

#### Lecture 9: Linear Regression

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Mathematics for Machine Learning https://yung-web.github.io/home/courses/mathml.html KAIST EE

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Problem Formulation

• Parameter Estimation: ML

Parameter Estimation: MAP

• Bayesian Linear Regression

• Maximum Likelihood as Orthogonal Projection

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#### Roadmap



Regression Problem



#### Problem Formulation

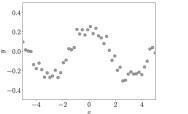
• Parameter Estimation: ML

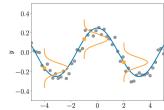
• Parameter Estimation: MAP

• Bayesian Linear Regression

• Maximum Likelihood as Orthogonal Projection







- For some input values  $x_n$ , we observe noisy function values  $y_n = f(x_n) + \epsilon$
- Goal: infer the function f that generalizes well to function values at new inputs
- Applications: time-series analysis, control and robotics, image recognition, etc.

- Linear regression, Gaussian noise
- Notation for simplification (this is how the textbook uses)

$$p(y|\mathbf{x}) = p_{Y|\mathbf{X}}(y|\mathbf{x}), \quad Y \sim \mathcal{N}(\mu, \sigma^2) \xrightarrow{\text{simplifies}} \mathcal{N}(y \mid f(\mathbf{x}), \sigma^2)$$

- Likelihood: for  $\mathbf{x} \in \mathbb{R}^D$  and  $\mathbf{y} \in \mathbb{R}$ ,  $p(\mathbf{y} \mid \mathbf{x}) = \mathcal{N}(\mathbf{y} \mid f(\mathbf{x}), \sigma^2)$
- $y = f(\mathbf{x}) + \epsilon$ , where  $\epsilon \sim \mathcal{N}(0, \sigma^2)$
- Linear regression with the parameter  $oldsymbol{ heta} \in \mathbb{R}^D$

$$p(y \mid \mathbf{x}) = \mathcal{N}(y \mid \mathbf{x}^{\mathsf{T}} \boldsymbol{\theta}, \sigma^2) \Longleftrightarrow y = \mathbf{x}^{\mathsf{T}} \boldsymbol{\theta} + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \sigma^2)$$

Prior with Gaussian nose:  $p(y \mid \mathbf{x}) = \mathcal{N}(y \mid \mathbf{x}^T \boldsymbol{\theta}, \sigma^2)$ 

• Training set  $\mathcal{D} = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_N, y_N)\}$ 



• Assuming iid of *N* data, the likelihood is factorized into:

$$p(\mathcal{Y} \mid \mathcal{X}, \boldsymbol{\theta}) = \prod_{n=1}^{N} p(y_n \mid \boldsymbol{x}_n, \boldsymbol{\theta}) = \prod_{n=1}^{N} \mathcal{N}(y_n \mid \boldsymbol{x}_n^{\mathsf{T}}, \sigma^2),$$

where  $\mathcal{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$  and  $\mathcal{Y} = \{y_1, \dots, y_n\}$ 

ML and MAP

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### Roadmap



MLE (Maximum Likelihood Estimation) (1)



- Problem Formulation
- Parameter Estimation: ML
- Parameter Estimation: MAP
- Bayesian Linear Regression
- Maximum Likelihood as Orthogonal Projection

- $\theta_{\mathsf{ML}} = \operatorname{arg\,max}_{\theta} p(\mathcal{Y} \mid \mathcal{X}, \theta) = \operatorname{arg\,min}_{\theta} \left( -\log p(\mathcal{Y} \mid \mathcal{X}, \theta) \right)$
- For Gaussian noise with  $\pmb{X} = [\pmb{x}_1, \dots, \pmb{x}_n]^\mathsf{T}$  and  $\pmb{y} = [y_1, \dots, y_n]^\mathsf{T},$

$$-\log p(\mathcal{Y} \mid \mathcal{X}, \boldsymbol{\theta}) = -\log \prod_{n=1}^{N} p(y_n \mid \boldsymbol{x}_n, \boldsymbol{\theta}) = -\sum_{n=1}^{N} \log p(y_n \mid \boldsymbol{x}_n, \boldsymbol{\theta})$$
$$= \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^{\mathsf{T}} \boldsymbol{\theta})^2 + \text{const} = \frac{1}{2\sigma^2} \|\boldsymbol{y} - \boldsymbol{X} \boldsymbol{\theta}\|^2 + \text{const}$$

