# Lecture 10: Interval estimation Statistical Methods for Data Science

#### Yinan Yu

Department of Computer Science and Engineering

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#### Today

- Central limit theorem
  - Terminology
  - Standardization
  - Central limit theorem
- 2 Interval estimation
  - Confidence interval
  - Credible interval
- Summary





#### Learning outcome

- Be able to explain the following terminology:
  - Sample statistic, sampling distribution, sample mean, sample variance, standardization, z-table, t-table
  - Point estimation, interval estimation
  - Confidence interval, credible interval
- Be able to explain the central limit theorem (CLT)
- Be able to construct the following interval estimates:
  - Confidence interval for
    - ullet sample mean of i.i.d. sample with unknown  $\sigma$
    - unknown sampling distribution using bootstrap
  - Credible interval for a given posterior function



### Today

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  - Standardization
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# Terminology





#### Terminology

- (Statistical) population: all items of interest (e.g., all ducks in the world)
- Sample: a random data set  $\{x_1, x_2, \cdots, x_N\}$ ; the corresponding random variables are denoted as  $X_1, X_2, \cdots, X_N$ ; a subset of the population (e.g., the 20 ducks you have weighed)
- i.i.d. sample:  $X_1, X_2 \cdots, X_N$  are i.i.d. random variables
- Sample statistic: a statistic computed from a sample For example,
  - Sample mean:

$$\bar{X} = \frac{1}{N} \sum_{i=1}^{N} X_i$$

• Sample variance:

$$S^{2} = \frac{1}{N-1} \sum_{i=1}^{N} (X_{i} - \bar{X})^{2}$$

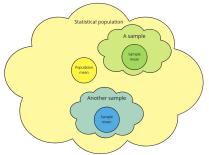
Note: capital letters and small letters are used to denote random variables and the values, respectively.





# Terminology (cont.)

 Sampling distribution: the probability distribution of a sample statistic that is computed from a random sample (of size N)



ullet Asymptotic: in this context, asymptotic means  ${\cal N} o \infty$ 

What's the difference between the mean of a Gaussian distribution is random (Bayesianist) vs the sample mean is random?





# Awesome properties of Gaussian random variables

- Let  $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$  be a Gaussian random variable, then the following random variables are also Gaussian (location scale family)
  - Scaling (scale):  $tX \sim \mathcal{N}(t\mu_X, t^2\sigma_X^2)$ ,  $t \neq 0$  is a constant
  - Translation (location):  $X + c \sim \mathcal{N}(\mu_X + c, \sigma_X^2)$ , c is a constant
  - $tX + c \sim \mathcal{N}(t\mu_X + c, t^2\sigma_X^2)$
- Let  $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$  and  $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$  be two independent Gaussian random variables, then the following random variables are also Gaussian
  - $X + Y \sim \mathcal{N}(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Y^2)$
  - $X Y \sim \mathcal{N}(\mu_X \mu_Y, \sigma_X^2 + \sigma_Y^2)$





Terminology
Standardization
Central limit theorem

#### Standardization

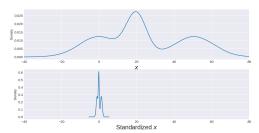




#### Standardization

- Why standardization? We want to translate and scale data into a standard shape so that
  we can use standard tools to compare and analyze it
- Let X be a random variable that follows any probability distribution with mean  $\mu$  and standard deviation  $\sigma$ . The standardization of X is

$$Y = \frac{X - \mu}{\sigma}$$



Question: what is the mean and standard deviation of Y? Random variable Y has mean 0 and standard deviation 1

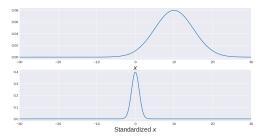




#### Standardization

• Let X be a random variable following a Gaussian distribution with mean  $\mu$  and standard deviation  $\sigma$ , i.e.  $X \sim \mathcal{N}(\mu, \sigma^2)$ ; the standardization of X is

$$Z = \frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1) \tag{1}$$



The distribution  $\mathcal{N}(0,1)$  is called a standard Gaussian (normal) distribution





