GRAPHICAL MODEL STRUCTURE LEAERNING

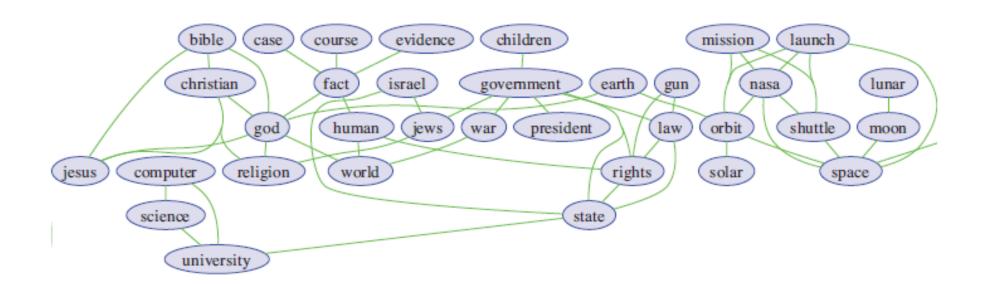
Presented by Zi Wang

- Structure learning for knowledge discovery
- Learning tree structures
- Learning DAG structures
- Learning DAG structure with latent variables
- Learning causal DAGs
- Learning undirected Gaussian graphical models
- Learning undirected discrete graphical models

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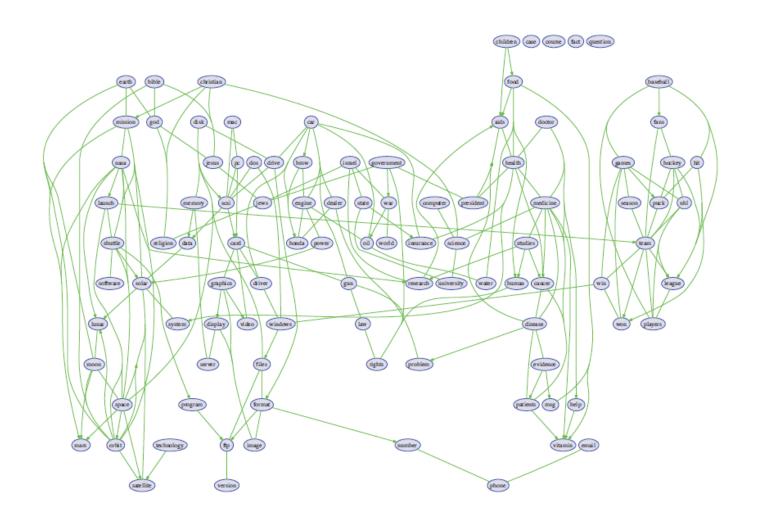
STRUCTURE LEARNING FOR KNOWLEDGE DISCOVERY

- Relevance networks
 - Visualizing mutual information(MI)
 - Dense graph because of disability to capture conditional independence
 - Solution: graphical model



STRUCTURE LEARNING FOR KNOWLEDGE DISCOVERY

- Dependency networks
 - $p(x_t|x_{\neg t})$
 - only its Markov blanket will be chosen as input
 - Sparse graph
- Application
 - Visualization
 - Inference with gibbs sampling



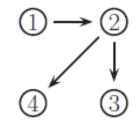
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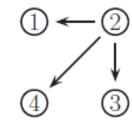
LEARNING TREE STRUCTURES

Equivalence between directed and undirected representation (for trees)

•
$$p(x_1, x_2, x_3, x_4|T) = p(x_1)p(x_2|x_1)p(x_3|x_2)p(x_4|x_2)$$

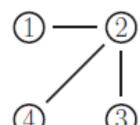
•
$$p(x_1, x_2, x_3, x_4|T) = p(x_2)p(x_1|x_2)p(x_3|x_2)p(x_4|x_2)$$





•
$$p(x_1, x_2, x_3, x_4|T) = p(x_1)p(x_2)p(x_3)p(x_4)\frac{p(x_1, x_2)}{p(x_1)p(x_2)}\frac{p(x_2, x_3)}{p(x_2)p(x_3)}\frac{p(x_2, x_4)}{p(x_2)p(x_4)}$$

