$MSE(f) = \frac{1}{N} \sum_{n=1}^{N} (y_n - f(x_n))^2$ $Convex: h(\lambda u + (1 - \lambda)v) \le \lambda h(u) + (1 - \lambda)h(v)$

Cvex f has a unique glob min w^* , sum of cvx = cvx

Convex set: \mathcal{C} is convex iff $\theta u + (1-\theta)v \in \mathcal{C}$ Proof convexity: show that Hess is positive semi-def (all eigvals are ≥ 0)

(1) nonneg weighted sums (2) compo with affine mapping (3) poitwise max of convex funcs (4) restric to a line (5)

Week2 Gradient Descent $\mathcal{O}(ND)$

1. Start arbitrary $w_0 \in \mathbb{R}$

2. For *i* do $w_{t+1} = w_t - \eta_t \nabla \mathcal{L}(w_t)$ MSE for linreg: $\mathcal{L}(w) = \frac{1}{2N} e^{T} e^{T}$

 $\nabla \mathcal{L}(w) = -\frac{1}{N} X^T e$

Stochastic Gradient Descent (SGD) $\mathcal{O}(D)$

NB: Watch out that your loss has indeed this form before applying SGD!

$$\mathcal{L}(w) = \frac{1}{N} \sum_{n=1}^{N} \mathcal{L}_n(w)$$

 \mathcal{L}_n = cost at n-th training example (unbiased)

1. Start at an arbitrary $w_0 \in \mathbb{R}^d$ 2. For t = 1, 2, ... do: Pick data point $(x', y') \in_{u,q,r} D$

 $w_{t+1} = w_t - \gamma \nabla \mathcal{L}_n(w_t; x', y')$

Mini-batch SGD

 $g = \frac{1}{|B|} \sum_{n \in B} \nabla \mathcal{L}_n(w_t)$

 $W_{t+1} = W_t - \gamma q$

Subgradient

 $q \in \mathbb{R}^D$ is a subgradient iff: $\forall u.\mathcal{L}(u) \geq$ $\mathcal{L}(w) + a^T(u-w)$

 $w_{t+1} = w_t - \gamma q$ g = subgradient to \mathcal{L} at w_t

NB: if \mathcal{L} is convex and differentiable at w, the only subgrad at w is $q = \nabla \mathcal{L}(w)$. SubGD => \mathcal{L}_n = subgrad

Projected Gradient Descent

add a projection ont convex set \mathcal{C}

 $P_{\mathcal{C}}(w') = argmin_{v \in \mathcal{C}} ||v - w'||$ $w_{t+1} = P_{\mathcal{C}}[w_t - \gamma \nabla \mathcal{L}(w_t)]$ Week3

Linear regression MSE

Normal equations: equations obt when setting the $\nabla = 0. X'(y - Xw) = 0$ for lin reg

$$\mathcal{L}(w) = \frac{1}{2N} ||Xw - y||^2$$

 $\nabla \mathcal{L}(w) = 0 \Rightarrow w^* = (X^T X)^{-1} X^T y$ We take $\hat{y} =$ projection of y onto span(X). Prediction: $\hat{q}_m = x_m^T w^*$ Gram Matrix X^TX only invertible if rank(X) = D(1) if D > N we have rank(X) < D (we can still solve for rank-defficiency)(2) if $D \le N$ but some columns are collinear then X is ill-conditionned *Rightarrow*

Closed form: $\mathcal{O}(D^2N)(X^TX) + \mathcal{O}(DN)(X^Ty) +$ $\mathcal{O}(D^3)$ (inverse X^TX) = $\mathcal{O}(D^2N)$

MLE for linear least-squares vs Gaussian

 $p(y|X,w) \sim \mathcal{N}(w^T x_i, \sigma^2)$: $u_i = w^T x_i + \epsilon_i, \epsilon_i \sim \mathcal{N}(0, \sigma^2)$

Maximizing the log likelihood:

= argmin $\sum_{i=1}^{n} (y_i - w^T x_i)^2$ Gaussian noise on

the data⇔ Least squares

Similarly for MAE and Laplace prior: p(y|X,w) =

 $\frac{1}{2b}e^{-\frac{1}{b}|y_n-x_n^-Tw|} \Rightarrow \min_{b} \sum_{i=1}^{n} |w^Tx_i-y_i|$ **Regularization** $\min_{w} \mathcal{L}(w) + \Omega w$ **Ridge Regression**

