cognition is not conclusive about why some scorers are more consistent than others, it offers some insight into variables that may account for individual differences in scoring competence [Pula & Huot 1993; Wolfe & Feltovich 1994].

The purpose of this study is to empirically identify characteristics that differentiate essay scorers of differing levels of proficiency. By identifying the characteristics that differentiate better from poorer scorers, developers of large-scale essay examinations may be able to make scorer training and monitoring efforts less costly and more effective.

## Theoretical Framework

Based on the literature on scoring cognition [Freedman & Calfee 1983; Huot 1993; Vaughan 1991; Wolfe & Feltovich 1994], we developed a model of scoring cognition that contains two components: *processing actions* and *content focus*. *Processing actions* are the mental operations that are used to compare an essay's content to a set of scoring criteria. Here is an example of how *processing actions* are applied during the scoring process: As they read an essay, scorers occasionally interrupt their reading to *monitor* how well the essay satisfies the scoring criteria. After reading the paper, scorers may *review* the essay to note its strengths and weaknesses. They may also *diagnose* problems with the essay and suggest ways to improve it. Finally, the scorer assigns a score and provides a *rationale* for that score.

We define *content focus* as the features of the essay upon which scoring decisions are based. The *content focus* for this study includes the ability to tell a *story*, essay *organization*, individual writing *style*, and control of *mechanics* as well as other less relevant features of the essay (e.g., textual *appearance*, how well the writing addresses the *assignment*, and *non-specific* comments).

Based on findings in similar domains of expertise, one might expect the knowledge structures of experts to be organized in a more sophisticated way, allowing them to perceive information in the form of large meaningful patterns and to access this information more quickly and with deeper understandings than novices [Voss & Post 1988]. Prior research on scorer cognition led us to formulate four hypotheses about scoring expertise. *Hypothesis 1* predicted that the more proficient scorers as a group would be more consistent in their use of *processing actions*. *Hypothesis 2* predicted that the more proficient scorers as a group would be more consistent

in their use of *content focus* categories. *Hypothesis 3* predicted that scorers of different proficiency levels would emphasize different *processing actions*. *Hypothesis 4* predicted that scorers of different proficiency levels would emphasize different *content focus* categories.

## Method

### Subjects

Subjects for this study were 36 scorers who took part in a large essay scoring project. Based on their demonstrated levels of proficiency with the scoring rubric, subjects were selected to equally (i.e., by 12's) represent three proficiency groups: *novices*, *intermediates*, and *experts*. Subjects performed a think aloud task as they scored 24 essays. Interviews were audiotaped and transcribed for analysis. Each statement made by a scorer was coded according to its *content focus* (i.e., *appearance*, *assignment*, *mechanics*, *non-specific*, *organization*, *storytelling*, or *style*) and its *processing action* (i.e., *diagnose*, *monitor*, *review*, or *rationale*). The proportions of statements that fell into each coding category across essays served as the data for making group comparisons. Cohen's  $\kappa$  was .85 and .93 for the *content focus* and *processing action* codes, respectively.

### Analyses

Multiple *t* tests were employed to investigate our four hypotheses. Our goal was to identify monotonic relationships between the cognitive habits of scorers and scoring proficiency. For *Hypotheses 1* and 2, two *a priori* orthogonal contrasts were applied to the variances of the proportions for each *processing action* and *content focus* category. The two contrasts compared *experts* to the combined group of *intermediates* and *novices* ( $\Psi_1: \sigma^2_{\text{experts}} - 1/2(\sigma^2_{\text{intermediates}} + \sigma^2_{\text{novices}})$ ) and *intermediates* to *novices* ( $\Psi_2: \sigma^2_{\text{intermediates}} - \sigma^2_{\text{novices}}$ ). For *Hypotheses 3* and 4, similar analyses were performed on the means of the proportions for each *processing action* and *content focus* category ( $\Psi_1: \mu^2_{\text{experts}} - 1/2(\mu^2_{\text{intermediates}} + \mu^2_{\text{novices}})$  and  $\Psi_2: \mu^2_{\text{intermediates}} - \mu^2_{\text{novices}}$ ). *Post hoc* analyses were performed on additional variables identified by protocol coders.