#### Standard Gaussian distribution

- Remember how much we love Gaussian distributions? We love the standard Gaussian distribution even more! We love it so much that we gave its CDF a special name:  $\Phi(z)$
- There is a table describing the quantiles of the standard Gaussian (the z-table)

```
z + 0.00 + 0.01 + 0.02 + 0.03 + 0.04 + 0.05 + 0.06 + 0.07 + 0.08 + 0.09
0.0 0.50000 0.50399 0.50798 0.51197 0.51595 0.51994 0.52392 0.52790 0.53188 0.53586
0.1 0.53983 0.54380 0.54776 0.55172 0.55567 0.55962 0.56360 0.56749 0.57142 0.57535
0.2 0.57926 0.58317 0.58706 0.59095 0.59483 0.59871 0.60257 0.60642 0.61026 0.61409
0.3 0.61791 0.62172 0.62552 0.62930 0.63307 0.63683 0.64058 0.64431 0.64803 0.65173
0.4 0.65542 0.65910 0.66276 0.66640 0.67003 0.67364 0.67724 0.68082 0.68439 0.68793
0.5 0.69146 0.69497 0.69847 0.70194 0.70540 0.70884 0.71226 0.71566 0.71904 0.72240
0.6 0.72575 0.72907 0.73237 0.73565 0.73891 0.74215 0.74537 0.74857 0.75175 0.75490
0.7 0.75804 0.76115 0.76424 0.76730 0.77035 0.77337 0.77637 0.77935 0.78230 0.78524
0.8 0.78814 0.79103 0.79389 0.79673 0.79955 0.80234 0.80511 0.80785 0.81057 0.81327
0.9 0.81594 0.81859 0.82121 0.82381 0.82639 0.82894 0.83147 0.83398 0.83646 0.83891
1.0 0.84134 0.84375 0.84614 0.84849 0.85083 0.85314 0.85543 0.85769 0.85993 0.86214
1.1 0.86433 0.86650 0.86864 0.87076 0.87286 0.87493 0.87698 0.87900 0.88100 0.88298
1.2 0.88493 0.88686 0.88877 0.89065 0.89251 0.89435 0.89617 0.89796 0.89973 0.90147
1.3 0.90320 0.90490 0.90658 0.90824 0.90988 0.91149 0.91308 0.91466 0.91621 0.91774
 1.4 0.91924 0.92073 0.92220 0.92364 0.92507 0.92647 0.92785 0.92922 0.93056 0.93189
1.5 0.93319 0.93448 0.93574 0.93699 0.93822 0.93943 0.94062 0.94179 0.94295 0.94408
 1.6 0.94520 0.94630 0.94738 0.94845 0.94950 0.95053 0.95154 0.95254 0.95352 0.95449
1.7 0.95543 0.95637 0.95728 0.95818 0.95907 0.95994 0.96080 0.96164 0.96246 0.96327
 1.8 0.96407 0.96485 0.96562 0.96638 0.96712 0.96784 0.96856 0.96926 0.96995 0.97062
1.9 0.97128 0.97193 0.97257 0.97320 0.97381 0.97441 0.97500 0.97558 0.97615 0.97670
```

- Each row represents the integer and the first decimal of z
- Each column represents the second decimal of z
- · Each cell is the

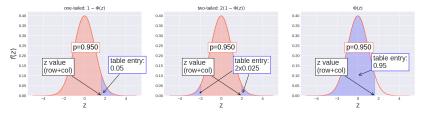
```
P(Z \le \text{row} + \text{column}) = \Phi(\text{row} + \text{column})
= stats.norm.cdf(x=row + column, loc=0, scale=1)
```





# Standard Gaussian distribution (cont.)

• There are different representations of the z-table; the difference is what is inside each cell, e.g.  $\Phi(\text{row} + \text{column})$ ,  $2(1 - \Phi(\text{row} + \text{column}))$ ,  $1 - \Phi(\text{row} + \text{column})$  or  $\frac{1}{2}(1 - \Phi(\text{row} + \text{column}))$ ; but the principle is the same; for now we use the version with  $\Phi(\text{row} + \text{column})$ 

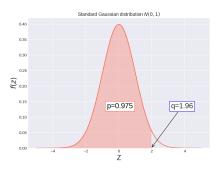


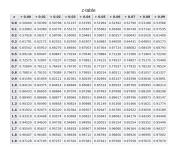
ullet Due to symmetry, there are only positive values for z in the z-table



#### Standard Gaussian distribution (cont.)

#### Exercise:





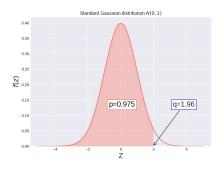
Try to find the corresponding pair (p, q) = (0.975, 1.96) in the z-table (60 secs).

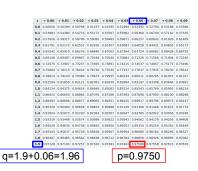




### Standard Gaussian distribution (cont.)

#### Answer:





Note: the table itself is not important (we use a computer these days); the point is to reflect on the meaning of z values (quantiles) and the related probabilities (CDFs)





Standardization
Central limit theorem

#### Central limit theorem





#### Motivation and use cases

So far, we have been looking at **distributions** (centrality and spread); the central limit theorem is about the **mean** (centrality only); do we care about the mean that much?

- Yes, we do!
- Example: we want to test the effectiveness of a drug; a patient can be either cured by this drug or not cured, i.e., we can model the data using a (2 secs) Bernoulli distribution with parameter (2 second) p (cure rate) and the maximum likelihood estimation of p is the (4 secs) sample mean
- In general, we are often interested in how things work "on average"





#### Distribution of the sample mean

- You have 1000 ducks
- Now, you take 30 of them and measure the sample mean of their weights  $x_i$ :

$$\hat{\mu}_1 = \frac{1}{30} \sum_{i=1}^{30} x_i$$

• Then you take another 30 ducks to measure the sample mean of their weights  $y_i$ :

$$\hat{\mu}_2 = \frac{1}{30} \sum_{i=1}^{30} y_i$$

- You do this experiment 100 times and plot the histogram of these 100 sample means  $\hat{\mu}_i$  for  $j=1,\cdots,100$
- ullet Then you realize these sample means  $\hat{\mu}_j$  seem to follow a Gaussian distribution







