

# Adaptive evolution with aneuploidy and mutation

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## Abstract

Aneuploidy is common in eukaryotes, often leading to decreased cell growth and fitness. However, evidence from yeast and fungi, as well as human tumour cells, suggests that aneuploidy can be beneficial under stressful conditions and lead to elevated growth rates and adaptation. Importantly, aneuploidy differs from point mutations in rate, fitness effect, and reversibility. Here, we develop evolutionary models for adaptive evolution with both mutation and aneuploidy. These models are used within an approximate Bayesian computation framework to estimate the formation rate and fitness effect of aneuploidy and mutation from results of evolutionary experiments in which *Saccharomyces cerevisiae* adapted to heat stress: the experimental populations first acquired chromosome duplications, only to later revert back to a euploid state. We also analyze our models to estimate the effect of the aneuploidy and mutation rates on the expected adaptation time and the probability for adaptation via aneuploidy. Our results suggest that aneuploidy can be a transient adaptive solution, which can decelerate adaptation in a non-intuitive manner. By creating an evolutionary conflict between the individual and the population, aneuploidy further complicates the process of adaptation in cell populations.

# Introduction

**Aneuploidy is common in eukaryotes.** Aneuploidy is an imbalance in the number of chromosomes in the cell: an incorrect karyotype. Evidence suggests aneuploidy is very common in eukaryotes, e.g. animals (Santaguida and Amon, 2015; Naylor and van Deursen, 2016; Bakhoun and Landau, 2017), and fungi (Pavelka et al., 2010; Zhu et al., 2016; Robbins et al., 2017; Todd et al., 2017). Aneuploidy has been implicated in cancer formation and progression (Boveri, 2008; Schwartzman et al., 2010): 90% of solid tumours and 50% of blood cancers are aneuploid (Santaguida and Amon, 2015). Aneuploidy is also linked to the emergence of drug resistance (Selmecki et al., 2009) and virulence (Möller et al., 2018) in fungal pathogens, which are under-studied (Rodrigues and Albuquerque, 2018) despite infecting close to a billion people per year, causing serious infections and significant morbidity in >150 million people per year and killing >1.5 million people per year (Selmecki et al., 2009; Rodrigues and Albuquerque, 2018). In addition, aneuploidy is common in protozoan pathogens of the *Leishmania* genus, a major global health concern (Mannaert et al., 2012).

**Aneuploidy is generally deleterious.** The molecular and genetic mechanisms involved in aneuploidy have been explored (Musacchio and Salmon, 2007; Sheltzer and Amon, 2011; Chen et al., 2012; Rancati and Pavelka, 2013; Gerstein et al., 2015; Shor and Perlin, 2015). Experiments with human and mouse embryos found that aneuploidy is usually lethal. It is also associated with developmental defects and lethality in other multicellular organisms (Sheltzer and Amon, 2011). For example, aneuploid mouse embryonic cells grow slower than euploid cells (Williams et al., 2008). Similarly, in unicellular eukaryotes growing in benign conditions, aneuploidy usually leads to slower growth and decreased overall fitness (Niwa et al., 2006; Torres et al., 2007; Pavelka et al., 2010; Sheltzer and Amon, 2011; Kasuga et al., 2016), in part due to proteotoxic stress caused by increased expression in aneuploid cells (Pavelka et al., 2010; Santaguida et al., 2015; Zhu et al., 2018) and hypo-osmotic-like stress (Tsai et al., 2019).

**Aneuploidy can lead to adaptation.** However, aneuploidy can be beneficial under stressful conditions due to the wide range of phenotypes it can produce, some of which are advantageous (Pavelka et al., 2010). Thus, aneuploidy can lead to rapid adaptation in unicellular eukaryotes (Gerstein et al., 2015; Torres et al., 2010; Hong and Gresham, 2014; Rancati et al., 2008), as well as to rapid growth of somatic tumour cells (Schwartzman et al., 2010; Sheltzer et al., 2017). For example, aneuploidy in *S. cerevisiae* facilitates adaptation to a variety of stressful conditions like heat and pH (Yona et al., 2012), copper (Covo et al., 2014; Gerstein et al., 2015), salt (Dhar et al., 2011), and nutrient limitation (Dunham et al., 2002; Gresham et al., 2008). Importantly, aneuploidy can also lead to drug

resistance in pathogenic fungi such as *Candida albicans* (Selmecki et al., 2008, 2010; Gerstein and Berman, 2018) and *Cryptococcus neoformans* (Sionov et al., 2010), which cause candidiasis and meningoencephalitis, respectively.

**Transient adaptive solution.** Aneuploidy differs from mutation due to its distinct properties. Chromosome duplication usually occurs more often than mutation and on average produces larger fitness effects. Yet, because it affects many genes on a whole chromosome or a chromosome fragment, aneuploidy also carries fitness costs. Thus, aneuploidy can be a *transient adaptive solution*: it can rapidly occur and fix in the population under stressful conditions, and can be rapidly lost when the cost outweighs the benefit—when stress is removed or after beneficial mutations occur. Experimental evidence of such a transient role of aneuploidy was demonstrated by Yona et al. (2012). They evolved populations of *S. cerevisiae* under strong heat or pH stress. The populations adapted to the heat and pH stress within 450 and 150 generations, and this adaptation was determined to be due to chromosome duplications. Much later, after more than 1500 and 750 generations, for the heat and pH stress, respectively, the populations reverted back to an euploid state, while remaining adapted to the stress and accumulating multiple mutations. However, under gradual heat stress, aneuploidy was not observed. Yona et al. (2012) concluded that aneuploidy serves as a transient adaptive solution, or a “quick fix”, which is expected to facilitate adaptation.

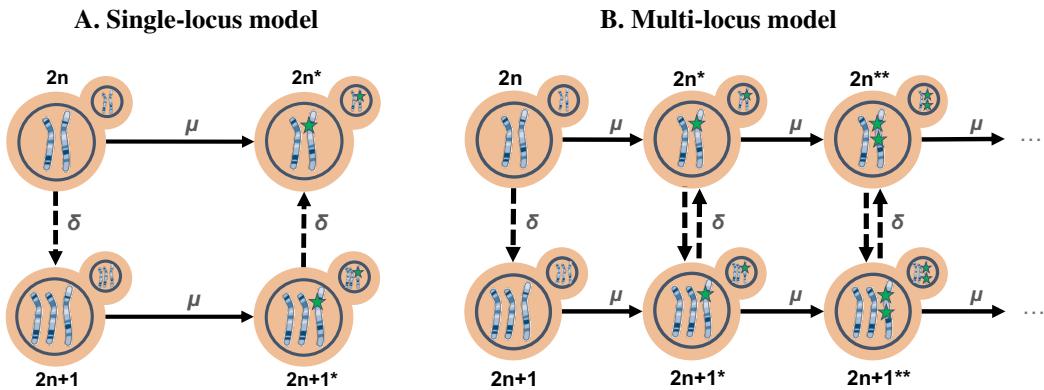
**The present study.** Here, we develop evolutionary-genetic models that include the effects of natural selection, genetic drift, aneuploidy, and mutation to examine the role of aneuploidy in adaptive evolution. These models follow a population of cells characterised by both their ploidy and their genotype. We fit these models to the experimental results of Yona et al. (2012) using an *approximate Bayesian computation* framework (Sisson et al., 2007) to infer model parameters, including selection coefficients and rates of aneuploidy and mutation, and to perform model selection between different models, thereby testing hypotheses about the evolutionary process. We analyze these evolutionary-genetic models to estimate the effects of parameters on the adaptation time and the probability for adaptation via aneuploidy. We find that

# Models and Methods

**Evolutionary Models.** We model the evolution of a population of cells using two models: a single-locus model and a multi-locus model. Both models are based on the Wright-Fisher model (Otto and Day, 2007), assuming non-overlapping generations and including the effects of natural selection, genetic drift, aneuploidy, and mutation. We focus on beneficial mutations, neglecting the effects of deleterious and neutral mutations. Both models allow for a single aneuploid karyotype (e.g., chromosome III duplication). While the single-locus model allows for only a single mutation to accumulate in the genotype, the multi-locus model allows for multiple mutations to accumulate (fig. 1), as well as for a fluctuating population size.

**Single-locus model.** This model assumes a constant effective population size  $N$  and follows four genotypes (fig. 1A): euploid wild-type,  $2n$ , the initial genotype; euploid mutant,  $2n^*$ , with the standard karyotype and a single beneficial mutation; aneuploid wild-type,  $2n+1$ , with an extra chromosome, i.e., following chromosome duplication; and aneuploid mutant,  $2n+1^*$ , with an extra chromosome and a beneficial mutation.

