

Homework4

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```
# import data
df = haven::read_dta("./data/umaru.dta")
```

1. Parametric/Accelerated Failure Time Models

a.

Variables: age, nonwhite, treat, site, ivdrug

Models: Exponential, Weibull, Log-logistic, Log-normal, Generalized Gamma

```
# define the survival object
surv_object <- Surv(time = df$time, event = df$censor)

# fit AFT models with different distributions
aft_exponential <-
  flexsurvreg(surv_object ~ age + nonwhite + treat + site + ivdrug, data = df, dist = "exponential")
aft_weibull <-
  flexsurvreg(surv_object ~ age + nonwhite + treat + site + ivdrug, data = df, dist = "weibull")
aft_llogis <-
  flexsurvreg(surv_object ~ age + nonwhite + treat + site + ivdrug, data = df, dist = "llogis")
aft_lognormal <-
  flexsurvreg(surv_object ~ age + nonwhite + treat + site + ivdrug, data = df, dist = "lognormal")
aft_gen_gamma <-
  flexsurvreg(surv_object ~ age + nonwhite + treat + site + ivdrug, data = df, dist = "gengamma")

# check the results
# aft_exponential
# aft_weibull
# aft_llogis
# aft_lognormal
# aft_gen_gamma
```

i. Values of the $-2 \log L$, the total number of parameters (including shape and scale for ϵ) and the AIC for each of these models

```
# extract the log-likelihood, total parameters, and AIC
results <- data.frame(
  Model = c("Exponential", "Weibull", "Log-logistic", "Log-normal", "Generalized Gamma"),
  "-2LogL" = c(
    -2 * aft_exponential$loglik,
```

```

-2 * aft_weibull$loglik,
-2 * aft_llogis$loglik,
-2 * aft_lognormal$loglik,
-2 * aft_gen_gamma$loglik
),
"Total Parameters" = c(
  aft_exponential$npars,
  aft_weibull$npars,
  aft_llogis$npars,
  aft_lognormal$npars,
  aft_gen_gamma$npars
),
AIC = c(
  aft_exponential$AIC,
  aft_weibull$AIC,
  aft_llogis$AIC,
  aft_lognormal$AIC,
  aft_gen_gamma$AIC
)
)

colnames(results) <- c("Model", "-2 Log L", "Total Parameters", "AIC")

results |>
  kable()

```

| Model | -2 Log L | Total Parameters | AIC |
|-------------------|----------|------------------|----------|
| Exponential | 6180.608 | 6 | 6192.608 |
| Weibull | 6179.805 | 7 | 6193.805 |
| Log-logistic | 6127.191 | 7 | 6141.191 |
| Log-normal | 6137.951 | 7 | 6151.951 |
| Generalized Gamma | 6137.219 | 8 | 6153.219 |

The Log-logistic model has the lowest AIC (6141.191), making it the best-fitting model based on AIC.

ii.

Given the AIC, Weibull model does not provide an improved fit compared to the exponential model. I will confirm this using likelihood ratio test.

```

# extract log-likelihoods
logL_exp <- aft_exponential$loglik
logL_weib <- aft_weibull$loglik

# compute LRT statistic
lrt_stat <- -2 * (logL_exp - logL_weib)

# df
degf <- aft_weibull$npars - aft_exponential$npars

# p-value

```

```

p_value <- pchisq(lrt_stat, df = degf, lower.tail = FALSE)

# output results
cat("LRT Statistic:", lrt_stat)
## LRT Statistic: 0.8022232
cat("Degrees of Freedom:", degf)
## Degrees of Freedom: 1
cat("p-value:", p_value)
## p-value: 0.3704295

```

Given $p\text{-value} > 0.05$, we fail to reject the null hypothesis and conclude that Weibull model does not improve fit compared to the exponential model.

iii.

Exponential model and Weibull model are nested within the generalized gamma model.
Generalized gamma model vs exponential model

```

# extract log-likelihoods
logL_exp <- aft_exponential$loglik
logL_ggamma <- aft_gen_gamma$loglik

# compute LRT statistic
lrt_stat <- -2 * (logL_exp - logL_ggamma)

# df
degf <- aft_gen_gamma$npars - aft_exponential$npars

# p-value
p_value <- pchisq(lrt_stat, df = degf, lower.tail = FALSE)

# output results
cat("LRT Statistic:", lrt_stat)
## LRT Statistic: 43.38879
cat("Degrees of Freedom:", degf)
## Degrees of Freedom: 2
cat("p-value:", p_value)
## p-value: 3.786546e-10

```

Given $p\text{-value} < 0.05$, we reject the null hypothesis and conclude that generalized gamma model provides a better fit compared to the exponential model.

Generalized gamma model vs Weibull model