 $\min_{w} \frac{1}{2N} \sum_{n=1}^{N} [y_n - x_n^T w]^2 + \lambda ||w||_2^2$

Gradient: $\nabla_w L(w) = -2 \sum_{n=1}^N (y_n - w^T x_n) x_n +$ $2\lambda w = 2[(X^TX + \lambda I)w - X^Ty];\mathcal{O}(D)$ for stoch

 $w_{ridge}^* = (X^T X + \lambda' I)^{-1} X^T y, \lambda' = 2N\lambda; \mathcal{O}(D^3 + ND^2)$

eigvals $(X^TX + \lambda'I) \ge \lambda'$ so the inverse always exists Lasso: MSE and L1 regularization, encourages spar-

 $\frac{\text{se } w^*}{\text{Ridge reg}} = \text{MAP estimate}$

Introduce bias by expressing assumption through a Bayesian prior $w_i \in \mathcal{N}(0, \beta^2)$

Bayes rule: $P(w|x,y) = \frac{P(w|x)P(y|x,w)}{P(y|x)}$

= argmin $\lambda ||w||_2^2 + \sum_{i=1}^n (y_i - w^T x_i)^2$, $\lambda = \frac{\sigma^2}{\varrho^2}$

Underfitting and Overfitting

Fight overfitting with (1)simpler model (2) more

Generalization, Model Selection

True Risk: $L_{\mathcal{D}}(t) = \mathbb{E}_{\mathcal{D}}[\ell(y, f(x))]$, where loss fiction $= \ell(y, f(x)) = \frac{1}{2}(y - f(x))^2$ for ridge.

The True risk cannot be computed since \mathcal{D} is not

Empirical Risk: $L_S(f) = \frac{1}{|S|} \sum_{(x_n, y_n) \in S} \ell(y_n, f(x_n))$ (for ridge, emp risk = MSE)

Training Risk: $L_S(f) = \frac{1}{|S|} \sum_{(x_n, y_n) \in S} \ell(y_n, f_S(x_n))$

 $L_S(f_S) \neq L_D(f_S)$ because of overfitting.

Validation Risk: $L_{S_{test}}(f_{S_{train}})$ in expectation = True

Generalization error: $G(t) = |L_D(t) - L_{S_{test}}(t)|$ Gen error decreases as $\mathcal{O}(\frac{1}{\sqrt{|S_{tend}|}})$

Model Selection When testing K hpyer-parameters, the gen error goes up by a factor proportional to $\mathcal{O}(\sqrt{(ln(K))})$. If we chose the "best"function according to the va-

lidation risk, then its true risk is not too far away from the true risk of the optimal choice. KFold

For simple models s: the bias is large but $Var(L_D(f_S))$ is small. $\mathbb{E}_{S_{train} \sim \mathcal{D}, \epsilon \sim \mathcal{D}}[(f(x_0) + \epsilon - f_{S_{train}}(x_0))^2] =$

 $Var_{\epsilon \sim \mathcal{D}_{\epsilon}}[\epsilon] + (f(x_0) - \mathbb{E}_{S'_{train} \sim \mathcal{D}}[f_{S'_{train}}(x_0))^2] +$ $\mathbb{E}_{S_{train} \sim \mathcal{D}}[(\mathbb{E}_{S'_{train} \sim \mathcal{D}}[f_{S'_{train}}(x_0)] - f_{S_{train}}(x_0))^2]$

Simpler: True Risk = Noise variance + $bias^2$ + variance Each of the term is a lower bound on the true error for any input x_0 . The noise imposes a strict lower bound on what we can achieve.

Classification
(1) Estimating the probability (2) Quantize the value to a discrete value according to some threshold.

Boundaries = decision boundaries. Goal: fraction of misclassified cases is small: MSE ≠good match:counts + and - deviation eq bad, although only one of them can lead to a missclass.

 $f_{S_{train},k}(x) = \frac{1}{k} \sum_{n:x_n \in nbh_{S_{train},k}(x)} y_n$ Large k = simple model, small k = complex model if N samples allow you to perform well in 1D you need $\mathcal{O}(N^D)$ samples in D dimensions.

Has a misclass rate at most 2x the one of the Bayes classifier if we have a lot of data.

MAP criterion for classification We take the label decision which maximize the proba: $\hat{q}(x) = arqmax_{u \in \mathcal{V}} p(y|x)$.