Transitions between the genotypes occur as follows (fig. 1A): Beneficial mutations from  $2n$  to  $2n^*$  and from  $2n+1$  to  $2n+1^*$  occur with probability  $\mu$ , the mutation rate. We neglect back-mutations (i.e., from  $2n^*$  to  $2n$  and from  $2n+1^*$  to  $2n+1$ ). Aneuploidy is formed by chromosome mis-segregation, so that cells transition from  $2n$  to  $2n+1$  and from  $2n+1^*$  to  $2n^*$  with probability  $\delta$ , the aneuploidy



**Figure 1: Model illustrations.** **(A)** In the single-locus model, the four genotypes are: euploid wild-type,  $2n$ ; euploid mutant,  $2n^*$ ; aneuploid wild-type,  $2n+1$ ; and aneuploid mutant,  $2n+1^*$ . Overall there are two possible trajectories from  $2n$  to  $2n^*$ . **(B)** In the multi-locus model, each genotype is characterized by its karyotype,  $2n$  or  $2n+1$ , and the number of accumulated beneficial mutations, denoted by stars. In both panels arrows denote transitions between genotypes, with transitions rates:  $\mu$ , beneficial mutation rate;  $\delta$ , aneuploidy rate.

rate. That is, we assume chromosomes are gained and lost at the same rate, and we neglect events that form a less-fit genotype (i.e.,  $2n+1$  to  $2n$  and  $2n^*$  to  $2n+1^*$ ). The fitness values of the four genotypes are given by table 1.

**Table 1: Single-locus model fitness values.**

Genotype $i$	$2n$	$2n + 1$	$2n + 1^*$	$2n^*$
Fitness $w_i$	1	$1 - c + b$	$(1 - c)(1 + s) + b$	$1 + s$

$s \geq 0$  is the selection coefficient of a beneficial mutation;  $0 \leq c \leq 1$  is the fitness cost of aneuploidy; and  $b \geq c$  is the selection coefficient, or fitness benefit, of aneuploidy.

The initial population has  $N$  cells with genotype  $2n$ . The effect of natural selection on the frequency  $f_i$  of genotype  $i = 2n, 2n + 1, 2n + 1^*$ , or  $2n^*$  is given by

$$f_i^s = \frac{f_i w_i}{\bar{w}}, \quad (1)$$

where the fitness values  $w_i$  are given in table 1 and  $\bar{w} = \sum_j f_j w_j$  is the population mean fitness. The effect of mutation and aneuploidy on genotype frequencies is given by

$$\begin{aligned} f_{2n}^m &= (1 - \delta - \mu) f_{2n}^s, \\ f_{2n+1}^m &= \delta f_{2n}^s + (1 - \mu) f_{2n+1}^s, \\ f_{2n+1^*}^m &= \mu f_{2n+1}^s + (1 - \delta) f_{2n+1^*}^s, \\ f_{2n^*}^m &= \mu f_{2n}^s + \delta f_{2n+1}^s + f_{2n^*}^s. \end{aligned} \quad (2)$$

Finally, random genetic drift is modeled using a multinomial distribution (Otto and Day, 2007),

$$\mathbf{f}' \sim \frac{1}{N} \cdot \text{Mult}(N, \mathbf{f}^m), \quad (3)$$

where  $\mathbf{f}^m = (f_{2n}^m, f_{2n+1}^m, f_{2n+1^*}^m, f_{2n^*}^m)$  are the frequencies of the genotypes after mutation and aneuploidy,  $\mathbf{f}'$  are the genotype frequencies in the next generation, and  $\text{Mult}(N, \mathbf{f})$  is a multinomial distribution parameterized by the population size  $N$  and the genotype frequencies  $\mathbf{f}$ . Overall, the change in genotype frequencies from one generation to the next is given by the transformation  $f_i \rightarrow f'_i$ .

**Multi-locus model.** This model expands the single-locus model by allowing for (i) the accumulation of beneficial mutations, and (ii) a fluctuating population size.

A genotype is characterized by its karyotype,  $2n$  or  $2n+1$ , and the number of accumulated beneficial mutations, which can be zero or more. The selection coefficient of the  $i$ -th accumulated mutation in each individual,  $s_i$ , is drawn from an exponential distribution with expected value  $s$ :  $s_i \sim \text{Exp}(s)$ . The

rest of the parameters ( $N$ ,  $\mu$ ,  $\delta$ ,  $b$ ,  $c$ ) are the same as in the single-locus model. However, since the multi-locus model allows several mutations with smaller fitness effects to accumulate, we expect the mutation rate to be higher compared to the single-locus model, which focuses on a single, large-effect mutation.

The fitness of the different genotypes is the same as in the single-locus model (table 1), except that the fitness contribution of  $k$  beneficial mutations is the product of their independent effects,  $\prod_{i=1}^k (1 + s_i)$ , instead of the contribution of the single mutation allowed in the single-locus model,  $(1 + s)$ , see table 2. Therefore, aneuploidy loss would be favored by selection only if there are enough beneficial mutations and/or the selection coefficients  $s_i$  are large enough. The intuition is that when the benefit of the accumulated beneficial mutations is small, then the benefit of aneuploidy has a large effect; when the benefit of the accumulated beneficial mutations is large, then aneuploidy is no longer beneficial due to its fitness cost.

In contrast to the single-locus model, in the multi-locus model the population size fluctuates to model serial-transfer experimental protocol (Yona et al., 2012): the population is serially diluted by transferring a fraction of the population (1/120) to a fresh medium approximately every seven generations. Thus, the population initial size is  $N_0 = N$ , and the population size is doubled every generation,  $N_1 = 2N, N_2 = 4N, \dots$ , and diluted back to  $N$  after seven generations such that  $N_8 = N$ .

The change in frequencies due to selection is exactly the same as in the single-locus model (eq. (1)), only applied using the fitness values in table 2. The change due to random genetic drift is also the same as in eq. (3), except that the frequencies vector is  $\mathbf{f} = (f_{2n}, f_{2n+1}, f_{2n^*}, f_{2n+1^*}, f_{2n^{**}}, f_{2n+1^{**}}, \dots)$  and that the population size changes between generations, as described above.

The effects of mutation and aneuploidy on genotype frequencies is more elaborate than in the single-locus model. Genotype  $i$  is classified according to its karyotype ( $2n$  or  $2n+1$ ), the number of accumulated beneficial mutations ( $k \geq 0$ ), and their fitness ( $w_i$ ). Each offspring cell inherits these properties from its mother cell. Then, a new beneficial mutation is accumulated with probability  $\mu$ , such that the number of mutations is  $k + 1$ , and its effect  $s_{k+1}$  is drawn from an exponential distribution with expected value  $s$ , such that the contribution of the mutations to the fitness is  $\prod_{j=0}^{k+1} (1 + s_j)$ . Next, euploid offspring become aneuploid with probability  $\delta$ , and aneuploid offspring become euploid with probability  $\delta_L$ .

**Table 2: Multi-locus model fitness values.**

<i>Genotype, g</i>	$2n$	$2n + 1$	$2n + 1^{*k}$	$2n^{*k}$
<i>Fitness, <math>w_g</math></i>	1	$1 - c + b$	$(1 - c) \prod_{j=1}^k (1 + s_i) + b$	$\prod_{j=1}^k (1 + s_i)$

$k$  is the number of accumulated beneficial mutations;  $s \geq 0$  is the selection coefficient of a beneficial mutation;  $0 \leq c \leq 1$  is the fitness cost of aneuploidy; and  $b \geq c$  is the selection coefficient, or fitness benefit, of aneuploidy.

**Empirical evidence.** We use the results of evolutionary experiments reported by Yona et al. (2012). In their heat-stress experiment, four populations of *S. cerevisiae* evolved under 39 °C. Aneuploidy fixed in all four population in the first 450 generations (hereafter, fixation or elimination of a genotype by *generation t* means that more than 95% or less than 5% of the population carry the genotype at generation  $t$ , and possibly earlier). From unpublished results, aneuploidy did not fix before at least 200 generations elapsed. The experiment continued with two populations, in which aneuploidy was eliminated by generation 1,700 and 2,350.

