CS229 Lecture notes

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Part XI

Principal components analysis

王风分分析法

无监督学习

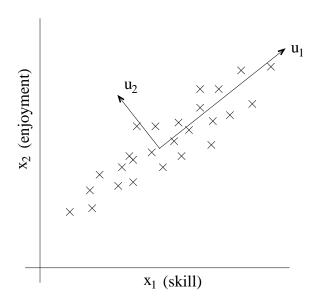
In our discussion of factor analysis, we gave a way to model data $x \in \mathbb{R}^n$ as "approximately" lying in some k-dimension subspace, where $k \ll n$. Specifically, we imagined that each point $x^{(i)}$ was created by first generating some $z^{(i)}$ lying in the k-dimension affine space $\{\Lambda z + \mu; z \in \mathbb{R}^k\}$, and then adding Ψ -covariance noise. Factor analysis is based on a probabilistic model, and parameter estimation used the iterative EM algorithm.

In this set of notes, we will develop a method, Principal Components Analysis (PCA), that also tries to identify the subspace in which the data approximately lies. However, PCA will do so more directly, and will require only an eigenvector calculation (easily done with the eig function in Matlab), and does not need to resort to EM.

Suppose we are given dataset $\{x^{(i)}; i=1,\ldots,m\}$ of attributes of m different types of automobiles, such as their maximum speed, turn radius, and so on. Lets $x^{(i)} \in \mathbb{R}^n$ for each i ($n \ll m$). But unknown to us, two different attributes—some x_i and x_j —respectively give a car's maximum speed measured in miles per hour, and the maximum speed measured in kilometers per hour. These two attributes are therefore almost linearly dependent, up to only small differences introduced by rounding off to the nearest mph or kph. Thus, the data really lies approximately on an n-1 dimensional subspace. How can we automatically detect, and perhaps remove, this redundancy?

For a less contrived example, consider a dataset resulting from a survey of pilots for radio-controlled helicopters, where $x_1^{(i)}$ is a measure of the piloting skill of pilot i, and $x_2^{(i)}$ captures how much he/she enjoys flying. Because RC helicopters are very difficult to fly, only the most committed students, ones that truly enjoy flying, become good pilots. So, the two attributes x_1 and x_2 are strongly correlated. Indeed, we might posit that that the

PCA要解决的问题就是 怎么检测并移除多余的属性。2 data actually likes along some diagonal axis (the u_1 direction) capturing the intrinsic piloting "karma" of a person, with only a small amount of noise lying off this axis. (See figure.) How can we automatically compute this u_1 direction?



We will shortly develop the PCA algorithm. But prior to running PCA per se, typically we first pre-process the data to normalize its mean and variance, as follows: 运行PCA算法之前, PCA一般都需要预处理

数据的标准化

1. Let
$$\mu = \frac{1}{m} \sum_{i=1}^{m} x^{(i)}$$
.
2. Replace each $x^{(i)}$ with $x^{(i)} - \mu$.

这两步将数据的均值变为0,当数据的均值 本来就是0的时候可以省略

3. Let
$$\sigma_j^2 = \frac{1}{m} \sum_i (x_j^{(i)})^2$$
 这一步是求方差

合起来这两步就是将数据在各个维度的方差变为1,5 样就使得各个维度在同一尺度下被衡量,根本目的是

4. Replace each $x_j^{(i)}$ with $x_j^{(i)}/\sigma_j$.

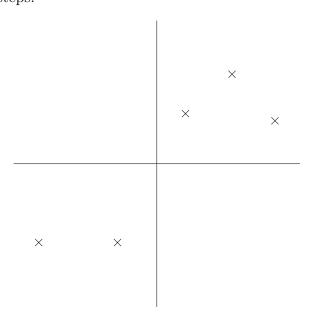
Steps (1-2) zero out the mean of the data, and may be omitted for data known to have zero mean (for instance, time series corresponding to speech or other acoustic signals). Steps (3-4) rescale each coordinate to have unit variance, which ensures that different attributes are all treated on the same "scale." For instance, if x_1 was cars' maximum speed in mph (taking values in the high tens or low hundreds) and x_2 were the number of seats (taking values around 2-4), then this renormalization rescales the different attributes to make them more comparable. Steps (3-4) may be omitted if we had apriori knowledge that the different attributes are all on the same scale. One

example of this is if each data point represented a grayscale image, and each $x_j^{(i)}$ took a value in $\{0, 1, \dots, 255\}$ corresponding to the intensity value of pixel j in image i.

