Introduction to Machine Learning

Lecture 11 - Unsupervised Learning: Principal Component Analysis
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Introduction to Continuous Latent Variables

- Continuous Latent Variables
- · Principal Component Analysis
- · Factor Analysis

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Continuous Latent Variables

- Previously discussed discrete latent variables, such as mixture of Gaussians.
- Sometimes, it is more appropriate to think in terms of continuous factors which control the data we observe.
- · This motivation for such models is:
- for many datasets, data points lie close to a manifold of much lower dimensionality compared to that of the original data space.
- Training continuous latent variable models often called dimensionality reduction, since there are typically many fewer latent dimensions.
- · Examples include:
 - · Principal Components Analysis,
 - · Factor Analysis,
 - · Independent Components Analysis

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Intrinsic Latent Dimensions

• In this dataset, there is only 3 degrees of freedom of variability, vertical and horizontal translations, and the rotations.











Copied from the book: Pattern Recognition and Machine Learning by Christopher M. Bishop

- · Each image undergoes a random displacement and rotation within some larger image field.
- The resulting images have 100 x 100 = 10,000 pixels.
- However, the data points live on a subspace of the data space whose intrinsic dimensionality is three.

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Generative View

- To generate the data points, first selecting a point from a distribution in the latent space, then sampling a point from the conditional distribution in the input space
- Taking Gaussian distribution for both latent and observed variables is the simplest latent variable model.
- This generative approach leads to probabilistic formulation of the **Principal** Component Analysis and Factor Analysis.
- Take look at the standard PCA, and then consider its probabilistic formation
- Mixture of PCAs, Bayesian PCA are advantages of using EM for parameters.

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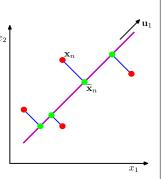
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Principal Component Analysis

- Widely used for data compression, visualization, feature extraction, and dimensionality reduction
- The goal is find M principal components underlying D-dimensional data
- Select the top M eigenvectors of S (data covariance matrix): $\{u_1, \dots u_M\}$
- project each input vector X into this subspace, e.g., $z_{n1} = X_n^T u_1$
- The full projection into M dimensions:

$$\begin{bmatrix} \mathbf{u}_1^\top \\ \cdots \\ \mathbf{u}_M^\top \end{bmatrix} [\mathbf{x}_1 \cdots \mathbf{x}_N] = [\mathbf{z}_1 \cdots \mathbf{z}_N]$$

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Maximum Variance

- Consider a dataset $\{x_1, \dots, x_N\}, x_n \in \mathbb{R}^D$. Our goal is to project data onto a space having dimensionality
- Consider the projection into M=1 dimensional space
- Define the direction of this space using a D-dimensional unit vector \mathbf{u}_1 , so that $u_1^T u_1 = 1$
- Objective is to maximize the variance of the projected data w.r.t. u_1

•
$$\frac{1}{N} \sum_{n=1}^{N} \{u_1^T x_n - u_1^T \bar{x}\}^2 = u_1^T S u_1,$$

• where sample mean and data covariance:

•
$$\bar{x} = \frac{1}{N} \sum_{n=1}^{N} x_n$$

• $S = \frac{1}{N} \sum_{n=1}^{N} (x_n - \bar{x})(x_n - \bar{x})^T$

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Maximum Variance

- To maximize the variance of the projected data, use Language multiplier with constrain $||u_1|| = 1$
- maximum: $u_1^T S u_1 + \lambda (1 u_1^T u_1)$
- setting the derivative w.r.t. u_1 to zero:
- $Su_1 = \lambda_1 u_1$
- Hence u_1 must be an eigenvector of S, and
- the maximum variance of the projected data is given by: $u_1^T S u_1 = \lambda_1$,
- Optimal u_1 is the principal component (eigenvector with maximal eigenvalue)