Negative-log likelihood for 
$$f(\mathbf{x}) = \mathbf{x}^\mathsf{T} \mathbf{\theta} + \mathcal{N}(0, \sigma^2)$$
:

$$-\log p(\mathcal{Y} \mid \mathcal{X}, \boldsymbol{\theta}) = \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|^2 + \text{ const}$$



• For Gaussian noise with  $\mathbf{X} = [\mathbf{x}_1, \dots, \mathbf{x}_n]^\mathsf{T}$  and  $\mathbf{y} = [y_1, \dots, y_n]^\mathsf{T}$ ,

$$m{ heta}_{\mathsf{ML}} = \arg\min_{m{ heta}} rac{1}{2\sigma^2} \left\| m{y} - m{X} m{ heta} 
ight\|^2, \quad L(m{ heta}) = rac{1}{2\sigma^2} \left\| m{y} - m{X} m{ heta} 
ight\|^2$$

- In this special case of Gaussian noise, finding MLE is equivalent to finding  $\theta$  that minimizes the empirical risk with squared loss function
  - Models as functions = Model as probabilistic models
- We find  $\theta$  such that  $\frac{dL}{d\theta} = 0$

$$\begin{split} &\frac{\mathrm{d} L}{\mathrm{d} \theta} = \frac{1}{2\sigma^2} \left( -2 (\boldsymbol{y} - \boldsymbol{X} \boldsymbol{\theta})^\mathsf{T} \boldsymbol{X} \right) = \frac{1}{\sigma^2} \left( -\boldsymbol{y}^\mathsf{T} \boldsymbol{X} + \boldsymbol{\theta}^\mathsf{T} \boldsymbol{X}^\mathsf{T} \boldsymbol{X} \right) = 0 \\ &\iff \boldsymbol{\theta}_\mathsf{ML}^\mathsf{T} \boldsymbol{X}^\mathsf{T} \boldsymbol{X} = \boldsymbol{y}^\mathsf{T} \boldsymbol{X} \\ &\iff \boldsymbol{\theta}_\mathsf{ML}^\mathsf{T} = \boldsymbol{y}^\mathsf{T} \boldsymbol{X} (\boldsymbol{X}^\mathsf{T} \boldsymbol{X})^{-1} \quad (\boldsymbol{X}^\mathsf{T} \boldsymbol{X} \text{ is positive definite if } \mathsf{rk}(\boldsymbol{X}) = D) \\ &\iff \boldsymbol{\theta}_\mathsf{ML} = \left( \boldsymbol{X}^\mathsf{T} \boldsymbol{X} \right)^{-1} \boldsymbol{X}^\mathsf{T} \boldsymbol{y} \end{split}$$

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- Linear regression: Linear in the parameters
  - $\phi(\mathbf{x})^{\mathsf{T}} \theta$  is also fine, where  $\phi(\mathbf{x})$  can be non-linear (we will cover this later)
  - $\phi(\mathbf{x})$  are the features
- Linear regression with the parameter  $\theta \in \mathbb{R}^K$ ,  $\phi(\mathbf{x}) : \mathbb{R}^D \mapsto \mathbb{R}^K$ :

$$p(y \mid \mathbf{x}) = \mathcal{N}(y \mid \phi(\mathbf{x})^{\mathsf{T}} \boldsymbol{\theta}, \sigma^2) \Longleftrightarrow y = \phi(\mathbf{x})^{\mathsf{T}} \boldsymbol{\theta} + \epsilon = \sum_{k=0}^{K-1} \theta_k \phi_k(\mathbf{x}) + \epsilon$$

• Example. Polynomial regression. For  $x \in \mathbb{R}$  and  $\theta \in \mathbb{R}^K$ , we lift the original 1-D input into K-D feature space with monomials  $x^k$ :

$$\phi(x) = \begin{pmatrix} \phi_0(x) \\ \vdots \\ \phi_{K-1}(x) \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ x^{K-1} \end{pmatrix} \in \mathbb{R}^K \implies f(x) = \sum_{k=0}^{K-1} \theta_k x^k$$

