

Single Parameter Model II:

Review:

$$\theta \in (0, 1)$$

prior : $\theta \sim \text{uniform}(0, 1) = \text{Beta}(1, 1)$

model: $y_i \stackrel{\text{iid}}{\sim} \text{Bernoulli}(\theta) \Rightarrow \text{likelihood } p(y_1, y_2, \dots, y_n | \theta)$

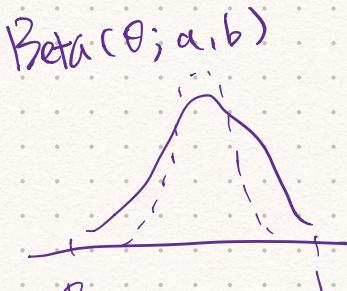
$$= \theta \sum_{i=1}^n y_i (1-\theta)^{n-i} \sum_{j=1}^i y_j$$

$$\text{posterior: Beta}\left(1 + \frac{\sum_{i=1}^n y_i}{\text{head}}, 1 + n - \frac{\sum_{i=1}^n y_i}{\text{tails}}\right)$$

Generalization: Conjugate priors

Extension: Binomial model with beta prior

$$\left\{ \begin{array}{l} \text{prior: } \theta \sim \text{Beta}(a, b) \iff P(\theta) \propto \theta^{a-1} (1-\theta)^{b-1} \\ \text{likelihood: } P(y_1, \dots, y_n | \theta) = \theta^{\sum_{i=1}^n y_i} (1-\theta)^{n - \sum_{i=1}^n y_i} \\ \text{posterior: } P(\theta | y_1, y_2, \dots, y_n) = \frac{P(\theta) P(y_1, y_2, \dots, y_n | \theta)}{P(y_1, \dots, y_n)} \end{array} \right.$$



a,b越大，动量越强。

pseudo-count
for head

$$= \theta^{a-1} (1-\theta)^{b-1} \theta^{\sum_{i=1}^n y_i} (1-\theta)^{n-\sum_{i=1}^n y_i}$$

$$\propto \text{Beta}(\theta; a + \sum_{i=1}^n y_i, b + n - \sum_{i=1}^n y_i)$$

head
(in $\text{tail}^{\text{left}}$)

tail

Pseudo-count
for tail.

When can we use " \propto "?

(But make sure
the constant is)

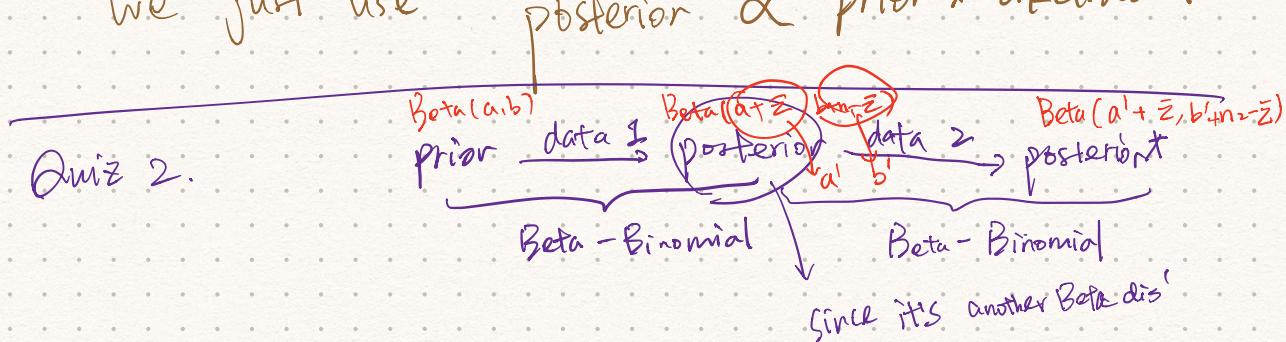
When we want to eliminate a "constant" positive! .

e.g. $p(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$, if $p(x) \propto \exp(-x^2/2)$
 $\rightarrow x \sim N(0, 1)$

if the term contains $\theta \Rightarrow$ cannot reduce it!

