

Introduction to Machine Learning

Lecture 10 - Unsupervised Learning: EM in General Guang Bing Yang, PhD

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EM in General

- Introduction to EM in General
- EM for K-means algorithm
- General EM revisit Gaussian mixture
- EM for mixture of Bernoulli distributions

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The General EM Algorithm

- In general, the expectation maximization algorithm, or EM algorithm, is a technique for finding maximum likelihood solutions for probabilistic models having latent variables (Dempster et al., 1977; McLachlan and Krishnan, 1997).

- The goal is to maximize the likelihood function $p(X|\theta)$ with respect to θ .

$$p(X|\theta) = \sum_z p(X, Z|\theta) \text{ assume } Z \text{ is discrete}$$

- Direct optimize $p(X|\theta)$ is difficult, but
- optimize the complete-data likelihood function $p(X, Z|\theta)$ is easier

Reference: Dempster, A. P., N. M. Laird, and D. B. Rubin (1977). Maximum likelihood from incomplete data via the EM algorithm. Journal of the Royal Statistical Society, B 39(1), 1-38.
McLachlan, G. J. and T. Krishnan (1997). The EM Algorithm and its Extensions. Wiley.

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The General EM Algorithm

- Initialize parameters θ^{old}
- In the E-step: compute posterior: $p(Z|X, \theta^{old})$ w.r.t. the latent variable Z
- In the M-step: search the new estimate of parameters θ^{new} , given
$$Q(\theta, \theta^{old}) = \sum_Z p(Z|X, \theta^{old}) \ln p(X, Z|\theta)$$
- evaluate the convergence of either log-likelihood or the parameter values:
 - $\theta^{new} \leftarrow \theta^{old}$,
- iterate till converged or the difference of $\theta^{new} - \theta^{old} \leq \text{threshold}$, or over the loop limit

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The General EM Algorithm

- The expected complete data log likelihood, also called **auxiliary function**, can be described as:

$$Q(\theta, \theta^{old}) = \mathbb{E} \left[\sum_i \log p(x_i, z_i | \theta) \right] = \sum_i \mathbb{E} \left[\log \left[\prod_{k=1}^K (\pi_k p(x_i, z_i | \theta_k))^{I(z_i=k)} \right] \right]$$

$$= \sum_i \sum_k \mathbb{E}[I(z_i = k)] \log[\pi_k p(x_i, z_i | \theta_k)] = \sum_i \sum_k p(z_i = k | x_i, \theta^{old}) \log[\pi_k p(x_i, z_i | \theta_k)]$$

$$= \sum_i \sum_k r_{ik} \log \pi_k + \sum_i \sum_k r_{ik} \log p(x_i | \theta_k)$$

- where $r_{ik} = p(z_i = k | x_i, \theta^{old})$, the posterior or responsibility for cluster k takes for data x_i .

- E-step: $r_{ik} = \frac{\pi_k p(x_i | \theta_k^{old})}{\sum_{k'} \pi_{k'} p(x_i | \theta_{k'}^{old})}$,
- M-step: $\pi_k = \frac{1}{N} \sum_i r_{ik} = \frac{r_k}{N}$, where $r_k = \sum_i r_{ik}$
- $\mu_k = \frac{\sum_i r_{ik} x_i}{r_k}$, $\Sigma_k = \frac{\sum_i r_{ik} (x_i - \mu_k)(x_i - \mu_k)^T}{r_k} = \frac{\sum_i r_{ik} x_i x_i^T}{r_k} - \mu_k \mu_k^T$

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General EM: Variational Bound

- Given a joint distribution $p(X, Z | \theta)$ over observed variables X and hidden variables Z , the goal is to
- to maximize the likelihood function $p(X | \theta)$ with respect to θ .

