Combining infinite sets of experts

Yoav Freund

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Freund: Predicting a binary Sequence almost as well the the optimal biased coin.

Risannen: Fisher Information and Stochastic Complexity.

Review

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The Universal prediction machine

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The biased coins set of experts
Laplace Approximation
Choosing the optimal prior
Kritchevski Trofimov Prediction Rule

Laplace Rule of Succession

Lower Bound

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Generalization to larger sets of distributions

Fisher Information

Exponential Families of Distribution

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- $\triangleright P(M_i)$ probability of message i
- Arithmetic coding defines a code of length $\lceil -\log_2 P(M_i) \rceil$ for message i

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Prediction of algorithm A

$$\mathbf{p}_{A}^{t} = \frac{\sum_{i=1}^{N} w_{i}^{t} \mathbf{p}_{i}^{t}}{\sum_{i=1}^{N} w_{i}^{t}}$$

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- $V(\vec{b}, \vec{X}, t)$ is computable (recursively enumerable).

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- technical details: On iteration t, $|\vec{X}| = t$. Use the predictions of programs \vec{b} such that $|\vec{b}| \le t$ and for which $V(\vec{b}, \vec{X}, 2^t) = 1$.

the unused algorithms predict 1/2 (insuring a loss of 1)

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- What is the two part code?

Bayes coding is better than two part codes

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- More generally, the regret is smaller if many of the experts perform well.

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- Can we still get a meaningful bound?

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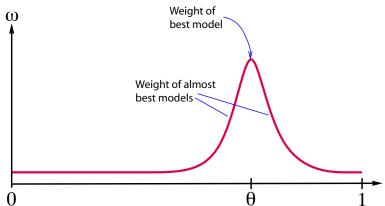
We need a new lower bound on the final total weight

Main Idea

If $\mathbf{w}^t(\theta)$ is large then $\mathbf{w}^t(\theta + \epsilon)$ is also large.

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Expanding the exponent around the peak

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The total loss scales with T

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$$\begin{array}{ll} \mathcal{L}_{A}-\mathcal{L}_{\min} & \leq & \ln\int_{0}^{1}w(\theta)e^{-\mathcal{L}_{\theta}}d\theta-\ln e^{\mathcal{L}_{\min}}\\ \\ & = & \ln\int_{0}^{1}w(\theta)e^{-(\mathcal{L}_{\theta}-\mathcal{L}_{\min})}d\theta\\ \\ & = & \ln\int_{0}^{1}w(\theta)e^{T(g(\hat{\theta},\theta)-g(\hat{\theta},\hat{\theta}))}d\theta \end{array}$$

Laplace Approximation

Laplace approximation (idea)

► Taylor expansion of $g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta})$ around $\theta = \hat{\theta}$.

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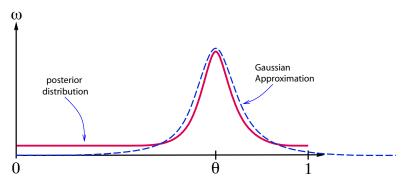
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$$\int_{0}^{1} w(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= w(\hat{\theta}) \sqrt{\frac{-2\pi}{T \frac{d^{2}}{d\theta^{2}} \Big|_{\theta = \hat{\theta}} (g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))}} + O(T^{-3/2})$$

Choosing the optimal prior

 \triangleright Choose $w(\theta)$ to maximize the worst-case final total weight

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▶ Make bound equal for all $\hat{\theta} \in [0, 1]$ by choosing

$$w^*(\hat{\theta}) = \frac{1}{Z} \sqrt{\frac{\frac{d^2}{d\theta^2}\Big|_{\theta=\hat{\theta}} (g(\hat{\theta},\theta) - g(\hat{\theta},\hat{\theta}))}{-2\pi}},$$

where Z is the normalization factor:

$$Z=\sqrt{rac{1}{2\pi}}\int_0^1\left.\sqrt{rac{d^2}{d heta^2}}
ight|_{ heta=\hat{ heta}}\left(g(\hat{ heta},\hat{ heta})-g(\hat{ heta}, heta)
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The bound for the optimal prior