Generally ignore $P(y_1, \dots, y_n)$

we just use posterior \propto prior * likelihood.



Gaussian Conjugate Model:

$\xrightarrow{\text{Binomial (Bernoulli)}}$

Conjugacy: if f_i is a class of sampling distributions $p(y|\theta)$

and P is a class of prior distribution for θ .

then class P is conjugate for f_i if $\underline{P(\theta|y) \in P}$

posterior

for all $p(y|\theta) \in \mathcal{F}$ & $p(\theta) \in \mathcal{P}$
 prior

prior	likelihood	posterior	
Beta (uniform)	Binomial	Beta	✓
Gaussian	Gaussian	Gaussian	(in class)
Gamma	Poisson	Gamma	(quiz)
:	:	:	

for exponential families \rightarrow natural conjugate prior

① General form $P(y|\theta)$ for exponential family
 & its natural conjugate prior:

② Use the Beta-Binomial & match \downarrow to the Beta distribution.

Conjugate Gaussian - Gaussian:

Example 2. normal with variance known

data $\{y_1, \dots, y_n\} \in \mathbb{R}$

parameter of interest: $\theta \in \mathbb{R}$

infer θ given y_1, y_2, \dots, y_n (consider mean)
 (unknown)

Likelihood: $y_i \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$ known constant (so far)
 ↓
 parameter of interest

△ prior distribution: $\theta \sim N(\mu_0, \tau_0^2)$ known

posterior distribution (p.d.f.):

$$P(\theta | y_1, \dots, y_n) \propto P(\theta) P(y_1, y_2, \dots, y_n | \theta)$$

$$\propto \exp\left(-\frac{(\theta - \mu_0)^2}{2\tau_0^2}\right) \prod_{i=1}^n \exp\left[-\frac{(y_i - \theta)^2}{2\sigma^2}\right]$$

$$\propto \exp\left[-\frac{(\theta - \mu_0)^2}{2\tau_0^2} - \frac{\sum_{i=1}^n (y_i - \theta)^2}{2\sigma^2}\right]$$

$$= \exp \left[-\frac{1}{2} \left(\frac{1}{T_0^2} + \frac{n}{\sigma^2} \right) \theta^2 + \left(\frac{\mu_0}{T_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2} \right) \theta \right.$$

$$\left. - \left(\frac{\mu_0^2}{T_0^2} + \frac{\sum_{i=1}^n y_i^2}{2\sigma^2} \right) \right] \quad \begin{aligned} & (\alpha\theta^2 + b\theta \\ & = \alpha(\theta + \frac{b}{2\alpha})^2 - \frac{b^2}{4\alpha}) \end{aligned}$$

throw! since not related with θ

$$\propto \exp \left[-\frac{1}{2} \left(\frac{1}{T_0^2} + \frac{n}{\sigma^2} \right) \left(\theta - \frac{\mu_0/T_0^2 + \sum y_i/\sigma^2}{1/T_0^2 + n/\sigma^2} \right)^2 \right]$$

$$\mathcal{N}(\mu_n, T_n^{-2})$$

$$(1-W_n)\mu_0 + W_n \bar{y}_n$$

$$\Rightarrow \mu_n = \frac{\mu_0/T_0^2 + \sum y_i/\sigma^2}{1/T_0^2 + n/\sigma^2} \quad \text{posterior mean}$$

$$T_n^{-2} = \frac{1}{1/T_0^2 + n/\sigma^2} \quad \text{posterior variance}$$

$$(T_n^{-2})^{-1} : \text{Posterior precision} = \frac{1}{\text{variance}}$$

$$= \frac{1}{T_0^2} + \frac{n}{\sigma^2} \quad \text{likelihood precision (since need to times n)}$$

prior precision

$$\text{for } \mu_n, W_n = \frac{n/\sigma^2}{1/T_0^2 + n/\sigma^2} \quad \rightarrow 1 - W_n \propto \frac{1}{T_0^2}$$

(KL prior)

$\propto \frac{n}{\sigma^2}$ (KL likelihood)