Combining infinite sets of experts

Yoav Freund

January 20, 2025

Freund: Predicting a binary Sequence almost as well the the optimal biased coin.

Risannen: Fisher Information and Stochastic Complexity.

Review

Review

The Universal prediction machine

Review

The Universal prediction machine

The biased coins set of experts
Laplace Approximation
Choosing the optimal prior
Kritchevski Trofimov Prediction Rule

Laplace Rule of Succession

Lower Bound

Review

The Universal prediction machine

The biased coins set of experts

Laplace Approximation Choosing the optimal prior Kritchevski Trofimov Prediction Rule

Laplace Rule of Succession

Lower Bound

Generalization to larger sets of distributions

Fisher Information

Exponential Families of Distribution

Review

Probabilities and codes

 $ightharpoonup M_1, \dots, M_n$ - possible messages

Probabilities and codes

- $ightharpoonup M_1, \dots, M_n$ possible messages
- $ightharpoonup P(M_i)$ probability of message *i*

Probabilities and codes

- $ightharpoonup M_1, \dots, M_n$ possible messages
- $\triangleright P(M_i)$ probability of message i
- Arithmetic coding defines a code of length $\lceil -\log_2 P(M_i) \rceil$ for message i

► Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Weight of expert i

$$w_i^t = w_i^1 e^{-L_i^{t-1}} = w_i^1 \prod_{s=1}^{t-1} p_i^s(c^s)$$

Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Weight of expert i

$$w_i^t = w_i^1 e^{-L_i^{t-1}} = w_i^1 \prod_{s=1}^{t-1} p_i^s(c^s)$$

Freedom to choose initial weights.

$$w_t^1 \ge 0, \sum_{i=1}^n w_i^1 = 1$$

► Total loss of expert i

$$L_i^t = -\sum_{s=1}^t \log p_i^s(c^s); \quad L_i^0 = 0$$

Weight of expert i

$$w_i^t = w_i^1 e^{-L_i^{t-1}} = w_i^1 \prod_{c=1}^{t-1} p_i^s(c^s)$$

Freedom to choose initial weights.

$$w_t^1 \ge 0, \sum_{i=1}^n w_i^1 = 1$$

Prediction of algorithm A

$$\mathbf{p}_{A}^{t} = \frac{\sum_{i=1}^{N} w_{i}^{t} \mathbf{p}_{i}^{t}}{\sum_{i=1}^{N} w_{i}^{t}}$$

Total weight: $W^t \doteq \sum_{i=1}^N w_i^t$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$W^{t+1} \qquad \sum_{i=1}^N w_i^t e^{\log p_i^t(c^t)}$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t}$$

Total weight:
$$\mathbf{W}^t \doteq \sum_{i=1}^N \mathbf{w}_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t}$$

Total weight: $W^t \doteq \sum_{i=1}^N w_i^t$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$

Total weight:
$$\mathbf{W}^t \doteq \sum_{i=1}^N \mathbf{w}_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$
$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

Total weight:
$$\mathbf{W}^t \doteq \sum_{i=1}^N \mathbf{w}_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$
$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$
$$-\log \frac{W^{T+1}}{W^1} = -\sum_{i=1}^{T} \log p_A^t(c^t)$$

Total weight:
$$W^t \doteq \sum_{i=1}^{N} w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$

$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log \frac{W^{T+1}}{W^1} = -\sum_{t=1}^{T} \log p_A^t(c^t) = L_A^T$$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$
$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log W^{T+1} = -\log \frac{W^{T+1}}{W^1} = -\sum_{t=1}^{T} \log p_A^t(c^t) = L_A^T$$

Total weight:
$$W^t \doteq \sum_{i=1}^N w_i^t$$

$$\frac{W^{t+1}}{W^t} = \frac{\sum_{i=1}^{N} w_i^t e^{\log p_i^t(c^t)}}{\sum_{i=1}^{N} w_i^t} = \frac{\sum_{i=1}^{N} w_i^t p_i^t(c^t)}{\sum_{i=1}^{N} w_i^t} = p_A^t(c^t)$$
$$-\log \frac{W^{t+1}}{W^t} = -\log p_A^t(c^t)$$

$$-\log W^{T+1} = -\log \frac{W^{T+1}}{W^1} = -\sum_{t=1}^{T} \log p_A^t(c^t) = L_A^T$$

EQUALITY not bound!

