

FINM 34000, Autumn 2023

Lecture 7

Reading: Notes, rest of Section 5

Exercise 1 Let X_1, X_2, \dots be independent, identically distributed random variables with

$$\mathbb{P}\{X_j = 2\} = \frac{1}{3}, \quad \mathbb{P}\{X_j = \frac{1}{2}\} = \frac{2}{3}.$$

Let $M_0 = 1$ and for $n \geq 1$, $M_n = X_1 X_2 \cdots X_n$.

1. Show that M_n is a martingale.

Note that $\mathbb{E}[X_j] = 2 \cdot \frac{1}{3} + \frac{1}{2} \cdot \frac{2}{3} = 1$.

$$E[M_{n+1} \mid \mathcal{F}_n] = E[M_n X_{n+1} \mid \mathcal{F}_n] = M_n E[X_{n+1} \mid \mathcal{F}_n] = M_n \mathbb{E}[X_{n+1}] = M_n.$$

2. Explain why M_n satisfies the conditions of the martingale convergence theorem.

Since $M_n \geq 0$,

$$\mathbb{E}[|M_n|] = \mathbb{E}[M_n] = 1.$$

3. Let $M_\infty = \lim_{n \rightarrow \infty} M_n$. Explain why $M_\infty = 0$. (Hint: there are at least two ways to show this. One is to consider $\log M_n$ and use the law of large numbers. Another is to note that with probability one M_{n+1}/M_n does not converge.)

We will discuss both ways.

For the first note that

$$\mathbb{E}[\log X_j] = \frac{1}{3} \log 2 + \frac{2}{3} \log(1/2) < 0.$$

If we write $Y_j = \log X_j$, **then the strong law of large numbers tells us that**

$$\frac{Y_1 + \cdots + Y_n}{n} \rightarrow \mathbb{E}[Y_j] < 0,$$

which shows that $Y_1 + \cdots + Y_n \rightarrow -\infty$. **Hence**

$$\lim_{n \rightarrow \infty} M_n = \lim_{n \rightarrow \infty} \exp\{Y_1 + \cdots + Y_n\} = 0.$$

For the other method, we know that M_n is converging. If it were converging to a strictly positive number, that would imply (just by the definition of a limit) that

$$\lim_{n \rightarrow \infty} \frac{M_{n+1}}{M_n} = 1.$$

However, we know that M_{n+1}/M_n always takes the value 2 or 1/2 which is a contradiction.

4. Use the optional sampling theorem to determine the probability that M_n ever attains a value as large as 64.

For any positive integer r , let T_r be the first time that $M_n = 64$ or $M_n = 2^{-r}$. Then $M_{n \wedge T_r}$ is a bounded martingale and we can use the optional sampling theorem,

$$1 = \mathbb{E}[M_{n \wedge T_r}] = 64 \mathbb{P}\{T_r = 64\} - 2^{-r} \mathbb{P}\{T_r = 2^{-r}\}.$$

Using $\mathbb{P}\{T_r = 2^{-r}\} = 1 - \mathbb{P}\{T_r = 64\}$, we get

$$\mathbb{P}\{T_r = 64\} = \frac{1 + 2^{-r}}{64 + 2^{-r}}.$$

The probability that one ever reaches 64 can be given as

$$\lim_{r \rightarrow \infty} \mathbb{P}\{T_r = 64\} = \lim_{r \rightarrow \infty} \frac{1 + 2^{-r}}{64 + 2^{-r}} = \frac{1}{64}.$$

5. Does there exist a $C < \infty$ such that $\mathbb{E}[M_n^2] \leq C$ for all n ?

If this were true we would have $\mathbb{E}[M_0] = \mathbb{E}[M_\infty]$ but we do not have this, so it better not be true!.

We can check it directly. By independence

$$\mathbb{E}[M_n^2] = \mathbb{E}[X_1^2] \mathbb{E}[X_2^2] \cdots \mathbb{E}[X_n^2].$$

But

$$\mathbb{E}[X_j^2] = 2^2 \cdot \frac{1}{3} + 2^{-2} \cdot 23 = \frac{3}{2} > 1$$

and hence $\mathbb{E}[M_n^2] = (3/2)^n$ and no such C exists.

Exercise 2 Consider the martingale betting strategy as discussed in Section 5. Let W_n be the “winnings” at time n , which for positive n equals either 1 or $1 - 2^n$.

1. Is W_n a square integrable martingale?

In lectures we showed it was a martingale. Since W_n is a bounded random variable (note that $|W_n| \leq 2^n$), we have $\mathbb{E}[W_n^2] < \infty$ and hence it is a square integrable martingale.

2. If $\Delta_n = W_n - W_{n-1}$ what is $\mathbb{E}[\Delta_n^2]$?

There are three possible values for $W_n - W_{n-1}$. It equals 0 if there has been a heads on the first $n - 1$ flips. Otherwise it is $\pm 2^{n-1}$. Therefore

$$\mathbb{E}[\Delta_n^2] = 2^{-(n-1)} [2^{(n-1)}]^2 = 2^{n-1}.$$

3. What is $\mathbb{E}[W_n^2]$?

There are two ways to do this. The easiest is to note that

$$\mathbb{P}\{W_n = 1\} = 1 - 2^{-n}, \quad \mathbb{P}\{W_n = 1 - 2^n\} = 2^{-n}.$$

$$\mathbb{E}[W_n^2] = 1^2 (1 - 2^{-n}) + (1 - 2^n)^2 2^{-n} = 2^n - 1$$

The other way is to use the “orthogonality of martingale increments” to see that

$$\mathbb{E}[W_n^2] = \sum_{j=1}^n \mathbb{E}[\Delta_j^2] = \sum_{j=1}^n 2^{j-1} = 2^n - 1.$$

4. What is $E(\Delta_n^2 \mid \mathcal{F}_{n-1})$?

This is just checking the definition. Using part 2 we see that

$$E(\Delta_n^2 \mid \mathcal{F}_{n-1}) = \begin{cases} 0 & \text{if } W_{n-1} = 1 \\ 2^{2(n-1)} & \text{if } W_{n-1} = 1 - 2^n \end{cases}$$

Since W_{n-1} is measurable with respect to \mathcal{F}_{n-1} , the right-hand side is also measurable with respect to \mathcal{F}_{n-1} .

Exercise 3 Here are some statements about martingales. Say whether they are always true. If always true give reason (citing a fact from the lecture or notes is fine). If it is not always true give an example to show this. Let $M_n, n = 0, 1, 2, \dots$ be a martingale with respect to $\{\mathcal{F}_n\}$ with $M_0 = 1$.

1. For all positive integers n , $\mathbb{E}[M_n] = 1$.

Yes, this is a fact from class.

2. With probability one, the limit

$$M_\infty := \lim_{n \rightarrow \infty} M_n \tag{1}$$

exists and is finite.

No, for example, let M_n be simple random walk starting at 1.

3. Suppose the limit M_∞ exists as in (1) and is finite. Then $\mathbb{E}[M_\infty] = 1$.

No. There are many examples. One is to let $M_n = 1 - W_n$ where W_n is the martingale betting strategy.

4. Suppose we assume know that with probability one $M_n \geq 0$ for all n ? Does this imply that the limit in (1) exists with probability one?

Yes, we discussed this in class. In this case, $\mathbb{E}[|M_n|] = \mathbb{E}[M_n] = \mathbb{E}[M_0] = 1$.

5. If we assume that $M_n \geq 0$ for all n does the answer to part 3 change?

No, since the example we gave has $M_n \geq 0$ for all n .