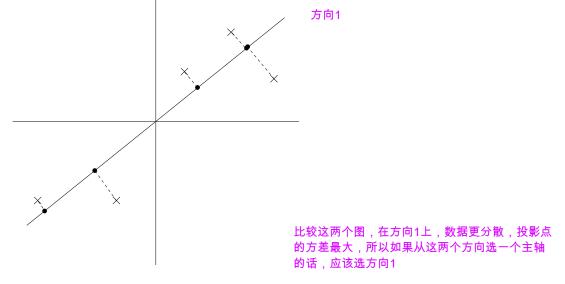
Now, having carried out the normalization, how do we compute the "major axis of variation" u—that is, the direction on which the data approximately lies? One way to pose this problem is as finding the unit vector u so that when the data is projected onto the direction corresponding to u, the variance of the projected data is maximized. Intuitively, the data starts off with some amount of variance/information in it. We would like to choose a direction u so that if we were to approximate the data as lying in the direction/subspace corresponding to u, as much as possible of this variance is still retained.

Consider the following dataset, on which we have already carried out the normalization steps:

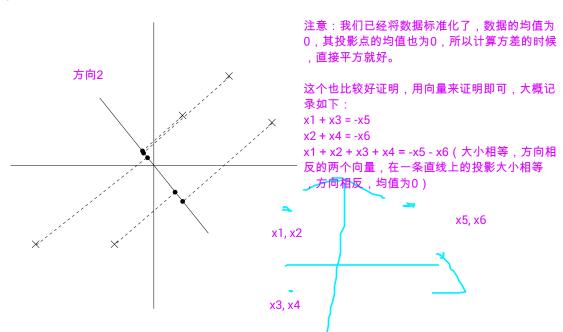
怎么去找到主轴呢?假设找到主轴方向了,向的不是数据在主轴方的区分开,自然的说也就是我们希望数据在主轴方向上尽量分的,自然在主轴方向上尽量分散。具体操作就是让所有的点的方差最大。



Now, suppose we pick u to correspond the the direction shown in the figure below. The circles denote the projections of the original data onto this line.



We see that the projected data still has a fairly large variance, and the points tend to be far from zero. In contrast, suppose had instead picked the following direction:



Here, the projections have a significantly smaller variance, and are much closer to the origin.

We would like to automatically select the direction u corresponding to the first of the two figures shown above. To formalize this, note that given a

unit vector u and a point x, the length of the projection of x onto u is given by $x^T u$. I.e., if $x^{(i)}$ is a point in our dataset (one of the crosses in the plot), then its projection onto u (the corresponding circle in the figure) is distance $x^{T}u$ from the origin. Hence, to maximize the variance of the projections, we would like to choose a unit-length u so as to maximize:

$$/$$
/ $/$ / $/$

$$\frac{1}{m} \sum_{i=1}^{m} (x^{(i)^T} u)^2 = \frac{1}{m} \sum_{i=1}^{m} u^T x^{(i)} x^{(i)^T} u$$
在投影方向上的
长度
$$= u^T \left(\frac{1}{m} \sum_{i=1}^{m} x^{(i)} x^{(i)^T} \right) u.$$

注意 u 是单位向量,||u|| = 1,这也是最

n阶对称阵有 n 个相互 正交的特征向量

怎么降维? 选取最大的k个特征值 对应的特征向量即可

这个矩阵实际上是一个 We easily recognize that the maximizing this subject to $||u||_2 = 1$ gives the 协方差矩阵 (也是个对 principal eigenvector of $\Sigma = \frac{1}{m} \sum_{i=1}^m x^{(i)} x^{(i)^T}$, which is just the empirical covariance matrix of the data (assuming it has zero mean).

