



Lecture 11: Density Estimation with Gaussian Mixture Models

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Mathematics for Machine Learning
https://yung-web.github.io/home/courses/mathml.html
KAIST EE

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Please watch this tutorial video by Luis Serrano on Gaussian Mixture Model.

https://www.youtube.com/watch?v=q71Niz856KE

April 7, 2021 1 / 26

April 7, 2021 2 / 26

Roadmap



Roadmap

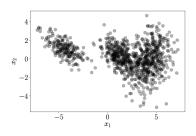


- (1) Gaussian Mixture Model
- (2) Parameter Learning: MLE
- (3) Latent-Variable Perspective for Probabilistic Modeling
- (4) EM Algorithm

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- Represent data compactly using a density from a parametric family, e.g., Gaussian or Beta distribution
- Parameters of those families can be found by MLE and MAPE
- However, there are many cases when simple distributions (e.g., just Gaussian) fail to approximate data.



L11(1) April 7, 2021 5 / 26

- More expressive family of distribution
- Idea: Let's mix! A convex combination of K "base" distributions

$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k p_k(\mathbf{x}), \quad 0 \leq \pi_k \leq 1, \quad \sum_{k=1}^K \pi_k = 1$$

- Multi-modal distributions: Can be used to describe datasets with multiple clusters
- Our focus: Gaussian mixture models
- Want to finding the parameters using MLE, but cannot have the closed form solution (even with the mixture of Gaussians) → some iterative methods needed

L11(1) April 7, 2021 6 / 26

Gaussian Mixture Model

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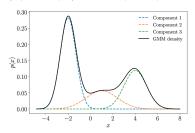
Roadmap



 $p(\mathbf{x}|\boldsymbol{\theta}) = \sum_{k=1}^K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad 0 \leq \pi_k \leq 1, \quad \sum_{k=1}^K \pi_k = 1,$

where the parameters $\boldsymbol{\theta} := \{ \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \pi_k : k = 1, \dots, K \}$

• Example. $p(x|\theta) = 0.5\mathcal{N}(x|-2,1/2) + 0.2\mathcal{N}(x|1,2) + 0.3\mathcal{N}(x|4,1)$



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• Given a iid dataset $\mathcal{X} = \{x_1, \dots, x_n\}$, the log-likelihood is:

$$\mathcal{L}(\theta) = \log p(\mathcal{X}|\theta) = \sum_{n=1}^{N} \log p(\mathbf{x}_n|\theta) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- $\theta_{\mathsf{ML}} = \operatorname{arg\,min}_{\boldsymbol{\theta}}(-\mathcal{L}(\boldsymbol{\theta}))$
- Necessary condition for θ_{ML} : $\frac{d\mathcal{L}}{d\theta}\Big|_{\theta_{ML}} = 0$
- However, the closed-form solution of $\theta_{\rm ML}$ does not exist, so we rely on an iterative algorithm (also called EM algorithm)
- We show the algorithm first, and then discuss how we get the algorithm.

• Definition. Responsibilities. Given *n*-th data point x_n and the parameters $(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \boldsymbol{\pi}_k : k = 1, \ldots, K),$

$$r_{nk} = rac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

- How much is each component k responsible, if the data x_n is sampled from the current mixture model?
- $\mathbf{r}_n = (r_{nk} : k = 1, ..., K)$ is a probability distribution, so $\sum_{k=1}^K r_{nk} = 1$
- Soft assignment of x_n to the K mixture components

L11(2) April 7, 2021 9 / 26

L11(2)

April 7, 2021 10 / 26

EM Algorithm: MLE in Gaussian Mixture Models

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Example: EM Algorithm



EM for MLE in Gaussian Mixture Models

- **S1.** Initialize μ_k, Σ_k, π_k
- **S2.** E-step: Evaluate responsibilities r_{nk} for every data point x_n using the current μ_k, Σ_k, π_k :

$$r_{nk} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}, \quad N_k = \sum_{n=1}^N r_{nk}$$

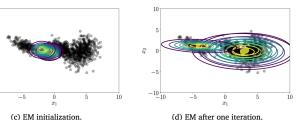
S3. M-step: Reestimate parameters μ_k , Σ_k , π_k using the current responsibilities r_{nk} :

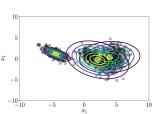
$$\mu_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} r_{nk} \mathbf{x}_{n}, \ \Sigma_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} r_{nk} (\mathbf{x}_{n} - \mu_{k}) (\mathbf{x}_{n} - \mu_{k})^{\mathsf{T}}, \ \pi_{k} = \frac{N_{k}}{N},$$

and go to \$2.