```

# extract log-likelihoods
logL_exp <- aft_exponential$loglik
logL_weib <- aft_weibull$loglik

# compute LRT statistic
lrt_stat <- -2 * (logL_weib - logL_ggamma)

# df

```

```

degf <- aft_gen_gamma$npars - aft_weibull$npars

# p-value
p_value <- pchisq(lrt_stat, df = degf, lower.tail = FALSE)

# output results
cat("LRT Statistic:", lrt_stat)
## LRT Statistic: 42.58657
cat("Degrees of Freedom:", degf)
## Degrees of Freedom: 1
cat("p-value:", p_value)
## p-value: 6.762209e-11

```

Given $p\text{-value} < 0.05$, we reject the null hypothesis and conclude that generalized gamma model provides a better fit compared to the Weibull model.

b.

Time ratio: $\phi = e^\beta$

```

# function to compute time ratio and confidence interval
tr_ci <- function(model) {
  # extract coefficient and standard error for ivdrug
  beta <- model$res["ivdrug", "est"]
  se <- model$res["ivdrug", "se"]

  # compute time ratio (phi) and 95% confidence interval
  phi <- exp(beta)
  ci_lower <- exp(beta - 1.96 * se)
  ci_upper <- exp(beta + 1.96 * se)

  # return as a named vector
  c("Time Ratio (phi)" = phi,
    "95% CI Lower" = ci_lower,
    "95% CI Upper" = ci_upper)
}

# apply the function to all models
results <- data.frame(
  Model = c("Exponential", "Weibull", "Log-logistic", "Log-normal", "Generalized Gamma"),
  rbind(
    tr_ci(aft_exponential),
    tr_ci(aft_weibull),
    tr_ci(aft_llogis),
    tr_ci(aft_lognormal),
    tr_ci(aft_gen_gamma)
  )
)

colnames(results) = c("Model", "Time Ratio (phi)", "95% CI Lower", "95% CI Upper")

results |>
  kable()

```

| Model | Time Ratio (ϕ) | 95% CI Lower | 95% CI Upper |
|-------------------|-----------------------|--------------|--------------|
| Exponential | 1.5210016 | 1.2386592 | 1.8677018 |
| Weibull | 0.6527657 | 0.5274755 | 0.8078157 |
| Log-logistic | 0.6610930 | 0.5254482 | 0.8317545 |
| Log-normal | 0.6547605 | 0.5161506 | 0.8305934 |
| Generalized Gamma | 0.6520866 | 0.5154833 | 0.8248900 |

In the Weibull model, the estimated ϕ for the `ivdrug` covariate is 0.653 (95%CI: 0.527-0.808). On average, individuals with IV drug use have survival times that are about 65.3% of those without IV drug use, after adjusting for other covariates. This association is statistically significant.

c.

Exponential model

$$\beta_{HR} = \beta_{AFT}$$

$$HR = e^{\beta_{HR}} = e^{\beta_{AFT}}$$

$$95\%CI = e^{\beta_{HR} \pm 1.96 \times SE_{HR}}$$

```
# extract AFT beta and SE for IV drug use (exponential model)
beta_aft_exp <- aft_exponential$res["ivdrug", "est"]
se_aft_exp <- aft_exponential$res["ivdrug", "se"]

# log-hazard ratio for exponential
beta_hr_exp <- beta_aft_exp

# HR and 95% CI
hr_exp <- exp(beta_hr_exp)
hr_exp_ci <- exp(c(beta_hr_exp - 1.96 * se_aft_exp, beta_hr_exp + 1.96 * se_aft_exp))
```

Weibull model

$$\beta_{HR} = -\beta_{AFT}/\alpha, \text{ where } \alpha = \text{shape parameter}$$

$$HR = e^{\beta_{HR}} = e^{\beta_{AFT}/\alpha}$$

$$95\%CI = e^{\beta_{HR} \pm 1.96 \times SE_{HR}}$$

```
# extract AFT beta and SE for IV drug use (Weibull model)
beta_aft_weib <- aft_weibull$res["ivdrug", "est"]
se_aft_weib <- aft_weibull$res["ivdrug", "se"]