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#### Feature Matrix and MLE



Polynomial Fit



• Now, for the entire training set  $\{x_1, \dots, x_N\}$ 

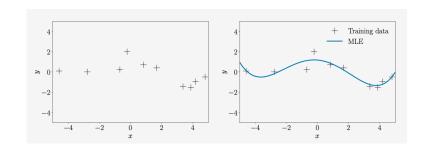
$$\Phi := \begin{pmatrix} \phi^{\mathsf{T}}(\mathbf{x}_1) \\ \vdots \\ \phi^{\mathsf{T}}(\mathbf{x}_N) \end{pmatrix} = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \cdots & \phi_{K-1}(\mathbf{x}_1) \\ \vdots & \cdots & \vdots \\ \phi_0(\mathbf{x}_N) & \cdots & \phi_{K-1}(\mathbf{x}_N) \end{pmatrix} \in \mathbb{R}^{N \times K}, \quad \Phi_{ij} = \phi_j(\mathbf{x}_i), \ \phi_j : \mathbb{R}^D \mapsto \mathbb{R}$$

- Negative log-likelihood: Similarly to the case of  $y = X\theta$ ,
  - $p(\mathcal{Y}|\mathcal{X}, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{y} \mid \boldsymbol{\Phi}\boldsymbol{\theta}, \sigma^2 \boldsymbol{I})$
  - Negative-log likelihood for  $f(\mathbf{x}) = \phi^{\mathsf{T}}(\mathbf{x})\theta + \mathcal{N}(0, \sigma^2)$ :

$$-\log p(\mathcal{Y} \mid \mathcal{X}, \boldsymbol{\theta}) = \frac{1}{2\sigma^2} \| \boldsymbol{y} - \boldsymbol{\Phi} \boldsymbol{\theta} \|^2 + \mathsf{const}$$

• MLE:  $\theta_{\text{MI}} = (\Phi^{\mathsf{T}}\Phi)^{-1}\Phi^{\mathsf{T}}\mathbf{y}$ 

- N = 10 data, where  $x_n \sim \mathcal{U}[-5, 5]$  and  $y_n = -\sin(x_n/5) + \cos(x_n) + \epsilon$ .  $\epsilon \sim \mathcal{N}(0, 0.2^2)$
- Fit with polovnomial with degree 4 using ML



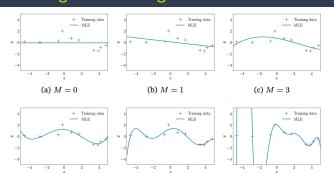
## Overfitting in Linear Regression

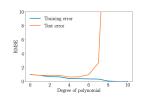
(d) M = 4

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Roadmap







Higher polynomial degree is better (training error always decreases)

(f) M = 9

• Test error increases after some polynomial degree

(e) M = 6

Problem Formulation

• Parameter Estimation: ML

Parameter Estimation: MAP

• Bayesian Linear Regression

• Maximum Likelihood as Orthogonal Projection

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## MAPE (Maximum A Posteriori Estimation)



MAPE for Gaussian Prior (1)



- MLE: prone to overfitting, where the magnitude of the parameters becomes large.
- a prior distribution  $p(\theta)$  helps: what  $\theta$  is plausible
- MAPE and Bayes' theorem

$$p(\theta \mid \mathcal{X}, \mathcal{Y}) = \frac{p(\mathcal{Y} \mid \mathcal{X}, \theta)p(\theta)}{p(\mathcal{Y} \mid \mathcal{X})} \implies \theta_{\mathsf{MAP}} \in \arg\min_{\theta} \Big( -\log p(\mathcal{Y} \mid \mathcal{X}, \theta) - \log p(\theta) \Big)$$

Gradient

$$-\frac{\mathsf{d}\log p(\boldsymbol{\theta}|\mathcal{X},\mathcal{Y})}{\mathsf{d}\boldsymbol{\theta}} = -\frac{\mathsf{d}\log p(\mathcal{Y}|\mathcal{X},\boldsymbol{\theta})}{\mathsf{d}\boldsymbol{\theta}} - \frac{\mathsf{d}\log p(\boldsymbol{\theta})}{\mathsf{d}\boldsymbol{\theta}}$$

• Example. A (conjugate) Gaussian prior  $p(\theta) \sim \mathcal{N}(0, b^2 I)$ 

For Gaussian likelihood, Gaussian prior ⇒ Gaussian posterior

Negative log-posterior

Negative-log posterior for 
$$f(\mathbf{x}) = \phi^{\mathsf{T}}(\mathbf{x})\boldsymbol{\theta} + \mathcal{N}(0, \sigma^2)$$
 and  $p(\boldsymbol{\theta}) \sim \mathcal{N}(0, b^2 \boldsymbol{I})$ :
$$-\log p(\boldsymbol{\theta}|\mathcal{X}, \mathcal{Y}) = \frac{1}{2\sigma^2}(\mathbf{y} - \boldsymbol{\Phi}\boldsymbol{\theta})^{\mathsf{T}}(\mathbf{y} - \boldsymbol{\Phi}\boldsymbol{\theta}) + \frac{1}{2b^2}\boldsymbol{\theta}^{\mathsf{T}}\boldsymbol{\theta} + \text{const}$$