•
$$p(\vec{x}|T) = \prod_{t \in V} p(x_t) \prod_{(s,t) \in E} \frac{p(x_s, x_t)}{p(x_s)p(x_t)}$$



CHOW-LIU ALGORITHM FOR FINDING THE ML TREE STRUCTURE

Log-likelihood function

$$\log p(\mathcal{D}|\boldsymbol{\theta}, T) = \sum_{t} \sum_{k} N_{tk} \log p(x_t = k|\boldsymbol{\theta})$$

$$+ \sum_{s,t} \sum_{j,k} N_{stjk} \log \frac{p(x_s = j, x_t = k|\boldsymbol{\theta})}{p(x_s = j|\boldsymbol{\theta})p(x_t = k|\boldsymbol{\theta})}$$

CHOW-LIU ALGORITHM FOR FINDING THE ML TREE STRUCTURE

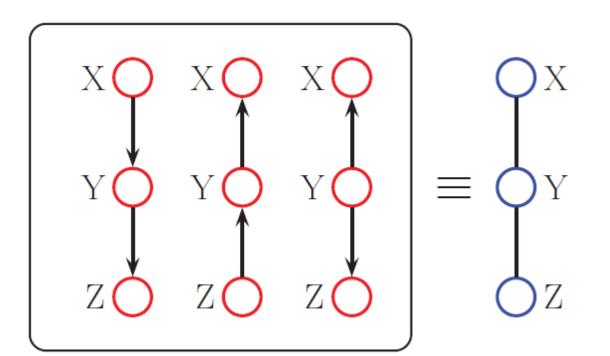
Log-likelihood function

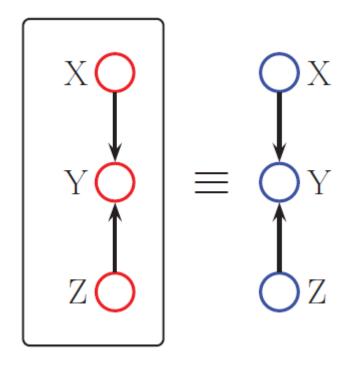
$$\frac{\log p(\mathcal{D}|\boldsymbol{\theta}, T)}{N} = \sum_{t \in \mathcal{V}} \sum_{k} p_{\text{emp}}(x_t = k) \log p_{\text{emp}}(x_t = k) + \sum_{(s,t) \in \mathcal{E}(T)} \mathbb{I}(x_s, x_t | \hat{\boldsymbol{\theta}}_{st})$$

$$\mathbb{I}(x_s, x_t | \hat{\boldsymbol{\theta}}_{st}) = \sum_{j} \sum_{k} p_{\text{emp}}(x_s = j, x_t = k) \log \frac{p_{\text{emp}}(x_s = j, x_t = k)}{p_{\text{emp}}(x_s = j) p_{\text{emp}}(x_t = k)}$$

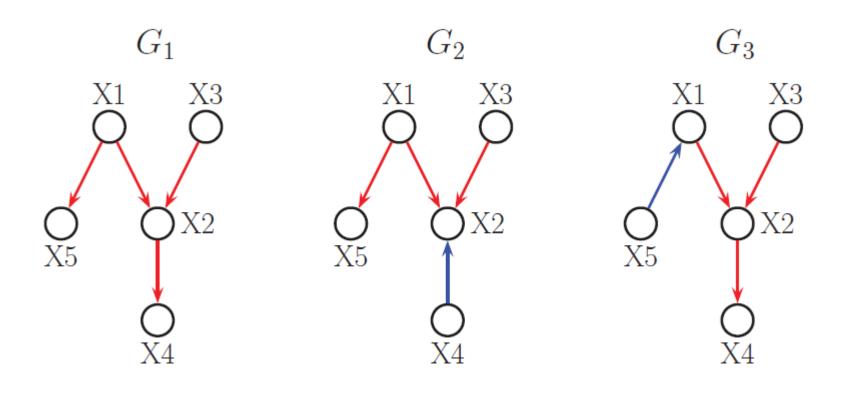
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• Markov Equivalence