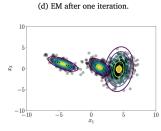
- The update equation in M-step is still mysterious, which will be covered later.







(e) EM after 10 iterations.





• Given $\mathcal X$ and r_{nk} from E-step, the new updates of μ_k , Σ_k , π_k should be made, such that the followings are satisfied:

$$\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}_k} = 0^{\mathsf{T}} \Longleftrightarrow \sum_{n=1}^{N} \frac{\partial \log p(\boldsymbol{x}_n | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} = 0^{\mathsf{T}}$$

$$\frac{\partial \mathcal{L}}{\partial \Sigma_k} = 0 \Longleftrightarrow \sum_{n=1}^N \frac{\partial \log p(\mathbf{x}_n | \boldsymbol{\theta})}{\partial \Sigma_k} = 0$$

$$\frac{\partial \mathcal{L}}{\partial \pi_k} = 0 \Longleftrightarrow \sum_{n=1}^{N} \frac{\partial \log p(\mathbf{x}_n | \boldsymbol{\theta})}{\partial \pi_k} = 0$$

- Nice thing: the new updates of μ_k , Σ_k , π_k are all expressed by the responsibilities $[r_{nk}]$
- Let's take a look at them one by one!

 $oldsymbol{\mu}_k^{\mathsf{new}} = rac{\sum_{n=1}^N r_{nk} oldsymbol{x}_n}{\sum_{n=1}^N r_{nk}}, k = 1, \dots, \mathcal{K}$

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L11(2) April 7, 2021 13 / 26 L11(2) April 7, 2021 14 / 26

M-Step: Update of Σ_k



M-Step: Update of π_k



$$oldsymbol{\Sigma}_k^{\mathsf{new}} = rac{1}{N_k} \sum_{n=1}^N r_{nk} (oldsymbol{x}_n - oldsymbol{\mu}_k) (oldsymbol{x}_n - oldsymbol{\mu}_k)^\mathsf{T}, k = 1, \dots, K$$

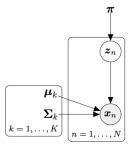
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$$\pi_k^{\mathsf{new}} = \frac{\sum_{n=1}^{N} r_{nk}}{N}, k = 1, \dots, K$$

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• Justify some ad hoc decisions made earlier

- Allow for a concrete interpretation of the responsibilities as posterior distributions
- Iterative algorithm for updating the model parameters can be derived in a principled manner



L11(3)

L11(3)

Joint distribution

April 7, 2021 18 / 26

Generative Process

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April 7, 2021

19 / 26

April 7, 2021

17 / 26

Joint Distribution, Likelihood, and Posterior (1)



- Latent variable z: One-hot encoding random vector $z = [z_1, \dots, z_K]^T$ consisting of K-1 many 0s and exactly one 1.
- An indicator rv $z_k = 1$ represents whether k-th component is used to generate the data sample x or not.
- $p(\mathbf{x}|z_k=1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$
- Prior for z with $\pi_k = p(z_k = 1)$

$$p(\mathbf{z}) = \mathbf{\pi} = [\pi_1, \dots, \pi_K]^\mathsf{T}, \quad \sum_{k=1}^K \pi_k = 1$$

- Sampling procedure
 - 1. Sample which component to use $z^{(i)} \sim p(z)$
 - 2. Sample data according to *i*-th Gaussian $\mathbf{x}^{(i)} \sim p(\mathbf{x}|z^{(i)})$

$$p(\mathbf{x}, \mathbf{z}) = \begin{pmatrix} p(\mathbf{x}, z_1 = 1) \\ \vdots \\ p(\mathbf{x}, z_K = 1) \end{pmatrix} = \begin{pmatrix} p(\mathbf{x}|z_1 = 1)p(z_1 = 1) \\ \vdots \\ p(\mathbf{x}|z_K = 1)p(z_K = 1) \end{pmatrix} = \begin{pmatrix} \pi_1 \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) \\ \vdots \\ \pi_K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K) \end{pmatrix}$$

