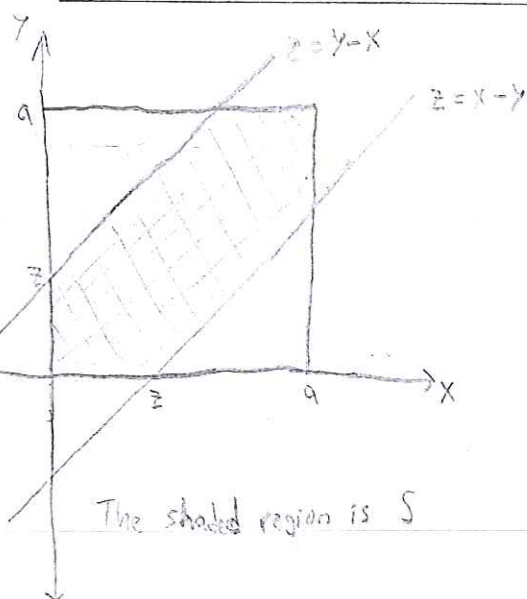


Use this space to show your work for Problem 1 only.



$$F_2(z) = P(z \leq z) = P(|x-y| \leq z)$$

$$= P(-z \leq x-y \leq z)$$

$$= \int_S \frac{1}{a^2} dS = \frac{1}{a^2} \int_S dS = \frac{1}{a^2} (\text{Area of } S)$$

$$= \frac{1}{a^2} \cdot 2 \left(\frac{a^2}{2} - \frac{(a-z)^2}{2} \right) = \frac{1}{a^2} (a^2 - (a-z)^2)$$

$$= \frac{1}{a^2} (2az - z^2) = \frac{2z}{a} - \frac{z^2}{a^2} \quad \text{when } 0 \leq z \leq a$$

when $z < 0$, $P(z \leq z) = 0$

when $z > a$, $P(z \leq z) = 1$

$$\Rightarrow F_2(z) = \begin{cases} 0 & , z < 0 \\ \frac{2z}{a} - \frac{z^2}{a^2} & , 0 \leq z \leq a \\ 1 & , \text{otherwise} \end{cases}$$

P1. (5 points)

Let X and Y be two independent drawings from the uniform distribution on $[0, a]$, with $a > 0$ a given constant. Let $Z = |X - Y|$ be the distance between the two points. Find the CDF $F_Z(z)$. Show your work by drawing a figure that explains how you calculate $F_Z(z)$.

$$F_Z(z) = \begin{cases} 0 & , \quad z < 0 \\ \frac{2z}{a} - \frac{z^2}{a^2} & , \quad 0 \leq z \leq a \\ 1 & , \quad z > a \end{cases}$$

Use this space to show your work for Problem 2 only.

$$X_i = \begin{cases} 1, & \text{if the } i\text{th digit drawn is 7.} \\ 0, & \text{otherwise.} \end{cases}$$

$$\Rightarrow Y = X_1 + X_2 + X_3 + \dots + X_{10000}$$

$$\Rightarrow P(Y < 968) = P(X_1 + X_2 + X_3 + \dots + X_{10000} < 968)$$

$$\Rightarrow \text{mean of } X_1 + X_2 + X_3 + \dots + X_{10000} = 10000 * (\text{mean of } X_i) = 10000 * \frac{1}{10} = 1000 = E[Y]$$

$$\Rightarrow \text{variance of } X_1 + X_2 + X_3 + \dots + X_{10000} = 10000 * \text{var}(X_i) = 10000 * \left(\frac{1}{10} - \frac{1}{10}^2\right) = 900 = \text{var}(Y)$$

\Rightarrow By CLT,

$$P(Y < 968) = P\left(\frac{X_1 + X_2 + X_3 + X_4 + \dots + X_{10000} - 1000}{\sqrt{900}} < \frac{968 - 1000}{\sqrt{900}}\right) = \Phi\left(\frac{968 - 1000}{30}\right) = \Phi\left(\frac{-32}{30}\right)$$

$$= 1 - \Phi\left(\frac{16}{15}\right) = 1 - \Phi(1.0\bar{6}) \simeq 1 - \Phi(1.07) = 1 - 0.8577 = \underline{\underline{0.1423}}$$

Note: we could use $\Phi\left(\frac{967.5 - 1000}{30}\right)$ (or even $\Phi\left(\frac{967.001 - 1000}{30}\right)$, etc.) for a better estimation, as the case is discrete.