### Distribution of the sample mean (cont.)

- ullet The colors of your 1000 ducks can be either red  $t_i=0$  or blue  $t_i=1$
- Now, you take 30 of them and measure the sample mean of their color  $t_i$ :

$$\hat{n}_1 = \frac{1}{30} \sum_{i=1}^{30} t_i = \frac{1}{30} (1 + 1 + 0 + 1 + \dots + 1)$$

Note: here  $t_i \in \{0,1\}$  is discrete

• You take another 30 ducks and measure the sample mean of their color  $t_i$ :

$$\hat{n}_2 = \frac{1}{30} \sum_{i=1}^{30} t_i = \frac{1}{30} (0 + 0 + 0 + 1 + \dots + 1 + 0)$$

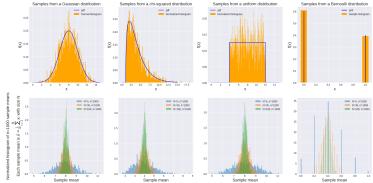
- $\bullet$  You do this experiment 100 times and plot the histogram of these 100 sample means  $\hat{n}_i$
- Then you realize these sample means  $\hat{n}_j$  also seem to follow a Gaussian distribution |||???||





### Distribution of the sample mean (cont.)

• In fact, this is true for i.i.d. samples drawn from ANY probability distribution



- The larger the sample size N (in the previous example N=30), the "more Gaussian" it becomes
- A rule of thumb: N > 30
- If the data distribution is Gaussian-like (bell-shaped, symmetric), only a small sample size is needed for the sample mean to be Gaussian





#### Central limit theorem

- One of the most important results in probability theory and statistics
- Given an i.i.d. sample  $X_1, X_2, \dots, X_N$  from ANY probability distribution with finite mean  $\mu$  and variance  $\sigma^2$  (most distributions satisfy this!), when the sample size N is sufficiently large, the sample mean approximately follows a Gaussian distribution with mean  $\mu$  and variance  $\frac{\sigma^2}{N}$ , i.e.,

$$\bar{X} \sim \mathcal{N}(\mu, \frac{\sigma^2}{N})$$
 (2)

where  $\bar{X} = \frac{1}{N} \sum_{i=1}^{N} X_i$  is the sample mean

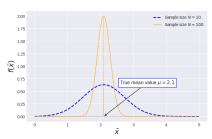




# Central limit theorem (cont.)

How to interpret this?

$$ar{X} \sim \mathcal{N}(\mu, \frac{\sigma^2}{N})$$



- ullet The sample mean  $ar{X}$  is around the true mean value  $\mu$
- The "deviation" of  $\bar{X}$  from  $\mu$  is  $\frac{\sigma^2}{N}$ ; the larger N, the smaller the deviation



# Estimation error $\bar{X} - \mu$

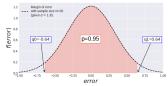
We are interested in estimating the mean value  $\mu$  We use the sample mean  $\bar{X}$  to estimate the mean value  $\mu$  We are interested in how good this estimation is





# Analysis of the estimation error $\bar{X} - \mu$

- In many use cases, we want to estimate the population mean μ using the sample mean from one sample {x<sub>1</sub>,···, x<sub>N</sub>} and we are interested in the statistics of the estimation error
- Random variable  $X_1, \dots, X_N$  from any distribution
- Assumption: i.i.d. with known standard deviation  $\sigma$  and unknown mean  $\mu$
- Random variable of interest:  $\bar{X} = \frac{1}{N} \sum_{i=1}^{N} X_i$ 
  - From CLT (page 21), we know that for a large N, the sample mean approximately follows a Gaussian distribution  $\bar{X} \sim \mathcal{N}(\mu, \frac{\sigma^2}{N})$ :  $\bar{X}$  is around the true mean  $\mu$
  - Let  $\mathcal{E} = \bar{X} \mu$  be the estimation **error**; what distribution does  $\mathcal{E}$  follow (awesome properties of Gaussian 30 secs)?  $\mathcal{E} \sim \mathcal{N}(0, \frac{\sigma^2}{N})$ ; can we plot the PDF of  $\mathcal{E}$ ? (5 secs) Yes!  $\sigma$  and N are both known!