MARKOV EQUIVALENCE - EXAMPLE



- $\theta_{tck} = p(x_t = k | x_{pa(t)} = c)$ where c = 1: C_t , $C_t = K^{d_t}$, $d_t = \dim(pa(t))$
- N_{tck} is the number of times node t is in state k and its parents are in state c
- Assumption: no missing data

$$p(\mathcal{D}|G, \boldsymbol{\theta}) = \prod_{i=1}^{N} \prod_{t=1}^{V} \operatorname{Cat}(x_{it}|\mathbf{x}_{i, \operatorname{pa}(t)}, \boldsymbol{\theta}_{t})$$

$$= \prod_{i=1}^{N} \prod_{t=1}^{V} \prod_{c=1}^{C_{t}} \operatorname{Cat}(x_{it}|\boldsymbol{\theta}_{tc})^{\mathbb{I}(\mathbf{x}_{i, \operatorname{pa}(t)} = c)}$$

$$= \prod_{i=1}^{N} \prod_{t=1}^{V} \prod_{c=1}^{C_{t}} \prod_{k=1}^{K_{t}} \theta_{tck}^{\mathbb{I}(x_{i,t} = k, \mathbf{x}_{i, \operatorname{pa}(t)} = c)}$$

$$= \prod_{i=1}^{V} \prod_{t=1}^{C_{t}} \prod_{c=1}^{K_{t}} \theta_{tck}^{N_{tck}}$$

t=1 c=1 k=1

- Constraints on priors:
 - $p(\theta) = \prod_t p(\theta_t)$ global prior parameter independence
 - $p(\theta_t) = \prod_c p(\theta_{tc})$ local prior parameter independence
 - => $p(\theta_{tc}) = Dir(\theta_{tc}|\alpha_{tc})$

Marginal likelihood

•
$$p(D|G) = \prod_{t=1}^{V} \prod_{c=1}^{C_t} \frac{B(N_{tc} + \alpha_{tc})}{B(\alpha_{tc})}$$

•
$$score(N_t, pa(t)) = \prod_{c=1}^{C_t} \frac{B(N_{tc} + \alpha_{tc})}{B(\alpha_{tc})}$$

• Recap Chapter 5:

•
$$p(\theta|D) = \frac{p(\theta)p(D|\theta)}{p(D)} = > \frac{1}{B(\alpha+N)} \prod_t \theta_t^{\alpha_t + N_t - 1} = \frac{\frac{1}{B(\alpha)} \prod_t \theta_t^{\alpha_t - 1} \prod_t \theta_t^{N_t}}{p(D)} = > p(D) = \frac{B(N+\alpha)}{B(\alpha)}$$

Here just compute the marginal likelihood directly

- Setting the prior
- $\alpha_{tck} = \alpha p_0(x_t = k, x_{pa(t)} = c) = \alpha/K_tC_t$
- Example:

•
$$N_1 = (5,3)$$
 $N_{21} = (4,1)$ $N_{22} = (1,2)$

•
$$G_1 = X_1 \to X_2 : \alpha_1 = (\frac{\alpha}{2}, \frac{\alpha}{2}) \alpha_{21} = (\frac{\alpha}{4}, \frac{\alpha}{4}) \alpha_{22} = (\frac{\alpha}{4}, \frac{\alpha}{4})$$

•
$$G_2 = X_1$$
 $X_2 : \alpha_1 = (\frac{\alpha}{2}, \frac{\alpha}{2}) \alpha_{21} = (\frac{\alpha}{2}, \frac{\alpha}{2}) \alpha_{22} = (\frac{\alpha}{2}, \frac{\alpha}{2})$

X_1	X_2
1	1
1	2
1	1
2	2
1	1
2	1
1	1
2	2

Scaling problem: too many possible graphs

$$f(D) = \sum_{i=1}^{D} (-1)^{i+1} \binom{D}{i} 2^{i(D-i)} f(D-i)$$

- Nodes number: 1 2 3 4 5 6
- Graphs number: 1 3 25 543 29281 3781503
- Solutions
 - Greedy hill climbing
 - Sample DAGs from the posterior, e.g. Metropolis Hasting with proposals in greedy search