► Use non-uniform initial weights $\sum_i w_i^1 = 1$

- ► Use non-uniform initial weights $\sum_i w_i^1 = 1$
- ► Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1}$$

- ► Use non-uniform initial weights $\sum_i w_i^1 = 1$
- ► Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1}$$

- ► Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1} = -\log \sum_{i=1}^N w_i^{T+1}$$

► Use non-uniform initial weights $\sum_{i} w_{i}^{1} = 1$

► Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1} = -\log \sum_{i=1}^N w_i^{T+1}$$
$$= -\log \sum_{i=1}^N w_i^1 e^{-L_i^T}$$

- ► Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_A^T = -\log W^{T+1} = -\log \sum_{i=1}^N w_i^{T+1}$$

$$= -\log \sum_{i=1}^N w_i^1 e^{-L_i^T} \le -\log \max_i \left(w_i^1 e^{-L_i^T} \right)$$

- ► Use non-uniform initial weights $\sum_i w_i^1 = 1$
- Total Weight is at least the weight of the best expert.

$$L_{A}^{T} = -\log W^{T+1} = -\log \sum_{i=1}^{N} w_{i}^{T+1}$$

$$= -\log \sum_{i=1}^{N} w_{i}^{1} e^{-L_{i}^{T}} \le -\log \max_{i} \left(w_{i}^{1} e^{-L_{i}^{T}} \right)$$

$$= \min_{i} \left(L_{i}^{T} - \log w_{i}^{1} \right)$$

_ The Universal prediction machine

The Universal prediction machine

► Fix a universal Turing machine *U*.

The Universal prediction machine

- Fix a universal Turing machine *U*.
- ► An online prediction algorithm *E* is a program that

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - A prediction for the next bit $p(\vec{X}) \in [0, 1]$.

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - ► To ensure *p* has a finite description. Restrict to rational numbers *n*/*m*

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - To ensure *p* has a finite description. Restrict to rational numbers *n/m*
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - ► To ensure *p* has a finite description. Restrict to rational numbers *n/m*
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.
- Most sequences do not correspond to valid prediction algorithms.

- Fix a universal Turing machine *U*.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - ▶ A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - To ensure p has a finite description. Restrict to rational numbers n/m
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.
- Most sequences do not correspond to valid prediction algorithms.
- $V(\vec{b}, \vec{X}, t) = 1$ if the program \vec{b} , given \vec{X} as input, halts within t steps and outputs a well-formed prediction. Otherwise $V(\vec{b}, \vec{X}, t) = 0$

- Fix a universal Turing machine U.
- An online prediction algorithm E is a program that
 - ▶ given as input The past $\vec{X} \in \{0, 1\}^t$
 - runs finite time and outputs
 - A prediction for the next bit $p(\vec{X}) \in [0, 1]$.
 - To ensure p has a finite description. Restrict to rational numbers n/m
- Any online prediction algorithm can be represented as code $\vec{b}(E)$ for U. The code length is $|\vec{b}(E)|$.
- Most sequences do not correspond to valid prediction algorithms.
- $V(\vec{b}, \vec{X}, t) = 1$ if the program \vec{b} , given \vec{X} as input, halts within t steps and outputs a well-formed prediction. Otherwise $V(\vec{b}, \vec{X}, t) = 0$
- $V(\vec{b}, \vec{X}, t)$ is computable (recursively enumerable).

► Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| - \log_2 |\vec{b}|}$.

- Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| \log_2 |\vec{b}|}$.
- The total initial weight over all finite binary sequences is one.

- Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| \log_2 |\vec{b}|}$.
- The total initial weight over all finite binary sequences is one.
- ► Run the Bayes algorithm over "all" prediction algorithms.

