

Lecture 11: Density Estimation with Gaussian Mixture Models

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Mathematics for Machine Learning
https://yung-web.github.io/home/courses/mathml.html
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Warm-Up



Please watch this tutorial video by Luis Serrano on Gaussian Mixture Model.

https://www.youtube.com/watch?v=q71Niz856KE

Roadmap



- (1) Gaussian Mixture Model
- (2) Parameter Learning: MLE
- (3) Latent-Variable Perspective for Probabilistic Modeling
- (4) EM Algorithm

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Roadmap



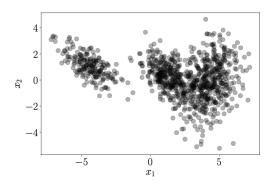
- (1) Gaussian Mixture Model
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Density Estimation



- Represent data compactly using a density from a parametric family, e.g., Gaussian or Beta distribution
- Parameters of those families can be found by MLE and MAPE
- However, there are many cases when simple distributions (e.g., just Gaussian) fail to approximate data.



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Mixture Models



- More expressive family of distribution
- Idea: Let's mix! A convex combination of K "base" distributions

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k p_k(\mathbf{x}), \quad 0 \le \pi_k \le 1, \quad \sum_{k=1}^{K} \pi_k = 1$$

- Multi-modal distributions: Can be used to describe datasets with multiple clusters
- Our focus: Gaussian mixture models
- Want to finding the parameters using MLE, but cannot have the closed form solution (even with the mixture of Gaussians) → some iterative methods needed

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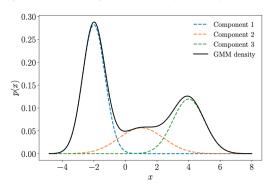
Gaussian Mixture Model



$$p(\mathbf{x}|\mathbf{\theta}) = \sum_{k=1}^K \mathcal{N}(\mathbf{x}|\mathbf{\mu}_k, \mathbf{\Sigma}_k), \quad 0 \leq \pi_k \leq 1, \quad \sum_{k=1}^K \pi_k = 1,$$

where the parameters $oldsymbol{ heta} := \{oldsymbol{\mu}_k, oldsymbol{\Sigma}_k, \pi_k : k = 1, \dots, K\}$

• Example. $p(x|\theta) = 0.5\mathcal{N}(x|-2,1/2) + 0.2\mathcal{N}(x|1,2) + 0.3\mathcal{N}(x|4,1)$



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Roadmap



- (1) Gaussian Mixture Model
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Parameter Learning: Maximum Likelihood



• Given a iid dataset $\mathcal{X} = \{ extbf{\emph{x}}_1, \dots, extbf{\emph{x}}_n \},$ the log-likelihood is:

$$\mathcal{L}(\boldsymbol{\theta}) = \log p(\mathcal{X}|\boldsymbol{\theta}) = \sum_{n=1}^{N} \log p(\boldsymbol{x}_n|\boldsymbol{\theta}) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\boldsymbol{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- $\theta_{\mathsf{ML}} = \operatorname{arg\,min}_{\boldsymbol{\theta}}(-\mathcal{L}(\boldsymbol{\theta}))$
- Necessary condition for $heta_{
 m ML}$: $\left. rac{d\mathcal{L}}{d heta} \right|_{ heta_{
 m ML}} = 0$
- However, the closed-form solution of $\theta_{\rm ML}$ does not exist, so we rely on an iterative algorithm (also called EM algorithm).
- We show the algorithm first, and then discuss how we get the algorithm.

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Responsibilities



• Definition. Responsibilities. Given *n*-th data point x_n and the parameters $(\mu_k, \Sigma_k, \pi_k : k = 1, ..., K)$,

$$r_{nk} = rac{\pi_k \mathcal{N}(\mathbf{x}_n | \mathbf{\mu}_k, \mathbf{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_n | \mathbf{\mu}_j, \mathbf{\Sigma}_j)}$$

- How much is each component k responsible, if the data \mathbf{x}_n is sampled from the current mixture model?
- $r_n = (r_{nk} : k = 1, ..., K)$ is a probability distribution, so $\sum_{k=1}^K r_{nk} = 1$
- Soft assignment of x_n to the K mixture components

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EM Algorithm: MLE in Gaussian Mixture Models



EM for MLE in Gaussian Mixture Models

- **S1.** Initialize μ_k, Σ_k, π_k
- **S2.** E-step: Evaluate responsibilities r_{nk} for every data point \mathbf{x}_n using the current $\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \boldsymbol{\pi}_k$:

$$r_{nk} = rac{\pi_k \mathcal{N}(\mathbf{x}_n | oldsymbol{\mu}_k, oldsymbol{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_n | oldsymbol{\mu}_j, oldsymbol{\Sigma}_j)}, \quad oldsymbol{N}_k = \sum_{n=1}^N r_{nk}$$