To summarize, we have found that if we wish to find a 1-dimensional subspace with with to approximate the data, we should choose u to be the principal eigenvector of Σ . More generally, if we wish to project our data into a k-dimensional subspace (k < n), we should choose u_1, \ldots, u_k to be the top k eigenvectors of Σ . The u_i 's now form a new, orthogonal basis for the $data.^2$

Then, to represent $x^{(i)}$ in this basis, we need only compute the corresponding vector

$$y^{(i)} = \left[egin{array}{c} u_1^T x^{(i)} \ u_2^T x^{(i)} \ dots \ u_k^T x^{(i)} \end{array}
ight] \in \mathbb{R}^k.$$

Thus, whereas $x^{(i)} \in \mathbb{R}^n$, the vector $y^{(i)}$ now gives a lower, k-dimensional, approximation/representation for $x^{(i)}$. PCA is therefore also referred to as a dimensionality reduction algorithm. The vectors u_1, \ldots, u_k are called the first k principal components of the data.

Remark. Although we have shown it formally only for the case of k=1, using well-known properties of eigenvectors it is straightforward to show that 为什么u就是这个矩阵 的特征向量的相关证

https://blog.csdn.net/s tdcoutzyx/article/detail s/37568225 基本步骤就是根据约 束写出拉格朗日方程 ,求导,让导数为 0,可证u就是这个矩 阵的特征向量

¹If you haven't seen this before, try using the method of Lagrange multipliers to maximize $u^T \Sigma u$ subject to that $u^T u = 1$. You should be able to show that $\Sigma u = \lambda u$, for some λ , which implies u is an eigenvector of Σ , with eigenvalue λ .

²Because Σ is symmetric, the u_i 's will (or always can be chosen to be) orthogonal to each other.

of all possible orthogonal bases u_1, \ldots, u_k , the one that we have chosen maximizes $\sum_i ||y^{(i)}||_2^2$. Thus, our choice of a basis preserves as much variability as possible in the original data.

In problem set 4, you will see that PCA can also be derived by picking the basis that minimizes the approximation error arising from projecting the data onto the k-dimensional subspace spanned by them.

PCA has many applications, our discussion with a small number of examples. First, compression—representing $x^{(i)}$'s with lower dimension $y^{(i)}$'s—is an obvious application. If we reduce high dimensional data to k=2 or 3 dimensions, then we can also plot the $y^{(i)}$'s to visualize the data. For instance, if we were to reduce our automobiles data to 2 dimensions, then we can plot it (one point in our plot would correspond to one car type, say) to see what cars are similar to each other and what groups of cars may cluster together.

Another standard application is to preprocess a dataset to reduce its dimension before running a supervised learning learning algorithm with the $x^{(i)}$'s as inputs. Apart from computational benefits, reducing the data's dimension can also reduce the complexity of the hypothesis class considered and help avoid overfitting (e.g., linear classifiers over lower dimensional input spaces will have smaller VC dimension).

Lastly, as in our RC pilot example, we can also view PCA as a noise reduction algorithm. In our example it, estimates the intrinsic "piloting karma" from the noisy measures of piloting skill and enjoyment. In class, we also saw the application of this idea to face images, resulting in **eigenfaces** method. Here, each point $x^{(i)} \in \mathbb{R}^{100 \times 100}$ was a 10000 dimensional vector, with each coordinate corresponding to a pixel intensity value in a 100x100 image of a face. Using PCA, we represent each image $x^{(i)}$ with a much lower-dimensional $y^{(i)}$. In doing so, we hope that the principal components we found retain the interesting, systematic variations between faces that capture what a person really looks like, but not the "noise" in the images introduced by minor lighting variations, slightly different imaging conditions, and so on. We then measure distances between faces i and j by working in the reduced dimension, and computing $||y^{(i)}-y^{(j)}||_2$. This resulted in a surprisingly good face-matching and retrieval algorithm.