- Assign to the code \vec{b} the initial weight $w_{\vec{b}}^1 = 2^{-|\vec{b}| \log_2 |\vec{b}|}$.
- The total initial weight over all finite binary sequences is one.
- ► Run the Bayes algorithm over "all" prediction algorithms.
- technical details: On iteration t, $|\vec{X}| = t$. Use the predictions of programs \vec{b} such that $|\vec{b}| \le t$ and for which $V(\vec{b}, \vec{X}, 2^t) = 1$.

the unused algorithms predict 1/2 (insuring a loss of 1)

▶ Using $L_A \leq \min_i (L_i - \log w_i^1)$

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ▶ We get that the loss of the Universal algorithm is at most $|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ► We get that the loss of the Universal algorithm is at most $|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$
- More careful analysis can reduce to $|\vec{b}(E)| + L_E$

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- ► We get that the loss of the Universal algorithm is at most $|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$
- ► More careful analysis can reduce to $|\vec{b}(E)| + L_E$
- ► How good is that?

- ▶ Using $L_A \leq \min_i (L_i \log w_i^1)$
- Assume E is a prediction algorithm which generates the tth prediction in time smaller than 2^t
- ▶ When $t \le |\vec{b}(E)|$ the algorithm is not used and thus it's loss is 1
- We get that the loss of the Universal algorithm is at most $|\vec{b}(E)| + \log_2 |\vec{b}(E)| + L_E$
- ► More careful analysis can reduce to $|\vec{b}(E)| + L_E$
- How good is that?
- What is the two part code?

Bayes coding is better than two part codes

Simple bound as good as bound for two part codes (MDL) but enables online compression

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- Two part code has to point to one of the KN experts $L_A \leq \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- Two part code has to point to one of the KN experts $L_A \leq \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$
- ▶ If we use Bayes predictor + arithmetic coding we get:

$$L_A = -\log W^{T+1} \le \log K \max_i \frac{1}{NK} e^{-L_i^T} = \log N + \min_i L_i^T$$

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- Two part code has to point to one of the KN experts $L_A \leq \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$
- ▶ If we use Bayes predictor + arithmetic coding we get:

$$L_{\mathcal{A}} = -\log W^{T+1} \leq \log K \max_{i} \frac{1}{NK} e^{-L_{i}^{T}} = \log N + \min_{i} L_{i}^{T}$$

We don't pay a penalty for copies.

- Simple bound as good as bound for two part codes (MDL) but enables online compression
- Suppose we have K copies of each expert.
- Two part code has to point to one of the KN experts $L_A \leq \log NK + \min_i L_i^T = \log NK + \min_i L_i^T$
- If we use Bayes predictor + arithmetic coding we get:

$$L_{A} = -\log W^{T+1} \leq \log K \max_{i} \frac{1}{NK} e^{-L_{i}^{T}} = \log N + \min_{i} L_{i}^{T}$$

- We don't pay a penalty for copies.
- More generally, the regret is smaller if many of the experts perform well.

► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.

- Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.
- ► Instead, we assign the experts a Density Measure.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.
- ► Instead, we assign the experts a Density Measure.
- ► $L_A \le \min_i (L_i \log w_i^1)$ is meaningless.

- ► Each expert corresponds to a biased coin, predicts with a fixed $\theta \in [0, 1]$.
- Set of experts is uncountably infinite.
- Only countably many experts can be assigned non-zero weight.
- ► Instead, we assign the experts a Density Measure.
- ► $L_A \le \min_i (L_i \log w_i^1)$ is meaningless.
- Can we still get a meaningful bound?

Bayes Algorithm for biased coins

Replace the initial weight by a density measure

$$w(\theta) = w^{1}(\theta), \int_{0}^{1} w(\theta) d\theta = 1$$

Bayes Algorithm for biased coins

- ► Replace the initial weight by a density measure $w(\theta) = w^{1}(\theta), \int_{0}^{1} w(\theta) d\theta = 1$
- Relationship between final total weight and total log loss remains unchanged:

$$L_A = \ln \int_0^1 w(\theta) e^{-L_{\theta}^{T+1}} d\theta$$

Bayes Algorithm for biased coins

- ► Replace the initial weight by a density measure $w(\theta) = w^{1}(\theta), \int_{0}^{1} w(\theta) d\theta = 1$
- Relationship between final total weight and total log loss remains unchanged:

$$L_A = \ln \int_0^1 w(\theta) e^{-L_{\theta}^{T+1}} d\theta$$

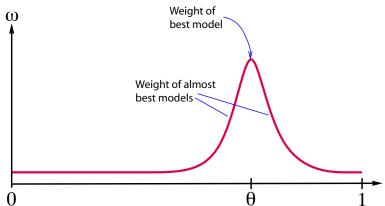
We need a new lower bound on the final total weight

Main Idea

If $\mathbf{w}^t(\theta)$ is large then $\mathbf{w}^t(\theta + \epsilon)$ is also large.