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LEARNING DAG STRUCTURES WITH LATENT VARIABLES

Learning DAG structure without complete data

$$p(\mathcal{D}|G) = \int \sum_{\mathbf{h}} p(\mathcal{D}, \mathbf{h}|\boldsymbol{\theta}, G) p(\boldsymbol{\theta}|G) d\boldsymbol{\theta} = \sum_{\mathbf{h}} \int p(\mathcal{D}, \mathbf{h}|\boldsymbol{\theta}, G) p(\boldsymbol{\theta}|G) d\boldsymbol{\theta}$$

Intractable to compute

APPROXIMATING THE MARGINAL LIKELIHOOD WITH MISSING DATA

BIC approximation

$$\mathrm{BIC}(G) \triangleq \log p(\mathcal{D}|\hat{\boldsymbol{\theta}}, G) - \frac{\log N}{2} \dim(G)$$

Cheeseman-Stutz approximation

$$p(\mathcal{D}|G) \approx p(\overline{\mathcal{D}}|G) = \int p(\overline{\mathcal{D}}|\boldsymbol{\theta}, G)p(\boldsymbol{\theta}|G)d\boldsymbol{\theta}$$

$$\log p(\mathcal{D}|G) \approx \log p(\overline{\mathcal{D}}|G) + \log p(\mathcal{D}|\hat{\boldsymbol{\theta}}, G) - \log p(\overline{\mathcal{D}}|\hat{\boldsymbol{\theta}}, G)$$

Variational Bayes EM

$$p(\boldsymbol{\theta}, \mathbf{z}_{1:N} | \mathcal{D}) \approx q(\boldsymbol{\theta}) q(\mathbf{z}) = q(\boldsymbol{\theta}) \prod_{i} q(\mathbf{z}_{i})$$

STRUCTURAL EM

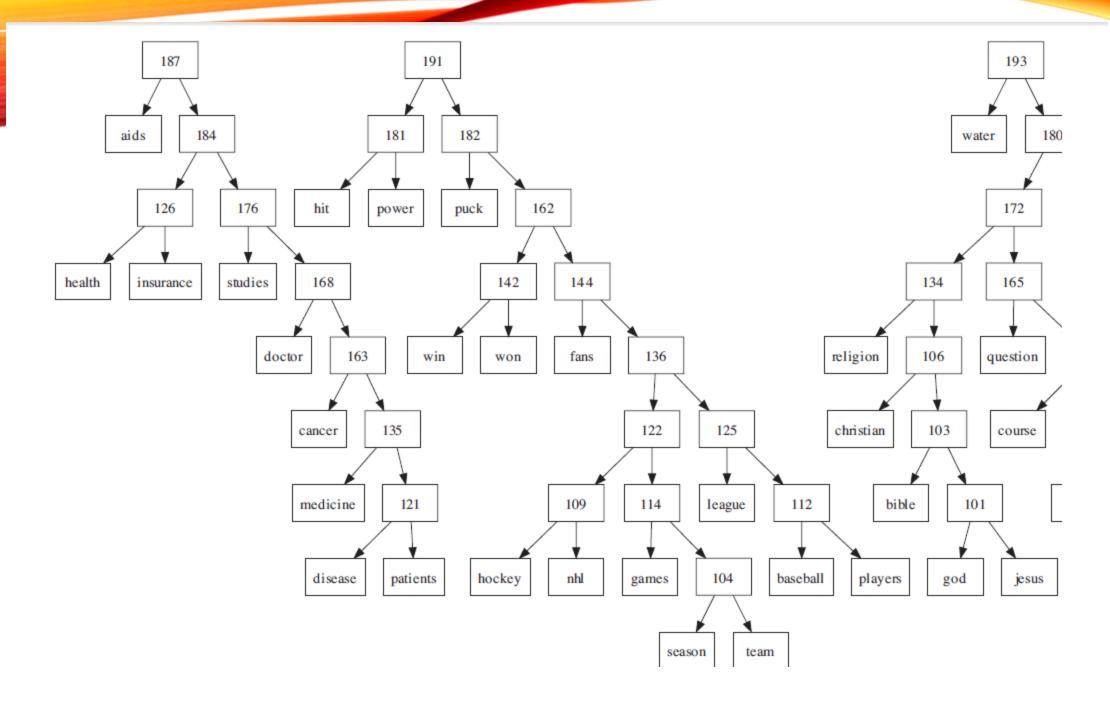
- 0) Initialize a model
- 1) Fill in the data with the current model
- 2) Use the filled-in data to evaluate the score of all the neighbors
- 3) Pick the best neighbor
- 4) Repeat 1),2),3)