 $\mathbb{P}(\hat{q}(x) = y) = \int p(x)p(\hat{q}(x)|x)dx$ Logistic regression

we use $y_n \in +-1$ for svm and $y_n \in 0,1$ for logistic $\phi(x) = \frac{1}{1 + e^{-x}}, \sigma'(x) = \sigma(x)(1 - \sigma(x))$

we bet that $p(y|x,w) \sim Ber(y,\sigma(w^Tx))$ Link function: $\sigma(w^Tx) = \frac{1}{1 + exp(-w^Tx)}$ (Sigmoid)

 $logisticLoss[0-1](w^Tx) = log(1 + e^{-yw^Tx})$ MLE for logistic regression

 $arqmax_w p(y,X|w) = arqmax_w p(y|X,w) =$ $\prod_{n=1}^{N} \sigma(x_n^T w)^{y_n} [1 - \sigma(x_n^T w)]^{1-y_n}$

We can define:

 $\mathcal{L}(w) = -\sum_{n=1}^{N} y_n ln \sigma(x_n^T w) + (1 - y_n) ln[1]$ $\sigma(x_n^T w)] = \sum_{n=1}^N \ln[1 + exp(x_n^T w)] - y_n x_n^T w$ so we chose: $w^* = argmin_w \mathcal{L}(w)$

 $\nabla \mathcal{L}(w) = X^T [\sigma(Xw) - y]$ Newton Method for Logistic Regression

Newtonupdate: $w_{(t+1)} = w_t - \gamma_t(H_t)^{-1} \nabla \mathcal{L}(w_t)$

Regularized Logistic Regression

When data is linsep logreg tend to inf. $w^* =$ $argmin_w - \sum_{n=1}^N lnp(y_n|x_n^Tw) + \frac{\lambda}{2} ||w||^2$ eq to MAP instd of MLE: $argmax_w p(w|X, y) =$ $argmax_w \frac{p(X,y|w)p(w)}{p(X,y)} = argmax_w p(X,y|w)p(w)$

Still for classif take the most likely label.

Week6 Exponential Families

Expo family:: $p(y|\eta) = h(y) exp[\eta^T \phi(y) - A(\eta)]$ $\phi(y)$ is the sufficient statistic: given an independent set of samples and the empirical average of $\phi(u)$ we can optimally estimate η .

Link function: q relates the mean of $\phi(y)$ to the

param η Constraint: $\int_{u} h(y) exp[\eta^{T} \phi(y)] dy = e^{A(\eta)}$. $A(\eta)$ ensures a proper normalization (cumulant) For every choice of h(y), $\phi(y)$, η we will get a member of the expo family. Then chose $A(\mu)$ s.t.

Lemma: $\nabla A(\eta) = \mathbb{E}[\phi(y)], \nabla^2 A(\eta) =$ $\mathbb{E}[\phi(y)\phi(y)^T] - \mathbb{E}[\phi(y)]\mathbb{E}[\phi(y)]^T$

the expression is well normalized.

Poisson as expo: $p(y|\mu) = \frac{\mu^y e^{-\mu}}{\mu!} = \frac{1}{\mu!} e^{y \ln(\mu) - \mu} =$ $h(y)e^{\eta\phi(y)-A(\eta)}$

So $\bullet h(y) = \frac{1}{\mu} \bullet \phi(y) = y \bullet \eta = g(\mu) = \ln(\mu) \bullet \mu =$ $q^{-1}(\eta) = e^{\eta}$

Max Likelihood param estimation for expo Goal: estimate the parameter η

 $L(\eta) = -ln(p(y|\eta)) = \sum_{n=1}^{N} [-ln(h(y_n)) - \eta^T \phi(y_n) +$ $A(\eta)$] then $\eta = g(\frac{1}{N} \sum_{n=1}^{N} \phi(y_n))$

Generalized Linear Models
Given an element of the expo family with a scalar $\phi(y)$ we can construct from this a data model by assuming that a sample (x,y) follows:

Generalized linear model: p(y|x, w) = $h(y)e^{x^Tw\phi(y)-A(x^Tw)}$ We consider the cost function: $L(w) = -\sum_{n=1}^{N} lnp(y_n|x_n^T w) = -\sum_{n=1}^{N} ln(h(y_n)) +$ $x_n^T w \phi(y_n) - A(x_n^T w)$