Main Idea

If $\mathbf{w}^t(\theta)$ is large then $\mathbf{w}^t(\theta + \epsilon)$ is also large.



Expanding the exponent around the peak

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

Laplace Approximation

Expanding the exponent around the peak

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

The total loss scales with T

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

Laplace Approximation

Expanding the exponent around the peak

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

The total loss scales with T

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

Expanding the exponent around the peak

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

The total loss scales with T

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

$$L_A - L_{\min} \le \ln \int_0^1 w(\theta) e^{-L_{\theta}} d\theta - \ln e^{L_{\min}}$$

Expanding the exponent around the peak

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

The total loss scales with T

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

$$L_{A} - L_{\min} \leq \ln \int_{0}^{1} w(\theta) e^{-L_{\theta}} d\theta - \ln e^{L_{\min}}$$
$$= \ln \int_{0}^{1} w(\theta) e^{-(L_{\theta} - L_{\min})} d\theta$$

Expanding the exponent around the peak

 \triangleright For log loss the best θ is empirical distribution of the seq.

$$\hat{\theta} = \frac{\#\{x^t = 1; \ 1 \le t \le T\}}{T}$$

The total loss scales with T

$$L_{\theta} = T \cdot (\hat{\theta}\ell(\theta, 1) + (1 - \hat{\theta})\ell(\theta, 0)) \doteq T \cdot g(\hat{\theta}, \theta)$$

$$\begin{array}{ll} \mathcal{L}_{A}-\mathcal{L}_{\min} & \leq & \ln\int_{0}^{1}w(\theta)e^{-\mathcal{L}_{\theta}}d\theta-\ln e^{\mathcal{L}_{\min}}\\ \\ & = & \ln\int_{0}^{1}w(\theta)e^{-(\mathcal{L}_{\theta}-\mathcal{L}_{\min})}d\theta\\ \\ & = & \ln\int_{0}^{1}w(\theta)e^{T(g(\hat{\theta},\theta)-g(\hat{\theta},\hat{\theta}))}d\theta \end{array}$$

Laplace Approximation

Laplace approximation (idea)

► Taylor expansion of $g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.

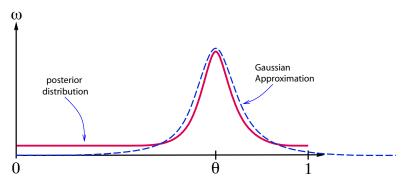
Laplace Approximation

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent
- ightharpoonup \Rightarrow a gaussian approximation of the posterior.

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent
- ightharpoonup \Rightarrow a gaussian approximation of the posterior.

- ► Taylor expansion of $g(\hat{\theta}, \theta) g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.
- First and second terms in the expansion are zero.
- Third term gives a quadratic expression in the exponent
- ightharpoonup \Rightarrow a gaussian approximation of the posterior.



Laplace Approximation, Watson's lemma

$$\int_0^1 w(\theta) e^{T(g(\hat{\theta},\theta) - g(\hat{\theta},\hat{\theta}))} d\theta$$

Laplace Approximation

Laplace Approximation, Watson's lemma

$$\int_{0}^{1} w(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \frac{d^{2}}{d\theta^{2}} \Big|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}} + O(T^{-3/2})$$

Choosing the optimal prior

 \triangleright Choose $w(\theta)$ to maximize the worst-case final total weight

$$\min_{\hat{\theta}} w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \frac{d^2}{d\theta^2} \Big|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}}$$

 \triangleright Choose $w(\theta)$ to maximize the worst-case final total weight

$$\min_{\hat{\theta}} w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \frac{d^2}{d\theta^2} \Big|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}}$$

▶ Make bound equal for all $\hat{\theta} \in [0, 1]$ by choosing

$$w^*(\hat{\theta}) = \frac{1}{Z} \sqrt{\frac{\frac{d^2}{d\theta^2}\Big|_{\theta=\hat{\theta}} (g(\hat{\theta},\theta) - g(\hat{\theta},\hat{\theta}))}{-2\pi}},$$

where Z is the normalization factor:

$$Z=\sqrt{rac{1}{2\pi}}\int_0^1\left.\sqrt{rac{d^2}{d heta^2}}
ight|_{ heta=\hat{ heta}}\left(g(\hat{ heta},\hat{ heta})-g(\hat{ heta}, heta)
ight)\;d\hat{ heta}$$

