



CSCI E-106: Data Modeling  
Assignment 0  
Due: January, 28 2019 at 7pm EST

**Instructions:** Students should submit their reports on Canvas. The report needs to contain copies of the questions, step-by-step walk-through solutions, and final answers clearly indicated. For the questions 1 to 7, please solve it by hand; for questions 8 to 10, please submit two files: (1) a R Markdown file (.Rmd extension) and (2) a PDF document generated using knitr.

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For the questions 1 to 3, calculate the mean and variance for each of the following probability distributions

1.  $f_X(x) = ax^{a-1}$ ,  $0 < x < 1$ ,  $a > 0$
2.  $f_X(x) = \frac{1}{n}$ ,  $x = 1, \dots, n$ ,  $n > 0$  an integer
3.  $f_X(x) = \frac{3}{2}(x-1)^2$ ,  $0 < x < 2$

Questions for Joint Distribution:

4. Let  $X_1$  and  $X_2$  be independent standard normal distribution random variables  $N(0,1)$ . Find the PDF of  $\frac{(X_1 - X_2)^2}{2}$
5. Let  $X_1$  and  $X_2$  be independent gamma distribution random variables with  $\text{gamma}(\alpha_1, 1)$  and  $\text{gamma}(\alpha_2, 1)$ . Find the marginal distributions of  $\frac{X_1}{X_1 + X_2}$  and  $\frac{X_2}{X_1 + X_2}$

For the following questions, find the Maximum Likelihood Estimation (MLE).

6. Let  $X_1, \dots, X_n$  be a random sample from a  $\text{gamma}(\alpha, \beta)$  population. Find the MLE of  $\beta$  assuming  $\alpha$  is known.
7. Let  $X_1, \dots, X_n$  be a random sample from the PDF. Find the MLE of  $\theta$ :

$$f(x|\theta) = \theta x^{-2}, \quad 0 < \theta \leq x < \infty$$

R Programming:

For the following questions, let matrix  $X = \begin{bmatrix} 10 & 1 & 9 \\ 3 & 8 & 7 \\ 5 & 2 & 4 \end{bmatrix}$   $Y = \begin{bmatrix} 2 & 8 & 3 \\ 5 & 1 & 12 \\ 13 & 4 & 7 \end{bmatrix}$

(Please note that  $A^T$  is a transpose of a matrix A, a matrix such that  $[A^T]_{ij} = A_{ji}$ )

8. For matrix  $X$  and  $Y$  above, calculate  $(X + Y)$
9. For matrix  $X$  and  $Y$  above, calculate  $(X^T X)^{-1} X^T Y$
10. Write R code to draw 10,000 random samples from uniform distribution and calculate the 99% percentile.

Note:  $\mu$  = mean  $\sigma^2$  = variance

①  $f_x(x) = ax^{a-1}$ ,  $0 < x < 1$ ,  $a > 0$

$$\mu = E(f_x(x)) = \int x f_x(x) dx$$

$$= \int_0^1 x \cdot a x^{a-1} dx = a \int_0^1 x^{a+1-1} dx = a \left. \frac{x^{a+1}}{a+1} \right|_{x=0}^1$$

$$\boxed{\mu = \frac{a}{a+1}} - 0$$

$$\sigma^2 = E[(X-\mu)^2] = E(x^2) - [E(x)]^2$$

$$\int_0^1 x^2 \cdot a x^{a-1} dx$$

$$= a \int_0^1 x^{a+1} dx = a \left. \frac{x^{a+2}}{a+2} \right|_{x=0}^1$$

$$= a \frac{1^{a+2}}{a+2} - a \frac{0^{a+2}}{a+2}$$

$$= \frac{a}{a+2}$$

$$E(x) = \frac{a}{a+1}$$

$$[E(x)]^2 = \left(\frac{a}{a+1}\right)^2$$

$$= \frac{a^2}{a^2+2a+1}$$

$$\boxed{\therefore \sigma^2 = E(x^2) - [E(x)]^2 = \frac{a}{a+2} - \frac{a^2}{a^2+2a+1}}$$

