## Math255 Probability and Statistics Midterm 2 1 Dec. 2016

**Problem 1.** [5 pts] Let  $X_A$ ,  $X_E$ ,  $X_K$  be the rainfalls (in cms) in a given year to Ankara, Eskişehir, and Konya provinces, respectively. Suppose the rainfalls are individually Gaussian random variables  $X_A \sim N(40,100)$ ,  $X_E \sim N(50,144)$ ,  $X_K \sim N(30,81)$ , with pairwise correlations  $\rho(X_A,X_E) = \rho(X_A,X_K) = \rho(X_E,X_K) = 0.5$ . Compute var(X) where  $X = X_A + X_E + X_K$  is the total rainfall in all three provinces. (Express your answer as a real number.)

**Solution.** Recall that, for any two RVs X, Y, var(X + Y) = var(X) + 2 cov(X, Y) + var(Y) and  $cov(X, Y) = \sqrt{var(X) var(Y)} \rho(X, Y)$ . Thus,

$$var(X) = var(X_A) + var(X_E) + var(X_K) + 2 cov(X_A, X_E) + 2 cov(X_A, X_K) + 2 cov(X_E, X_K)$$

$$= 100 + 144 + 81 + 2(\sqrt{100 \cdot 144}) \cdot 0.5 + 2(\sqrt{100 \cdot 81}) \cdot 0.5 + 2(\sqrt{144 \cdot 81}) \cdot 0.5$$

$$= 325 + 120 + 90 + 108 = 643.$$

**Problem 2.** [5 pts] Let X be exponentially distributed with parameter  $\mu > 0$ ,

$$f_X(x) = \mu e^{-\mu x}, \quad x \ge 0;$$

and Y be Poisson with parameter x conditional on X = x,

$$p_{Y|X}(k|x) = \frac{x^k e^{-x}}{k!}, \quad k = 0, 1, \dots$$

Evaluate  $\mathbf{E}[X|Y=1]$  for  $\mu=2$ . (A real number is required as an answer.)

**Solution.** By definition,

$$\mathbf{E}[X|Y=1] = \int x f_{X|Y}(x|1) dx = \int x \frac{f_X(x) p_{Y|X}(1|x)}{p_Y(1)} dx = \int_0^\infty x \frac{\mu e^{-\mu x} x e^{-x}}{p_Y(1)} dx,$$

where

$$p_Y(1) = \int f_X(x) p_{Y|X}(1|x) dx = \int_0^\infty \mu e^{-\mu x} x e^{-x} dx = \mu \int_0^\infty x e^{-(1+\mu)x} dx = \mu (1+\mu)^{-2}.$$

Thus,

$$\mathbf{E}[X|Y=1] = \int_0^\infty x \, \frac{\mu e^{-\mu x} x e^{-x}}{\mu (1+\mu)^{-2}} dx = \int_0^\infty \frac{x^2 e^{-(1+\mu)x}}{(1+\mu)^{-2}} dx = 2(1+\mu)^{-1}.$$

For  $\mu = 2$ , we have  $\mathbf{E}[X|Y = 1] = 2/3$ .

**Problem 3.** [5 pts] Let (X,Y) have the joint distribution

$$f_{X,Y}(x,y) = \frac{1}{2\pi\sigma_1\sigma_2} \exp\left[-\frac{(x-\mu_1)^2}{2\sigma_1^2} - \frac{(y-\mu_2)^2}{2\sigma_2^2}\right], -\infty < x, y < \infty.$$

Let Z=2X+Y. Compute the CDF  $F_Z(z)$  in terms of the CDF of unit normal,  $\Phi(u)=\int_{-\infty}^u \frac{1}{\sqrt{2\pi}}e^{-\frac{v^2}{2}}dv$ , and the parameters  $\mu_1, \mu_2, \sigma_1$  and  $\sigma_2$ .

**Solution.** Since  $f_{X,Y}(x,y)$  has product form, we understand that X and Y are independent with  $X \sim N(\mu_1, \sigma_1^2)$  and  $Y \sim N(\mu_2, \sigma_2^2)$ . Thus, from class notes, Z = 2X + Y is Gaussian with mean  $\mu_z = 2\mu_1 + \mu_2$  and variance  $\sigma_z^2 = 4\sigma_1^2 + \sigma_2^2$ . Hence,

$$F_Z(z) = P(Z \le z) = P\left(\frac{Z - \mu_z}{\sigma_z} \le \frac{z - \mu_z}{\sigma_z}\right) = \Phi\left(\frac{z - \mu_z}{\sigma_z}\right) = \Phi\left(\frac{z - 2\mu_1 - \mu_2}{\sqrt{4\sigma_1^2 + \sigma_2^2}}\right).$$

**Problem 4.** [5 pts] Let X be uniform on  $[0, \frac{\pi}{2}]$ , i.e.,  $f_X(x) = \frac{2}{\pi}$ ,  $0 \le x \le \frac{\pi}{2}$ , and let  $Y = \sin(X)$ . Compute  $f_Y(y)$  at  $y = \sqrt{3}/2$ . (A real number is required as an answer, with full justification.)