We use the fact that $\frac{dA(\eta)}{d\rho} = \mathbb{E}[\phi(y)] = g^{-1}(\eta)$ to obtain

 $\nabla_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} x_n \phi(y_n) - x_n g^{-1}(x_n^T \mathbf{w})$ matrix notation: $\nabla \mathcal{L}(w) = X^T[q^{-1}(Xw) - \phi(y)] = 0$

Example for Logistic reg: $\nabla \mathcal{L}(w) = X^T [\sigma(Xw)$ y = 0

Curse of Dimensionality

Bayes Class: $f_*(x) = \mathbb{1}_{\{P(y=1|x)>1/2\}}$

Claim 1 : number of sample = $\mathcal{O}(N^D)$ the risk to stau constant.

Claim 2: As dimension increases the choice of KNN becomes random.

Lemma: $\mathbb{E}_{S_{train}}[\mathcal{L}(t_{S_{train}})] \leq 2\mathcal{L}(t_*) + 4c\sqrt{d}N^{-\frac{1}{d+1}}$

 $w^* = argmin_w \sum_{n=1}^{N} [a - y_n x_n^T w]_+ + \frac{\lambda}{2} ||w||^2$ using Hinge loss: $[z]_+ := max\{0,z\}$

 $\mathcal{L}(w) = max_{\alpha}G(w, \alpha)$ So since $[1 - y_nx_n^Tw]_+ =$ $\max_{\alpha_n \in [0,1]} \alpha_n (1 - y_n x_n^T w)$ then: $G(w,\alpha)$:

 $\min_{w} \max_{\alpha \in [0,1]^N} \sum_{n=1}^N \alpha_n (1 - y_n x_n^T w) + \frac{\lambda}{2} ||w||^2 =$ $min_w max_\alpha G(w,\alpha)$

In general: $max_x min_u f(x,y) \le min_u max_x f(x,y)$

Minimax Theorem: if f(x,y) is convex in x and concave in y then

 $max_{II}min_{x}f(x,y) = min_{x}max_{II}f(x,y)$ Y = diag(y)

$$\begin{array}{ll} \max_{\alpha \in [0,1]^N} & \sum_{n=1}^N & \alpha_n (1 & -\frac{1}{\lambda} y_n x_n^T X^T Y \alpha) \\ \frac{\lambda}{2} | \left\| \frac{1}{\lambda} X^T Y \alpha \right| \right\| & + \\ = \\ \max_{\alpha \in [0,1]^N} \alpha^T 1 - \frac{1}{2\lambda} \alpha^T Y X X^T Y \alpha \end{array} + \left\{ \begin{array}{ll} \pi_k > 0 \forall k \text{ and } \sum_{k=1}^K \pi_k = 1 \\ \text{Gaussian Mixture Model} \\ \text{Prior: } p(z_n = k) = \pi_k \\ \text{Marginal likelihood}(\theta) : p(z_n = k) = \pi_k \\ \text{Marginal likelihood}($$

The dual is kenerlized with $K = XX^T$ and the solution $\alpha = \text{nnz}$ only for the training ex that are important for the decision boundary (inside margin, wrong side or on margin)

Coordinate descent

1. init $\alpha^{(0)} \in \mathcal{R}^N$

2. For t = 0:maxIter do: sample a rand coord $n \in$

optimize g w.r.t n

$$u^* \leftarrow armin_{u \in \mathcal{R}} g(\alpha_1^{(t)}, ..., u, ..., \alpha_N^{(t)})$$
update $\alpha_n^{(t+1)} = u^*, \alpha_{n'}^{(t+1)} = \alpha_{n'}^{(t)} \forall n \neq n'$
Kernel Trick

Solving Ridge:
$$\mathcal{O}(D^3 + ND^2)$$
, via kernel: $w^* = X^T \alpha^*$ with $\alpha^* = (XX^T + \lambda I_N)^{-1} y$: $\mathcal{O}(N^3 + DN^2)$

Representer Theorem: for all problem of the form $\min_{w} \sum_{n=1}^{N} \mathcal{L}_n(x_n^T w, y_n) + \frac{\lambda}{2} ||w||^2$ there exist α^* s.t.