P2. (5 points)

Let Y be the number of times the digit 7 appears among $n = 10,000$ digits drawn from the set $\{0, 1, 2, 3, 4, 5, 6, 7, 8, 9\}$ independently at random (using the uniform distribution). Use the Central Limit Theorem (CLT) to approximate the probability $P(Y < 968)$ in terms of the function $\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$. Use the table given below to give a numerical estimate of probability. (Show clearly the expectation and variance of Y . Show how you standardize the Y to obtain the CLT approximation. An answer without sufficient justification will receive no credit even if correct.)

$$P(Y < 968) \approx 0.1423$$

Table 1: Standard Normal Table $\Phi(x)$.

[illegible]

Use this space to show your work for Problem 3 only.

a) $H_0: \theta = 0, H_1: \theta = 1$

$$p_0(0) \cdot p_{Y|0}(y|0) \sum_{H_0} p_0(1) \cdot p_{Y|1}(y|1)$$

$$\frac{1}{2} \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-A)^2/2\sigma^2} \sum_{H_1} \frac{1}{2} \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(y+A)^2/2\sigma^2}$$

$$\Rightarrow e^{-(y-A)^2/2\sigma^2} \sum_{H_1} e^{-(y+A)^2/2\sigma^2}$$

$$\Rightarrow -(y-A)^2 \sum_{H_1} -(y+A)^2$$

$$\Rightarrow y^2 - 2yA + A^2 \sum_{H_1} y^2 + 2yA + A^2$$

$$\Rightarrow -y \sum_{H_1} y \Rightarrow 0 \sum_{H_1} y$$

$$\text{Decision rule} = \begin{cases} \theta = 0 & \text{if } Y > 0 \\ \theta = 1 & \text{if } Y < 0 \end{cases}$$

$$P_e = p_0(0) \cdot P(Y < 0 | \theta = 0) + p_0(1) P(Y > 0 | \theta = 1)$$

$$= \frac{1}{2} P(Z < -A) + \frac{1}{2} P(Z > A)$$

$$= \frac{1}{2} P(Z < -A) + \frac{1}{2} (1 - P(Z < A))$$

$$= 1 - P(Z < A) \quad \text{with } A=4, \sigma^2=4, \text{ using standardization:}$$

$$P_e = 1 - P(Z < 4) = 1 - P\left(\frac{Z}{2} < 2\right) = 1 - \Phi(2) = 0.0228$$

$$P(Z < -A) = 1 - P(Z < A)$$

As Z is normal with $\mu=0$

b) $\hat{\theta} = E[\theta] + \frac{\text{cov}(\theta, x)}{\text{var}(x)} (x - E[x])$, $E[\theta] = \frac{4+10}{2} = 7$, $E[x] = E[\theta + w] = E[\theta] + E[w] = \frac{4+10}{2} + \frac{1+1}{2} = 7$

$\text{var}(x) = \text{var}(\theta + w) = \text{var}(\theta) + \text{var}(w)$ as θ and w are independent.

$$\Rightarrow \text{var}(x) = \frac{(10-4)^2}{12} + \frac{(1-(-1))^2}{12} = 3 + \frac{4}{12} = \frac{10}{3}$$