**Likelihood function.** Because our model, just like the Wright-Fisher model, is non-linear and stochastic, computing the distribution of fixation time  $T(g)$  of genotype  $g$  for use in the likelihood function is intractable (it is even hard to use a diffusion-equation approximation due to the model having multiple genotypes, rather than just two). We overcome this problem by approximating the likelihood using simulations. We simulate 1,000 experiments per parameter vector  $\theta = (\mu, \delta, s, b, c)$ , resulting in a set of simulated observations  $\tilde{\mathbf{X}} = \{\tilde{X}_i\}_{i=1}^{1000}$ . We then compute the approximate likelihood,

$$\begin{aligned} \mathcal{L}(\theta) = P^4(200 \leq T(2n+1) \leq 450) \cdot & \left[ 1 - \right. \\ & P_{\tilde{\mathbf{X}}}^4(\{\{T(2n^*) < 1700\} \mid 200 \leq T(2n+1) \leq 450\}) - \\ & P_{\tilde{\mathbf{X}}}^4(\{\{1700 < T(2n^*) < 2350\} \mid 200 \leq T(2n+1) \leq 450\}) + \\ & \left. P_{\tilde{\mathbf{X}}}^4(\{\{T(2n^*) < 1700\} \wedge \{\{1700 < T(2n^*) < 2350\} \mid 200 \leq T(2n+1) \leq 450\}\} \right], \end{aligned} \quad (4)$$

where  $\{\dots\}$  is the "logical not" operator,  $P^4(\dots)$  is the 4th power of  $P(\dots)$ , and all probabilities  $P_{\tilde{\mathbf{X}}}(\dots)$  are approximated from the results of the simulations  $\tilde{\mathbf{X}}$ . For example,  $P_{\tilde{\mathbf{X}}}(\{\{T(2n^*) < 1700\} \mid 200 \leq T(2n+1) \leq 450\})$  is approximated by taking simulations in which  $2n+1$  fixed before generation 450 but not before generation 200, and computing the fraction of such simulations in which  $2n^*$  did not fix by generation 1,700, and hence aneuploidy did not extinct before generation 1,700. fig. S4 compares results with less and more simulated experiments, demonstrating that 1,000 simulations are likely enough.

For a model without aneuploidy (that is, when the aneuploidy rate is fixed at zero,  $\delta = 0$ ), the likelihood

is similarly approximated by

$$\begin{aligned}\mathcal{L}_!(\theta) = 1 - P_{\tilde{\mathbf{X}}}^4(!\{T(2n^*) < 1700\}) - \\ P_{\tilde{\mathbf{X}}}^4(!\{1700 < T(2n^*) < 2350\}) + \\ P_{\tilde{\mathbf{X}}}^4(!\{T(2n^*) < 1700\} \wedge !\{1700 < T(2n^*) < 2350\}) .\end{aligned}\quad (5)$$

**Parameter inference.** To infer model parameters, we use approximate Bayesian computation with a sequential Monte-Carlo scheme, or ABC-SMC (Sisson et al., 2007), implemented in the pyABC Python package (Klinger et al., 2018, [pyabc.readthedocs.io](https://pyabc.readthedocs.io)). This approach uses numerical stochastic simulations of the model to infer a posterior distribution over the model parameters. It is a method of likelihood-free, simulation-based inference (Cranmer et al., 2020), that is, for estimating a posterior distribution when a likelihood function cannot be directly computed. It is therefore suitable in our case, in which the likelihood function can only be approximated from simulations, and cannot be directly computed.

The ABC-SMC algorithm employs sequential importance sampling over multiple iterations (Toni et al., 2009; Klinger and Hasenauer, 2017; Syga et al., 2021). In iteration  $t$  of the algorithm, a set of parameter vectors,  $\{\theta_{i,t}\}_{i=1}^{n_t}$ , also called *particles*, are constructed in the following way. A proposal particle,  $\theta^*$ , is sampled from a proposal distribution, and is either accepted or rejected, until  $n_t$  particles are accepted. The number of particles,  $n_t$ , is adapted at every iteration  $t$  using the adaptive population strategy (Klinger et al., 2018, [pyabc.readthedocs.io](https://pyabc.readthedocs.io)). For  $t = 0$ , the proposal particle is sampled from the prior distribution,  $p(\theta)$ . For  $t > 0$ , the proposal particle is sampled from the particles accepted in the previous iteration,  $\{\theta_{i,t-1}\}_{i=1}^{n_{t-1}}$ , each with a probability relative to its weight  $W_{t-1}(\theta_{i,t-1})$  (see below). The proposal particle is then perturbed using a kernel perturbation kernel,  $K_t(\theta^* | \theta)$  where  $\theta$  is the sample from the previous iteration. Then, a set of synthetic observations  $\tilde{\mathbf{X}}^*$  is simulated, and the proposal particle  $\theta^*$  is accepted if its approximate likelihood (eq. (4)) is high enough,  $\mathcal{L}(\theta^*) > 1 - \epsilon_t$  (or more commonly, if  $1 - \mathcal{L}(\theta^*) < \epsilon_t$ ), where  $\epsilon_t > 0$  is the *acceptance threshold*, as higher values of  $\epsilon_t$  allow more particles to be accepted. The acceptance threshold  $\epsilon_t$  is chosen as the median of the  $1 - \mathcal{L}(\theta)$  of the particles accepted in the previous iteration,  $t - 1$ , and  $\epsilon_0 = 0.01$ . For each accepted particle  $\theta_{i,t}$  a weight  $W_t(\theta_{i,t})$  is assigned: for  $t = 0$ ,  $W_0(\theta_{i,0}) = 1$ , and for  $t > 0$ ,  $W_t(\theta_{i,t}) = p(\theta_{i,t}) / \sum_{i=1}^{n_{t-1}} W_{t-1}(\theta_{i,t-1}) K_t(\theta_{i,t}, \theta_{i,t-1})$ , where  $p(\theta)$  is the prior density of  $\theta$  and  $K_t(\theta', \theta)$  is the probability of a perturbation from  $\theta$  to  $\theta'$ .  $K_t(\theta' | \theta)$  is a multivariate normal distribution, fitted at iteration  $t$  to the particles from the previous iteration,  $\{\theta_{i,t-1}\}_{i=1}^{n_{t-1}}$ , and their weights,  $\{W(\theta_{i,t-1})\}_{i=1}^{n_{t-1}}$ .

Acceptance is determined according to the approximate likelihood (eq. (4)), which has a maximum

value of 0.875. Thus, we terminated the inference when  $\epsilon \leq 0.13$  after six iterations, with  $n_6 = 982$  accepted parameter vectors and effective sample size ESS=651 (fig. S3). Running the inference algorithm with different initialization seeds and less or more simulations for approximating the likelihood produced similar posterior distributions (fig. S4).

After producing a set of weighted particles from the the posterior distribution using the above ABC-SMC algorithm, we approximate the posterior using kernel density estimation (KDE) with Gaussian kernels, from which we find the MAP (maximum a posteriori) estimate as the maximum of the KDE function. We then draw 50,000 samples from the posterior KDE to compute the HDI (highest density interval) and visualize the posterior distribution with histograms.

**Model comparison.** We perform model selection using WAIC, the widely applicable information criterion (Gelman et al., 2013),

$$WAIC(\theta) = -2 \log \mathbb{E}[\mathcal{L}(\theta)] + 2\mathbb{V}[\log \mathcal{L}(\theta)] \quad (6)$$

where  $\theta$  is a parameter vector, and  $\mathbb{E}[\cdot]$  and  $\mathbb{V}[\cdot]$  are the expectation and variance taken over the posterior distribution, which in practice are approximated using 50,000 samples from the posterior KDE. WAIC values are scaled as a deviance measure: lower values imply higher predictive accuracy (Kass and Raftery, 1995).

**Prior distributions.** We used informative prior distributions for  $w_{2n+1} = 1 - c + b$ ,  $w_{2n+1^*} = (1 + s)(1 - c) + b$  and  $w_{2n^*} = 1 + s$ , which we estimated from growth curves data from mono-culture growth experiments previously reported by Yona et al. (2012, Figs. 3C, 4A, and S2). We used Curveball, a method for predicting results of competition experiments from growth curve data (Ram et al., 2019, [curveball.yoavram.com](http://curveball.yoavram.com)). Briefly, Curveball takes growth curves of two strains growing separately in mono-culture and predicts how they would grow in a mixed culture, that is, it predicts the results of a competition assay. From these predictions, relative fitness values can be computed. Because Curveball uses a maximum-likelihood approach to estimate model parameters, we were able to estimate a distribution of relative fitness values by sampling from a truncated multivariate normal distribution defined by the maximum-likelihood covariance matrix. We sampled 10,000 samples to use as a prior distribution (fig. S1). We used growth curves of  $2n$  and  $2n+1$  in 39 °C to estimate a prior distribution for  $w_{2n+1}$  (fig. S1-D). In lieu of a better prior, we used the same prior for  $w_{2n+1^*}$  and  $w_{2n^*}$ . To increase computational efficiency, we also assumed  $w(2n^*) > w(2n+1^*) > w(2n+1) > w(2n)$ ; running the inference without this assumption produced similar results.