Kernel Ridge: $\alpha^* = argmax_{\alpha} - \frac{1}{2}\alpha^T(XX^T + \lambda I_N)\alpha +$

Kernels A kernel function is associated with a feature map: $\kappa(x,x') = \phi(x)' \phi(x')$

Prove f=kernel: (1) find the associated feature map $\psi(x)$ show that $K(x,y) = \psi(x)^T \psi(y)$. (2) Verify Menger's therorem

Menger's Theorem:

1. K should be symmetric $\kappa(x,x') = \kappa(x',x)$

2. K should be positive semi-definite

NB: positive-definite functions give rise to ∞-dim feature space.

New prediction with kernel:

 $\hat{y} = w^{*T} x_{test} = \sum_{n=1}^{N} k(x_{test}, x_n) \alpha_n$ (here for id kernel) cost: $\mathcal{O}(N*kernel)$ Week8

$$\hat{R}(\mu) = \hat{R}(\mu_1, ..., \mu_k) = \sum_{i=1}^n \min_{j \in \{1, ..., k\}} ||x_i - \mu_j||_2^2$$

$$\hat{\mu} = \operatorname{argmin} \hat{R}(\mu)$$

(Lloud's heuristic):

Initialize cluster centers $\mu^{(0)} = [\mu_1^{(0)}, \dots, \mu_k^{(0)}]$ While not converged

$$z_i \leftarrow arg \min_{j \in \{1, ..., k\}} ||x_i - \mu_j^{(t-1)}||_2^2; \ \mu_j^{(t)} \leftarrow \frac{1}{n_j} \sum_{i: z_i = j} x_i$$

NB: kmean enforces spherical clusters.

Gaussian Mixture Models No spherical cluster:

$$p(X|\mu,\Sigma,z) = \prod_{n=1}^{N} \prod_{k=1}^{K} [\mathcal{N}(x_n|\mu_k,\Sigma_k)]^{z_{nk}}$$

No hard assignment: $p(z_n = k) = \pi_k$ where

Gaussian Mixture Models in action Prior: $p(z_n = k) = \pi_k$

Marginal likelihood(θ): $p(x_n | \theta)$

 $p(x_n|\theta) = \sum_{k=1}^K p(x_n, z_n = k|\theta) =$

 $\sum_{k=1}^{K} p(z_n = k | \theta) p(x_n | z_n = k, \theta) =$ $\sum_{k=1}^{K} \pi_k \mathcal{N}(x_n | \mu_k, \Sigma_k)$

Without latent variable: $\mathcal{O}(N)$ parameters, after marginalization: $\mathcal{O}(D^2K)$ (ass $D,K \ll N$)

 $Q(\theta, \theta^{(t)}) = \mathbb{E}_{z|x,\tilde{\theta}}[log P(x,z|\theta)|x,\tilde{\theta}]$ we want $O(\theta^{(t)}, \theta^{(t-1)}) > O(\theta^{(t-1)}, \theta^{(t-1)})$

(1) E-step: For each k and n calculate $q_{kn}^{(t)}$

$$q_{kn}^{(t)} = P(z_n = k|x_n, \theta^{(t-1)}) = \frac{\pi_k^{(t-1)} \mathcal{N}(x_n | \mu_k^{(t-1)}, \Sigma_k^{(t-1)})}{\sum_{k=1}^K \pi_k^{(t-1)} \mathcal{N}(x_n | \mu_k^{(t-1)}, \Sigma_k^{(t-1)})} = \frac{1}{2} \left(\frac{\pi_k^{(t-1)} \mathcal{N}(x_n | \mu_k^{(t-1)}, \Sigma_k^{(t-1)})}{\pi_k^{(t-1)} \mathcal{N}(x_n | \mu_k^{(t-1)}, \Sigma_k^{(t-1)})} \right)$$

(2) Compute marginal likelihood (cost): $\mathcal{L}(\theta_t) =$ $\sum_{n=1}^{N} log \sum_{k=1}^{K} \pi_k^{(t-1)} \mathcal{N}(x_n | \mu_k^{(t-1)}, \Sigma_k^{(t-1)})$ (3) M-step: Fit clusters to weighted data points

(MLE marginal):

 $\theta^{(t)} = aramax_{\theta} O(\theta | \theta^{(t-1)})$

$$\pi_k^{(t)} \leftarrow \frac{1}{n} \sum_{n=1}^n q_{kn}^{(t)}; \mu_k^{(t)} \leftarrow \frac{\sum_{n=1}^n q_{kn}^{(t)} x_n}{\sum_{n=1}^n q_{kn}^{(t)}}$$
$$\sum_k^{(t)} \leftarrow \frac{\sum_{n=1}^n q_{kn}^{(t)} (x_n - \mu_k^{(t)}) (x_n - \mu_k^{(t)})^T}{\sum_{n=1}^n q_{kn}^{(t)}}$$

SVD: $X = USV^T$ U,T unitary. Unitary matrices rotate, don't change norm.