Gradient

$$-\frac{\mathsf{d} \log p(\boldsymbol{\theta}|\mathcal{X}, \mathcal{Y})}{\mathsf{d} \boldsymbol{\theta}} = \frac{1}{\sigma^2} (\boldsymbol{\theta}^\mathsf{T} \boldsymbol{\Phi}^\mathsf{T} \boldsymbol{\Phi} - \boldsymbol{y}^\mathsf{T} \boldsymbol{\Phi}) + \frac{1}{b^2} \boldsymbol{\theta}^\mathsf{T}$$



• MAP vs. ML

$$oldsymbol{ heta}_{\mathsf{MAP}} = \underbrace{\left(oldsymbol{\Phi}^\mathsf{T}oldsymbol{\Phi} + rac{\sigma^2}{b^2}oldsymbol{I}
ight)}^{-1}oldsymbol{\Phi}^\mathsf{T}oldsymbol{y}, \quad oldsymbol{ heta}_{\mathsf{ML}} = \left(oldsymbol{\Phi}^\mathsf{T}oldsymbol{\Phi}
ight)^{-1}oldsymbol{\Phi}^\mathsf{T}oldsymbol{y}$$

- The term  $\frac{\sigma^2}{b^2}$ 
  - Ensures that (\*) is symmetric, strictly positive definite
  - Role of regularizer

- Example. A (conjugate) Gaussian prior  $p(\theta) \sim \mathcal{N}(\mathbf{m}_0, \mathbf{S}_0)$
- Negative log-posterior

Negative-log posterior for 
$$f(\mathbf{x}) = \phi^{\mathsf{T}}(\mathbf{x})\theta + \mathcal{N}(0, \sigma^2)$$
 and  $p(\theta) \sim \mathcal{N}(\mathbf{m}_0, \mathbf{S}_0)$ :  

$$-\log p(\theta|\mathcal{X}, \mathcal{Y}) = \frac{1}{2\sigma^2}(\mathbf{y} - \Phi\theta)^{\mathsf{T}}(\mathbf{y} - \Phi\theta) + \frac{1}{2}(\theta - \mathbf{m}_0)^{\mathsf{T}}\mathbf{S}_0^{-1}(\theta - \mathbf{m}_0) + \text{const}$$

• We will use this later for computing the parameter posterior distribution in Bayesian linear regression.

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#### Regularization: MAPE vs. Explicit Regularizer



Roadmap



• Explicit regularizer in regularized least squares (RLS)

$$\|\mathbf{y} - \mathbf{\Phi} \boldsymbol{\theta}\|^2 + \lambda \|\boldsymbol{\theta}\|^2$$

- MAPE wth Gaussian prior  $p(\theta) \sim \mathcal{N}(0, b^2 I)$ 
  - $\circ$  Negative log-Gaussian prior

$$-\log p(\theta) = \frac{1}{2b^2}\theta^{\mathsf{T}}\theta + \mathsf{const}$$

- $\lambda = 1/2b^2$  is the regularization term
- Not surprising that we have

$$oldsymbol{ heta}_{\mathsf{RLS}} = \left(oldsymbol{\Phi}^\mathsf{T}oldsymbol{\Phi} + \lambda oldsymbol{I}
ight)^{-1}oldsymbol{\Phi}^\mathsf{T}oldsymbol{y}$$

- Problem Formulation
- Parameter Estimation: ML
- Parameter Estimation: MAP
- Bayesian Linear Regression
- Maximum Likelihood as Orthogonal Projection

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- Earlier, ML and MAP. Now, fully Bayesian
- Model

$$\begin{aligned} & \text{prior} \quad p(\boldsymbol{\theta}) \sim \mathcal{N}(\boldsymbol{m}_0, \boldsymbol{S}_0) \\ & \text{likelihood} \quad p(y|\boldsymbol{x}, \boldsymbol{\theta}) \sim \mathcal{N}\big(y \mid \boldsymbol{\phi}^\mathsf{T}(\boldsymbol{x})\boldsymbol{\theta}, \sigma^2\big) \\ & \text{joint} \quad p(y, \boldsymbol{\theta}|\boldsymbol{x}) = p(y \mid \boldsymbol{x}, \boldsymbol{\theta})p(\boldsymbol{\theta}) \end{aligned}$$