## Results

Table 1 shows the variance of the *processing action* proportions for the think aloud data by proficiency group (*Hypothesis 1*). These data show that *experts* were more consistent in their use of *monitor* actions than were *intermediates* and *novices*;  $t(33) = -4.21, p < .001$ . Although the difference between *intermediates* and *novices* for *monitoring* was not statistically significant ( $t(33) = -1.00, p = .16$ ), the data suggest that there may be an increase in the consistency with which *monitor* actions are used as scoring proficiency increases. The *intermediates* were significantly more consistent than *novices* for the *review* and *rationale processing actions*;  $t(33) = -2.99, p = .004$  and  $t(33) = -2.53, p = .008$ , respectively. Although the contrasts did not afford a comparison of the differences between *experts* and *intermediates*, inspection of the data seems to indicate that the variances of *experts* and *intermediates* are similar for both of these variables [Table 1].

Processing Action	Expert $\sigma^2$	Intermediate $\sigma^2$	Novice $\sigma^2$
<i>Monitor</i>	0.0034	0.0318	0.0649
<i>Review</i>	0.0326	0.0204	0.0717
<i>Rationale</i>	0.0182	0.0138	0.0636
<i>Diagnose</i>	0.0055	0.0013	0.0036

Table 1: Group Variances on Processing Actions

Similar analyses compared the variances of the proportions for the *content focus* codes (recall *Hypothesis 2*). Table 2 shows the group variance of the proportions for the content focus data. Only one of the contrasts in these data was statistically significant--*experts* were less variable in their use of *storytelling* than were *intermediates* and *novices*;  $t(33) = -2.18, p = .02$ . However, this difference cannot be easily interpreted from a learning perspective because the relationship between scoring proficiency and *storytelling* use is non-monotonic. Furthermore, most of these trends are opposite those predicted by *Hypothesis 2*. That is, scorers who were more proficient were typically less consistent in their use of *content focus* categories than were less proficient scorers [Table 2].

<i>Content Focus</i>	<i>Expert</i> $\sigma^2$	<i>Intermediate</i> $\sigma^2$	<i>Novice</i> $\sigma^2$
<i>Appearance</i>	0.0010	0.0038	0.0012
<i>Assignment</i>	0.0006	0.0004	0.0004
<i>Mechanics</i>	0.0016	0.0025	0.0019
<i>Non-Specific</i>	0.0064	0.0066	0.0024
<i>Organization</i>	0.0068	0.0094	0.0020
<i>Storytelling</i>	0.0024	0.0131	0.0043
<i>Style</i>	0.0026	0.0056	0.0018

**Table 2: Group Variances on Content Focus**

The analyses for *Hypothesis 3* were based on the data in Table 3 which shows the descriptive statistics of the proportions for the *processing action* codes for each proficiency group. Two of the group comparisons were statistically significant. *Experts* were less likely to use *monitor* actions than were *intermediates* and *novices*;  $t(21) = -4.45, p = .001$ . On the other hand, *experts* were more likely to use *review* actions than were *intermediate* and *novice* scorers;  $t(33) = 3.23, p = .005$ . Statistical tests did not indicate any differences between *intermediates* and *novices* for any of these variables [Table 3].