The bound for the optimal prior

We get that the regret is

$$L_{A} - L_{\min} \leq \ln \int_{0}^{1} w^{*}(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= \ln \left(\sqrt{\frac{2\pi Z}{T}} + O(T^{-3/2}) \right)$$

$$= \frac{1}{2} \ln \frac{T}{2\pi} - \frac{1}{2} \ln Z + O(1/T) .$$

Solving for log-loss

The exponent in the integral is

$$g(\hat{ heta}, heta) - g(\hat{ heta}, \hat{ heta}) = \hat{ heta} \ln \frac{\hat{ heta}}{ heta} + (1 - \hat{ heta}) \ln \frac{1 - \hat{ heta}}{1 - heta} = D_{KL}(\hat{ heta}|| heta)$$

Solving for log-loss

The exponent in the integral is

$$g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}) = \hat{\theta} \ln \frac{\hat{\theta}}{\theta} + (1 - \hat{\theta}) \ln \frac{1 - \hat{\theta}}{1 - \theta} = D_{KL}(\hat{\theta}||\theta)$$

The second derivative

$$\left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} D_{KL}(\hat{\theta}||\theta) = \frac{1}{\hat{\theta}(1 - \hat{\theta})}$$

Is called the Fisher information

Solving for log-loss

The exponent in the integral is

$$g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}) = \hat{\theta} \ln \frac{\hat{\theta}}{\theta} + (1 - \hat{\theta}) \ln \frac{1 - \hat{\theta}}{1 - \theta} = D_{KL}(\hat{\theta}||\theta)$$

The second derivative

$$\left. \frac{d^2}{d\theta^2} \right|_{\theta = \hat{\theta}} D_{KL}(\hat{\theta}||\theta) = \frac{1}{\hat{\theta}(1 - \hat{\theta})}$$

Is called the Fisher information

The optimal prior:

$$w^*(\hat{\theta}) = \frac{1}{\pi \sqrt{\hat{\theta}(1-\hat{\theta})}}$$

Known in as Jeffrey's prior. And, in this case, the Dirichlet-(1/2, 1/2) prior.

The regret of Bayes using Jeffrey's prior

The regret (for coding: redundancy) is

$$L_A - L_{\min} \le \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} + O(1/T)$$

As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.

Kritchevski Trofimov Prediction Bule

- As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.
- Observed t bits, n of which were 1. The posterior is:

$$\frac{1}{Z\sqrt{\theta(1-\theta)}}\theta^{n}(1-\theta)^{t-n} = \frac{1}{Z}\theta^{n-1/2}(1-\theta)^{t-n-1/2}$$

- As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.
- Observed t bits, n of which were 1. The posterior is:

$$\frac{1}{Z\sqrt{\theta(1-\theta)}}\theta^{n}(1-\theta)^{t-n} = \frac{1}{Z}\theta^{n-1/2}(1-\theta)^{t-n-1/2}$$

► The posterior average is:

$$\frac{\int_0^1 \theta^{n+1/2} (1-\theta)^{t-n-1/2} d\theta}{\int_0^1 \theta^{n-1/2} (1-\theta)^{t-n-1/2} d\theta} = \frac{n+1/2}{t+1}$$

- As luck would have it the Dirichlet prior is the conjugate prior for the Binomial distribution.
- Observed *t* bits, *n* of which were 1. The posterior is:

$$\frac{1}{Z\sqrt{\theta(1-\theta)}}\theta^{n}(1-\theta)^{t-n} = \frac{1}{Z}\theta^{n-1/2}(1-\theta)^{t-n-1/2}$$

The posterior average is:

$$\frac{\int_0^1 \theta^{n+1/2} (1-\theta)^{t-n-1/2} d\theta}{\int_0^1 \theta^{n-1/2} (1-\theta)^{t-n-1/2} d\theta} = \frac{n+1/2}{t+1}$$

This is called the Trichevsky Trofimov prediction rule.