2.  $f_x(x) = \frac{1}{n}$ ,  $x=1, \dots, n$ ,  $n > 0$  integer

$$\mu = E(f_x(x)) = \int x \cdot f_x(x) dx$$

$$\mu = \int_1^n x \cdot \frac{1}{n} dx$$

$$= \frac{1}{n} \int_1^n x dx = \frac{1}{n} \left( \frac{x^2}{2} \right)_{x=1}^n$$

$$= \frac{1}{n} \left( \frac{n^2}{2} - \frac{1^2}{2} \right) = \frac{n}{2} - \frac{1}{2n}$$

$$\boxed{\mu = \frac{n}{2} - \frac{1}{2n}}$$

$$\sigma^2 = E(x^2) - [E(x)]^2$$

$$E(x) = \mu = \frac{n}{2} - \frac{1}{2n}$$

$$[E(x)]^2 = \left( \frac{n}{2} - \frac{1}{2n} \right)^2$$

$$E(x^2) = \int_1^n x^2 \cdot \frac{1}{n} dx$$

$$= \frac{1}{n} \int_1^n x^2 dx = \frac{1}{n} \left( \frac{x^3}{3} \right)_{x=1}^n$$

$$= \frac{1}{n} \left( \frac{n^3}{3} - \frac{1}{3} \right)$$

$$= \frac{n^2}{3} - \frac{1}{3n}$$

$$\boxed{\therefore \sigma^2 = \frac{n^2}{3} - \frac{1}{3n} - \left( \frac{n}{2} - \frac{1}{2n} \right)^2}$$

$$(3) f_x(x) = \frac{3}{2} (x-1)^2, 0 < x < 2$$

$$\mu = E[f_x(x)] = \int_0^2 x \cdot f_x(x) dx$$

$$\mu = \int_0^2 x \cdot \frac{3}{2} (x-1)^2 dx = \frac{3}{2} \int_0^2 x (x^2 - 2x + 1) dx$$

$$= \frac{3}{2} \int_0^2 x^3 - 2x^2 + x dx$$

$$= \frac{3}{2} \left( \frac{x^4}{4} - \frac{2x^3}{3} + \frac{x^2}{2} \right)_{x=0}^2$$

$$= \frac{3}{2} \left( \frac{2^4}{4} - \frac{2(2^3)}{3} + \frac{2^2}{2} \right)$$

$$\boxed{\mu = 1}$$

$$\sigma^2 = E[(x-\mu)^2] = E(x^2) - [E(x)]^2$$

$\leftarrow [E(x)]^2 = \mu^2 = 1$

$$E(x^2) = \int_0^2 x^2 \cdot \frac{3}{2} (x-1)^2 dx$$

$$= \frac{3}{2} \int_0^2 x^2 (x^2 - 2x + 1) dx$$

$$= \frac{3}{2} \int_0^2 x^4 - 2x^3 + x^2 dx$$

$$= \frac{3}{2} \left( \frac{x^5}{5} - \frac{2x^4}{4} + \frac{x^3}{3} \right)_{x=0}^2$$

$$= \frac{3}{2} \left( \frac{2^5}{5} - \frac{2(2^4)}{4} + \frac{2^3}{3} \right)$$

$$= 6.83$$

$$\sigma^2 = 6.83 - 1$$

$$\boxed{\sigma^2 = 5.83}$$

4.  $X_1, X_2 \sim N(0, 1)$ . Find PDF of  $\frac{(X_1 - X_2)^2}{2}$

$$f(x_1) = f(x_2) = \frac{e^{-x^2/2}}{\sqrt{2\pi}}$$

$$f(x_1, x_2) = \frac{(x_1 - x_2)^2}{2} = \frac{\left( \frac{e^{-x_1^2/2} - e^{-x_2^2/2}}{\sqrt{2\pi}} \right)^2}{2}$$

$$= \frac{e^{-x_1^2} - e^{-x_2^2}}{2\pi^2 4}$$

$$\text{PDF} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{e^{-x_1^2} - e^{-x_2^2}}{4\pi^2} dx_1 dx_2$$

$$= \frac{1}{4\pi^2} \int_{-\infty}^{\infty} \sqrt{\pi} - x_1 e^{-x_2^2} dx_2$$

$$= \frac{1}{4\pi^2} (\sqrt{\pi} x_2 - x_1 \sqrt{\pi}) = \frac{x_2}{4\pi^{3/2}} - \frac{x_1}{4\pi^{3/2}}$$