<i>Processing Action</i>	<i>Expert</i>		<i>Intermediate</i>		<i>Novice</i>	
	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>
<i>Monitor</i>	.06	0.06	.31	0.18	.24	0.25
<i>Review</i>	.57	0.18	.34	0.14	.33	0.27
<i>Rationale</i>	.27	0.13	.28	0.12	.33	0.25
<i>Diagnose</i>	.10	0.07	.07	0.04	.10	0.06

**Table 3: Group Proportions on Think Aloud Processing Actions**

Table 4 presents the descriptive statistics for the group proportions for the *content focus* codes (recall *Hypothesis 4*). Two of the group comparisons were statistically significant. *Intermediates* were more likely to make *organization* content statements than were *novices*;  $t(16) = 2.60, p = .01$ . However, as was true for the variance comparison for *organization*, a visual inspection of the three group means renders this finding uninterpretable in a learning context. The second statistically significant difference revealed that *intermediate* scorers were less likely to make *storytelling* comments than were *novices*;  $t(33) = -2.74, p = .01$ . *Experts* were similar to *intermediates* in their use of this content focus category as shown by the data [Table 4].

<i>Content Focus</i>	<i>Expert</i>		<i>Intermediate</i>		<i>Novice</i>	
	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>	<i>M</i>	<i>SD</i>
<i>Appearance</i>	.04	0.03	.06	0.06	.05	0.03
<i>Assignment</i>	.03	0.03	.02	0.02	.02	0.02
<i>Mechanics</i>	.09	0.04	.06	0.05	.10	0.04
<i>Non-Specific</i>	.16	0.08	.16	0.08	.13	0.05
<i>Organization</i>	.19	0.08	.24	0.10	.16	0.04
<i>Storytelling</i>	.34	0.05	.32	0.11	.41	0.07
<i>Style</i>	.15	0.05	.15	0.07	.14	0.04

**Table 4: Group Proportions on Think Aloud Content Focus**

Based on observations of the protocol coders, *post hoc* analyses were performed on group differences regarding the extent to which scorers tended to withhold judgment during the decision making process. An examination of the think aloud protocols suggested that *experts* were more likely to approach the scoring task as a two-phased process. First, the *expert* scorer would read the essay, typically from beginning to end, without interrupting the reading to comment on the essay's content (i.e., without *monitor* comments). Second, the *expert* scorer would *review* the essay's contents and announce a score for the essay. Some *experts* made very few comments about the essay's content, seeming to arrive at an implicit decision rather than feeling the need to explicitly *review* the essay's content or provide a *rationale* for the assigned score.

*Novices* and *intermediates*, on the other hand, seemed to be more likely to *monitor* the essay's content, a process that interrupts the reading process. Some of the less proficient scorers even performed *monitoring* actions prior to reading any portion of the essay. These scorers took a cursory glance at the essay or read the first sentence of the essay and described what they expected the remainder of the essay to contain. These less proficient scorers were also more likely to interrupt the reading process to predict how they would eventually score the essay or to update a prior prediction. One of the *novice* scorers even assigned scores to most of the essays without reading the entire essay. Other scorers mentioned that they typically had a score in mind well before completing the essay, but none of them failed to read the entire essay during their think-aloud interviews.

To see whether these observations could be supported statistically, the number of *early decisions* for each scorer (i.e., scores that were announced during the reading of the essay) were summed across all essays. Table 5 shows the descriptive statistics for each proficiency group. Because these were *post hoc* analyses, Hartley's  $F_{max}$  test was used to test the homogeneity of variances assumption. This test revealed that the group variances were not equal,  $F_{max}(3,11) = 115.90$ ,  $p < .05$ , so the groups were compared with a studentized range statistic ( $q$ ) to test all pairwise comparisons, controlling type I errors at  $\alpha = .05$ . The differences between *experts* versus *intermediates* as well as the difference between *experts* versus *novices* were statistically significant ( $q(11) = -3.99$ ,  $p = .04$  and  $q(11) = -5.92$ ,  $p = .03$ , respectively) and the difference between *intermediates* and *novices* was not;  $q(19) = 0.85$ ,  $p > .10$  [Table 5].

<i>Group</i>	<i>M</i>	<i>SD</i>
<i>Experts</i>	0.42	0.90
<i>Intermediates</i>	8.33	9.69
<i>Novices</i>	6.33	6.17

**Table 5: Group Frequencies for Early Decisions**

## Discussion

These results are consistent with findings of similar studies of scorer cognition [Vaughan 1991; Huot 1993; Wolfe & Feltovich 1994]. However, our study suggests three trends in the development of scoring expertise that have not been offered by other researchers.