- $p(X | \theta) = \sum_Z p(X, Z | \theta)$

- assume Z is discrete (change the summation to integral if Z is continuous, others are the same)
- For any distribution $q(Z)$, there is following variational lower bound:

- $\ln p(X | \theta) = \ln \sum_Z p(X, Z | \theta) = \ln \sum_Z q(Z) \frac{p(X, Z | \theta)}{q(Z)}$

- Logarithm is concave, so Jensen's inequality exists: so, $\ln p(X | \theta) \geq \sum_Z q(Z) \ln \frac{p(X, Z | \theta)}{q(Z)}$

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General EM: Variational Bound

$$\ln p(X | \theta) = \ln \sum_Z p(X, Z | \theta) = \ln \sum_Z q(Z) \frac{p(X, Z | \theta)}{q(Z)}$$

$$\geq \sum_Z q(Z) \ln \frac{p(X, Z | \theta)}{q(Z)}$$

- $= \sum_Z q(Z) \ln p(X, Z | \theta) + \sum_Z q(Z) \ln \frac{1}{q(Z)}$
- $= E_{q(Z)}[\ln p(X, Z | \theta)] + \mathbb{H}(q(Z)) = \mathbb{L}(q, \theta)$

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General EM: Variational Bound

- There are two components in the log likelihood function:
- $\ln p(X | \theta) \geq E_{q(Z)}[\ln p(X, Z | \theta)] + \mathbb{H}(q(Z)) = \mathbb{L}(q, \theta)$
- The first part is the Expected complete log-likelihood, the second part is the Entropy of the distribution $q(Z)$.
- $\mathbb{L}(q, \theta)$ is the variational lower-bound.
- For a discrete random variable Z , the entropy is defined as:

- $\mathbb{H}(p) = - \sum_i p(z_i) \log p(z_i)$, or $\mathbb{H}(p) = - \int p(z) \log p(z) dz$ for continuous random variables

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General EM: Variational Bound

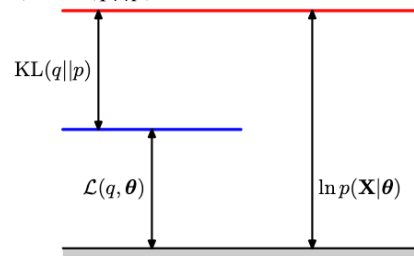
- Having:
 - $\ln p(X|\theta) \geq E_{q(Z)}[\ln p(X, Z|\theta)] + \mathbb{H}(q(Z)) = \mathbb{L}(q, \theta)$, and
 - $\ln p(X|\theta) = \mathbb{L}(q, \theta) + \text{KL}(q||p)$, where
 - $\mathbb{L}(q, \theta) = \sum_Z q(Z) \ln \frac{p(X, Z|\theta)}{q(Z)}$, the lower bound
 - $\text{KL}(q||p) = - \sum_Z q(Z) \ln \frac{p(Z|X, \theta)}{q(Z)}$, a relative entropy
- Since $\ln p(X, Z|\theta) = \ln p(Z|X, \theta) + \ln p(X|\theta)$, and substitute the $\mathbb{L}(q, \theta)$

General EM: Variational Bound

- Note that variational bound becomes tight iff $q(Z) = p(Z|X, \theta)$.
- In other words the distribution $q(Z)$ is equal to the true posterior distribution over the latent variables, so that $\text{KL}(q||p) = 0$.
- As $\text{KL}(q||p) = 0$, it immediately follows that:
 - $\ln p(X|\theta) \geq \mathbb{L}(q, \theta)$,
 - which is also showed using Jensen's inequality,

General EM: Decomposition of $q(Z)$

- To illustrate the decomposition of the distribution $q(Z)$:
- $\ln p(X|\theta) = \mathbb{L}(q, \theta) + \text{KL}(q||p)$

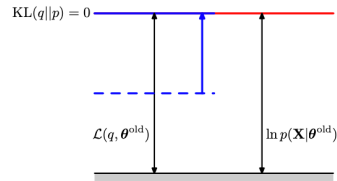


General EM: Summary

- uses the decomposition to define the EM algorithm, and
- shows that it maximizes the log-likelihood function:
 - $\ln p(X|\theta) = \mathbb{L}(q, \theta) + \text{KL}(q||p)$
- In the E-step, the lower bound $\mathbb{L}(q, \theta)$ is maximized w.r.t. the distribution q with fixed parameters θ
- In the M-step, the lower bound $\mathbb{L}(q, \theta)$ is maximized w.r.t. the parameters θ with the distribution q fixed.
- These steps increase the corresponding log-likelihood

General EM: E-step

- Let the current value of parameters as θ^{old}
- In the E-step, maximize the lower bound w.r.t. q with θ^{old} fixed:
 - $\mathbb{L}(q, \theta^{old}) = \ln p(X|\theta^{old}) - \text{KL}(q||p)$
 - θ^{old} does not depend on q
- the lower bound $\mathbb{L}(q, \theta)$ is maximized when $\text{KL} = 0$
- In other words, $q(Z) = p(Z|X, \theta^{old})$
- The lower bound becomes equal to the log-likelihood



The E-step of the EM algorithm.