Laplace Rule of Succession

Laplace suggested using the uniform prior, which is also a conjugate prior.

Laplace Rule of Succession

- Laplace suggested using the uniform prior, which is also a conjugate prior.
- In this case the posterior average is:

$$\frac{\int_0^1 \theta^{n+1} (1-\theta)^{t-n} d\theta}{\int_0^1 \theta^n (1-\theta)^{t-n} d\theta} = \frac{n+1}{t+2}$$

Laplace Rule of Succession

- Laplace suggested using the uniform prior, which is also a conjugate prior.
- In this case the posterior average is:

$$\frac{\int_{0}^{1} \theta^{n+1} (1-\theta)^{t-n} d\theta}{\int_{0}^{1} \theta^{n} (1-\theta)^{t-n} d\theta} = \frac{n+1}{t+2}$$

The bound on the cumulative log loss is worse:

$$L_A - L_{\min} = \ln T + O(1)$$

- Laplace suggested using the uniform prior, which is also a conjugate prior.
- In this case the posterior average is:

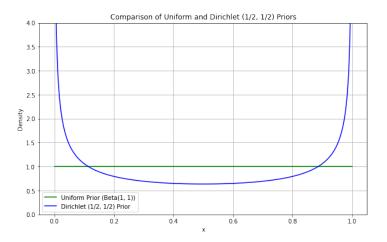
$$\frac{\int_{0}^{1} \theta^{n+1} (1-\theta)^{t-n} d\theta}{\int_{0}^{1} \theta^{n} (1-\theta)^{t-n} d\theta} = \frac{n+1}{t+2}$$

The bound on the cumulative log loss is worse:

$$L_A - L_{\min} = \ln T + O(1)$$

Suffers larger regret when $\hat{\theta}$ is far from 1/2

Comparing the priors



The biased coins set of experts

Lower Bound

Shtarkov's Normalized Maximal Likelihood (NML)

What is the optimal prediction when T is know in advance?

The biased coins set of experts

Lower Bound

Shtarkov's Normalized Maximal Likelihood (NML)

- What is the optimal prediction when T is know in advance?
- Equalize the reducedancy over all sequences.

Shtarkov's Normalized Maximal Likelihood (NML)

- What is the optimal prediction when T is know in advance?
- Equalize the reducndancy over all sequences.
- Assign to each sequence the Maximal likelihood:

$$P(x) = \frac{1}{Z} \max_{p} P_{p}(x)$$

Lower Bound

Shtarkov's Normalized Maximal Likelihood (NML)

- ▶ What is the optimal prediction when *T* is know in advance?
- Equalize the reducedancy over all sequences.
- Assign to each sequence the Maximal likelihood:

$$P(x) = \frac{1}{Z} \max_{p} P_p(x)$$

 $\triangleright \log Z$ is the regret (redundancy).

Shtarkov's Normalized Maximal Likelihood (NML)

- What is the optimal prediction when T is know in advance?
- Equalize the reducndancy over all sequences.
- Assign to each **sequence** the **Maximal** likelihood:

$$P(x) = \frac{1}{Z} \max_{p} P_{p}(x)$$

- $ightharpoonup \log Z$ is the regret (redundancy).
- For Bernoulli distributions:

$$L_*^T - \min_{\theta} L_{\theta}^T \geq \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} - O(\frac{1}{\sqrt{T}})$$

Shtarkov's Normalized Maximal Likelihood (NML)

- ▶ What is the optimal prediction when T is know in advance?
- Equalize the reducedancy over all sequences.
- Assign to each sequence the Maximal likelihood: $P(x) = \frac{1}{7} \max_{D} P_{D}(x)$
- $\triangleright \log Z$ is the regret (redundancy).
- For Bernoulli distributions:

$$L_*^T - \min_{\theta} L_{\theta}^T \geq \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} - O(\frac{1}{\sqrt{T}})$$

► The regret (for coding: redundancy) is

$$L_A - L_{\min} \le \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} + O(1/T)$$

Generalization to larger sets of distributions

Generalization to larger sets of distributions

Multinomial Distributions

► For a distribution over *k* elements (Multinomial) [Xie and Barron]

Multinomial Distributions

- ► For a distribution over *k* elements (Multinomial) [Xie and Barron]
- ► Use the add 1/2 rule (KT).