(for future use, note $\text{var}(\theta) = \frac{(10-4)^2}{12} = 3$)

$$\text{cov}(\theta, x) = E[\theta x] - E[\theta]E[x] = E[\theta(\theta + w)] - 49 = E[\theta^2] + E[\theta w] - 49$$

$$= E[\theta^2] + E[\theta]E[w] - 49 \quad (\text{as } \theta \text{ and } w \text{ are independent})$$

$$= E[\theta^2] - 49 \quad (\text{as } E[w] = \frac{1+1}{2} = 0)$$

$$E[\theta^2] = \int_4^{10} \theta^2 f_\theta(\theta) d\theta = \int_4^{10} \theta^2 \frac{1}{6} d\theta = \left[\frac{\theta^3}{18} \right]_{\theta=4}^{\theta=10} = \frac{976}{18} = 52$$

$$\Rightarrow \text{cov}(\theta, x) = 52 - 49 = 3$$

$$\Rightarrow \hat{\theta} = 7 + \frac{3}{10} (x - 7) = \frac{3x}{10} + \frac{7}{10} = \frac{3x+7}{10}$$

$$\text{MSE of LLMS} = (1 - \rho^2) \sigma_\theta^2 = \left(1 - \left(\frac{\text{cov}(\theta, x)}{\sigma_\theta \sigma_x}\right)^2\right) \sigma_\theta^2 = \left(1 - \left(\frac{3^2}{3 \cdot \frac{10}{3}}\right)\right) 3 = \left(1 - \frac{3}{10}\right) 3 = \frac{3}{10} = 0.3$$

P3. (10 points) The two parts of this problem are independent.

(a) (5 pts) Consider a binary communication system in which the message is modeled as an unknown parameter Θ that takes values as

$$p_{\Theta}(\theta) = \begin{cases} \frac{1}{2}, & \theta = 0; \\ \frac{1}{2}, & \theta = 1. \end{cases}$$

The transmitted signal X is a function of the message Θ :

$$X = \begin{cases} +A, & \text{if } \Theta = 0; \\ -A, & \text{if } \Theta = 1, \end{cases}$$

where $A > 0$ is a constant. The received signal is given by

$$Y = X + Z$$

where $Z \sim N(0, \sigma^2)$ represents channel noise. Let $\hat{\Theta}$ denote the receiver's estimate of Θ and $P_e = P(\hat{\Theta} \neq \Theta)$ the probability of error. Construct a Bayesian estimator that minimizes P_e . Give the decision rule in as simple a form as possible. Calculate the resulting P_e as a function of $\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$, the CDF of standard normal. Evaluate the result for $A = 4$, $\sigma^2 = 4$. (Use the table in Problem 2.)

$$P_e = 0.0228$$

(b) (5 pts) Let Θ be uniformly distributed over the interval $[4, 10]$ and let $X = \Theta + W$ be a measurement of Θ where the measurement noise W is uniformly distributed over $[-1, 1]$. Find the Linear Least Mean Squares (LLMS) estimator of Θ and the associated mean square error MSE.

$$\hat{\Theta} = \frac{3x+7}{10}$$

$$\text{MSE} = 0.3$$

Use this space to show your work for P4 only.

$$a) f_{x_1, x_2, x_3, \dots, x_n}(x_1, x_2, x_3, \dots, x_n) = (\theta e^{-\theta x_1}) (\theta e^{-\theta x_2}) \dots (\theta e^{-\theta x_n})$$

$$= \theta^n e^{-\theta(x_1 + x_2 + \dots + x_n)}$$

To maximize this, we take the derivative with respect to θ ,

$$n \theta^{n-1} e^{-\theta(x_1 + x_2 + x_3 + \dots + x_n)} - (x_1 + x_2 + x_3 + \dots + x_n) e^{-\theta(x_1 + x_2 + x_3 + \dots + x_n)} \theta^n = 0$$

$$\Rightarrow n \theta^{n-1} = (x_1 + x_2 + x_3 + \dots + x_n) \theta^n \Rightarrow n = (x_1 + x_2 + x_3 + \dots + x_n) \theta$$