 Goal: For an input x<sub>\*</sub>, we want to compute the following posterior predictive distribution of y<sub>\*</sub>:

$$p(y_*|x_*,\mathcal{X},\mathcal{Y}) = \int \overbrace{p(y_*|\mathbf{x}_*,\boldsymbol{\theta})}^{\text{likelihood}} \overbrace{p(\boldsymbol{\theta}|\mathcal{X},\mathcal{Y})}^{(*)} d\boldsymbol{\theta}$$

• (\*): parameter posterior distribution that needs to be computed

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• Parameter posterior distribution

$$p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y}) = \mathcal{N}(\boldsymbol{\theta} \mid \boldsymbol{m}_N, \boldsymbol{S}_N), \quad \text{where}$$

$$\boldsymbol{S}_N = \left(\boldsymbol{S}_0^{-1} + \sigma^2 \boldsymbol{\Phi}^\mathsf{T} \boldsymbol{\Phi}\right)^{-1}, \quad \boldsymbol{m}_N = \boldsymbol{S}_N \left(\boldsymbol{S}_0^{-1} \boldsymbol{m}_0 + \sigma^{-2} \boldsymbol{\Phi}^\mathsf{T} \boldsymbol{y}\right)$$

(Proof of Sketch)

• From the negative-log posterior for general Gaussian prior,

$$-\log p(\boldsymbol{\theta}|\mathcal{X},\mathcal{Y}) = \frac{1}{2\sigma^2}(\boldsymbol{y} - \boldsymbol{\Phi}\boldsymbol{\theta})^\mathsf{T}(\boldsymbol{y} - \boldsymbol{\Phi}\boldsymbol{\theta}) + \frac{1}{2}(\boldsymbol{\theta} - \boldsymbol{m}_0)^\mathsf{T}\boldsymbol{S}_0^{-1}(\boldsymbol{\theta} - \boldsymbol{m}_0) + \mathsf{const}$$

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#### Parameter Posterior Distribution (2)



Posterior Predictions (1)



$$= \frac{1}{2} \left( \sigma^{-2} \mathbf{y}^{\mathsf{T}} \mathbf{y} - 2 \sigma^{-2} \mathbf{y}^{\mathsf{T}} \boldsymbol{\Phi} \boldsymbol{\theta} + \boldsymbol{\theta}^{\mathsf{T}} \sigma^{-2} \boldsymbol{\Phi}^{\mathsf{T}} \boldsymbol{\Phi} \boldsymbol{\theta} + \boldsymbol{\theta}^{\mathsf{T}} \mathbf{S}_{0}^{-1} \boldsymbol{\theta} - 2 \mathbf{m}_{0}^{\mathsf{T}} \mathbf{S}_{0}^{-1} \boldsymbol{\theta} + \mathbf{m}_{0}^{\mathsf{T}} \mathbf{S}_{0}^{-1} \mathbf{m}_{0} \right)$$

$$= \frac{1}{2} \left( \boldsymbol{\theta}^{\mathsf{T}} (\sigma^{-2} \boldsymbol{\Phi}^{\mathsf{T}} \boldsymbol{\Phi} + \mathbf{S}_{0}^{-1}) \boldsymbol{\theta} - 2 (\sigma^{-2} \boldsymbol{\Phi}^{\mathsf{T}} \mathbf{y} + \mathbf{S}_{0}^{-1} \mathbf{m}_{0})^{\mathsf{T}} \boldsymbol{\theta} \right) + \text{const}$$

- cyan color: quadratic term, orange color: linear term
- $p(\theta|\mathcal{X},\mathcal{Y}) \propto \exp(\text{ quadratic in }\theta) \implies \text{Gaussian distribution}$
- Assume that  $p(\theta|\mathcal{X},\mathcal{Y}) = \mathcal{N}(\theta|\mathbf{m}_N,\mathbf{S}_N)$ , and find  $\mathbf{m}_N$  and  $\mathbf{S}_N$ .