Compared to other priors we tested, this prior produced lower WAIC, better posterior prediction plots,

and more stable parameter estimates, as follows. We tried to use additional growth curves. We used growth curves of  $2n^*$  (*refined* strain from Yona et al. (2012)) and  $2n+1$  in 39 °C to estimate  $w_{2n^*}/w_{2n+1}$  (fig. S1-L). The same prior was used for  $w_{2n^*}/w_{2n+1*}$ . This prior resulted in WAIC 71.91, compared to 67.87 with the above informative prior. We also tried to use growth curves of  $2n+1$  and  $2n$  in 30 °C to estimate  $1 - c$  (fig. S1-H). This estimation assumes that the cost of aneuploidy is the same in 39 °C and 30 °C; this might be incorrect, but we only assumed this to generate a prior distribution for the fitness values. The prior for  $b$  was taken the same as for  $c$ . This prior did not converge: no parameter sets were found with an approximate likelihood greater than zero. Finally, we also tested a uninformative uniform prior with  $U(1, 6)$ , for (i) all  $w_{2n+1}$ ,  $w_{2n+1*}$ ,  $w_{2n^*}$  and (ii) only for  $w_{2n+1*}$ ,  $w_{2n^*}$ , using the above informative prior for  $w_{2n+1}$ . In both cases the algorithm failed to converge.

For the mutation rate,  $\mu$ , and aneuploidy rate,  $\delta$ , we used uninformative uniform priors,  $\mu \sim U(10^{-9}, 10^{-5})$  and  $\delta \sim U(10^{-6}, 10^{-2})$ . A wider mutation rate prior,  $\mu \sim U(10^{-9}, 10^{-3})$ , produced similar results.

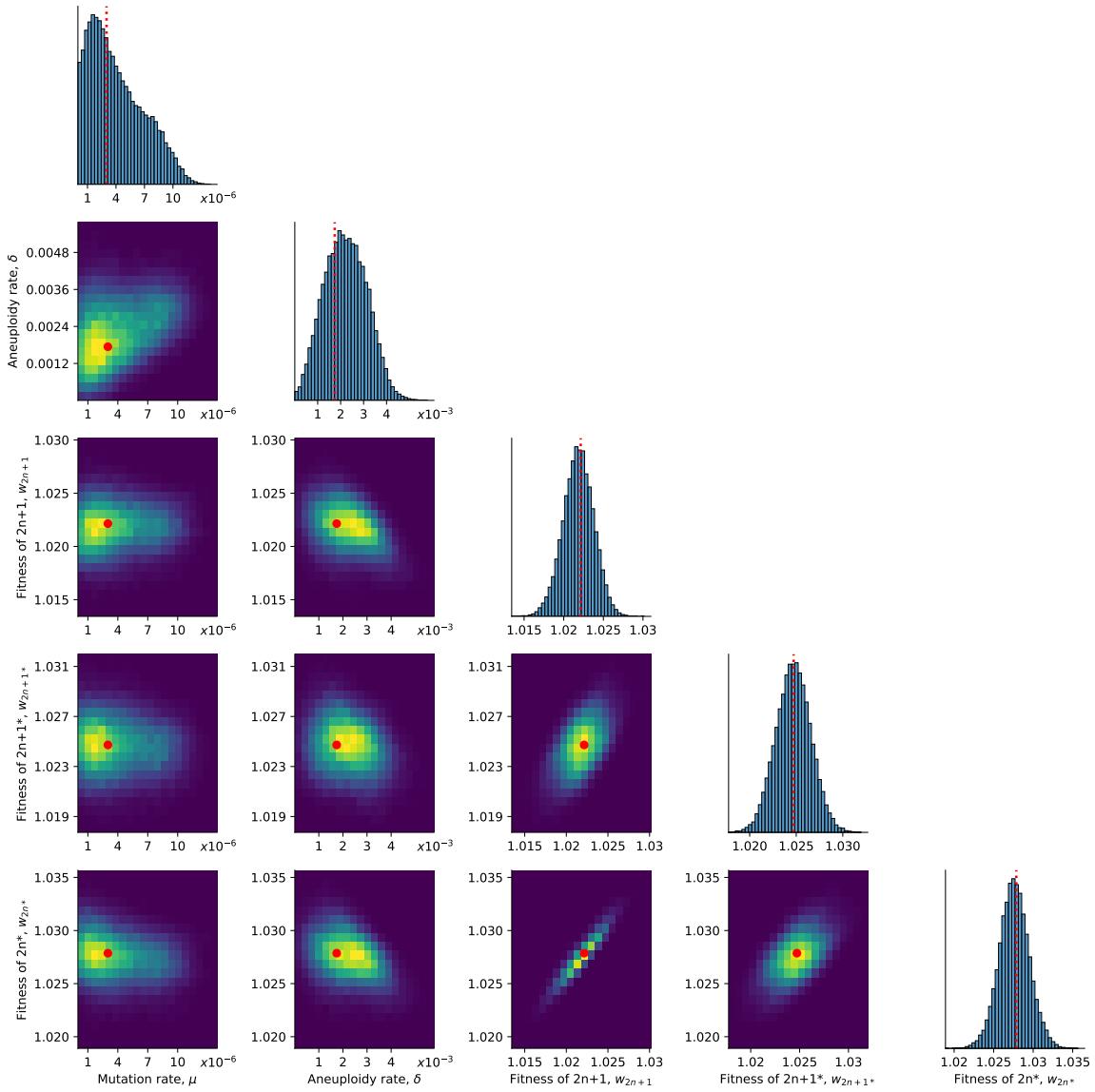
## Results

### Inference: single-locus model

**Parameter estimation.** We used ABC-SMC to infer the posterior distribution of the parameters of the single-locus model. The MAP (maximum a posteriori) and 50% HDI (highest density interval) of the parameters posterior are (fig. 2): mutation rate,  $\mu = 2.942 \times 10^{-6} (2.017 \times 10^{-7} - 4.15 \times 10^{-6})$ , aneuploidy rate,  $\delta = 1.722 \times 10^{-3} (1.394 \times 10^{-3} - 2.754 \times 10^{-3})$ , and fitness values,  $w_{2n+1} = 1.022 (1.021 - 1.023)$ ,  $w_{2n+1*} = 1.025 (1.024 - 1.026)$ ,  $w_{2n^*} = 1.028 (1.026 - 1.029)$ , all relative to the fitness of  $2n$ , which is set to  $w_{2n} = 1$ .

This estimated aneuploidy rate agrees with previous estimates. The mutation rate corresponds to a mutation target size of  $10^4$ , assuming the mutation rate per base pair is roughly  $2 \cdot 10^{-10}$  (Zhu et al., 2014) or  $3.3 \cdot 10^{-10}$  (Lynch et al., 2008).

**Model checking and comparison.** The single-locus model fits the data well: in simulations using the best-fit parameters (MAP estimate)  $2n^*$  fixed in 61% of simulations by generation 1,700 and in 100% of simulations by generation 2,350 (fig. 3-B). Interestingly, the genotype frequency dynamics in these simulations demonstrate that  $2n+1^*$  never reaches substantial frequency (fig. 4). Furthermore, sensitivity analysis shows that changing a single parameter while keeping the rest fixed at the MAP



**Figure 2: Posterior distribution of single-locus model.** On the diagonal, the inferred posterior distribution of each model parameter. Below the diagonal, the inferred joint posterior distribution of pairs of model parameters (dark purple and bright yellow for low and high density, respectively). Red markers and orange lines for the joint MAP estimate (which may differ from the marginal MAP, as the marginal distribution integrates over all other parameters).