Lemma: For any DxN matrix X and any DxN rank-K matrix \hat{X} :

$$\|X - \hat{X}\|_{F}^{2} \ge \|X - U_{k}U_{k}^{T}X\|_{F}^{2} = \sum_{i \ge K+1} s_{i}^{2}$$

where U_k is the D xK matrix consisting of the first

K columns of U We project on the first left singvecs of U.

We have: $U_K U_K^T X = U_K U_K^T U S V^T = U S^{(K)} V^T$.

Weeklu Matrix Factorization
$$min_{W,Z}\mathcal{L}(W,Z) = \frac{1}{2|\Omega|} \sum_{(d,n)\in\Omega} [x_{dn} - (WZ^T)_{dn}]^2$$
:

 $W \in \mathcal{R}^{DxK}$ $Z \in \mathcal{R}^{NxK}$ K << D.Nlarge K facilitates overfitting.

Regul: $min_{W,Z}\mathcal{L}(W, Z) = \frac{1}{2} \sum_{(d,u)\in\Omega} [x_{dn} - y_{du}]$

 $(WZ^{T})_{dn}]^{2} + \frac{\lambda_{w}}{2} ||W||_{F}^{2} + \frac{\lambda_{z}}{2} ||Z||_{F}^{2}$ **Identifiable**: a model is identifiable iff $\theta_1 \neq \theta_2 \Rightarrow$ $P_{\theta_1} \neq P_{\theta_2}$ i.e. 2 sets of different parameters imply 2 different proba distribs.

Alternating Least Squares 1. $Z^{*T} = (W^T W + \lambda_z I_K)^{-1} W^T X$ 2. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 2. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 3. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 3. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 3. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 4. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 5. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 6. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 7. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 7. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 7. $W^{*T} = (Z^T Z + I_K)^{-1} W^T X$ 9. $W^{*T} = (Z^T Z + I_K)$ $\lambda_W I_K$)⁻¹ $Z^T X^T$ cost as ridge regression per iteration.

 $\mathcal{O}(ND^2+D^3)$ **Handling Text**

Vocabulary: $\mathcal{V} = \{w_1..w_D\}$

Context: $C = \{w_1...w_N\}$ (document, paragraph, window)

GloVe, (Global Vectors)unsup learn algo for obtaining vect repr for words; achieved by mapping words into space where the distance between w is prop to semantic similarity. train perf on co-occ. Combines matfact and local context window me-

NB: we replace $x_{dn} := log(n_{dn})$ and obtain the count matrix X

$$\min_{W,Z} \mathcal{L}(W, Z) = \frac{1}{2} \sum_{(d,n) \in \Omega} f_{dn}[x_{dn}]$$
$$(WZ^T)_{dn}]^2; f_{dn} = \min\{1, (\frac{n_{dn}}{n_{max}})^{\alpha}\}, \alpha \in [0,1]$$

row(W) = rep(voc-w); row(Z) = rep(cont-w)

SG; Word2vec: uses logreg to sep real wp (w_d, w_n) from fake. Unsupervised. Streams through text, on the fly. Use neg-samplings; (NB: word2vec approximates glove)

textcolorpurpleLanguage models: predict next word, multi-class(soft-max loss) Problem: you need to do the dot product with all words of the vocabularu.