- Given q0 and q1, how to interpret this plot (5 secs)? **95% of the time, the error**  $\bar{X} \mu$  is within q0 = -0.64 and q1 = 0.64
- Now it's pretty cool because not only can we estimate the mean (using the sample mean), but we also get
  the margin of error!
- This 95% is called the confidence level; for a given confidence level, we can find a corresponding interval
   (q0, q1)





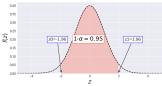
### Calculate the margin of error

• For a given confidence level, denoted as  $1-\alpha$ , how do we find this interval for the error term in Python? We can use the function **ppf** from **scipy.stats** 



# Find a standardized expression for the margin of error

- Standardize (cf. page 11)  $\mathcal E$  by  $\frac{\mathcal E}{\sigma/\sqrt{N}}=\frac{\bar X-\mu}{\sigma/\sqrt{N}}\sim \mathcal N(0,1)$
- We just learned that there is a special name for the standard Gaussian distributed random variable  $Z \sim \mathcal{N}(0,1)$  let  $Z = \frac{\bar{X} \mu}{\sigma/\sqrt{N}}$
- Now we have an expression for the error term in terms of  $Z\colon \mathcal{E}=\bar{X}-\mu=\mathbf{Z}\frac{\sigma}{\sqrt{N}}$
- The only random variable here is  $Z \sim \mathcal{N}(0,1)$



- $\bullet$  We can use a two-tailed z-table (cf. page 13) to find the values for z0 and z1
- $\bullet$  In order to find an interval for  ${\cal E},$  we just need to look at

$$(z0\frac{\sigma}{\sqrt{N}},z1\frac{\sigma}{\sqrt{N}})$$

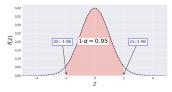




# Find a standardized expression for the margin of error (cont.)

 $\bullet$  For example, with  $1-\alpha=95\%$  confidence level, the error is within

$$\left(-1.96\frac{\sigma}{\sqrt{N}}, \ 1.96\frac{\sigma}{\sqrt{N}}\right)$$



- Generally speaking, the value z1 (denoted by  $z_{\alpha/2}$ ) is the quantile at  $1-\alpha/2$ ; the value of  $z_{\alpha/2}$  is called the (right) critical value;  $\frac{\sigma}{\sqrt{N}}$  is called the standard error; in this example, we have  $z_{\alpha/2}=z1=-z0=1.96$
- ullet Why two-tailed z-table: there are two tails  $z \leq -z_{\alpha/2}$  and  $z \geq z_{\alpha/2}$



# Find a standardized expression for the margin of error (cont.)

```
• In Python
    std = 1.8
    N = 30
    alpha = 0.05
    confidence_level = 1 - alpha # 95% confidence level
    z0 = stats.norm.ppf(alpha/2, 0, 1)
    z1 = stats.norm.ppf(confidence_level+alpha/2, 0, 1)
    print(z0*std/math.sqrt(N), z1*std/math.sqrt(N))
    >> (-0.6441098917381766, 0.6441098917381766)
```





# Find a standardized expression for the margin of error (cont.)

• For a given sample with an estimate  $\bar{x}$  (note: here the small letter  $\bar{x}$  denotes the value of the estimate itself instead of a random variable), it's more convenient to have this margin of error around  $\bar{x}$  instead - so that we can say: the estimated mean is  $\bar{x}$  with this uncertainty:

$$\left(\bar{x}-z_{\alpha/2}\frac{\sigma}{\sqrt{N}},\ \bar{x}+z_{\alpha/2}\frac{\sigma}{\sqrt{N}}\right)$$

- This is called the confidence interval
- The confidence interval for the sample mean is exact when the data distribution is Gaussian, otherwise it is an approximation under the central limit theorem
- This calculation is called **interval estimation**, because it gives an interval estimate  $\left(\bar{x}-z_{\alpha/2}\frac{\sigma}{\sqrt{N}},\ \bar{x}+z_{\alpha/2}\frac{\sigma}{\sqrt{N}}\right)$  instead of a single value estimate as in MAP or MLE





# Today

- Central limit theorem
- 2 Interval estimation
  - Confidence interval
  - Credible interval
- Summary





#### One quick question

- What does it mean by something being random?
  - We can use a random variable to model the data
  - The random variable has an underlying probability distribution
  - We can use distribution functions (e.g. CDF, PDF/PMF) to describe the probability distribution
- Actually another quick question... What does it mean by unknown probability distribution?
  - The family of the probability distribution is unknown, e.g. Gaussian? Uniform?
  - Or given a family of probability distributions, some parameters of the probability distribution are unknown, e.g. a Gaussian distribution with unknown mean or unknown variance





### More questions

- How to estimate an unknown probability distribution?
  - Q-Q plot to estimate the family of probability distributions,
     e.g. data distribution vs Gaussian distribution
  - Parameter estimation techniques (e.g. MLE, MAP) to estimate the unknown parameters
- Given a probability distribution with unknown mean value  $\mu$ , is  $\mu$  random?
  - In Bayesian approaches (e.g. MAP), it is assumed random
  - In Frequentist approaches (e.g. MLE), it is not assumed random; it is an unknown constant
- Is the sample mean random?
  - Yes, it is random. The sample mean is a sample statistic; a sample statistic is computed from a sample; a sample is random and hence the sample statistic is random.