# extract Weibull shape parameter
alpha <- aft_weibull$res["shape", "est"]

# log-hazard ratio for Weibull
beta_hr_weib <- -beta_aft_weib / alpha
se_hr_weib <- se_aft_weib / alpha

# HR and 95% CI
hr_weib <- exp(beta_hr_weib)
hr_weib_ci <- exp(c(beta_hr_weib - 1.96 * se_hr_weib, beta_hr_weib + 1.96 * se_hr_weib))
```

Results:

```

# summarize results
results <- data.frame(
  Model = c("Exponential", "Weibull"),
  "Log-HR" = c(beta_hr_exp, beta_hr_weib),
  "HR" = c(hr_exp, hr_weib),
  "95% CI Lower" = c(hr_exp_ci[1], hr_weib_ci[1]),
  "95% CI Upper" = c(hr_exp_ci[2], hr_weib_ci[2])
)

colnames(results) = c("Model", "Log-HR", "HR", "95% CI Lower", "95% CI Upper")

results |>
  kable()

```

| Model | Log-HR | HR | 95% CI Lower | 95% CI Upper |
|-------------|-----------|----------|--------------|--------------|
| Exponential | 0.4193691 | 1.521002 | 1.238659 | 1.867702 |
| Weibull | 0.4412566 | 1.554660 | 1.247057 | 1.938136 |

In the exponential model, the HR for IV drug use is 1.52 (95% CI: 1.24–1.87), indicating that individuals with IV drug use have a 52% higher hazard of the event compared to those without IV drug use, holding other variables constant. Similarly, in the Weibull model, the HR is 1.55 (95% CI: 1.25–1.94), showing consistent evidence of increased hazard associated with IV drug use.

d. Cox PH Model

Covariates: age, nonwhite, treat, site, ivdrug

```

# fit cox ph model with five covariates
cox_model = coxph(surv_object ~ age + nonwhite + treat + site + ivdrug, data = df)

# check the cox ph model
cox_model
## Call:
## coxph(formula = surv_object ~ age + nonwhite + treat + site +
##       ivdrug, data = df)
##
##              coef exp(coef)  se(coef)      z      p
## age          -0.024241  0.976051  0.008001 -3.030 0.002449
## nonwhite     -0.239605  0.786939  0.114377 -2.095 0.036182
## treat        -0.229902  0.794612  0.093462 -2.460 0.013900
## site         -0.148612  0.861903  0.104461 -1.423 0.154836
## ivdrug        0.365906  1.441820  0.105511  3.468 0.000524
##
## Likelihood ratio test=32.65  on 5 df, p=4.423e-06
## n= 575, number of events= 464

# extract HR and 95% CI
cox_hr <- exp(coef(cox_model)["ivdrug"])
cox_ci <- exp(confint(cox_model)["ivdrug", ])

# combine results for comparison

```

```

results_ivdrug <- data.frame(
  Covariate = "IV drug use",
  Cox_HR = cox_hr,
  Cox_CIL = cox_ci[1],
  Cox_CIU = cox_ci[2],
  Exp_HR = results[1, 3],
  Exp_CIL = results[1, 4],
  Exp_CIU = results[1, 5],
  Wei_HR = results[2, 3],
  Wei_CIL = results[2, 4],
  Wei_CIU = results[2, 5]
)

# iterate for other variables
# age
# extract AFT beta and SE for age (exponential model)
beta_aft_exp <- aft_exponential$res["age", "est"]
se_aft_exp <- aft_exponential$res["age", "se"]

# log-hazard ratio for exponential
beta_hr_exp <- beta_aft_exp

# HR and 95% CI
hr_exp <- exp(beta_hr_exp)
hr_exp_ci <- exp(c(beta_hr_exp - 1.96 * se_aft_exp, beta_hr_exp + 1.96 * se_aft_exp))

# extract AFT beta and SE for age (Weibull model)
beta_aft_weib <- aft_weibull$res["age", "est"]
se_aft_weib <- aft_weibull$res["age", "se"]

# extract Weibull shape parameter
alpha <- aft_weibull$res["shape", "est"]

# log-hazard ratio for Weibull
beta_hr_weib <- -beta_aft_weib / alpha
se_hr_weib <- se_aft_weib / alpha

# HR and 95% CI
hr_weib <- exp(beta_hr_weib)
hr_weib_ci <- exp(c(beta_hr_weib - 1.96 * se_hr_weib, beta_hr_weib + 1.96 * se_hr_weib))

results <- data.frame(
  Model = c("Exponential", "Weibull"),
  "Log-HR" = c(beta_hr_exp, beta_hr_weib),
  "HR" = c(hr_exp, hr_weib),
  "95% CI Lower" = c(hr_exp_ci[1], hr_weib_ci[1]),
  "95% CI Upper" = c(hr_exp_ci[2], hr_weib_ci[2])
)