oops...  $\mu = \frac{x_2 - x_1}{4\pi^{3/2}}$

$$\text{PDF} = \frac{x_2 - x_1}{4\pi^{3/2}}$$

⑤. Let  $X_1 = \Gamma(\alpha_1, 1)$ ,  $X_2 = \Gamma(\alpha_2, 2)$

Find marginal distributions of  $\frac{X_1}{X_1+X_2}$  and  $\frac{X_2}{X_1+X_2}$

$$X_1 = \frac{1}{\Gamma(\alpha)} X_1^{\alpha-1} e^{-X_1}$$

$$X_2 = \frac{1}{\Gamma(\alpha)} X_2^{\alpha-1} e^{-X_2}$$

$$\frac{X_1}{X_1+X_2} = \frac{\frac{1}{\Gamma(\alpha)} X_1^{\alpha-1} e^{-X_1}}{\frac{1}{\Gamma(\alpha)} X_1^{\alpha-1} e^{-X_1} + \frac{1}{\Gamma(\alpha)} X_2^{\alpha-1} e^{-X_2}}$$

m.d.f.  $\frac{X_1}{X_1+X_2} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \downarrow dx_1 dx_2$

vice versa  
for  
 $\frac{X_2}{X_1+X_2}$  ?

Not sure how to continue, and how I began...

⑥  $X_1, \dots, X_n$  random sample from  $\Gamma(a, \beta)$  population.

Find  $\beta$  assuming  $\alpha$  is known

$$f(x) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}$$

$$L(\beta|x) = (x_1, \dots, x_n)^{n-1} \left( \frac{1}{\Gamma(\alpha)\beta^\alpha} \right)^n e^{-\frac{\sum_{i=1}^n x_i}{\beta}}$$

$$\log L(\beta|x) = n(-\alpha) \log \Gamma(\alpha) - n(\alpha) \log(\beta) + (\alpha-1) \sum_{i=1}^n \log(x_i) - \frac{\sum_{i=1}^n x_i}{\beta}$$

$$\frac{\partial \log L(\beta|x)}{\partial \beta} = \frac{-n\alpha}{\beta} + \frac{\sum_{i=1}^n x_i}{\beta^2} = 0$$

$$\hat{\beta} = \frac{\sum_{i=1}^n x_i / n}{\alpha} = \boxed{\frac{\bar{x}}{\alpha}}$$

Note: MLE of  $\beta = \hat{\beta}$



7.  $f(x|\theta) = \theta x^{-2}$ ,  $0 < \theta \leq x < \infty$ ,  $x_1, \dots, x_n$   
Find MLE  $\theta$  ( $\hat{\theta}$ )

$$f(x|\theta) = \theta^n (x_1, \dots, x_n)^{-2}$$

$$L(x|\theta) = \theta^n (x_1, \dots, x_n)^{-2}$$

Given  $0 < \theta \leq x < \infty$  constraint,  $L(x|\theta)$  is greatest when  $n$  is at smallest value (as  $x^n$  sits in denominator of  $f(x|\theta)$ ).

$$\therefore \hat{\theta} = x_{n_{\min}} = x_1$$

# YK\_Assignment0

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## R Markdown

```
X <- matrix(c(10,1,9,3,8,7,5,2,4), nrow=3, ncol=3, byrow = TRUE)

Y <- matrix(c(2,8,3,5,1,12,13,4,7), nrow=3, ncol=3, byrow = TRUE)

print(X)
```

```
##      [,1] [,2] [,3]
## [1,]   10    1    9
## [2,]    3    8    7
## [3,]    5    2    4
```

```
print(Y)
```

```
##      [,1] [,2] [,3]
## [1,]    2    8    3
## [2,]    5    1   12
## [3,]   13    4    7
```

## Problem 8

```
print(X+Y)
```

```
##      [,1] [,2] [,3]
## [1,]   12    9   12
## [2,]    8    9   19
## [3,]   18    6   11
```

## Problem 9

```
print(solve(t(X)*X)*t(X)*Y)
```

```
##              [,1]              [,2]              [,3]
## [1,] -0.38365304 -0.32360300  0.9862385
## [2,] -0.06741729  0.07876934  0.7033639
## [3,]  7.69266055  0.82059123 -4.1457696
```

## Problem 10

```
sample <- runif(10000)
quantile(sample, .99)
```

```
##          99%
## 0.9905336
```

The 99th percentile in this uniform distribution of random variables is 0.9905336