*Conclusion 1: Experts are more consistent than non-experts in the way they approach the task of scoring essays.* The data supporting *Hypothesis 1* indicate that *expert* essay scorers are more likely to use similar strategies for scoring than are *intermediates* and *novices*. One explanation of the dissimilarities in the three groups' approaches to scoring may be that these processing differences are due to differences in the knowledge structures that underlie expert-like performance, leading experts to use that knowledge more effectively. Similar conclusions have been drawn from the results of numerous studies of expert performance in a variety of problem-solving domains.

*Conclusion 2: Expert scorers are more likely to use more fluent methods of scoring essays than are non-experts.* That is, the data supporting *Hypothesis 3* suggest that *expert* scorers seem to utilize a more holistic strategy for scoring that uses a less iterative decision making pattern than that of non-experts. *Experts* seem to use strategies in which the scorer interprets the student writing through reading and reacting to the text, thus creating an image of the text. They then map the features of this text image onto the their mental representations of the scoring criteria. Through this process, judgments are made about how well the writer has demonstrated the various aspects of the scoring criteria, and a decision is formulated about the score to assign to the essay.

On the other hand, *intermediate* and *novice* scorers seem to use less fluent strategies for scoring. That is, non-experts seem to go through an alternating cycle of reading and monitoring portions of the essay. During each iteration, the scoreable features of that section of the essay are mapped onto the scorer's mental representation of the scoring criteria, and a preliminary score may be assigned to the essay. After completing this process for the entire essay, non-experts may review the essay prior to assigning a final score. However, they are less likely to do so than are *experts*.

This interpretation does not necessarily suggest that the approach taken by non-experts is inferior to that taken by *experts*. It is quite plausible that similar processing is occurring during the reading phase of an *expert's* scoring. However, more emphasis seems to be placed on the reading and comprehension process by *experts* than by *intermediates* and *novices* [Huot 1993]. It may be that experts have simply automated these procedures.

*Conclusion 3: There is little evidence to suggest that the content focus adopted by essay scorers is related to scoring proficiency.* Scorers in this study demonstrated similar emphases in their *content foci* as has been observed in other studies of scoring [Huot 1993; Vaughan 1991]. That is, primary attention has been given to *storytelling*, *organization*, and *style*. Unfortunately, the analyses associated with *Hypotheses 2* and *4* provided little evidence that *content focus* has any relationship with scoring proficiency.

The most appealing explanation for the observed differences in *processing action* use by the proficiency groups is that these differences are caused by structural differences in the knowledge upon which those actions operate.

Given the myriad of studies of expertise in other domains of human performance that have indicated that the primary difference between experts and non-experts lies in the manner in which domain knowledge is structured, a possible explanation for the failure to detect differences in the *content focus* of the proficiency groups in this study may be that the coding system used in this study was not sensitive enough to the kinds of differences in knowledge structures that make expert-like processing possible. With hindsight, it seems unlikely that this coding system could have identified such differences because it emphasizes the focus of the comments that a scorer makes rather than the structure of the knowledge (i.e., the relationships between concepts) that underlies those comments. Future studies using other means of analysis should aim to determine whether there are differences in these knowledge structures.

One application of the findings of this investigation concerns studies of scorer recruitment and training. This study has shown that scoring proficiency is highly evident in the manner in which an essay's contents are processed during evaluation. Previous efforts to train scorers have focused on developing their understandings of scoring criteria. Little, if any, attention has been directed toward developing frameworks of scoring. Future training studies should aim to determine whether non-expert scorers can be trained to use expert-like approaches to scoring essays and whether the adoption of these strategies leads to improved scoring accuracy. Even if these training studies fail to improve scoring performance, it may be possible to use the findings of future studies to make better and more efficient evaluative decisions about which scorer candidates are more likely to perform well on a scoring project.

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