# extract HR and 95% CI
cox_hr <- exp(coef(cox_model)["age"])
cox_ci <- exp(confint(cox_model)["age", ])

```

```

# combine results for comparison
results_age <- data.frame(
  Covariate = "age",
  Cox_HR = cox_hr,
  Cox_CIL = cox_ci[1],
  Cox_CIU = cox_ci[2],
  Exp_HR = results[1, 3],
  Exp_CIL = results[1, 4],
  Exp_CIU = results[1, 5],
  Wei_HR = results[2, 3],
  Wei_CIL = results[2, 4],
  Wei_CIU = results[2, 5]
)

# nonwhite race
# extract AFT beta and SE for nonwhite race (exponential model)
beta_aft_exp <- aft_exponential$res["nonwhite", "est"]
se_aft_exp <- aft_exponential$res["nonwhite", "se"]

# log-hazard ratio for exponential
beta_hr_exp <- beta_aft_exp

# HR and 95% CI
hr_exp <- exp(beta_hr_exp)
hr_exp_ci <- exp(c(beta_hr_exp - 1.96 * se_aft_exp, beta_hr_exp + 1.96 * se_aft_exp))

# extract AFT beta and SE for nonwhite race (Weibull model)
beta_aft_weib <- aft_weibull$res["nonwhite", "est"]
se_aft_weib <- aft_weibull$res["nonwhite", "se"]

# extract Weibull shape parameter
alpha <- aft_weibull$res["shape", "est"]

# log-hazard ratio for Weibull
beta_hr_weib <- -beta_aft_weib / alpha
se_hr_weib <- se_aft_weib / alpha

# HR and 95% CI
hr_weib <- exp(beta_hr_weib)
hr_weib_ci <- exp(c(beta_hr_weib - 1.96 * se_hr_weib, beta_hr_weib + 1.96 * se_hr_weib))

results <- data.frame(
  Model = c("Exponential", "Weibull"),
  "Log-HR" = c(beta_hr_exp, beta_hr_weib),
  "HR" = c(hr_exp, hr_weib),
  "95% CI Lower" = c(hr_exp_ci[1], hr_weib_ci[1]),
  "95% CI Upper" = c(hr_exp_ci[2], hr_weib_ci[2])
)

# extract HR and 95% CI
cox_hr <- exp(coef(cox_model)["nonwhite"])
cox_ci <- exp(confint(cox_model)["nonwhite", ])

```



```

# combine results for comparison
results_nonwhite <- data.frame(
  Covariate = "nonwhite",
  Cox_HR = cox_hr,
  Cox_CIL = cox_ci[1],
  Cox_CIU = cox_ci[2],
  Exp_HR = results[1, 3],
  Exp_CIL = results[1, 4],
  Exp_CIU = results[1, 5],
  Wei_HR = results[2, 3],
  Wei_CIL = results[2, 4],
  Wei_CIU = results[2, 5]
)

# treatment
# extract AFT beta and SE for treatment (exponential model)
beta_aft_exp <- aft_exponential$res["treat", "est"]
se_aft_exp <- aft_exponential$res["treat", "se"]

# log-hazard ratio for exponential
beta_hr_exp <- beta_aft_exp

# HR and 95% CI
hr_exp <- exp(beta_hr_exp)
hr_exp_ci <- exp(c(beta_hr_exp - 1.96 * se_aft_exp, beta_hr_exp + 1.96 * se_aft_exp))

# extract AFT beta and SE for treatment (Weibull model)
beta_aft_weib <- aft_weibull$res["treat", "est"]
se_aft_weib <- aft_weibull$res["treat", "se"]

# extract Weibull shape parameter
alpha <- aft_weibull$res["shape", "est"]

# log-hazard ratio for Weibull
beta_hr_weib <- -beta_aft_weib / alpha
se_hr_weib <- se_aft_weib / alpha

# HR and 95% CI
hr_weib <- exp(beta_hr_weib)
hr_weib_ci <- exp(c(beta_hr_weib - 1.96 * se_hr_weib, beta_hr_weib + 1.96 * se_hr_weib))

results <- data.frame(
  Model = c("Exponential", "Weibull"),
  "Log-HR" = c(beta_hr_exp, beta_hr_weib),
  "HR" = c(hr_exp, hr_weib),
  "95% CI Lower" = c(hr_exp_ci[1], hr_weib_ci[1]),
  "95% CI Upper" = c(hr_exp_ci[2], hr_weib_ci[2])
)