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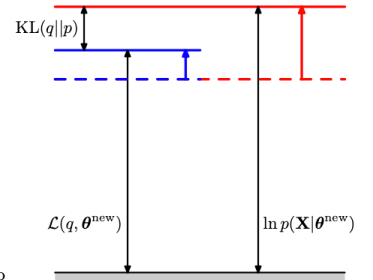
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General EM: M-step

- the lower bound is maximized w.r.t. parameters θ with q fixed.
- In the E-step, maximize the lower bound w.r.t. q with θ^{old} fixed:

$$\mathbb{L}(q, \theta) = \sum_Z p(Z|X, \theta^{old}) \ln p(X, Z|\theta)$$
 - $+ \sum_Z p(Z|X, \theta^{old}) \ln \frac{1}{p(Z|X, \theta^{old})}$
 - $\mathbb{L}(q, \theta) = Q(\theta, \theta^{old}) + \text{const}$
 - the last part of the $\mathbb{L}(q, \theta)$ does not depend on q
- Hence the M-step maximizes the expected complete log-likelihood
 - $\theta^{new} = \arg\max_{\theta} Q(\theta, \theta^{old})$
- Because the KL is non-negative, this makes the log-likelihood $p(X|\theta)$ to increase by at least as much as the lower bound does.



The M-step of EM algorithm

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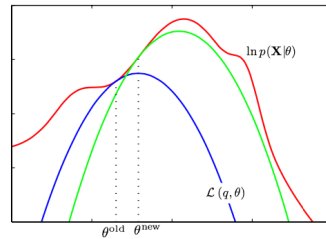
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General EM: Bound Optimization

- The EM algorithm belongs to the general class of bound optimization methods:
- At each step, we compute:
 - E-step: a lower bound on the log-likelihood function for the current parameter values. The bound is concave with unique global optimum.
 - M-step: maximize the lower-bound to obtain the new parameter values.



The EM algorithm maximizes this bound to obtain the new parameter values

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General EM: Extensions

- For some complex cases, either the E-step or the M-step or both remain intractable
- Two possible approaches:
 - The generalized EM (GEM) — deals with the intractable in the M-step
 - generalized the E-step by performing a partial optimization of the lower-bound w.r.t. q
- In GEM, using nonlinear optimization, conjugate gradient, etc to change parameters so as to increase its value.
- use an incremental form of EM, in which at each EM step only one data point is processed at a time
- In the E-step, instead of recomputing the responsibilities for all the data points, we just re-evaluate the responsibilities for one data point, and proceed with the M-step

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Maximizing the Posterior using EM

- There is a way to use EM to maximize the posterior $p(\theta|X)$ for models having the prior defined as $p(\theta)$
- Because: $\ln p(\theta|X) = \ln p(X|\theta) + \ln p(\theta) - \ln p(X)$
- Decomposing the log-likelihood into lower-bound and KL terms:
 - $\ln p(X|\theta) = \mathbb{L}(q, \theta) + \text{KL}(q||p)$
- $\ln p(\theta|X) = \mathbb{L}(q, \theta) + \text{KL}(q||p) + \ln p(\theta) - \ln p(X)$
 - where $\ln p(X)$ is a constant.
- The E-step is the same as for the standard EM algorithm
- The M-step equations are modified through introduction of the prior term, which typically amounts to only a small modification to the standard ML M-step equations.

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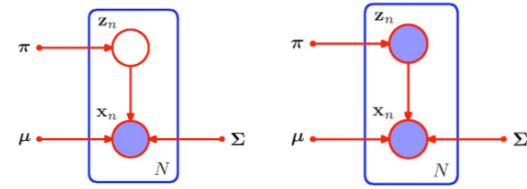
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General EM: Gaussian Mixture Revisited

- Recall the maximize likelihood of the Gaussian mixture is given as:

$$\ln p(X|\pi, \mu, \Sigma) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k N(x_n | \mu_k, \Sigma_k) \right\}$$



$\{X\}$ -- incomplete dataset. $\{X, Z\}$ -- complete dataset.