#### I can't promise but one last question...

- Is the sample mean always the MLE for the mean?
  - It is the MLE for the mean value of Gaussian distributions, but it is not always the MLE for the mean value of any distribution

Now we are done with questions
UNLESS...
Nah we are done





#### Interval estimation

- MLE and MAP are point estimation techniques since they only return one single value, i.e., a point, for the parameter estimation
- However, we are often interested in the uncertainty
   associated with the point estimate; a point estimate +
   uncertainty is called an interval estimate since they return an
   interval instead a single value



#### Confidence interval





### Confidence interval (CI)

- Data:  $x_1, \dots, x_N$
- Random variable:  $X_1, \dots, X_N$  with i.i.d. assumption
- Parameter of interest:  $\theta$ , e.g. the mean  $\mu$
- Estimate:  $\hat{\theta}$ , e.g. the sample mean  $\bar{x}$
- Confidence interval for a given confidence level  $1 \alpha$  (e.g. 95%)
  - Definition:

confidence interval =  $(\hat{\theta} - \text{margin of error}, \hat{\theta} + \text{margin of error})$ where

margin of error = critical value  $\times$  standard error of  $\hat{\theta}$ 

Distribution of $X_i$	Scenario	θ	$\hat{ heta}$ (sampling distribution)	Critical value	Standard error	Confidence interval	Note
i.i.d. Gaussian	<b>☑</b> σ known	mean	sample mean $\bar{x}$	$z_{\alpha/2}$	$\frac{\sigma}{\sqrt{N}}$	$\left(\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{N}}, \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{N}}\right)$	
	? σ unknown		(Gaussian distribution)	$t_{\alpha/2}$	<u>s</u> √N	$\left(\bar{x} - t_{\alpha/2} \frac{s}{\sqrt{N}}, \bar{x} + t_{\alpha/2} \frac{s}{\sqrt{N}}\right)$	exact
i.i.d.	<b>☑</b> σ known	incun	sample mean $\bar{x}$	$z_{\alpha/2}$	$\frac{\sigma}{\sqrt{N}}$	$\left(\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{N}}, \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{N}}\right)$	approximate
	? σ unknown		(approximately Gaussian under CLT)	$t_{\alpha/2}$	<u>s</u> √N	$\left(\bar{x} - t_{\alpha/2} \frac{s}{\sqrt{N}}, \bar{x} + t_{\alpha/2} \frac{s}{\sqrt{N}}\right)$	for large N
i.i.d.	<b>A</b> -	any	MLE (asymptotically Gaussian)	$z_{\alpha/2}$	$\frac{1}{\sqrt{NI_N(\hat{\theta})}}$	$\left(\hat{\theta} - z_{\alpha/2} \frac{1}{\sqrt{NI_N(\hat{\theta})}}, \hat{\theta} + z_{\alpha/2} \frac{1}{\sqrt{NI_N(\hat{\theta})}}\right)$	asymptotic
i.i.d.	? -	any	any statistic (any distribution)	bootstrap the error quantile		$(\hat{\theta} - \epsilon_{1-\alpha/2}, \hat{\theta} - \epsilon_{\alpha/2})$	approximate

Calculation:

where  $\sigma$  is the standard deviation of the  $X_i$  and s the sample standard deviation





#### Calculation of the confidence interval

Data:  $x_1, \dots, x_N$ 

**Random variable**:  $X_1, \dots, X_N$  i.i.d. with standard deviation  $\sigma$ 

- CI for Gaussian sampling distribution (exact, approximate, asymptotic):
  - Parameter of interest: mean value Estimation method: sample mean  $\bar{x}$

$$\sigma$$
 known:  $\left(\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{N}}, \ \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{N}}\right)$  (cf. page 29)

? 
$$\sigma$$
 unknown:  $\left(ar{x}-t_{lpha/2}rac{\sigma}{\sqrt{N}},\;ar{x}+t_{lpha/2}rac{\sigma}{\sqrt{N}}
ight)$ 

Parameter of interest: any statistic
 Estimation method: MLE (cf. lecture 3 properties of MLE)

$$\widehat{\mathbf{p}} \quad \text{[not required]} \ \left( \hat{\theta} - z_{\alpha/2} \frac{1}{\sqrt{N I_N(\hat{\theta})}}, \hat{\theta} + z_{\alpha/2} \frac{1}{\sqrt{N I_N(\hat{\theta})}} \right)$$

- CI for unknown sampling distribution
  - Parameter of interest: any parameter, e.g. median Estimation method: any method
    - ? Bootstrap  $\left(\hat{\theta} \epsilon_{1-\alpha/2}, \hat{\theta} \epsilon_{\alpha/2}\right)$





#### CI for unknown $\sigma$

- When the standard deviation  $\sigma$  is known, we have shown the standardization of the error term  $\frac{\mathcal{E}}{\sigma/\sqrt{N}} = \frac{X-\mu}{\sigma/\sqrt{N}} \sim \mathcal{N}(0,1)$  (cf. page. 26).
- When  $\sigma$  is unknown, which is the most common case, we replace  $\sigma$  by its estimate  $\hat{\sigma}$  - the sample standard deviation s - random variable S

$$\hat{\sigma} = S = \sqrt{\frac{1}{N-1} \sum_{i=1}^{N} (X_i - \bar{X})^2}$$

• Now the standardization of the error term becomes (random, constant):

$$\frac{\mathcal{E}}{\sigma/\sqrt{N}} o \frac{\mathcal{E}}{S/\sqrt{N}} = \frac{\bar{X} - \mu}{S/\sqrt{N}} \sim t(N-1)$$

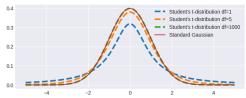
• Compared to the case with known  $\sigma$ ,  $\frac{\bar{X}-\mu}{\sigma/\sqrt{N}}\sim\mathcal{N}(0,1)$ , the distribution of  $\frac{\bar{X}-\mu}{c/\sqrt{N}}$  is no longer the standard Gaussian ( $\frac{\mu}{S/\sqrt{N}}$  is no longer a constant because S is a random variable). Instead, it follows a **Student's t-distribution** t. The Student's t-distribution has one parameter df = N - 1 (degrees of freedom).