$$p(i) = \frac{n_i + 1/2}{t + k/2}$$

Multinomial Distributions

- ► For a distribution over *k* elements (Multinomial) [Xie and Barron]
- ► Use the add 1/2 rule (KT).

$$p(i) = \frac{n_i + 1/2}{t + k/2}$$

Bound is

$$L_A - L_{\min} \leq \frac{k-1}{2} \ln T + C + o(1)$$

Multinomial Distributions

- ► For a distribution over k elements (Multinomial) [Xie and Barron]
- ► Use the add 1/2 rule (KT).

$$p(i) = \frac{n_i + 1/2}{t + k/2}$$

Bound is

$$L_A - L_{\min} \leq \frac{k-1}{2} \ln T + C + o(1)$$

The constant C is optimal.

Generalization to larger sets of distributions

Fisher Information

The Fisher Information Matrix



$$\mathbf{I}(\theta) = \nabla_{\theta'}^2 D_{\mathsf{KL}}(p(x;\theta) \| p(x;\theta')) \Big|_{\theta' = \theta}$$

The Fisher Information Matrix

$$\mathbf{I}(\theta) = \nabla_{\theta'}^2 D_{\mathsf{KL}}(p(x;\theta) \| p(x;\theta')) \bigg|_{\theta' = \theta}$$

$$\mathbf{I}(\theta) = \begin{bmatrix} \frac{\partial^{2}}{\partial \theta_{1}^{2}} D_{\mathsf{KL}} & \frac{\partial^{2}}{\partial \theta_{1} \partial \theta_{2}} D_{\mathsf{KL}} & \cdots & \frac{\partial^{2}}{\partial \theta_{1} \partial \theta_{n}} D_{\mathsf{KL}} \\ \frac{\partial^{2}}{\partial \theta_{2} \partial \theta_{1}} D_{\mathsf{KL}} & \frac{\partial^{2}}{\partial \theta_{2}^{2}} D_{\mathsf{KL}} & \cdots & \frac{\partial^{2}}{\partial \theta_{2} \partial \theta_{n}} D_{\mathsf{KL}} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^{2}}{\partial \theta_{n} \partial \theta_{1}} D_{\mathsf{KL}} & \frac{\partial^{2}}{\partial \theta_{n} \partial \theta_{2}} D_{\mathsf{KL}} & \cdots & \frac{\partial^{2}}{\partial \theta_{n}^{2}} D_{\mathsf{KL}} \end{bmatrix}_{\theta' = \theta}$$

The Fisher Information Matrix

$$\mathbf{I}(\theta) = \nabla_{\theta'}^2 D_{\mathsf{KL}}(p(x;\theta) \| p(x;\theta')) \bigg|_{\theta' = \theta}$$

$$\mathbf{I}(\theta) = \begin{bmatrix} \frac{\partial^2}{\partial \theta_1^2} D_{\mathsf{KL}} & \frac{\partial^2}{\partial \theta_1 \partial \theta_2} D_{\mathsf{KL}} & \cdots & \frac{\partial^2}{\partial \theta_1 \partial \theta_n} D_{\mathsf{KL}} \\ \frac{\partial^2}{\partial \theta_2 \partial \theta_1} D_{\mathsf{KL}} & \frac{\partial^2}{\partial \theta_2^2} D_{\mathsf{KL}} & \cdots & \frac{\partial^2}{\partial \theta_2 \partial \theta_n} D_{\mathsf{KL}} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2}{\partial \theta_n \partial \theta_1} D_{\mathsf{KL}} & \frac{\partial^2}{\partial \theta_n \partial \theta_2} D_{\mathsf{KL}} & \cdots & \frac{\partial^2}{\partial \theta_n^2} D_{\mathsf{KL}} \end{bmatrix}_{\theta' = \theta}$$

Jeffrey's prior

$$\pi_J(\boldsymbol{ heta}) \propto \sqrt{\det(\mathbf{I}(\boldsymbol{ heta}))}$$

Properties of Jeffrey's prior

Known as "least informative prior" in Bayesian statistics.