$$\begin{split} -\log \mathcal{N}(\boldsymbol{\theta}|\boldsymbol{m}_{N},\boldsymbol{S}_{N}) &= \frac{1}{2}(\boldsymbol{\theta}-\boldsymbol{m}_{N})^{\mathsf{T}}\boldsymbol{S}_{N}^{-1}(\boldsymbol{\theta}-\boldsymbol{m}_{N}) + \text{const} \\ &= \frac{1}{2}\Big(\boldsymbol{\theta}^{\mathsf{T}}\boldsymbol{S}_{N}^{-1}\boldsymbol{\theta} - 2\boldsymbol{m}_{N}^{\mathsf{T}}\boldsymbol{S}_{N}^{-1}\boldsymbol{\theta} + \boldsymbol{m}_{N}^{\mathsf{T}}\boldsymbol{S}_{N}^{-1}\boldsymbol{m}_{N}\Big) + \text{const} \end{split}$$

Thus,

$$\mathbf{S}_{N}^{-1} = \sigma^{-2} \Phi^{\mathsf{T}} \Phi + \mathbf{S}_{0}^{-1}$$
 and  $\mathbf{m}_{N}^{\mathsf{T}} \mathbf{S}_{N}^{-1} = (\sigma^{-2} \Phi^{\mathsf{T}} \mathbf{y} + \mathbf{S}_{0}^{-1} \mathbf{m}_{0}^{\mathsf{T}})$ 

Posterior predictive distribution

$$p(y_*|x_*, \mathcal{X}, \mathcal{Y}) = \int p(y_*|\mathbf{x}_*, \boldsymbol{\theta}) p(\boldsymbol{\theta}|\mathcal{X}, \mathcal{Y}) d\boldsymbol{\theta}$$

$$= \int \mathcal{N}(y_*|\phi^{\mathsf{T}}(\mathbf{x}_*)\boldsymbol{\theta}, \sigma^2) \mathcal{N}(\boldsymbol{\theta}|\mathbf{m}_N, \mathbf{S}_N) d\boldsymbol{\theta}$$

$$= \mathcal{N}(y_*|\phi^{\mathsf{T}}(\mathbf{x}_*)\mathbf{m}_N, \phi^{\mathsf{T}}(\mathbf{x}_*)\mathbf{S}_N \phi(\mathbf{x}_*) + \sigma^2)$$

• The mean  $\phi^{\mathsf{T}}(\mathbf{x}_*)\mathbf{m}_N$  coincides with the MAP estimate

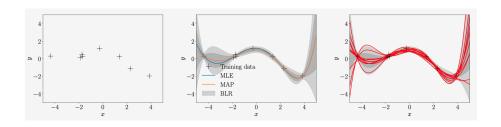
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#### Posterior Predictions (2)



## Computing Marginal Likelihood





• BLR: Bayesian Linear Regression

- Likelihood:  $p(\mathcal{Y}|\mathcal{X}, \theta)$ , Marginal likelihood:  $p(\mathcal{Y}|\mathcal{X}) = \int p(\mathcal{Y}|\mathcal{X}, \theta) p(\theta) d\theta$
- Recall that the marginal likelihood is important for model selection via Bayes factor:

$$(\text{Posterior odds}) = \frac{\mathbb{P}(M_1 \mid \mathcal{D})}{\mathbb{P}(M_2 \mid \mathcal{D})} = \frac{\frac{\mathbb{P}(\mathcal{D}|M_1)\mathbb{P}(M_1)}{\mathbb{P}(\mathcal{D})}}{\frac{\mathbb{P}(\mathcal{D}|M_2)\mathbb{P}(M_2)}{\mathbb{P}(\mathcal{D})}} = \underbrace{\frac{\mathbb{P}(M_1)}{\mathbb{P}(M_2)}}_{\text{Prior odds}} \underbrace{\frac{\mathbb{P}(\mathcal{D} \mid M_1)}{\mathbb{P}(\mathcal{D} \mid M_2)}}_{\text{Prior odds}} \underbrace{\frac{\mathbb{P}(\mathcal{D} \mid M_1)}{\mathbb{P}(\mathcal{D} \mid M_2)}}_{\text{Prior odds}}$$

$$\begin{split} \rho(\mathcal{Y}|\mathcal{X}) &= \int p(\mathcal{Y}|\mathcal{X}, \boldsymbol{\theta}) p(\boldsymbol{\theta}) \mathrm{d}\boldsymbol{\theta} = \int \mathcal{N}(\boldsymbol{y}|\boldsymbol{\Phi}\boldsymbol{\theta}, \sigma^2 \boldsymbol{I}) \mathcal{N}(\boldsymbol{\theta}|\boldsymbol{m}_0, \boldsymbol{S}_0) \, \mathrm{d}\boldsymbol{\theta} \\ &= \mathcal{N}(\boldsymbol{y} \mid \boldsymbol{\Phi}\boldsymbol{m}_0, \boldsymbol{\Phi}\boldsymbol{S}_0 \boldsymbol{\Phi}^\mathsf{T} + \sigma^2 \boldsymbol{I}) \end{split}$$