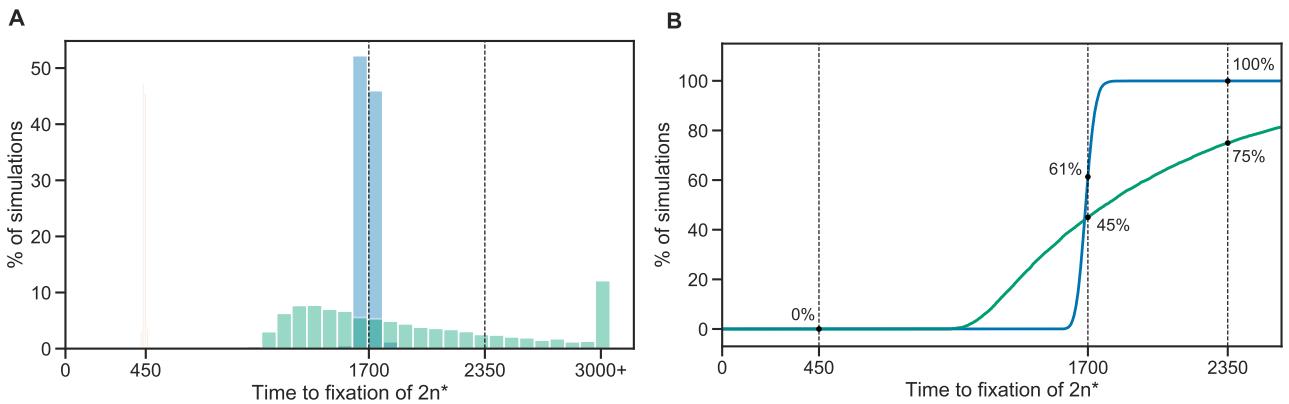
estimate produces a worse fit to the data (fig. S2).

However, a model without aneuploidy where the aneuploidy rate is fixed at zero,  $\delta = 0$ , cannot explain the experimental observations (Figure 3). The estimated parameter values without aneuploidy are: mutation rate,  $\mu = 7.979^{+0.156}_{-0.076} \times 10^{-9}$ , fitness of mutant,  $w_{2n^*} = 1.013^{+0}_{-0.001}$ . The estimated mutation rate without aneuploidy is much lower compared to a model with aneuploidy, suggesting a target size of just 40, because a higher mutation rate will lead to faster fixation of  $2n^*$  than in the experimental observations (fig. 3). Thus, even disregarding the appearance of trisomic cells in the experiment, the

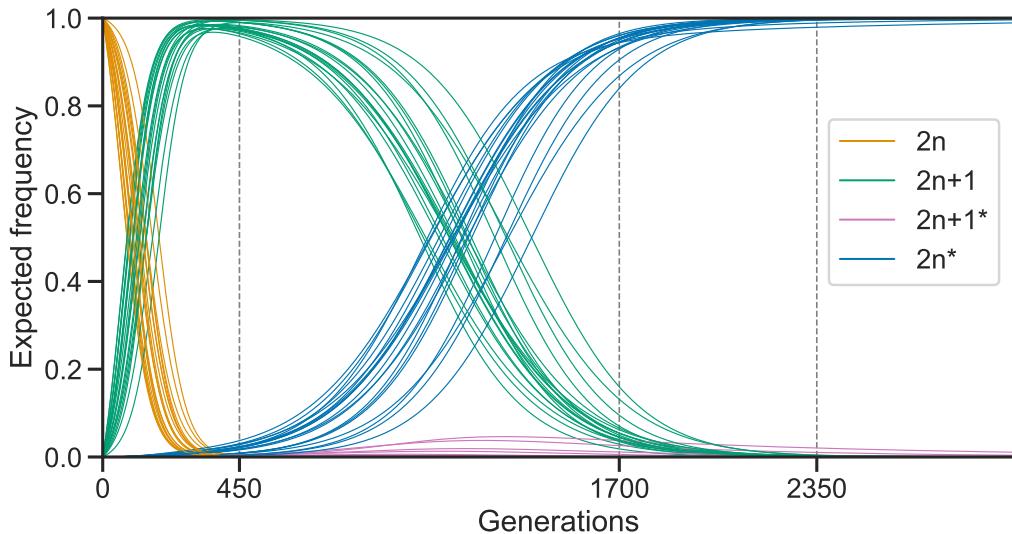
evidence supports the inclusion of aneuploidy in the model.

We also inferred model parameters under the assumption that the mutation rate in aneuploid cells, compared to euploid cells, is increased by a factor  $\tau=1$ , 33/32 (due to an additional chromosome), 2, 5, 10 or 100 (due to genetic instability). We found that the posterior distribution was similar for  $\tau = 1$ , 33/32, 2, and 5 (fig. S5). With  $\tau = 10$ , the estimated mutation rate was slightly lower ( $\mu = 1.674 \times 10^{-6} (2.501 \times 10^{-7} - 1.741 \times 10^{-6})$ ). With  $\tau = 100$ , the estimated mutation rate was considerably lower ( $\mu = 3.81 \times 10^{-7} (1.508 \times 10^{-7} - 4.995 \times 10^{-7})$ ) and the aneuploidy rate had higher variance ( $\delta = 7.824 \times 10^{-4} (6.616 \times 10^{-4} - 1.462 \times 10^{-3})$ ). We cannot rule out an increased mutation rate in aneuploid cells, but we find no evidence that it does, and unless the effect is strong, it does not seem to affect our results. We also checked the differences in fixation/loss times of  $2n$ ,  $2n+1$  and  $2n^*$  for the models with different  $\tau$ . We saw that the models can be distinguished using these parameters, especially  $\tau = 100$  (fig. 5).

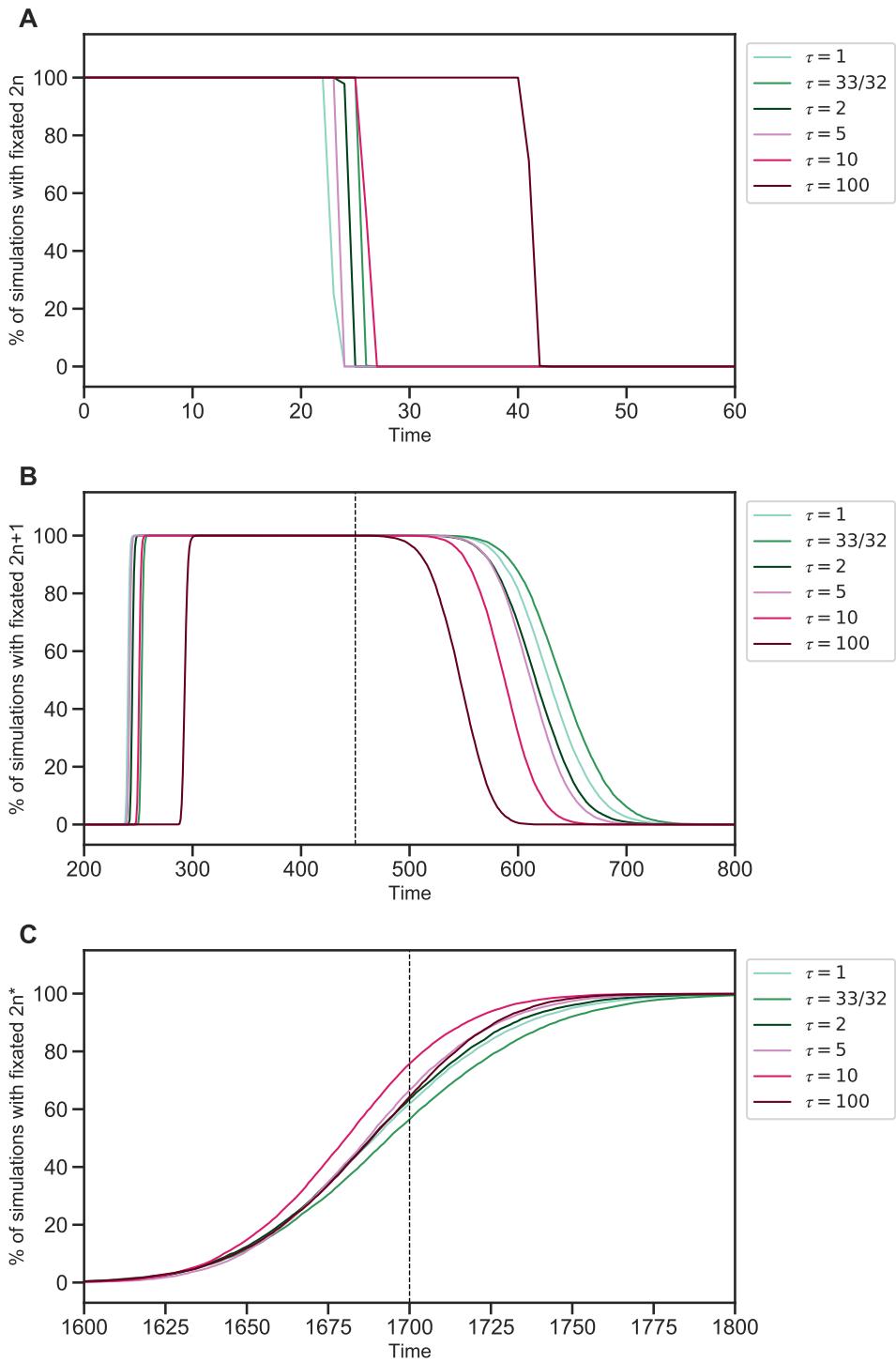
Next we tried the model with fixed mutation rate  $\mu$ . Hence  $\mu$  is not well identifiable by the model (the inferred value has high variance with no strong correlation between it and other parameters), we tried the models where mutation rate is fixed:  $\mu=10^{-5}$ ,  $10^{-6}$  and  $10^{-7}$ . We observed similar parameters inference comparing to the model with free  $\mu$  parameter. However WAIC was better for the models with fixed  $\mu$  (Table S1). The lowest WAIC got the model with  $\mu=10^{-5}$  - 0.6, it is lower than *no aneuploidy* model - 0.87, that implies higher predictive accuracy.