$$score_{\mathrm{BIC}}(G,\mathcal{D}) \triangleq \log p(\mathcal{D}|\hat{\boldsymbol{\theta}},G) - \frac{\log N}{2}\dim(G) + \log p(G) + \log p(\hat{\boldsymbol{\theta}}|G)$$

- Pro: efficient
- App: learn the phylogenetic tree structure, learn sparse mixture models

DISCOVERING HIDDEN VARIABLES

- Introduce hidden variables to structural signatures, e.g. sets of densely connected nodes
- Latent class model
 - introduce hidden variables with high mutual information with their children
 - Hierarchical latent class model
 - greedy local search algorithm
 - A faster greedy algorithm based on agglomerative hierarchical clustering (Harmeling and Williams 2011)
 - Another approach: Chow-Liu Tree on observed data + add hidden variable



STRUCTURAL EQUATION MODELS

SEM: a statistical technique for testing and estimating causal relations using a combination of statistical data and qualitative causal assumptions.

$$x_i = \mu_i + \sum_{j \neq i} w_{ij} x_j + \epsilon_i \qquad \epsilon \sim \mathcal{N}(0, \Psi)$$

$$\mathbf{x} = \mathbf{W}\mathbf{x} + \boldsymbol{\mu} + \boldsymbol{\epsilon} \Rightarrow \mathbf{x} = (\mathbf{I} - \mathbf{W})^{-1}(\boldsymbol{\epsilon} + \boldsymbol{\mu})$$

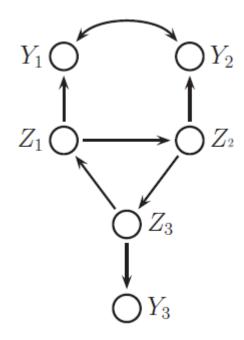
$$p(\mathbf{x}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$\Sigma = (\mathbf{I} - \mathbf{W})^{-1} \Psi (\mathbf{I} - \mathbf{W})^{-T}$$

SEM EXAMPLE

$$\begin{pmatrix} X_1 \\ X_2 \\ X_3 \\ X_4 \\ X_5 \\ X_6 \end{pmatrix} = \begin{pmatrix} Z_1 \\ Z_2 \\ Z_3 \\ Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} = \begin{pmatrix} 0 & 0 & w_{13} & 0 & 0 & 0 \\ w_{21} & 0 & 0 & 0 & 0 & 0 \\ 0 & w_{32} & 0 & 0 & 0 & 0 \\ w_{41} & 0 & 0 & 0 & 0 & 0 \\ w_{41} & 0 & w_{52} & 0 & 0 & 0 & 0 \\ 0 & w_{52} & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & w_{63} & 0 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} Z_1 \\ Z_2 \\ Z_3 \\ Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \\ \epsilon_6 \end{pmatrix}$$

$$\Psi = \begin{pmatrix} \Psi_{11} & 0 & 0 & 0 & 0 & 0 \\ 0 & \Psi_{22} & 0 & 0 & 0 & 0 \\ 0 & 0 & \Psi_{33} & 0 & 0 & 0 \\ 0 & 0 & 0 & \Psi_{44} & \Psi_{45} & 0 \\ 0 & 0 & 0 & \Psi_{54} & \Psi_{55} & 0 \\ 0 & 0 & 0 & 0 & 0 & \Psi_{66} \end{pmatrix}$$