# r CI for unknown $\sigma$ (cont.)

- The Student's t-distribution is a function of the sample size: df = N 1
- Think of it as a standard Gaussian compensated for the small sample size. For a large N, they become very similar.







## r CI for unknown $\sigma$ (cont.)

- t-table: similar to the z-table for the standard Gaussian distribution, there is a t-table for the Student's t-distribution (image from http://www.ttable.org/).
- each cell = stats.t.ppf(q=cum.prob, df=N-1, loc=0, scale=1)
- $\alpha = \text{two-tails}$  and confidence level =  $1 \alpha$

t.50	t 25	t.00	t.ss	t.50	t.ss	t.975	t.99	t.995	t.999	t.9995
0.50	0.25	0.20	0.15	0.10	0.05	0.025	0.01	0.005	0.001	0.0005
1.00	0.50	0.40	0.30	0.20	0.10	0.05	0.02	0.01	0.002	0.001
0.000	1.000	1.376	1.963	3.078	6.314	12.71	31.82	63.66	318.31	636.62
0.000	0.816	1.061	1.386	1.886	2.920	4.303	6.965	9.925	22.327	31.599
0.000	0.765	0.978	1.250	1.638	2.353	3.182	4.541	5.841	10.215	12.924
0.000	0.741	0.941	1.190	1.533	2.132	2.776	3.747	4.604	7.173	8.610
0.000	0.727	0.920	1.156	1.476	2.015	2.571	3.365	4.032	5.893	6.869
0.000	0.718	0.906	1.134	1.440	1.943	2.447	3.143	3.707	5.208	5.959
0.000	0.711	0.896	1.119	1.415	1.895	2.365	2.998	3.499	4.785	5.408
0.000	0.706	0.889	1.108	1.397	1.860	2.306	2.896	3,355	4.501	5.041
0.000	0.703	0.883	1.100	1.383	1.833	2.262	2.821	3.250	4.297	4.781
0.000	0.700	0.879	1.093	1.372	1.812	2.228	2.764	3.169	4.144	4.587
0.000	0.697	0.876	1.088	1.363	1.796	2.201	2.718	3.106	4.025	4.437
0.000	0.695	0.873	1.083	1.356	1.782	2.179	2.681	3.055	3.930	4.318
		0.870	1.079	1.350	1.771	2.160		3.012	3.852	4.221
0.000	0.692	0.868	1.076	1.345	1.761	2.145	2.624	2.977	3.787	4.140
0.000	0.691	0.865	1.074	1.341	1.753	2.131	2.583	2.947	3.733	4.073
0.000	0.690	0.863	1.069	1.337	1.740	2.120	2.583	2.921	3.646	3.965
0.000	0.688	0.862	1.069	1.333	1.734	2.110	2.552	2.878	3.610	3.922
0.000	0.688	0.861	1.066	1.328	1.729	2.093	2.539	2.861	3.579	3.883
0.000	0.687	0.860	1.064	1.325	1.725	2.093	2.539	2.845	3.552	3.850
0.000	0.686	0.859	1.063	1.323	1.721	2.080	2.518	2.831	3.527	3.819
0.000	0.686	0.858	1.061	1.323	1.717	2.074	2.508	2.819	3.505	3.792
0.000	0.685	0.858	1.060	1.319	1.714	2.069	2.500	2.807	3.485	3.768
0.000	0.685	0.857	1.059	1.318	1.711	2.064	2.492	2.797	3.467	3.745
0.000	0.684	0.856	1.058	1.316	1.708	2.060	2.485	2.787	3,450	3.725
0.000	0.684	0.856	1.058	1.315	1,706	2.056	2,479	2,779	3,435	3,707
0.000	0.684	0.855	1.057	1.314	1.703	2.052	2.473	2.771	3,421	3.690
0.000	0.683	0.855	1.056	1.313	1.701	2.048	2.467	2.763	3,408	3.674
0.000	0.683	0.854	1.055	1.311	1.699	2.045	2.462	2.756	3.396	3,659
0.000	0.683	0.854	1.055	1.310	1.697	2.042	2.457	2.750	3.385	3.646
0.000	0.681	0.851	1.050	1.303	1.684	2.021	2.423	2.704	3.307	3.551
0.000	0.679	0.848	1.045	1.296	1.671	2.000	2.390	2.660	3.232	3,460
0.000	0.678	0.846	1.043	1.292	1.664	1.990	2.374	2.639	3.195	3.416
0.000	0.677	0.845	1.042	1.290	1.660	1.984	2.364	2.626	3,174	3,390
0.000	0.675	0.842	1.037	1.282	1.646	1.962	2.330	2.581	3.098	3.300
0.000	0.674	0.842	1.036	1.282	1.645	1.960	2.326	2.576	3.090	3.291
0%	50%	60%	70%	80%	90%	95%	98%	99%	99.8%	99.9%
0.000		0.674	0.674 0.842	0.674 0.842 1.036	0.674 0.842 1.036 1.282 50% 60% 70% 80%	0.674 0.842 1.036 1.282 1.645 50% 60% 70% 80% 90%	0.674 0.842 1.036 1.282 1.645 1.960	0.674 0.842 1.036 1.282 1.645 1.960 2.326 50% 60% 70% 80% 90% 95% 98%	0.674         0.842         1.036         1.282         1.645         1.960         2.326         2.576           50%         60%         70%         80%         90%         95%         98%         99%	0.674 0.842 1.036 1.282 1.645 1.960 2.326 2.576 3.090 50% 60% 70% 80% 90% 95% 98% 99% 99.8%