**Figure 3: Single-locus model fit with and without aneuploidy.** (A) The distribution of time to fixation of  $2n^*$  (i.e., adaptation time) in 10,000 simulations of the single-locus model with aneuploidy (blue; MAP parameters) compared to two models without aneuploidy: a model with the same parameter values except  $\delta = 0$  (orange), and a model fitted to the data assuming  $\delta = 0$  (green). In the experiment by Yona et al. (2012), one population lost aneuploidy by generation 1,700 and another by generation 2,350 (dashed lines) but not before generation 450. Thus, the blue distribution is a better fit compared to the green, and the yellow histogram has a very poor fit. The last bin contains all the simulations with time equal or greater than 3,000. (B) Cumulative distribution of the time to fixation of  $2n^*$  in 10,000 simulations using the MAP estimate with and without aneuploidy in blue and green, respectively, and corresponding to the blue and green bars in panel A. The MAP likelihood (eq. (4)) is 0.84 and 0.67 with and without aneuploidy, respectively.



**Figure 4: Posterior genotype frequency dynamics for the single-locus model.** The posterior prediction for the frequencies of the four genotypes over time. Each of the 20 curves is the average of 10,000 simulations of the single-locus model using parameters drawn from the posterior distribution.



**Figure 5: Genotype fixations for single-locus models with different  $\tau$ .** For each model, 10,000 simulations using MAP parameters were evaluated and a percentage of simulations with fixated genotypes  $2n$ ,  $2n+1$  and  $2n^*$  (plots A, B and C respectively) were measured in time.  $2n+1^*$  didn't reach fixation, therefore it doesn't appear in figures. We can see that  $\tau = 100$  can be distinguished using data of loss of  $2n$  (plot A) and fixation/loss of  $2n+1$  (plot B). It is harder to distinguish  $1 \leq \tau \leq 10$ , however the correlation exists.

### Analysis: single-locus model.

Here, we aim to approximate the adaptation time (i.e., time to fixation of  $2n^*$ ) in the single-locus model and how it is affected by the model parameters. As a stochastic model, there are two events

that can occur in the  $2n$  population. The first is a mutation that produces a  $2n^*$  cell. The second is a missegregation event that produces a  $2n+1$  cell. The probability with which these events occur is significantly different. We are interested in the average waiting time until one of the two events occurs and then lead to adaptation (in the second case, it will be followed by a mutation to  $2n+1^*$  and loss of aneuploidy to  $2n^*$ ). We assume the waiting time for appearance of  $2n^*$  from  $2n$  and appearance of  $2n+1$  from  $2n$  are geometrically distributed with rates  $p_{2n^*}$  and  $p_{2n+1}$ , respectively, where

$$p_{2n^*} = 1 - (1 - 2\mu\Delta w_{2n^*})^N, \quad (7a)$$

$$p_{2n+1} = 1 - (1 - 2\delta\Delta w_{2n+1})^N, \quad (7b)$$

where  $\Delta w_i$  denotes the relative difference between the fitness of genotype  $i$  and the population mean fitness. However, we have two competing events: if one event occurs (e.g., mutation) then the other (aneuploidy) is unlikely to occur, we need to find the minimum of these waiting times. The minimum of two geometric random variables with parameters  $p_{2n^*}$  and  $p_{2n+1}$  is also a geometric random variable with parameter  $1 - (1 - p_{2n^*})(1 - p_{2n+1})$ , which is the inverse of the expected waiting time until the first event occurs. The probability that aneuploidy will be the first event to occur is

$$p_A = \frac{p_{2n+1}}{p_{2n^*} + p_{2n+1}}. \quad (8)$$

We are interested in the expected time to fixation of  $2n^*$ , and the events leading to this fixation depend on the event that occurs first. If  $2n^*$  appears first and avoids extinction by drift, then it will fix. If  $2n+1$  appears first and avoids extinction, then several other events must occur for fixation of  $2n^*$ . Thus, in this aneuploid trajectory, we need to approximate the expected waiting time to reach  $2n+1^*$  and finally  $2n^*$ . In general, we assume that the waiting time until a new genotype occurs is much longer than the time to its fixation. Nonetheless, these fixation times must be included in a rigorous analysis. There are two ways to compute the fixation times and they differ in the assumption of the underlying process.

A deterministic-process approach is given by the classical bi-allelic single-locus equation, which provides an approximation for the expected time for genotype  $i$  to reach frequency  $x$  when starting from frequency  $1 - x$  in a population of genotype  $j$  is

$$\tilde{T}_{i \rightarrow j} = 2 \frac{\log(\frac{x}{1-x} N)}{\log(w_i/w_j)}, \quad (9)$$

A stochastic-process approach for approximating  $\tilde{T}_i$  is given by the diffusion-equation approximation (Kimura and Ohta, 1969, eq. 17). Thus, the expected adaptation time via the aneuploidy trajectory is given by

$$T_A = \frac{1}{1 - (1 - p_{2n^*})(1 - p_{2n+1})} + \frac{1}{p_{2n+1^*}} + \frac{1}{p_{2n^*}} + \tilde{T}_{2n \rightarrow 2n+1} + \tilde{T}_{2n+1 \rightarrow 2n+1^*} + \tilde{T}_{2n+1^* \rightarrow 2n^*}, \quad (10)$$

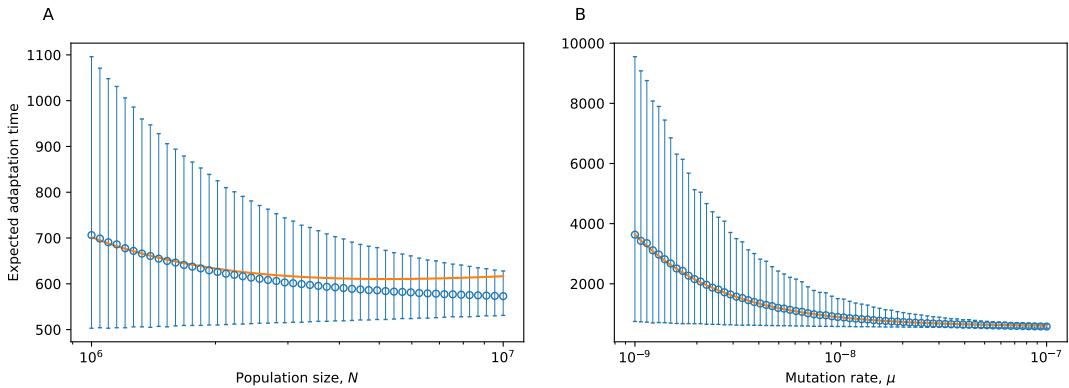
the expected adaptation time via the mutation trajectory is given by

$$T_M = \frac{1}{1 - (1 - p_{2n^*})(1 - p_{2n+1})} + \tilde{T}_{2n \rightarrow 2n^*}, \quad (11)$$

and the expected adaptation time is given by

$$T = p_A T_A + (1 - p_A) T_M. \quad (12)$$

To validate the approximation in eq. (12), we performed simulations of the full model (eqs. (1) to (3)) without aneuploidy,  $\delta = 0$  (fig. 6), and with the other parameters as inferred above. The approximation provides a good fit to simulation results except when the effective population size is larger than  $5 \cdot 10^6$  cells. In that case, the expected number of new mutants is 0.5, and therefore we can expect that in many generations the fixation event will start with more than a single  $2n^*$  cell appearing in the same generation, which is in contrast to our assumption that an adaptation event starts from a single mutant cell.