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General EM: Gaussian Mixture Revisited

- Use complete-data (log-)likelihood, and expectation given as:
- $p(X, Z|\pi, \mu, \Sigma) = \prod_{n=1}^N \prod_{k=1}^K [\pi_k N(x_n | \mu_k, \Sigma_k)]^{z_{nk}}$, taking the logarithm, obtain:
- $\ln p(X, Z|\pi, \mu, \Sigma) = \sum_{k=1}^K \left[\sum_{n=1}^N z_{nk} \{ \ln \pi_k + \ln N(x_n | \mu_k, \Sigma_k) \} \right]$
- Maximizing w.r.t. mixing proportions given: $\pi_k = \frac{1}{N} \sum_{n=1}^N z_{nk}$
- Similarly for the means and covariances.

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General EM: K-means Revisited

- Consider a Gaussian mixture model in which covariances are shared and are given by ϵ .
- $p(x|\mu_k, \Sigma_k) = \frac{1}{(2\pi\epsilon)^{D/2}} \exp[-\frac{1}{2\epsilon} ||x - \mu_k||^2]$
- Consider EM algorithm for a mixture of K Gaussians, in which let ϵ as a fixed constant. The posterior responsibilities take form:
 - $\gamma(z_{nk}) = \frac{\pi_k \exp(-||x_n - \mu_k||^2/2\epsilon)}{\sum_{j=1}^K \pi_j \exp(-||x_n - \mu_j||^2/2\epsilon)}$
- Consider the limit $\epsilon \rightarrow 0$
- $\gamma(z_{nk}) \rightarrow r_{nk}$, while $r_{nk} = 1$ if $k = \text{argmax}_j ||x_n - \mu_j||^2$ otherwiser $r_{nk} = 0$.

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Mixture of Bernoulli Distributions

- Let's look at mixture of discrete binary variables described by Bernoulli distributions.
- Consider a set of binary random variables $x_i, i=1, \dots, D$, each of which is governed by a Bernoulli distribution with μ_i :

$$p(x|\mu) = \prod_{i=1}^D \mu_i^{x_i} (1 - \mu_i)^{1-x_i}$$

- The mean and covariance of this distribution are:

$$\mathbb{E}[x] = \mu, \text{cov}[x] = \text{diag}(\mu_i(1 - \mu_i))$$

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Mixture of Bernoulli Distributions

- Given a finite (K) mixture of Bernoulli distributions:

$$p(x|\pi, \mu) = \sum_{k=1}^K \pi_k p(x|\mu_k)$$

$$p(x|\mu_k) = \prod_{i=1}^D \mu_{ki}^{x_i} (1 - \mu_{ki})^{1-x_i}$$

- The mean and covariance of this mixture distribution are:

$$\mathbb{E}[x] = \sum_{k=1}^K \pi_k \mu_k, \text{ and } \text{cov}[x] = \sum_{k=1}^K \pi_k (\Sigma_k + \mu_k \mu_k^T) - \mathbb{E}[x] \mathbb{E}[x]^T,$$

- where $\Sigma_k = \text{diag}(\mu_{ki}(1 - \mu_{ki}))$

- The covariance matrix is no longer diagonal, so the mixture distribution can capture correlations between the variables, unlike a single Bernoulli distribution.

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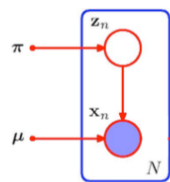
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Mixture of Bernoulli Distributions: Maximum Likelihood

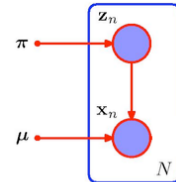
- Given a dataset X , the log-likelihood:

$$\ln p(x|\pi, \mu) = \sum_{n=1}^N \ln \left[\sum_{k=1}^K \pi_k p(x|\mu_k) \right]$$

- this is intractable, need to apply EM algorithm for maximizing this log-likelihood function



$\{X\}$ -- incomplete dataset.