Plugging in we get

$$L_{A} - L_{\min} \leq \ln \int_{0}^{1} w^{*}(\theta) e^{T(g(\hat{\theta}, \theta) - g(\hat{\theta}, \hat{\theta}))} d\theta$$

$$= \ln \left(\sqrt{\frac{2\pi Z}{T}} + O(T^{-3/2}) \right)$$

$$= \frac{1}{2} \ln \frac{T}{2\pi} - \frac{1}{2} \ln Z + O(1/T) .$$

Solving for log-loss

The exponent in the integral is

$$g(\hat{ heta}, heta) - g(\hat{ heta}, \hat{ heta}) = \hat{ heta} \ln \frac{\hat{ heta}}{ heta} + (1 - \hat{ heta}) \ln \frac{1 - \hat{ heta}}{1 - heta} = D_{KL}(\hat{ heta}|| heta)$$

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Known in as Jeffrey's prior. And, in this case, the Dirichlet-(1/2, 1/2) prior.

The cumulative log loss of Bayes using Jeffrey's prior



$$L_A - L_{\min} \le \frac{1}{2} \ln(T+1) + \frac{1}{2} \ln \frac{\pi}{2} + O(1/T)$$

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Kritchevski Trofimov Prediction Bule

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This is called the Trichevsky Trofimov prediction rule.

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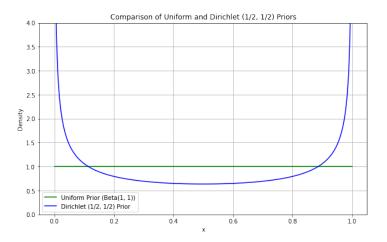
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Suffers larger regret when $\hat{\theta}$ is far from 1/2

Comparing the priors



The biased coins set of experts

Lower Bound

Shtarkov Lower bound

▶ What is the optimal prediction when *T* is know in advance?

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Shtarkov Lower bound

What is the optimal prediction when T is know in advance?

$$L_*^T - \min_{ heta} L_{ heta}^T \geq rac{1}{2} \ln(T+1) + rac{1}{2} \ln rac{\pi}{2} - O(rac{1}{\sqrt{T}})$$

Generalization to larger sets of distributions

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Fisher Information

The Fisher Information Matrix



$$\mathbf{I}(\theta) = \nabla_{\theta'}^2 D_{\mathsf{KL}}(p(x;\theta) \| p(x;\theta')) \Big|_{\theta' = \theta}$$

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Jeffrey's prior

$$\pi_J(\boldsymbol{ heta}) \propto \sqrt{\det(\mathbf{I}(\boldsymbol{ heta}))}$$

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- ▶ Often improper (integral = ∞).

Fisher Information

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- Normal: $p(y|\mu,\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{(y-\mu)^2}{2\sigma^2}\right)$ $\eta(\mu, \sigma^2) = \left[\frac{\mu^2}{2\sigma^2} - \frac{1}{2}\log 2\pi\sigma^2, \frac{\mu}{\sigma^2}, -\frac{1}{2\sigma^2}\right]$ $T(v) = (1, v, v^2)$

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- Many more: 1D: Poisson, Exponential, Gamma ... Multi-Variate: Gaussian, Dirichlet, Multivariate t-distribution

Online learning for Exponential Families

► For any set of distributions from the exponential family defined by *k* parameters, where the observed values come from a bounded set.

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next Class

Variable-length markov models - a set of distributions with increasing number of parameters.

next Class

- Variable-length markov models a set of distributions with increasing number of parameters.
- The context algorithm: An efficient implementation of the Bayes algorithm which achieves close-to-optimal worst case bounds.