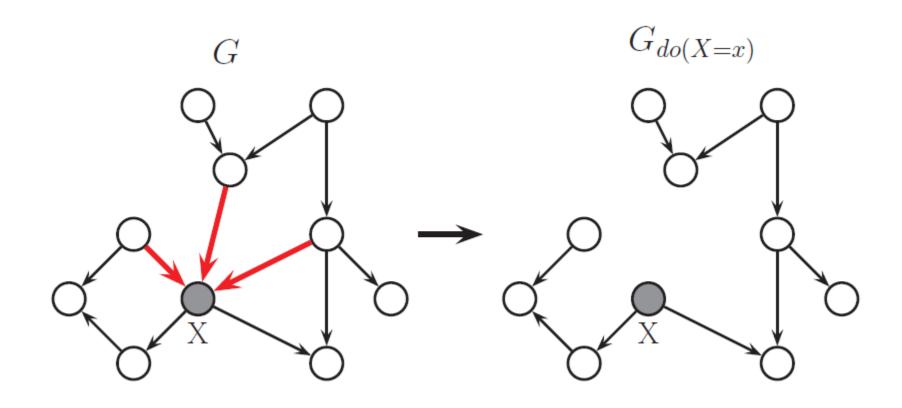


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LEARNING CAUSAL DAGS

- Causal models: models which can predict the effects of interventions to, or manipulations of, a system
- Assumptions:
 - Causal Markov assumption
 - Causal sufficiency assumption
- Notations:
 - $do(X_i = x_i)$ set X_i to be x_i
 - Difference between conditioning on observation and manipulation:
 - P(S=1 | Y=1)>P(S=1)
 - P(S=1 | do(Y=1))=P(S=1)

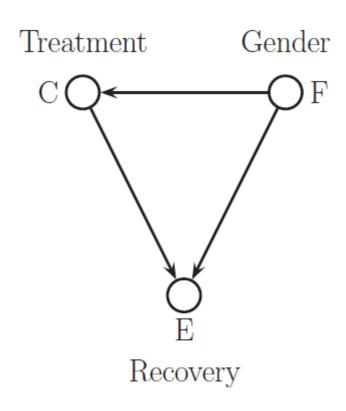
GRAPH SURGERY



SIMPSON'S PARADOX

- Any statistical relationship between two variables can be reversed by including additional factors the analysis
- $P(E | C) > P(E | \neg C)$
- Include additional factor F:
- $P(E \mid C, F) < P(E \mid \neg C, F)$
- $P(E | C, \neg F) < P(E | \neg C, \neg F)$
- The statement is not interpreted properly!
- Should be $P(E|do(C)) > P(E|do(\neg C))$

SIMPSON'S PARADOX



Suppose
$$p(E|do(C),F) < p(E|do(\neg C),F)$$

$$p(E|do(C),\neg F) < p(E|do(\neg C),\neg F)$$
 We show
$$p(E|do(C)) < p(E|do(\neg C))$$

SIMPSON'S PARADOX

• Proof:

- $p(F|do(C)) = p(F|do(\neg C)) = p(F)$
- $p(E|do(C)) = p(E|do(C), F)p(F|do(C)) + p(E|do(C), \neg F)p(\neg F|do(C))$ = $p(E|do(C), F)p(F) + p(E|do(C), \neg F)p(\neg F)$
- $p(E|do(\neg C)) = p(E|do(\neg C), F)p(F) + p(E|do(\neg C), \neg F)p(\neg F)$
- Given $p(E \mid do(C),F) < p(E \mid do(\neg C),F)$ and $p(E \mid do(C),\neg F) < p(E \mid do(\neg C),\neg F)$
- Conclusion: $p(E|do(C)) < p(E|do(\neg C))$

LEARNING CAUSAL DAG STRUCTURES

- Learning from observational data
 - Learn an PDAG from data
 - Enumerate all the DAGs from the PDAG equivalence class
 - Apply Pearl's do-calculus to compute the magnitude of each causal effect pair
 - Take the minimum of these effects as the lower bound
- Learning from interventional data
 - Control some variables and measure the consequences
 - First skipping over intervention cases
 - Then adding the intervention nodes with constraints