#### Summary

Data:  $x_1, \dots, x_N$ 

Random variable:  $X_1, \dots, X_N$  i.i.d. with standard deviation  $\sigma$  CI for unknown  $\sigma$  with Gaussian sampling distribution

$$\left(\bar{x}-t_{\alpha/2}\frac{s}{\sqrt{N}},\bar{x}+t_{\alpha/2}\frac{s}{\sqrt{N}}\right)$$



## CI for unknown sampling distribution

- Sample mean approximately follows a Gaussian distribution under the central limit theorem, but most other statistics do not have such luxury
- When the sampling distribution is unknown, we cannot use the t-table or z-table to find the critical values
- Recall the definition of CI: confidence interval =  $(\hat{\theta} \text{margin of error})$
- One way to find this CI is to approximate the margin of error using bootstrap





#### Bootstrap

- Data:  $x_1, \dots, x_N$
- Random variables:  $X_1, \dots, X_N$  i.i.d. from any distribution
- Parameter of interest: any  $\theta$
- Estimation method: any method
- Confidence interval:  $(\hat{\theta} \epsilon_{1-\alpha/2}, \hat{\theta} \epsilon_{\alpha/2})$ , where  $\epsilon_p$  denotes the quantile of the error term at p
- Where does it come from?
  - Confidence interval:  $100 * (1 \alpha)\%$  of the time, the error term  $\hat{\theta} - \theta$  is within interval  $(\epsilon_{\alpha/2}, \epsilon_{1-\alpha/2})$
  - $P(\epsilon_{\alpha/2} \leq \hat{\theta} \theta \leq \epsilon_{1-\alpha/2}) = 1 \alpha$
  - $P(\epsilon_{\alpha/2} \leq \hat{\theta} \theta \leq \epsilon_{1-\alpha/2}) \iff P(\hat{\theta} \epsilon_{1-\alpha/2} \leq \theta \leq \hat{\theta} \epsilon_{\alpha/2})$
- But the error term  $\epsilon_p$  is unknown we had CLT for  $\mu$ , but not for any  $\theta$

The idea of bootstrap is to approximate the error  $\epsilon_p$  directly from data





Lecture 10: Interval estimation

# Bootstrap example

Given a data set  $\mathcal{X} = \{1, 2, 3, 4, 5\}$  with size N = 5 and  $\hat{\theta} = median(\mathcal{X}) = 3$  estimated from this data set, construct CI with 95% confidence level Intuition:

- $oldsymbol{\hat{ heta}}=3$  is approximating the true median heta
- Construct statistic  $m_i$  to approximate  $\hat{\theta}$
- ullet Then we can use the "mock up" error  $m_i \hat{ heta}$  to approximate the actual error  $\hat{ heta} heta$

#### Steps:

#### Sample with replacement

```
Step 1.1: Randomly choose 5 elements from \mathcal{X}: \mathcal{X}_1^* = \{1,2,1,1,4\}
```

Step 1.2: Compute the median from 
$$\mathcal{X}_1^*$$
:  $m_1 = 1.0$ 

Step 2.1: Randomly choose 5 elements from 
$$\mathcal{X}$$
:  $\mathcal{X}_2^* = \{2,5,2,4,4\}$ 

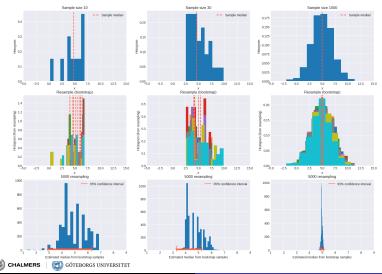
Step 2.2: Compute the median from 
$$\mathcal{X}_2^*$$
:  $\textit{m}_2 = 4.0$ 

- Repeat this 100 times and get the set  $\{m_1, \cdots, m_{100}\}$
- Compute  $\epsilon^i = m_i 3$  for  $i = 1, \dots, 100$
- ullet Compute 0.025-quantile  $\epsilon_{0.025}$  and 0.975-quantile  $\epsilon_{0.975}$  from the set  $\{\epsilon^1,\cdots,\epsilon^{100}\}$
- The 95% CI is constructed as  $(3 \epsilon_{0.975}, 3 \epsilon_{0.025})$