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#### Roadmap



ML as Orthogonal Projection



Problem Formulation

• Parameter Estimation: ML

• Parameter Estimation: MAP

- Bayesian Linear Regression
- Maximum Likelihood as Orthogonal Projection

• For 
$$f(\mathbf{x}) = \mathbf{x}^\mathsf{T} \mathbf{\theta} + \mathcal{N}(0, \sigma^2), \ \mathbf{\theta}_\mathsf{ML} = (\mathbf{X}^\mathsf{T} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \mathbf{y} = \frac{\mathbf{X}^\mathsf{T} \mathbf{y}}{\mathbf{X}^\mathsf{T} \mathbf{X}} \in \mathbb{R}$$

$$oldsymbol{X}oldsymbol{ heta}_\mathsf{ML} = rac{oldsymbol{X}oldsymbol{X}^\mathsf{T}}{oldsymbol{X}^\mathsf{T}oldsymbol{X}}oldsymbol{y}$$

 $\circ$  Orthogonal projection of y onto the one-dimensional subspace spanned by X

• For 
$$f(\mathbf{x}) = \phi^{\mathsf{T}}(\mathbf{x})\boldsymbol{\theta} + \mathcal{N}(0, \sigma^2), \ \boldsymbol{\theta}_{\mathsf{ML}} = (\boldsymbol{\Phi}^{\mathsf{T}}\boldsymbol{\Phi})^{-1}\boldsymbol{\Phi}^{\mathsf{T}}\mathbf{y} = \frac{\boldsymbol{\Phi}^{\mathsf{T}}\mathbf{y}}{\boldsymbol{\Phi}^{\mathsf{T}}\boldsymbol{\Phi}} \in \mathbb{R}$$

$$\boldsymbol{\Phi}\boldsymbol{\theta}_{\mathsf{ML}} = \frac{\boldsymbol{\Phi}\boldsymbol{\Phi}^{\mathsf{T}}}{\boldsymbol{\Phi}^{\mathsf{T}}\boldsymbol{\Phi}}\mathbf{y}$$

 $\circ$  Orthogonal projection of  ${m y}$  onto the K-dimensional subspace spanned by columns of  ${m \Phi}$ 

## Summary and Other Issues (1)



## Summary and Other Issues (2)



- Linear regression for Gaussian likelihood and conjugate Gaussian priors. Nice analytical results and closed forms
- Other forms of likelihoods for other applications (e.g., classification)
- GLM (generalized linear model):  $y = \sigma \circ f$  ( $\sigma$ : activation function)
  - $\circ$  No longer linear in heta
  - Logistic regression:  $\sigma(f) = \frac{1}{1 + \exp(-f)} \in [0, 1]$  (interpreted as the probability of becoming 1)
  - Building blocks of (deep) feedforward neural nets
  - $\mathbf{y} = \sigma(\mathbf{A}\mathbf{x} + \mathbf{b})$ . **A**: weight matrix, **b**: bias vector
  - K-layer deep neural nets:  $\mathbf{x}_{k+1} = f_k(\mathbf{x}_k), f_k(\mathbf{x}_k) = \sigma_k(\mathbf{A}_k \mathbf{x}_k + \mathbf{b}_k)$

- Gaussian process
  - $\circ$  A distribution over parameters o a distribution over functions
  - Gaussian process: distribution over functions without detouring via parameters
  - Closely related to BLR and support vector regression, also interpreted as Bayesian neural network with a single hidden layer and the infinite number of units
- Gaussian likelihood, but non-Gaussian prior
  - When  $N \ll D$  (small training data)
  - Prior that enforces sparsity, e.g., Laplace prior
  - A linear regression with teh Laplace prior = linear regression with LASSO (L1 regularization)

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**Review Questions** 



Questions?

1)

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