# extract HR and 95% CI
cox_hr <- exp(coef(cox_model)["treat"])
cox_ci <- exp(confint(cox_model)["treat", ])

```

```

# combine results for comparison
results_treat <- data.frame(
  Covariate = "treat",
  Cox_HR = cox_hr,
  Cox_CIL = cox_ci[1],
  Cox_CIU = cox_ci[2],
  Exp_HR = results[1, 3],
  Exp_CIL = results[1, 4],
  Exp_CIU = results[1, 5],
  Wei_HR = results[2, 3],
  Wei_CIL = results[2, 4],
  Wei_CIU = results[2, 5]
)

# site
# extract AFT beta and SE for site (exponential model)
beta_aft_exp <- aft_exponential$res["site", "est"]
se_aft_exp <- aft_exponential$res["site", "se"]

# log-hazard ratio for exponential
beta_hr_exp <- beta_aft_exp

# HR and 95% CI
hr_exp <- exp(beta_hr_exp)
hr_exp_ci <- exp(c(beta_hr_exp - 1.96 * se_aft_exp, beta_hr_exp + 1.96 * se_aft_exp))

# extract AFT beta and SE for site (Weibull model)
beta_aft_weib <- aft_weibull$res["site", "est"]
se_aft_weib <- aft_weibull$res["site", "se"]

# extract Weibull shape parameter
alpha <- aft_weibull$res["shape", "est"]

# log-hazard ratio for Weibull
beta_hr_weib <- -beta_aft_weib / alpha
se_hr_weib <- se_aft_weib / alpha

# HR and 95% CI
hr_weib <- exp(beta_hr_weib)
hr_weib_ci <- exp(c(beta_hr_weib - 1.96 * se_hr_weib, beta_hr_weib + 1.96 * se_hr_weib))

results <- data.frame(
  Model = c("Exponential", "Weibull"),
  "Log-HR" = c(beta_hr_exp, beta_hr_weib),
  "HR" = c(hr_exp, hr_weib),
  "95% CI Lower" = c(hr_exp_ci[1], hr_weib_ci[1]),
  "95% CI Upper" = c(hr_exp_ci[2], hr_weib_ci[2])
)

# extract HR and 95% CI
cox_hr <- exp(coef(cox_model)["site"])
cox_ci <- exp(confint(cox_model)["site", ])

```

```

# combine results for comparison
results_site <- data.frame(
  Covariate = "site",
  Cox_HR = cox_hr,
  Cox_CIL = cox_ci[1],
  Cox_CIU = cox_ci[2],
  Exp_HR = results[1, 3],
  Exp_CIL = results[1, 4],
  Exp_CIU = results[1, 5],
  Wei_HR = results[2, 3],
  Wei_CIL = results[2, 4],
  Wei_CIU = results[2, 5]
)

# data summary
combined <- rbind(
  results_ivdrug,
  results_age,
  results_nonwhite,
  results_treat,
  results_site
) |>
mutate(
  Cox_CI = paste0("(", round(Cox_CIL, 2), ", ", round(Cox_CIU, 2), ")"),
  Exp_CI = paste0("(", round(Exp_CIL, 2), ", ", round(Exp_CIU, 2), ")"),
  Wei_CI = paste0("(", round(Wei_CIL, 2), ", ", round(Wei_CIU, 2), ")")
) |>
select(Cox_HR, Cox_CI, Exp_HR, Exp_CI, Wei_HR, Wei_CI)