Properties of Jeffrey's prior

- Known as "least informative prior" in Bayesian statistics.
- Min/max: Equalizes the risk for all parameter setting

Properties of Jeffrey's prior

- Known as "least informative prior" in Bayesian statistics.
- Min/max: Equalizes the risk for all parameter setting
- invariant under re-parametrization.

Properties of Jeffrey's prior

- Known as "least informative prior" in Bayesian statistics.
- Min/max: Equalizes the risk for all parameter setting
- invariant under re-parametrization.
- ▶ Often improper (integral = ∞).

Fisher Information

Exponential Distributions

▶ The canonical form of an exponential distribution is

$$p(x|\theta) = \exp\left[\eta(\theta) \cdot T(x)\right]$$

for some fixed functions η , T

Exponential Distributions

▶ The canonical form of an exponential distribution is

$$p(x|\theta) = \exp\left[\eta(\theta) \cdot T(x)\right]$$

for some fixed functions η , T

► Multinomial: $p(i| < p_1, ..., p_k >) = p_i$ $\eta(\theta) = < \log p_1, ..., \log p_k >$ $T: i \to (0, ..., 1, 0, ..., 0)$

Exponential Distributions

The canonical form of an exponential distribution is

$$p(x|\theta) = \exp \left[\eta(\theta) \cdot T(x)\right]$$

for some fixed functions η , T

- ightharpoonup Multinomial: $p(i| < p_1, \ldots, p_k >) = p_i$ $\eta(\theta) = \langle \log p_1, \dots, \log p_k \rangle$ $T: i \to (0, ..., 1, 0, ..., 0)$
- Normal: $p(y|\mu,\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{(y-\mu)^2}{2\sigma^2}\right)$ $\eta(\mu, \sigma^2) = \left[\frac{\mu^2}{2\sigma^2} - \frac{1}{2}\log 2\pi\sigma^2, \frac{\mu}{\sigma^2}, -\frac{1}{2\sigma^2}\right]$ $T(v) = (1, v, v^2)$

Exponential Distributions

The canonical form of an exponential distribution is

$$p(x|\theta) = \exp\left[\eta(\theta) \cdot T(x)\right]$$

for some fixed functions η , T

- ightharpoonup Multinomial: $p(i| < p_1, \dots, p_k >) = p_i$ $\eta(\theta) = \langle \log p_1, \dots, \log p_k \rangle$ $T: i \to (0, ..., 1, 0, ..., 0)$
- Normal: $p(y|\mu,\sigma) = \frac{1}{\sqrt{2\sigma^2}} \exp\left(\frac{(y-\mu)^2}{2\sigma^2}\right)$ $\eta(\mu, \sigma^2) = \left[\frac{\mu^2}{2\sigma^2} - \frac{1}{2} \log 2\pi \sigma^2, \frac{\mu}{\sigma^2}, -\frac{1}{2\sigma^2} \right]$ $T(y) = (1, y, y^2)$
- Many more: 1D: Poisson, Exponential, Gamma ... Multi-Variate: Gaussian, Dirichlet, Multivariate t-distribution

Online learning for Exponential Families

For any set of distributions from the exponential family defined by k parameters, the observations and the parameters come from compact sets.

Generalization to larger sets of distributions

Exponential Families of Distribution

Online learning for Exponential Families

- ► For any set of distributions from the exponential family defined by *k* parameters, the observations and the parameters come from compact sets.
- Use Bayes Algorithm with Jeffrey's prior:

$$w^*(\theta) = \frac{1}{Z} \sqrt{\det(I(\theta))}$$

 $I(\theta)$, Z do not depend on T.

Exponential Families of Distribution

Online learning for Exponential Families

- ► For any set of distributions from the exponential family defined by *k* parameters, the observations and the parameters come from compact sets.
- Use Bayes Algorithm with Jeffrey's prior:

$$w^*(\theta) = \frac{1}{Z} \sqrt{\det(I(\theta))}$$

 $I(\theta)$, Z do not depend on T.

Redundancy bound

$$L_A - L_{\min} \leq \frac{k-1}{2} \ln T - \ln Z + o(1)$$

next Class

Variable-length markov models - a set of distributions with increasing number of parameters.

next Class

- Variable-length markov models a set of distributions with increasing number of parameters.
- The context algorithm: An efficient implementation of the Bayes algorithm which achieves close-to-optimal worst case bounds.