S3. M-step: Reestimate parameters μ_k , Σ_k , π_k using the current responsibilities r_{nk} :

$$\mu_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} r_{nk} \mathbf{x}_{n}, \ \Sigma_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} r_{nk} (\mathbf{x}_{n} - \mu_{k}) (\mathbf{x}_{n} - \mu_{k})^{\mathsf{T}}, \ \pi_{k} = \frac{N_{k}}{N},$$

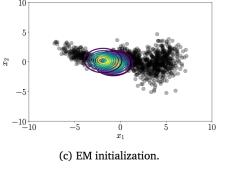
and go to \$2.

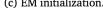
- The update equation in M-step is still mysterious, which will be covered later.

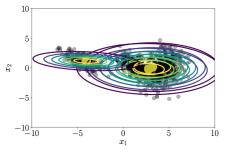
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Example: EM Algorithm

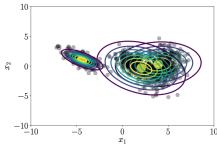




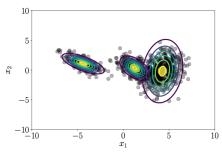




(d) EM after one iteration.



(e) EM after 10 iterations.



(f) EM after 62 iterations.

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M-Step: Towards the Zero Gradient



• Given \mathcal{X} and r_{nk} from E-step, the new updates of μ_k , Σ_k , π_k should be made, such that the followings are satisfied:

$$\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}_k} = \mathbf{0}^\mathsf{T} \Longleftrightarrow \sum_{n=1}^N \frac{\partial \log p(\boldsymbol{x}_n | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} = \mathbf{0}^\mathsf{T}$$

$$\frac{\partial \mathcal{L}}{\partial \Sigma_k} = 0 \Longleftrightarrow \sum_{n=1}^N \frac{\partial \log p(\mathbf{x}_n | \boldsymbol{\theta})}{\partial \Sigma_k} = 0$$

$$\frac{\partial \mathcal{L}}{\partial \pi_k} = 0 \Longleftrightarrow \sum_{n=1}^{N} \frac{\partial \log p(\mathbf{x}_n | \boldsymbol{\theta})}{\partial \pi_k} = 0$$

- Nice thing: the new updates of μ_k , Σ_k , π_k are all expressed by the responsibilities $[r_{nk}]$
- · Let's take a look at them one by one!

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M-Step: Update of μ_k



$$\mu_k^{\text{new}} = \frac{\sum_{n=1}^{N} r_{nk} x_n}{\sum_{n=1}^{N} r_{nk}}, k = 1, \dots, K$$

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M-Step: Update of Σ_k



$$\boldsymbol{\Sigma}_k^{\mathsf{new}} = \frac{1}{N_k} \sum_{n=1}^N r_{nk} (\boldsymbol{x}_n - \boldsymbol{\mu}_k) (\boldsymbol{x}_n - \boldsymbol{\mu}_k)^\mathsf{T}, k = 1, \dots, K$$

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M-Step: Update of π_k



$$\pi_k^{\mathsf{new}} = \frac{\sum_{n=1}^N r_{nk}}{N}, k = 1, \dots, K$$

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Roadmap



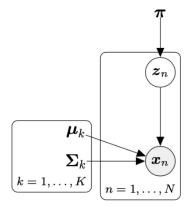
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Latent-Variable Perspective



- · Justify some ad hoc decisions made earlier
- Allow for a concrete interpretation of the responsibilities as posterior distributions
- Iterative algorithm for updating the model parameters can be derived in a principled manner



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Generative Process



- Latent variable z: One-hot encoding random vector $z = [z_1, \dots, z_K]^T$ consisting of K-1 many 0s and exactly one 1.
- An indicator rv $z_k = 1$ represents whether k-th component is used to generate the data sample x or not.
- $p(\mathbf{x}|z_k=1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$
- Prior for z with $\pi_k = p(z_k = 1)$

$$p(\mathbf{z}) = \boldsymbol{\pi} = [\pi_1, \dots, \pi_K]^\mathsf{T}, \quad \sum_{k=1}^K \pi_k = 1$$

- Sampling procedure
 - 1. Sample which component to use $z^{(i)} \sim p(z)$
 - 2. Sample data according to *i*-th Gaussian $\mathbf{x}^{(i)} \sim p(\mathbf{x}|z^{(i)})$

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Joint Distribution, Likelihood, and Posterior (1)