# Bootstrap example (cont.)



# CI for unknown sampling distribution using bootstrap

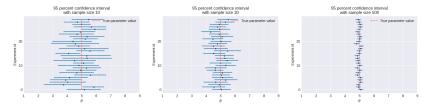
- Intuition:
  - $\hat{\theta}$  is approximating  $\theta$
  - $\hat{\theta}_i$  is approximating  $\hat{\theta}$
  - We can use  $\hat{\theta}_i \hat{\theta}$  to approximate  $\hat{\theta} \theta$
- Steps Given a data set  $\mathcal X$  with size N and a statistic  $\hat{\theta}$  computed from this data set, construct CI with  $1-\alpha$  confidence level:
  - Choose a large n
  - For  $i = 1, \dots, n$ , repeat
    - Sample N elements from  $\mathcal{X}$  with replacement:  $\mathcal{X}_i^*$
    - Estimate the parameter of interest from  $\mathcal{X}_{i}^{*}$ :  $\hat{\theta}_{i}$
    - Compute  $\epsilon^i = \hat{\theta}_i \hat{\theta}$
  - Compute  $\alpha/2$ -quantile  $\epsilon_{\alpha/2}$  and  $1-\alpha/2$ -quantile  $\epsilon_{1-\alpha/2}$  from the set  $\{\epsilon^1,\cdots,\epsilon^n\}$
  - ullet The 1-lpha CI is constructed as  $\left(\hat{ heta}-\epsilon_{1-lpha/2},\hat{ heta}-\epsilon_{lpha/2}
    ight)$
- Note: there are many alternative methods for bootstrap; the exact method needs to be described when you talk about bootstrap





#### Confidence interval interpretation

- Confidence interval is random (data is random; statistic is random); the true parameter value  $\theta$  is not random (illustrated in the image)
- ullet A 95% confidence interval means that 95% of the time, the interval covers the true value heta



- Question 1: with the same problem setup, the larger the confidence level,
  - A. the wider the confidence interval
  - B. the narrower the confidence interval

Answer: A

- Question 2: for a given confidence level, a good estimate has
  - A. a wide confidence interval
  - B. a narrow confidence interval

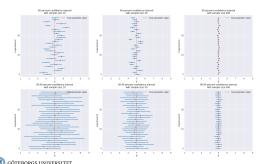
Answer: B





#### Confidence level interpretation

- If we compare 60% CI with 99.99% CI, the 60% CI does not always cover the true value  $\theta=5$  (it only covers it 60% of the time). On the other hand, the 99.99% CI covers the true value pretty much all the time. From this perspective, 99.99% CI is more meaningful to use as a quality measure.
- However, 99.99% CI can be very wide of course since it promises to cover the true value 99.99% of the time. A wide interval might not be meaningful sometimes, e.g. if you claim that you have estimated  $\hat{\theta}=4.3$  and you are 100% sure that the interval  $(4.3-\infty,4.3+\infty)$  contains the true value, your client might get mad.





#### Credible interval





# Credible interval for Bayesian approach

- In maximum a posteriori estimation, the parameter of interest  $\theta$  is modeled as a random variable  $\theta$  is generated from an underlying probability distribution described by  $f(\theta)$
- Technically, any interval (a, b) with  $P(a \le \Theta \le b) = 0.95$  is a 95% credible interval, but not all of them make sense, e.g.





• There are different techniques for choosing this interval



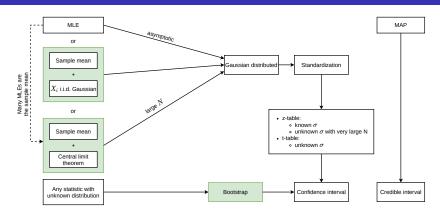
## Credible interval for Bayesian approach (cont.)

• In Python, for a given posterior (e.g. a standard Gaussian distribution  $\mathcal{N}(0,1)$ ), the .interval method computes the interval with equal areas around the median:

```
posterior = stats.norm(loc=0, scale=1)
credible_interval = posterior.interval(0.95)
```



## Recap





## Today

- Central limit theorem
- 2 Interval estimation
- Summary





#### Summary

#### So far:

- Data types and data containers
- Descriptive data analysis: descriptive statistics, visualization
- Probability distributions, events, random variables, PMF, PDF, parameters
- CDF, Q-Q plot, how to compare two distributions (data vs theoretical, data vs data)
- Modeling
- Parameter estimation: maximum likelihood estimation (MLE) and maximum a posteriori estimation (MAP)
- Classification, multinomial naive Bayes classifier, Gaussian naive Bayes classifier
- Clustering, cluster tendency
- Centroid clustering, k-means, parameter estimation, SSE, Silhouette score
- Gaussian Mixture Models, AIC/BIC
- The EM algorithm
- Central limit theorem, interval estimation

#### Next:

Hypothesis testing





#### See you next week!