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Minimum Error

- Introduce a complete orthonormal set of D-dimensional basis vectors:
- $\{u_1,\ldots,u_D\},\ u_i^Tu_i=\delta_{ii},$
- let $x_n = \sum_{i=1}^{D} \alpha_{ni} u_i$, $\alpha_{ni} = x_n^T u_i$, rotation of the coordinate system to a new system defined by u_i
- This enables the data to be represented by the projection into M-dimensional subspace as (represent M-dimensional linear subspace by the first M of the basis vectors):

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$$\tilde{x}_n = \sum_{i=1}^M z_{ni} u_i + \sum_{i=M+1}^D b_i u_i$$

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Minimum Error

- For the $\tilde{x}_n = \sum_{i=1}^M z_{ni}u_i + \sum_{i=M+1}^D b_iu_i$, where z_{ni} depend on the particular data point and b_i
- To minimize the distortion w.r.t. u_i , z_{ni} , and b_i

$$J = \frac{1}{N} \sum_{n=1}^{N} ||x_n - \tilde{x}_n||^2$$

- minimize J w.r.t. z_{ni} and b_i : $z_{nj} = x_n^T u_i$, $b_j = \bar{x}^T u_j$
- Then the objective J reduces to:, $J = \frac{1}{N} \sum_{n=1}^{N} \sum_{i=M+1}^{D} (x_n^T u_i \bar{x}^T u_i)^2 = \sum_{i=M+1}^{D} u_i^T S u_i$

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Applications of PCA

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• Apply PCA on 2429 19 x 19 grayscale images from CBCL database

Minimum Error

- The general solution is obtained by choosing u_i to be eigenvectors of the covariance matrix:
- $Su_i = \lambda_i u_i$
- . The distortion is then given by: $J = \sum_{i=1}^{D} \lambda_{i}$
- The M components are the eigenvectors of S with lowest eigenvalues when objective J is minimized.

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• with only 3 components

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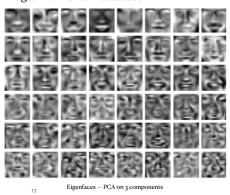
• For pre-processing, PCA with 3 components obtains 79% accuracy on face/non-face

discrimination test vs. 76.8% for mixture of Gaussian with 84 components.

Applications of PCA

• Apply PCA on 2429 19 x 19 grayscale images from CBCL database

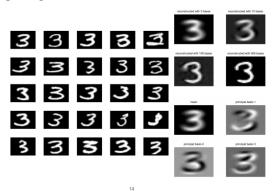
• good for "eigenfaces"



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Applications of PCA

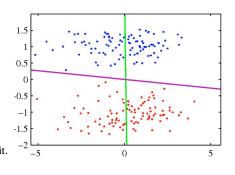
• Apply PCA on digit images



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PCA vs. Fisher's LDA

- · Both PCA and Fisher's LDA can be apply for linear dimensionality reduction.
- PCA chooses direction of maximum variance using unsupervised approach.
- The magenta curve shows a strong class overlap
- LDA uses supervised way with the class labels.
- The green curve shows a projection into it.



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PCA for High-Dimensional Data

- In some applications of PCA, the number of data points is smaller than the dimensionality of the data space, i.e. N<D.
- To find the eigenvectors of the D x D data covariance matrix S, the computation expense is $O(D^3)$.
- Thus, direct application of PCA is often computationally infeasible.
- To solve this disadvantage, here is a solution:
- Let X be the N x D centered data matrix. The corresponding eigenvector equation becomes: $\frac{1}{N}X^TXu_i = \lambda_i u_i$

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PCA for High-Dimensional Data

- Pre-multiply by X: $\frac{1}{N}XX^TXu_i = \lambda_i(Xu_i)$
- Let $v_i = Xu_i$, hence $\frac{1}{N}XX^Tv_i = \lambda_i v_i$
- This is an eigenvector equation for the N x N matrix.
 - $O(N^3) \ll O(D^3)$.
- To determine eigenvectors, multiply by X^T : $(\frac{1}{N}XX^T)(X^Tv_i) = \lambda_i X^Tv_i$
- Hence, $X^T v_i$ is an eigenvector of S with eigenvalue λ_i