• Likelihood for an arbitrary single data x: By summing out all latent variables¹,

$$p(\mathbf{x}|\boldsymbol{\theta}) = \sum_{\mathbf{z}} p(\mathbf{x}|\boldsymbol{\theta}, \mathbf{z}) p(\mathbf{z}|\boldsymbol{\theta}) = \sum_{k=1}^{K} p(\mathbf{x}|\boldsymbol{\theta}, z_k = 1) p(z_k = 1|\boldsymbol{\theta}) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

ullet For all the data samples ${\mathcal X},$ the log-likelihood is:

$$\log p(\mathcal{X}|\boldsymbol{\theta}) = \sum_{n=1}^{N} \log p(\mathbf{x}_n|\boldsymbol{\theta}) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

Compare: Page 7

L11(3) April 7, 2021 20 / 26

 $^{^{1}}$ In probabilistic PCA, z was continuous, so we integrated them out.



• Posterior for the k-th z_k , given an arbitrary single data x:

$$p(z_k = 1 | \boldsymbol{x}) = \frac{p(z_k = 1)p(\boldsymbol{x}|z_k = 1)}{\sum_{j=1}^{K} p(z_j = 1)p(\boldsymbol{x}|z_j = 1)} = \frac{\pi_k \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

• Now, for all data samples $\mathcal{X},$ each data \pmb{x}_n has $\pmb{z}_n = [z_{n1}, \dots, z_{nK}]^\mathsf{T},$ but with the same prior π .

$$p(z_{nk} = 1 | \mathbf{x}_n) = \frac{p(z_{nk} = 1)p(\mathbf{x}_n | z_{nk} = 1)}{\sum_{j=1}^{K} p(z_{nj} = 1)p(\mathbf{x}_n | z_{nj} = 1)} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)} = r_{nk}$$

Responsibilities are mathematically interpreted as posterior distributions.

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L11(4)

L11(3) April 7, 2021 21 / 26

Revisiting EM Algorithm for MLE



Other Issues



April 7, 2021

22 / 26

- **S1.** Initialize μ_k, Σ_k, π_k
- S2. E-step:

$$r_{nk} = rac{\pi_k \mathcal{N}(\mathbf{x}_n | \mathbf{\mu}_k, \mathbf{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_n | \mathbf{\mu}_j, \mathbf{\Sigma}_j)}$$

- **S3.** M-step: Update μ_k, Σ_k, π_k using r_{nk} and go to S2.
- E-step. Expectation over $z|x, \theta^{(t)}$: Given the current $\theta^{(t)} = (\mu_k, \Sigma_k, \pi_k)$, calculates the expected log-likelihood

$$Q(\theta|\theta^{(t)}) = \mathbb{E}_{z|x,\theta^{(t)}}[\log p(x,z|\theta)]$$
$$= \int \log p(x,z|\theta)p(z|x,\theta^{(t)})dz$$

- M-step. Maximization of the computation results in E-step for the new model parameters.
- Only guarantee of just local-optimum because the original optimization is not necessarily a convex optimization. L7(4)

- Model selection for finding a good K, e.g., using nested cross-validation
- Application: Clustering
 - K-means: Treat the means in GMM as cluster centers and ignore the covariances.
 - K-means: hard assignment, GMM: soft assignment
- EM algorithm: Highly generic in the sense that it can be used for parameter learning in general latent-variable models
- Standard criticism for MLE exists such as overfitting. Also, fully-Bayesian approach assuming some priors on the parameters is possible, but not covered in this notes.
- Other density estimation methods
 - Histogram-based method: non-parametric method
 - Kernel-density estimation: non-parametric method





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Questions?

L11(4)

April 7, 2021 25 / 26

L11(4)

April 7, 2021 26 / 26