**Figure 6: Expected adaptation time without aneuploidy.** The expected time to appearance and fixation of  $2n^*$  as a function of the population size,  $N$  (panel A) and as a function of the mutation rate,  $\mu$  (panel B). Mean and 95% CI from simulations in blue. Analytic approximation (eq. (12)) in orange. Here,  $w_{2n+1} = 1.021$ ,  $w_{2n+1^*} = 1.025$  and  $w_{2n^*} = 1.028$ . In panel A,  $\mu = 9.6 \cdot 10^{-8}$  and in panel B,  $N = 6 \cdot 10^6$ .

## Discussion

**Aneuploidy is not just another type of mutation.** The published data indicate that, like mutation, aneuploidy can be both deleterious and beneficial (Pavelka et al., 2010; Sheltzer and Amon, 2011). Nevertheless, there are important and fundamental differences between adaptation by aneuploidy and adaptation by beneficial mutations (Yona et al., 2015), which make aneuploidy a unique mechanism for generating genetic variation. First, the aneuploidy rate (i.e. the frequency of mis-segregation events) is significantly higher than the mutation rate (Santaguida and Amon, 2015). Thus, everything else being equal, adaptation by aneuploidy will be faster and more frequent. Second, fitness effects of aneuploidy are larger than those of the majority of mutations, on average, and are rarely neutral (Pavelka et al., 2010; Yona et al., 2012; Sunshine et al., 2015), allowing selection to quickly sort deleterious and beneficial genotypes. Third, the number of different karyotypes is considerably smaller than the number of different genotypes, and different karyotypes are likely to have different phenotypes (Pavelka et al., 2010). Therefore, exploration of the phenotype space by aneuploidy requires smaller populations and a shorter time span. Fourth, aneuploidy is a reversible state, as the rate of chromosome loss is high and the cost of aneuploidy is significant (Niwa et al., 2006). Indeed, aneuploidy often provides a transient solution: under short-term stress conditions, aneuploidy reverts (chromosome number returns to normal) when the stress subsides; under long-term stress conditions, aneuploidy reverts when refined solutions, generated by beneficial mutations, take over (Yona et al., 2012). Finally, aneuploidy results in increased genome instability, potentially increasing genetic variation by a positive feedback loop (Rancati and Pavelka, 2013; Bouchonville et al., 2009; Zhu et al., 2012), while also increasing its own transience.

**Evolutionary theory of aneuploidy.** The role of aneuploidy in adaptation has only recently been observed (Sionov et al., 2010; Yona et al., 2012; Gerstein et al., 2015), and is largely missing from the literature on evolution and adaptation: the introductory textbook *Evolution* by Bergstrom and Dugatkin (2012) does not mention the word aneuploidy, and the graduate-level book *Mutation-Driven Evolution* by Nei (2013) only briefly mentions aneuploidy in the context of speciation, but not adaptation. In recent reviews of the literature, aneuploidy is suggested to play an important role in fungal adaptation (Robbins et al., 2017; Todd et al., 2017) and cancer evolution (Santaguida and Amon, 2015; Naylor and van Deursen, 2016; Sansregret and Swanton, 2017), yet these reviews cite no theoretical studies nor any quantitative models. Indeed, evolutionary, ecological, and epidemiological studies mostly assume adaptation occurs via beneficial mutations, recombination, and sex. Therefore, there is a critical need to develop an evolutionary theory of aneuploidy like the evolutionary theories of other

mechanisms for generation of genetic variation, e.g. mutation (Lynch, 2010), recombination (Hartfield and Keightley, 2012), and sex (Otto, 2009). An evolutionary theory of aneuploidy will be central to the interpretation of experimental and clinical observations and design of new hypotheses, experiments, and treatments (Carja et al., 2014). For example, despite the lack of theoretical models, aneuploidy has been invoked in a new strategy to combat pathogens and tumour cells by setting “evolutionary traps” (Gerstein et al., 2015; Chen et al., 2015), in which a condition that predictably leads to emergence of aneuploidy is applied, followed by a condition that specifically selects against aneuploid cells.

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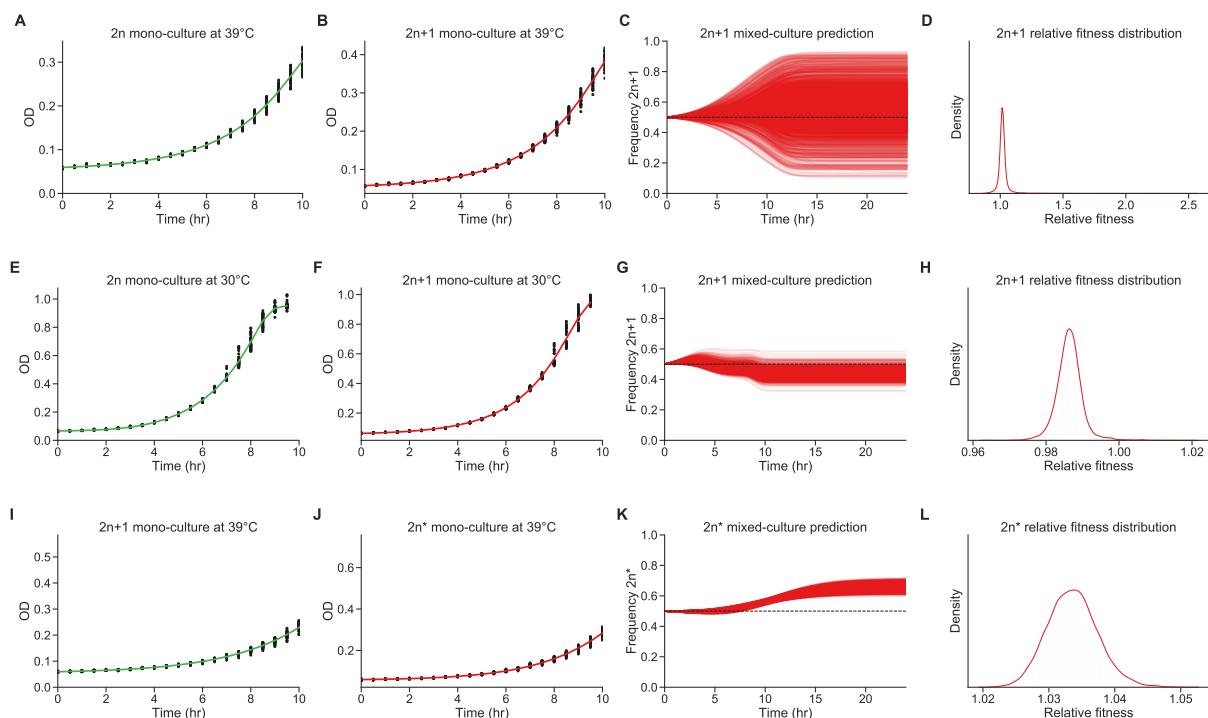
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## Supplementary Material