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Probabilistic PCA

- · Many advantages of Probabilistic PCA (PPCA):
- It represents a constrained form of the Gaussian distribution.
- · Able to derive EM algorithm for PCA which is computationally efficient.
- PPCA can deal with missing values in the data set.
- Mixture of PPCAs can be formulated in a principled way.
- PPCA forms the basis for a Bayesian PCA, in which the dimensionality of the principal subspace can be determined from the data.
- The existence of a likelihood function allows direct comparisons with other probabilistic density models
- PPCA can be used to model class conditional densities and hence it can be applied to classification problems.

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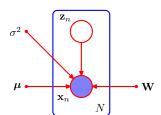
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Probabilistic PCA

- · Assumptions to formulate the PPCA:
- underlying latent M-dim variable z has a Gaussian distribution.
- linear relationship between M-dim latent z and D-dim observed x
- · isotropic Gaussian noise in observed
- p(z) = N(z | 0,I)
- $p(x|z) = N(x|Wz + \mu, \sigma^2 I)$
- The mean of x is a linear function of z governed by the D x M matrix W and the D-dim vector μ .
- The columns of W span the principal subspace of the data space (Columns of W are the principal components, σ² is sensor noise).

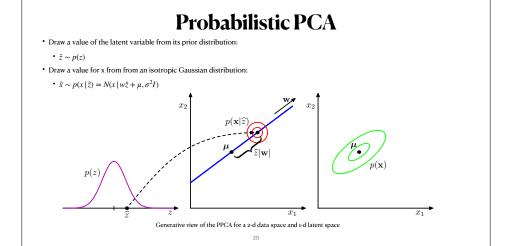


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Marginal Data Density

• The joint p(z, x), the marginal data distribution p(x) and the posterior p(z|x) are Gaussians

•
$$p(X) = \int p(z)p(x|z)dz = N(x|\mu, WW^T + \sigma^2 I), \ x = Wz + \mu + \epsilon$$

- · This is the marginal data density, also known as predictive distribution.
- · By computing mean and covariance of Gaussian:
- $E[x] = E[\mu + Wz + \epsilon] = \mu + WE[z] + E[\epsilon] = \mu + W0 + 0 = \mu$

$$\begin{split} C &= \operatorname{Cov}[x] \\ &= E[(x - \mu)(x - \mu)^T] \\ &= E[(\mu + Wz + \epsilon - \mu)(\mu + Wz + \epsilon - \mu)^T] \\ &= E[(Wz + \epsilon)(Wz + \epsilon)^T] \\ &= WW^T + \sigma^2 I \end{split}$$

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Redundancy in Parameterization

• Given the marginal distribution:

$$p(X) = \int_{z} p(z)p(x|z)dz = N(x|\mu, WW^{T} + \sigma^{2}I), \quad x = Wz + \mu + \epsilon$$

- Let R be an orthogonal matrix, then define a new matrix:
- $\hat{W} = WR$, $RR^T = I$
- $\hat{W}\hat{W}^T = WRR^TW^T = WW^T$.
- the redundancy in parameterization as if rotating the latent space coordinates.
- Hence, there is a whole family of matrices that all of which give rise to the same marginal distribution when rotating within the latent space.

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Joint Density for PPCA

• The joint density for PPCA is given as:

$$p(\begin{bmatrix}\mathbf{z}\\\mathbf{x}\end{bmatrix}) = \mathcal{N}(\begin{bmatrix}\mathbf{z}\\\mathbf{x}\end{bmatrix} | \begin{bmatrix}\mathbf{0}\\\boldsymbol{\mu}\end{bmatrix}, \begin{bmatrix}\boldsymbol{I} & \mathbf{W}^\top\\\mathbf{W} & \mathbf{W}\mathbf{W}^\top + \sigma^2\mathbf{I}\end{bmatrix})$$