Joint distribution

$$p(\mathbf{x}, \mathbf{z}) = egin{pmatrix}
ho(\mathbf{x}, z_1 = 1) \ dots \
ho(\mathbf{x}, z_K = 1) \end{pmatrix} = egin{pmatrix}
ho(\mathbf{x}|z_1 = 1)
ho(z_1 = 1) \ dots \
ho(\mathbf{x}|z_K = 1)
ho(z_K = 1) \end{pmatrix} = egin{pmatrix} \pi_1 \mathcal{N}(\mathbf{x}|oldsymbol{\mu}_1, oldsymbol{\Sigma}_1) \ dots \ \pi_K \mathcal{N}(\mathbf{x}|oldsymbol{\mu}_K, oldsymbol{\Sigma}_K) \end{pmatrix}$$

• Likelihood for an arbitrary single data x: By summing out all latent variables¹,

$$p(\mathbf{x}|\boldsymbol{\theta}) = \sum_{\mathbf{z}} p(\mathbf{x}|\boldsymbol{\theta}, \mathbf{z}) p(\mathbf{z}|\boldsymbol{\theta}) = \sum_{k=1}^{K} p(\mathbf{x}|\boldsymbol{\theta}, z_k = 1) p(z_k = 1|\boldsymbol{\theta}) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

• For all the data samples \mathcal{X} , the log-likelihood is:

$$\log p(\mathcal{X}|\boldsymbol{\theta}) = \sum_{n=1}^{N} \log p(\mathbf{x}_n|\boldsymbol{\theta}) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

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 $^{^{1}}$ In probabilistic PCA, z was continuous, so we integrated them out.

Joint Distribution, Likelihood, and Posterior (2)



• Posterior for the k-th z_k , given an arbitrary single data x:

$$p(z_k = 1 | \mathbf{x}) = \frac{p(z_k = 1)p(\mathbf{x}|z_k = 1)}{\sum_{i=1}^{K} p(z_i = 1)p(\mathbf{x}|z_i = 1)} = \frac{\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{i=1}^{K} \pi_j \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_j)}$$

• Now, for all data samples \mathcal{X} , each data \mathbf{x}_n has $\mathbf{z}_n = [z_{n1}, \dots, z_{nK}]^\mathsf{T}$, but with the same prior π .

$$p(z_{nk} = 1 | \mathbf{x}_n) = \frac{p(z_{nk} = 1)p(\mathbf{x}_n | z_{nk} = 1)}{\sum_{j=1}^{K} p(z_{nj} = 1)p(\mathbf{x}_n | z_{nj} = 1)} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} = r_{nk}$$

• Responsibilities are mathematically interpreted as posterior distributions.

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Roadmap



- (1) Gaussian Mixture Model
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Revisiting EM Algorithm for MLE



- **S1.** Initialize μ_k, Σ_k, π_k
- S2. E-step:

$$r_{nk} = rac{\pi_k \mathcal{N}(m{x}_n | m{\mu}_k, m{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(m{x}_n | m{\mu}_j, m{\Sigma}_j)}$$

- **S3.** M-step: Update μ_k, Σ_k, π_k using r_{nk} and go to **S2**.
- **E-step.** Expectation over $\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^{(t)}$: Given the current $\boldsymbol{\theta}^{(t)} = (\mu_k, \Sigma_k, \pi_k)$, calculates the expected log-likelihood

$$\begin{aligned} Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) &= \mathbb{E}_{\boldsymbol{z}|\boldsymbol{x},\boldsymbol{\theta}^{(t)}}[\log p(\boldsymbol{x},\boldsymbol{z}|\boldsymbol{\theta})] \\ &= \int \log p(\boldsymbol{x},\boldsymbol{z}|\boldsymbol{\theta})p(\boldsymbol{z}|\boldsymbol{x},\boldsymbol{\theta}^{(t)})d\boldsymbol{z} \end{aligned}$$

- M-step. Maximization of the computation results in E-step for the new model parameters.
- Only guarantee of just local-optimum because the original optimization is not necessarily a convex optimization.

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Other Issues



- Model selection for finding a good K, e.g., using nested cross-validation
- Application: Clustering
 - K-means: Treat the means in GMM as cluster centers and ignore the covariances.
 - K-means: hard assignment, GMM: soft assignment
- EM algorithm: Highly generic in the sense that it can be used for parameter learning in general latent-variable models
- Standard criticism for MLE exists such as overfitting. Also, fully-Bayesian approach assuming some priors on the parameters is possible, but not covered in this notes.
- Other density estimation methods
 - Histogram-based method: non-parametric method
 - Kernel-density estimation: non-parametric method

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Questions?

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Review Questions



1)

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