$$\Rightarrow \hat{\theta}_n = \frac{n}{(x_1 + x_2 + \dots + x_n)} = \frac{n}{\sum_{i=1}^n x_i}$$

$$\lim_{n \rightarrow \infty} \hat{\theta}_n = \lim_{n \rightarrow \infty} \left(\frac{n}{\sum_{i=1}^n x_i} \right) = \lim_{n \rightarrow \infty} \left(\frac{\frac{1}{n} \sum_{i=1}^n x_i}{1} \right)^{-1} = \left(\lim_{n \rightarrow \infty} \frac{\frac{1}{n} \sum_{i=1}^n x_i}{1} \right)^{-1}$$

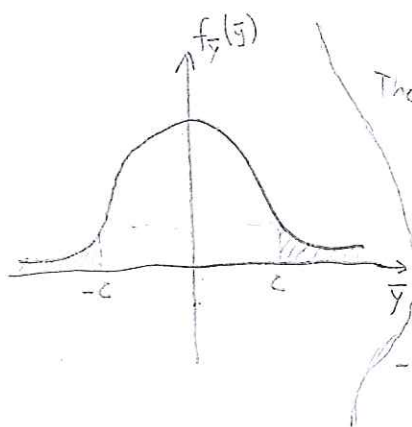
By weak law of large numbers,

$$\left(\lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n x_i}{n} \right)^{-1} \stackrel{(1)}{=} (E[x])^{-1} = \frac{1}{E[x]} = \theta \Rightarrow \text{consistent}$$

$$\hookrightarrow \left(\lim_{n \rightarrow \infty} P(|\hat{\theta}_n - \theta| > \epsilon) = 0 \text{ for all } \epsilon \right)$$

In step (1), $M_n = E[x]$ is used

b) Let \bar{X} be the mean for 100 samples and $\bar{y} = \frac{\bar{X} - 5}{1/10}$. (\bar{y} is standard normal, because $\mu = 5$, $\sigma = \frac{1}{10}$ for M_{100})



The probability of shaded area is to be 5% (2.5% each)

$$\Phi^{-1}(1 - 0.025) = \Phi^{-1}(0.975) = 1.96 = c$$

\Rightarrow Acceptance region: $-1.96 \leq \bar{y} \leq 1.96$

$$-1.96 \leq \frac{\bar{X} - 5}{1/10} \leq 1.96 \Rightarrow -0.196 \leq \bar{X} - 5 \leq 0.196 \Rightarrow \bar{X} \in [4.804, 5.196]$$

P4. (10 points) The two parts of this problem are independent.

(a) (5 pts) A certain type of light bulb has a lifetime X that is exponential with parameter $\theta > 0$: $f_X(x; \theta) = \theta e^{-\theta x}$, $x \geq 0$. Let x_1, x_2, \dots, x_n be the measured lifetimes of n such light bulbs. We model these measurements as samples of random variables (X_1, \dots, X_n) with a joint PDF such that

$$f_{X_1, \dots, X_n}(x_1, \dots, x_n; \theta) = \prod_{i=1}^n f_X(x_i; \theta).$$

Determine the ML (maximum-likelihood) estimator $\hat{\theta}_n(X_1, \dots, X_n)$. Is the estimator consistent? Justify.

$$\hat{\theta}_n(X_1, \dots, X_n) = \frac{n}{\sum_{i=1}^n x_i}$$

Consistent? Yes, justification is given in the solution page.

(b) (5 pts) Let X be a normal random variable with mean μ and unit variance. We want to test the hypothesis $H_0: \mu = 5$ at the 5% significance level, using $n = 100$ independent samples of X . What is the range of values of the sample mean $M_n = (X_1 + \dots + X_n)/n$ for which the hypothesis is accepted? Show your work in detail. In particular show how you are using the CDF of the unit normal $\Phi(x)$.

$$\text{Range of } M_n = [4.804, 5.196]$$