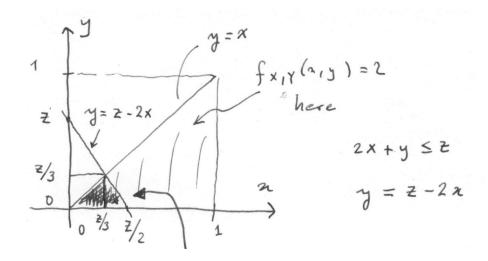
**Solution.** Let  $y = g(x) \stackrel{\triangle}{=} \sin(x)$ . This function is one-to-one and increasing over  $[0, \pi/2]$ , which is the range of the RV X. Hence, we have

$$f_Y(y) = \frac{1}{|g'(x)|} f_X(x) = \frac{2/\pi}{\cos(x)} = \frac{2/\pi}{\sqrt{1 - y^2}}, \quad 0 \le x \le \pi/2.$$

Thus,  $f_Y(\sqrt{3}/2) = 4/\pi$ .

**Problem 5.** [5 pts] Let (X,Y) be jointly distributed uniformly over the triangle in the (x,y) plane with corners at (0,0), (1,1), (1,0). Let  $Z \stackrel{\Delta}{=} 2X + Y$  and compute  $f_Z(z)$  for z=1.

**Solution.** The joint density equals 2 over the region  $0 \le y \le x \le 1$ , and 0 elsewhere. The range of Z = 2X + Y is 0 to 3. We will compute  $f_Z(1)$  by first computing  $F_Z(z)$  as a function of z and then computing  $f_Z(1) = \mathrm{d}F(1)/\mathrm{d}z$ . The figure shows how to calculate  $F_Z(z)$  as a function of z.



 $F_Z(z)$  equals the shaded area times the density (2) for  $0 \le z \le 2$ , in other words,

$$F_Z(z) = \frac{1}{2} \times \frac{z}{2} \times \frac{z}{3} \times 2 = \frac{z^2}{6}, \quad 0 \le z \le 2.$$

Differentiating, we get

$$f_Z(1) = (2z/6) \Big|_{z=1} = 1/3.$$

**Problem 6.** [5 pts] In 100 independent tosses of a fair die (with faces numbered 1 through 6), let  $X_i$  be the outcome (number on the up face) in the *i*th trial and  $M = \frac{1}{100} \sum_{i=1}^{100} X_i$  be the empirical mean. Estimate the P(3.4 < M < 3.6) using the central limit theorem and the table below (use linear interpolation if the number you are seeking is not in the table). A numerical answer is required with 4 digit precision. Show details on the solution page.

**Solution.** Observe that  $\mathbf{E}[X_i] = 3.5$  and  $\text{var}(X_i) = \sum_{i=1}^6 \frac{1}{6}(i - \frac{1}{6}\sum_{i=1}^6 i)^2 = 35/12$ . Thus,

 $\mathbf{E}[M] = \mathbf{E}[X_i] = 3.5$  and  $\text{var}(M) = \text{var}(X_i)/100 = 35/1200$ . Now, we can use the CLT.

$$\begin{split} P(3.4 < M < 3.6) &= P\bigg(\frac{3.4 - 3.5}{\sqrt{35/1200}} < \frac{M - 3.5}{\sqrt{35/1200}} < \frac{3.6 - 3.5}{\sqrt{35/1200}}\bigg) \\ &= P\bigg(\frac{-0.1}{\sqrt{35/1200}} < \frac{M - 3.5}{\sqrt{35/1200}} < \frac{0.1}{\sqrt{35/1200}}\bigg) \\ &= P\bigg(-0.5855 < \frac{M - 3.5}{\sqrt{35/1200}} < 0.5855\bigg) \\ &\approx 2\,\Phi\bigg(\sqrt{\frac{12}{35}}\bigg) - 1 = 2\,\Phi(0.5855) - 1 = 2\times0.7209 - 1 = 0.4418 \end{split}$$