combined |>
kable(digits = 2)

```

| | Cox_HR | Cox_CI | Exp_HR | Exp_CI | Wei_HR | Wei_CI |
|----------|--------|--------------|--------|--------------|--------|--------------|
| ivdrug | 1.44 | (1.17, 1.77) | 1.52 | (1.24, 1.87) | 1.55 | (1.25, 1.94) |
| age | 0.98 | (0.96, 0.99) | 0.98 | (0.96, 0.99) | 0.97 | (0.96, 0.99) |
| nonwhite | 0.79 | (0.63, 0.98) | 0.78 | (0.62, 0.97) | 0.77 | (0.6, 0.98) |
| treat | 0.79 | (0.66, 0.95) | 0.79 | (0.66, 0.95) | 0.78 | (0.64, 0.95) |
| site | 0.86 | (0.7, 1.06) | 0.88 | (0.71, 1.08) | 0.87 | (0.7, 1.08) |

Comparison of the HR for ivdrug in the three models:

- Cox Model
The HR for ivdrug is 1.44 (95%CI: 1.17, 1.77). This indicates that individuals with IV drug use have a 44% higher hazard of relapse compared to those without IV drug use, adjusting for other covariates. This is statistically significant.
- Exponential Model
The HR for ivdrug is 1.52 (95%CI: 1.24, 1.87). This suggests that individuals with IV drug use have a 53% higher hazard of relapse than those without IV drug use, adjusting for other covariates. This is statistically significant and the direction of HR is similar to that of Cox PH model.
- Weibull Model
The HR for ivdrug 1.55 (95%CI: 1.25, 1.94). This indicates a 55% higher hazard of relapse for

individuals with IV drug use, adjusting for other covariates. This is statistically significant and the direction of HR is similar to that of Cox PH model.

Overall, three models showed a consistent direction of the hazard ratio, with individuals with IV drug use having a higher hazard of relapse compared to those without IV drug use.

2. Hazard Rates and Survival from Parametric Models

a. Exponential Model

Under the exponential model, $\hat{\lambda} = \frac{\text{Number of events}}{\text{Total follow-up time}}$.

Therefore,

$$\hat{\lambda}_0 = \frac{165}{277.031 \times 223} \approx 0.00267$$

$$\hat{\lambda}_1 = \frac{299}{219.097 \times 352} \approx 0.00388$$

Based on the AFT model, $\ln(\hat{\lambda}_1) = \ln(\hat{\lambda}_0) - \beta_{AFT} \Rightarrow \hat{\lambda}_1 = \hat{\lambda}_0 \times e^{-\beta_{AFT}}$

Given the SAS output, we obtain:

$$\hat{\lambda}_0 = e^{-5.9254} \approx 0.00267$$

$$\hat{\lambda}_1 = 0.00267 \times e^{0.3727} \approx 0.00388$$

Both methods yield the same result.

b.

In the exponential model, $S(t) = e^{-\lambda t}$.

Therefore, the predicted probability of no relapse by 6 months (182 days) for each IV drug use group are:

$$S_0(182) = e^{-0.00267 \times 182} \approx 0.615$$

$$S_1(182) = e^{-0.00388 \times 182} \approx 0.494$$

Now, let t^* be the time where $S_1(t^*) = S_0(182)$.

$$e^{-\lambda_1 \times t^*} = e^{-\lambda_0 \times 182} \Rightarrow t^* = \frac{\lambda_0 \times 182}{\lambda_1} \approx 125$$

Individuals with prior IV drug use reach the same survival probability as the 6-month survival of those without IV drug use after approximately 125 days.

c. Weibull Model

Under the Weibull model, $S(t) = e^{-(\lambda t)^\kappa}$, where κ is shape parameter.

Given the SAS output,

$$\lambda_0 = e^{-\beta_0} = e^{-5.9255} \approx 0.00267$$

$$\lambda_1 = e^{-(\beta_0 + \beta_{ivdrug})} = e^{-(5.9255 - 0.3813)} = e^{-5.5442} \approx 0.00391$$

$$\kappa = 0.9520$$

Therefore,

$$S_0(182) = e^{-(0.00267 \times 182)^{0.9520}} \approx 0.68$$

$$S_1(182) = e^{-(0.00391 \times 182)^{0.9520}} \approx 0.57$$

At 6 months, individuals without IV drug use have a higher survival probability (68%) compared to those with IV drug use (57%).

d. Log-logistic Model

Under the log-logistic model, $S(t) = \frac{1}{1 + (\frac{t}{\rho})^\gamma}$, where ρ is scale parameter and γ is shape parameter.

Given the SAS output,

$$\rho_0 = e^{\beta_0} = e^{5.3830} \approx 217.674$$

$$\rho_1 = e^{\beta_0 + \beta_{\text{ivdrug}}} = e^{5.3830 - 0.3389} \approx 155.105$$

$$\gamma = 0.7288$$

Therefore,

$$S_0(182) = \frac{1}{1 + (\frac{182}{217.674})^{0.7288}} \approx 0.53$$

$$S_1(182) = \frac{1}{1 + (\frac{182}{155.105})^{0.7288}} \approx 0.47$$

At 6 months, individuals without IV drug use have a higher survival probability (53%) compared to those with IV drug use (47%).

All three models are consistent with the increased relapse risk associated with prior IV drug use.