$\{X, Z\}$ -- complete dataset.

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Mixture of Bernoulli Distributions: Maximum Likelihood

- Consider the complete log-likelihood:

$$p(z|\pi) = \prod_{k=1}^K \pi_k^{z_k}, \text{ and } p(x|z, \mu) = \prod_{k=1}^K p(x|\mu_k)^{z_k}$$

- the complete log-likelihood given as:

$$\ln p(X, Z|\pi, \mu) = \sum_{i=1}^N \sum_{k=1}^K z_{nk} [\ln \pi_k + \sum_{i=1}^D [x_{ni} \ln \mu_{ki} + (1 - x_{ni}) \ln (1 - \mu_{ki})]]$$

- The expected complete-data log-likelihood:

$$\mathbb{E}_Z[\ln p(X, Z|\pi, \mu)] = \sum_{i=1}^N \sum_{k=1}^K \gamma(z_{nk}) [\ln \pi_k + \sum_{i=1}^D [x_{ni} \ln \mu_{ki} + (1 - x_{ni}) \ln (1 - \mu_{ki})]],$$

- where $\mathbb{E}[z_{nk}] = \gamma(z_{nk})$

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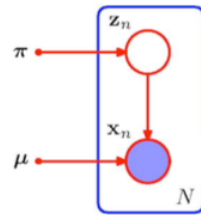
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Mixture of Bernoulli Distributions: E-step

- Similar to the mixture of Gaussians, in the E-step, using Bayes' rule to evaluate responsibilities:

$$\mathbb{E}[z_{nk}] = \frac{\sum_{Z_n} z_{nk} \prod_k [\pi_k p(x_n | \mu_k)]^{z_{nk}}}{\sum_{Z_n} \prod_j [\pi_j p(x_n | \mu_j)]^{z_{nj}}}$$

$$= \frac{\pi_k p(x_n | \mu_k)}{\sum_{j=1}^K \pi_j p(x_n | \mu_j)} = \gamma(z_{nk})$$



Mixture of Bernoulli Distributions: M-step

- The expected complete-data log-likelihood:

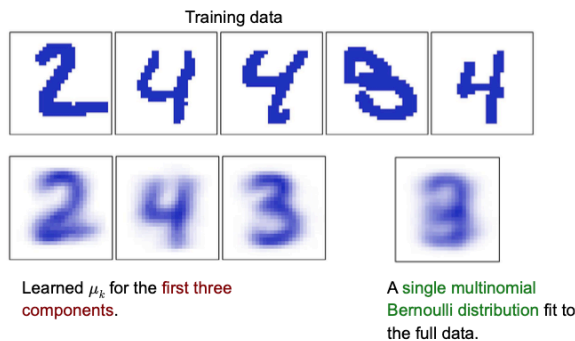
$$\mathbb{E}_Z[\ln p(X, Z | \pi, \mu)] = \sum_{i=1}^N \sum_{k=1}^K \gamma(z_{nk}) [\ln \pi_k + \sum_{i=1}^D [x_{ni} \ln \mu_{ki} + (1 - x_{ni}) \ln (1 - \mu_{ki})]]$$

- Maximizing the expected complete-data log-likelihood:

$$\mu_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) x_n, \pi_k = \frac{N_k}{N}, \text{ and } N_k = \sum_{n=1}^N \gamma(z_{nk}),$$

- where N_k is the effective number of data points associated with component k .
- Note that the mean of component k is equal to the weighted mean of the data, with weights given by the responsibilities that component k takes for explaining the data points.

Mixture of Bernoulli Distributions: Example



Recap

- The general EM algorithm, is a technique for finding maximum likelihood solutions for probabilistic models having latent variables (Dempster et al., 1977; McLachlan and Krishnan, 1977)
- The goal is to maximize the likelihood function $p(X | \theta)$ with respect to θ .
- But easy to use complete data, complete log-likelihood by given a joint distribution $p(X, Z | \theta)$ over observed variables X and hidden variables Z
- By decompose the distribution $q(Z)$, applying for the EM to maximize the log-likelihood over q and θ in a two-stage process.
- Iterate the E-step and M-step many times until converged, then the parameters are optimized values.

Questions?

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