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LEARNING UNDIRECTED GAUSSIAN GRAPHICAL MODELS

- MLE for a GGM covariance estimation
- $l(\Omega) = \log|\Omega| tr(S\Omega)$
- $\nabla l(\Omega) = \Omega^{-1} S$
- Constraints:
 - $\Omega_{st} = 0$ if $G_{st} = 0$
 - Ω is positive definite
- Property: $\Sigma_{st} = S_{st}$ if $G_{st} = 1$ or s = t

GRAPHICAL LASSO

•
$$J(\Omega) = -\log|\Omega| + tr(S\Omega) + \lambda ||\Omega||_1$$

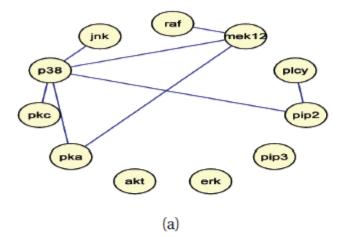
- Convex but non-smooth
- Coordinate descent algorithm

• Shooting algorithm for lasso $\underset{x}{\operatorname{arg}\min} \|Ax - y\|_2^2 + \lambda \|x\|_1$

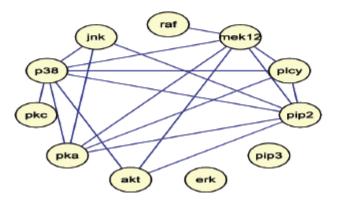
$$S_{j} = A_{(:,j)}^{T}(Ax) - 2(y^{T}A)_{j} + (A^{T}A)_{j,j}x_{j}^{t-1}$$

$$x_{j}^{t} \leftarrow \operatorname{sign}(S_{j})(|S_{j}| - \lambda)_{+}$$

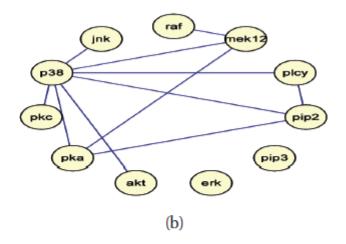
lambda=36.00, nedges=8



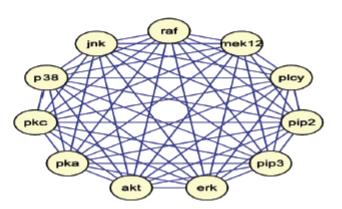
lambda=7.00, nedges=18



lambda=27.00, nedges=11



lambda=0.00, nedges=55



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LEARNING UNDIRECTED DISCRETE GRAPHICAL MODELS

Grapical lasso for CRF

$$\psi_t(y_t, \mathbf{x}) = \begin{pmatrix} \mathbf{v}_{t1}^T \mathbf{x} \\ \mathbf{v}_{t2}^T \mathbf{x} \\ \mathbf{v}_{t3}^T \mathbf{x} \end{pmatrix}, \ \psi_{st}(y_s, y_t, \mathbf{x}) = \begin{pmatrix} \mathbf{w}_{t11}^T \mathbf{x} & \mathbf{w}_{st12}^T \mathbf{x} & \mathbf{w}_{st13}^T \mathbf{x} \\ \mathbf{w}_{st21}^T \mathbf{x} & \mathbf{w}_{st22}^T \mathbf{x} & \mathbf{w}_{st23}^T \mathbf{x} \\ \mathbf{w}_{st31}^T \mathbf{x} & \mathbf{w}_{st32}^T \mathbf{x} & \mathbf{w}_{st33}^T \mathbf{x} \end{pmatrix}$$

$$J = -\sum_{i=1}^{N} \left[\sum_{t=1}^{N} \log \psi_{t}(y_{it}, \mathbf{x}_{i}, \mathbf{v}_{t}) + \sum_{s=1}^{V} \sum_{t=s+1}^{V} \log \psi_{st}(y_{is}, y_{it}, \mathbf{x}_{i}, \mathbf{w}_{st}) \right] + \lambda_{1} \sum_{s=1}^{V} \sum_{t=s+1}^{V} ||\mathbf{w}_{st}||_{p} + \lambda_{2} \sum_{t=1}^{V} ||\mathbf{v}_{t}||_{2}^{2}$$

Thin junction tree – bound the treewidth

SUMMARY

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•Thanks!