FastText: Supervised sent classification:uses mat-fact: (sup and unsup aspects): Given $\operatorname{sent} s_n = (w_1..w_m)$ let $x_n \in \mathbb{R}^V$ be its bow re-

$$min_{W,Z}\mathcal{L}(W,Z) = \sum_{s_n asentence} f(y_n W Z^T x_n)$$

 $Z^T x_n$ sums the words vectors (columns of Z) x to create a sentence rep: W classifies it. $f = \log \log s$

Neural Networks Repr Lemma: for subsetof cont diff fctions, NN approx 1hidlayer, n nodes approx well in av any fction on bounded domain: $\int_{|x|< r} (f(x) - f_n(x))^2 dx \le \frac{(2Cr)^2}{n}$

; $w_{i,i}^{(l)}$ = edge from node i in layer l-1 to node j in layer I; $z_i^{(l)} = \sum_i w_{i,i}^{(l)} x_i^{(l-1)} + b_i^{(l)}$; $\phi(x) = \frac{1}{1 + e^{-x}}$; Rect in 3d: $g(x_1, x_2) = \phi(w(x_1 + a_1)) - \phi(w(x_1 - b_1)) +$ $\phi(w(x_2+a_2)) \stackrel{?}{\rightarrow} \phi(w(x_2-b_2)) \Rightarrow \phi(w(g(x_1,x_2)-3/2))$ **Backpropagation**

$$\begin{array}{l} \frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} = \delta_j^{(l)} x_i^{(l-1)} \text{ Forward} \\ \text{pass: } z^{(l)} = (W^{(l)})^T x^{(l-1)} + b^{(l)},; x^{(l)} = \phi(z^{(l)}); \\ \delta_j^{(l)} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \text{ Bckward: Set } \delta^{(l+1)} = -2(y_n - x^{(l+1)})\phi'(z^{(l+1)}); \quad \delta^{(l)} = (W^{(l+1)}\delta^{(l+1)}) * \phi'z^{(l)}; \\ \text{Final comp: } \frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \delta_j^{(l)} x_i^{(l-1)}, \frac{\partial \mathcal{L}_n}{\partial b_i^{(l)}} = \delta_j^{(l)} \end{array}$$

Regularization ⇒ weight decay.

Augmentation

Cropping, resizing, rotating, compressing (PCA), addition of noise

 $p_i^{(l)}$ = proba to "keep"the node i in layer l

Pred: (1) generate K subnets, average output (2) sca-

le the output of node i at layer l by $p_i^{(l)}$. Advantage: (1) limits overfitting (2) advantage of

model averaging (bagging, ensemble methods) CNN:consider input of size: N1xN2.run filter withoug padding, outputsize $(N_1 + 2(K_1 - 1))x(N_2 +$

2(K₂-1)) (valid padding)

Week12

White Box attack

Have access to the details of the algo i.e. we need access to the NN that implemented the prediction in order to compute the gradients.

$$\tilde{x} = x - \epsilon h(x) \frac{\nabla_x g(x)}{\|\nabla_x g(x)\|_2}$$

where $g(x) = \hat{p}(y|x), g(x) = \text{ground truth.}$

Black Box attack Cannot look inside the box. Week13

Goal: Recognize independence relationship given a Bayes Net. The graph needs to be acyclic to correspond to a valid factorization.

Lemma: X are cond indep of Y cond on Z if X and Y are D-separated by Z. (converse !nec true)

If X and Y are not D-separated, then there is distrib compat with the bayes net s.t. X and Y are cond dep given Z.

D-sep \Leftrightarrow every path from X to Y is blocked by Z Blocked: a path from X to Y is blocked by Z iff it contains a variable/node s.t either (1) a node $\in \mathbb{Z}$ and is h2t or t2t (2)a node is h2h and neither this node nor any of its descendants are in Z.

NB: 2 variables are indep (given the empty set) if

there is a h2h node between them.

NB: a h2h node blocks if none of its children or himself are in Z.

Markov Blanket: $X_i \perp X_i \mid blanket(X_i)$

Factor graphs

 $\sum_{z} f(z,..) = \sum_{z} \prod_{k=1}^{K} g_{k}(z,..) = \prod_{k=1}^{K} \sum_{z} g_{k}(z,..)$ **NB**:Bayes net is a tree \Rightarrow F graph is a tree. converse

NB: a cyclic graph cannot use the message passing algorithm!

Message passing Algorithm

Initialization:variable node $\mu(x) = 1$, function node: $\mu(x) = f(x)$

f Node rule: $\mu(x) = \sum_{x_1...x_j} h(x,x_1,...,x_j) \prod_{j=1}^J \mu_j(x_j) \mathbf{x}$ is the higher variable the factor is attached to

Variable Node rule: $\mu(x) = \prod_{k=1}^{K} \mu_k(x)$ all the μ_k depend on the variable x

Max complexity = $\mathcal{O}(|\mathcal{X}|^p)$, $p = \max_n deg(f_n)$