· and covariance:

$$Cov[\mathbf{z}, \mathbf{x}] = E[(\mathbf{z} - 0)(\mathbf{x} - \mu)^T] = E[\mathbf{z}(\mu + \mathbf{W}\mathbf{z} + \epsilon - \mu)^T]$$
$$= E[\mathbf{z}(\mathbf{W}\mathbf{z} + \epsilon)^T] = \mathbf{W}^T$$

Reduce O(D^3) to O(M^3) by applying matrix inversion lemma:

$$C^{-1} = \sigma^{-1}I - \sigma^{-2}W(W^TW + \sigma^2I)^{-1}W^T$$

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Posterior Distribution for PPCA

- The posterior distribution is about the inference problem in PPCA:
- p(z|x) = N(z|m, V)
- $m = M^{-1}W^{T}(x \mu)$
- $V = \sigma^2 M^{-1}$
- $M = W^TW + \sigma^2I$
- Mean of inferred z is projection of centred x, a linear operation.
- The posterior variance does not depend on the input x at all
- Since $C^{-1} = \sigma^{-1}I \sigma^{-2}W(W^TW + \sigma^2I)^{-1}W^T$, and $C = WW^T + \sigma^2I$.

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Maximum Likelihood

• Using maximum likelihood (by integrating out latent variables) to determine the model parameters:

$$\begin{split} L(\theta; \mathbf{X}) &= \log p(\mathbf{X}|\theta) = \sum_{n} \log p(\mathbf{x}_{n}|\theta) \\ &= -\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} \sum_{n} (\mathbf{x}_{n} - \mu) \mathbf{C}^{-1} (\mathbf{x}_{n} - \mu)^{T} \\ &= -\frac{N}{2} \log |\mathbf{C}| - \frac{1}{2} Tr[\mathbf{C}^{-1} \sum_{n} (\mathbf{x}_{n} - \mu) (\mathbf{x}_{n} - \mu)^{T}] + \text{const.} \end{split}$$

- Let $\mu_{ML} = \bar{x}$, then
- $\log p(X \mid \theta) = -\frac{N}{2} \log |C| \frac{1}{2} Tr[C^{-1}S] + \text{const}$

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Maximum Likelihood

Maximum Likelihood

- Maximizing w.r.t. W: $W_{ML} = U_M (L_M \sigma^2 I)^{1/2} R$
- where
- $U_M \in R^{D \times M}$ matrix whose columns are given by the M principal eigenvectors of the data covariance matrix S.
- $L_M \in \mathbb{R}^{M \times M}$ diagonal matrix containing M largest eigenvalues.
- $R \to M \times M$ an arbitrary orthogonal matrix.

If the eigenvectors have been arranged in the order of decreasing values of the corresponding eigenvalues, then the columns of W define the principal subspace of standard PCA.

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• Maximizing w.r.t. σ^2 : $\sigma^2_{ML} = \frac{1}{D-M} \sum_{i=M+1}^{D} \lambda_i$

• which is the average variance associated with the discarded dimensions

EM for PPCA

- Instead of solving directly, we can use EM. The EM can be scaled to very large highdimensional datasets.
- The complete-data log-likelihood takes form:

$$\log p(X, Z \mid \mu, W, \sigma^2) = \sum_n \left[\log p(x_n \mid z_n) + \log p(z_n) \right]$$

- In E-step:
 - compute expectation of complete log likelihood w.r.t. z, using the current parameters.
 - Need to derive $E[z_n]$, $E[z_n z_n^T]$ w.r.t. the true posterior: $p(z \mid x)$

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EM for PPCA

- In M-step:
- Maximize w.r.t. parameters W and σ^2
- EM avoids direct $O(ND^2)$ construction of covariance matrix.
- Instead EM involves sums over data cases: O(NDM). It can also be implemented online, without storing data.

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EM for PPCA

- It is able to derive standard PCA as a limit of probabilistic PCA as the noise term goes to zero: $\sigma^2 \to 0$
- · Maximum likelihood parameters are the same.
- Inferring the distribution over latent variables is easier: The posterior mean reduces to:
- $\lim_{\sigma^2 \to 0} (W^T W + \sigma^2 I)^{-1} W^T (x \mu) = (W^T W)^{-1} W^T (x \mu)$
- which represents an orthogonal projection of the data point onto the latent space standard PCA.
- · Posterior covariance goes to zero

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Bayesian PCA

- A Bayesian viewpoint and place priors over model parameters is given as:
- $\bullet\,$ define an independent Gaussian prior over each column of W.
- $\bullet\,$ employ the evidence approximation (empirical Bayes) framework.
- Such Gaussian has an independent variance

$$p(W|\alpha) = \prod_{i=1}^{M} (\frac{\alpha_i}{a\pi}) \exp[\frac{1}{2}\alpha_i W_i^T W_i]$$

• where w; is the with column of W.

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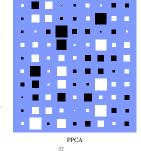
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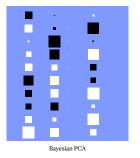
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- The values of α_i are re-estimated during training by maximizing the marginal likelihood.
- $p(X \mid \alpha, \mu, \sigma^2) = \int p(X \mid W, \mu, \sigma^2) p(W \mid \alpha) dW$
- Hinton diagram of the matrix W: each element of W is depicted as a square (white for positive and black for negative)
- 2. Bayesian PCA discovers appropriate dimensionality





Factor Analysis

- Use a linear Gaussian latent variable model which is related to PPCA.
- Assume:
- underlying latent M-dim variable z has a Gaussian distribution
- linear relationship between M-dim latent z and D-dim observed X variables
- · diagonal Gaussian noise in observed dimensions given as:
- p(z) = N(z | 0,I)
- $p(x | z) = N(x | Wz + \mu, \Psi)$
- W is a D x M factor loading matrix
- $\Psi \to M \times M$ diagonal matrix
- The only difference between PPCA and FA is that in Factor Analysis the 20 conditional distribution of the observed variable x has diagonal rather than isotropic covariance.

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Factor Analysis

- Same as the PPCA, the joint p(z, x), the marginal p(x) and the posterior p(z|x) are Gaussians.
- Marginal distribution: $p(X) = \int_{z} p(z)p(x \mid z)dz = N(x \mid \mu, WW^{T} + \Psi)$
- The joint distribution:

$$p(\begin{bmatrix}\mathbf{z}\\\mathbf{x}\end{bmatrix}) = \mathcal{N}(\begin{bmatrix}\mathbf{z}\\\mathbf{x}\end{bmatrix} \mid \begin{bmatrix}0\\\mu\end{bmatrix}, \begin{bmatrix}I&\mathbf{W}^\top\\\mathbf{W}&\mathbf{W}\mathbf{W}^\top + \Psi\end{bmatrix})$$

Use EM to solve the parameters.

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PCA and FA

Recap

- Introduced continuous latent variables and the most important applications: the principal component analysis, or PCA
- PCA is widely applied in dimensionality reduction, lossy data compression, feature extraction, and data visualization, and more.
- PCA can be defined as the orthogonal projection of the data onto a lower dimensional linear space, known as the principal subspace, such that the variance of the projected data is maximized, Or,
- It can be defined as the linear projection that minimizes the average project cost, defined as the mean squared distance between the data points and their projections.

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PCA and FA

Recap

- By maximize the projected variance, the data have reduced dimensionality M < D
- Where the eigenvector with the highest eigenvalue is the first principal component, and hence the second one having the second largest eigenvalue, and so on so forth
- These eigenvectors with highest values of eigenvalues are representation of the observed data X in D- dimension, the representatives have M- dimension
- Hence, PCA can reduce the data dimensionality, and compress the data with loss.
- Factor analysis is a linear-Gaussian latent variable model that is closed related to probabilistic PCA.

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