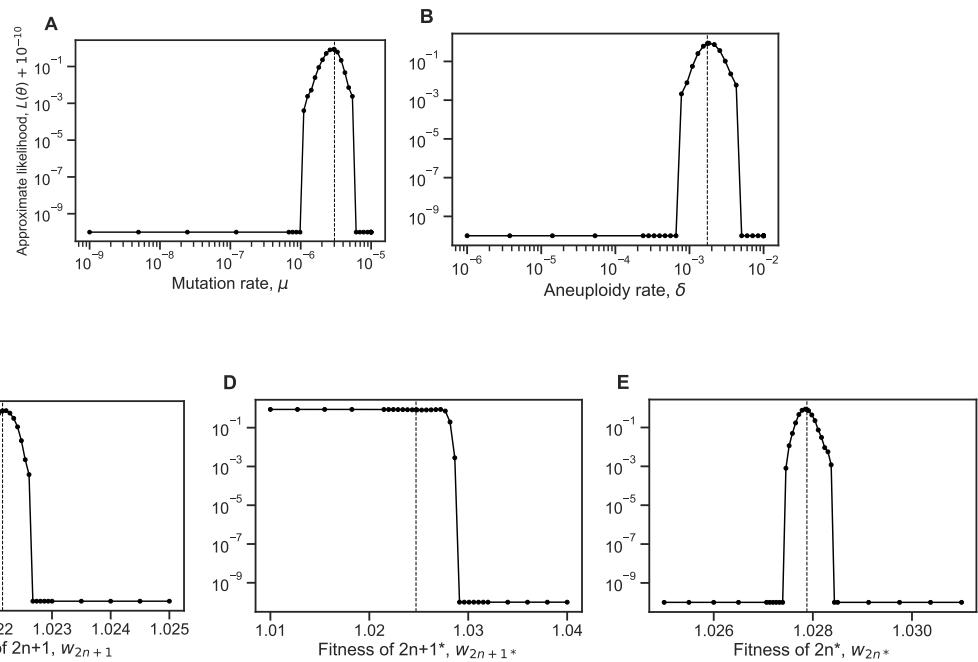


**Figure S1: Fitness estimation from growth curves.** (A-D) Fitness estimation from growth curves of  $2n$  and  $2n+1$  at  $39^{\circ}\text{C}$ :  $\hat{w}_{2n+1}/w_{2n}=1.024$  (95% CI: 0.959 - 1.115). (E-H) Fitness estimation from growth curves of  $2n$  and  $2n+1$  at  $30^{\circ}\text{C}$ :  $\hat{w}_{2n+1}/w_{2n}=0.986$  (95% CI: 0.979 - 0.994). (I-L) Fitness estimation from growth curves of  $2n+1$  and  $2n^*$  at  $39^{\circ}\text{C}$ :  $\hat{w}_{2n^*}/w_{2n+1}=1.033$  (95% CI: 1.027 - 1.041).

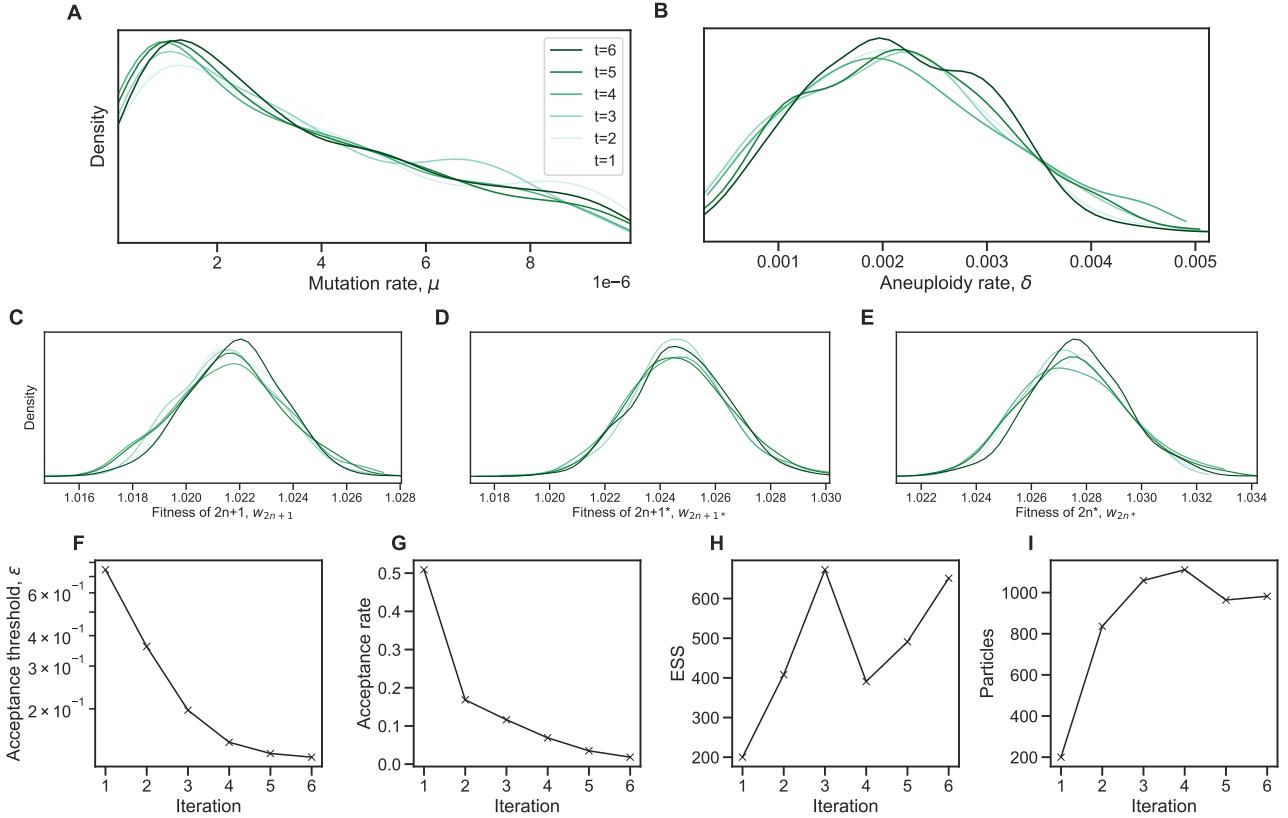
Growth curves previously described in Yona et al. (2012, Figs. 3C, 4A, and S2). Fitness estimated from growth curves using *Curveball*, a method for predicting results of competition experiments from growth curve data (Ram et al., 2019, [curveball.yoavram.com](http://curveball.yoavram.com)). See *Models and Methods, Prior distributions* for more details.

**Table S1: WAIC values for the different models.**

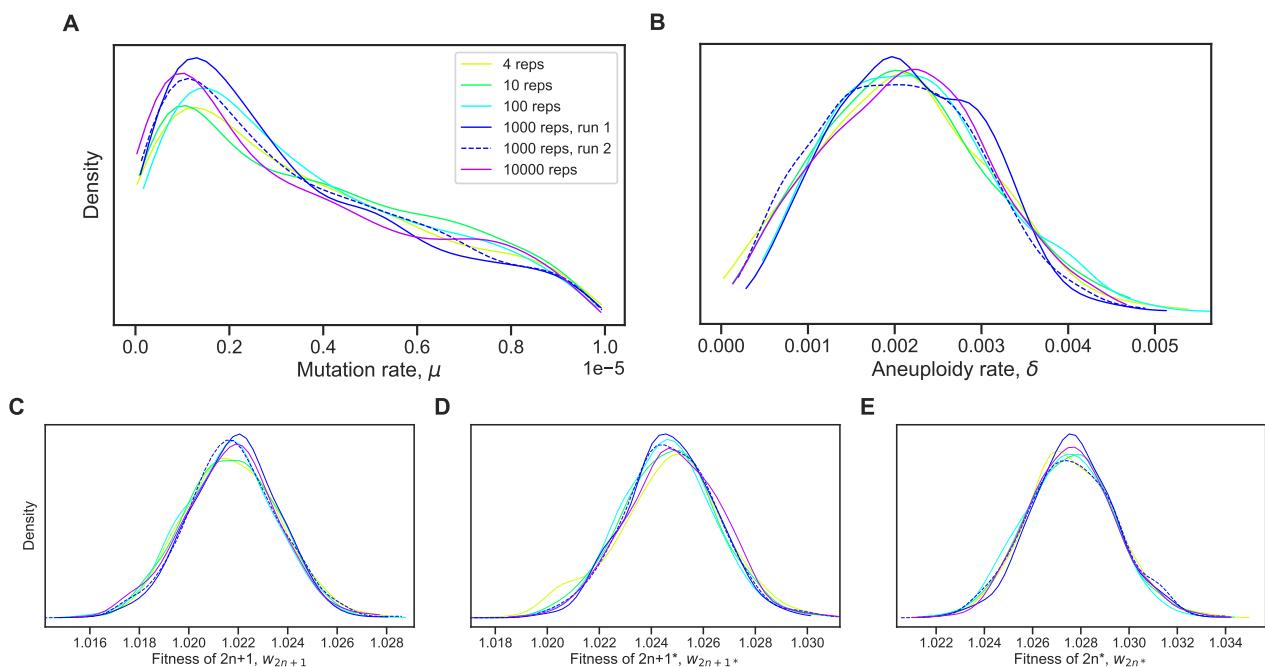
id	Model	WAIC
1	Without aneuploidy	0.83
2	Fixed mutation rate, $\mu = 10^{-5}$	0.60
3	Fixed mutation rate, $\mu = 10^{-6}$	33.89
4	Fixed mutation rate, $\mu = 10^{-7}$	29.58
5	Free mutation rate, $\tau = 1$	78.52
6	Free mutation rate, $\tau = 33/32$	65.92
7	Free mutation rate, $\tau = 2$	85.32
8	Free mutation rate, $\tau = 5$	71.89
9	Free mutation rate, $\tau = 10$	77.38
10	Free mutation rate, $\tau = 100$	57.67



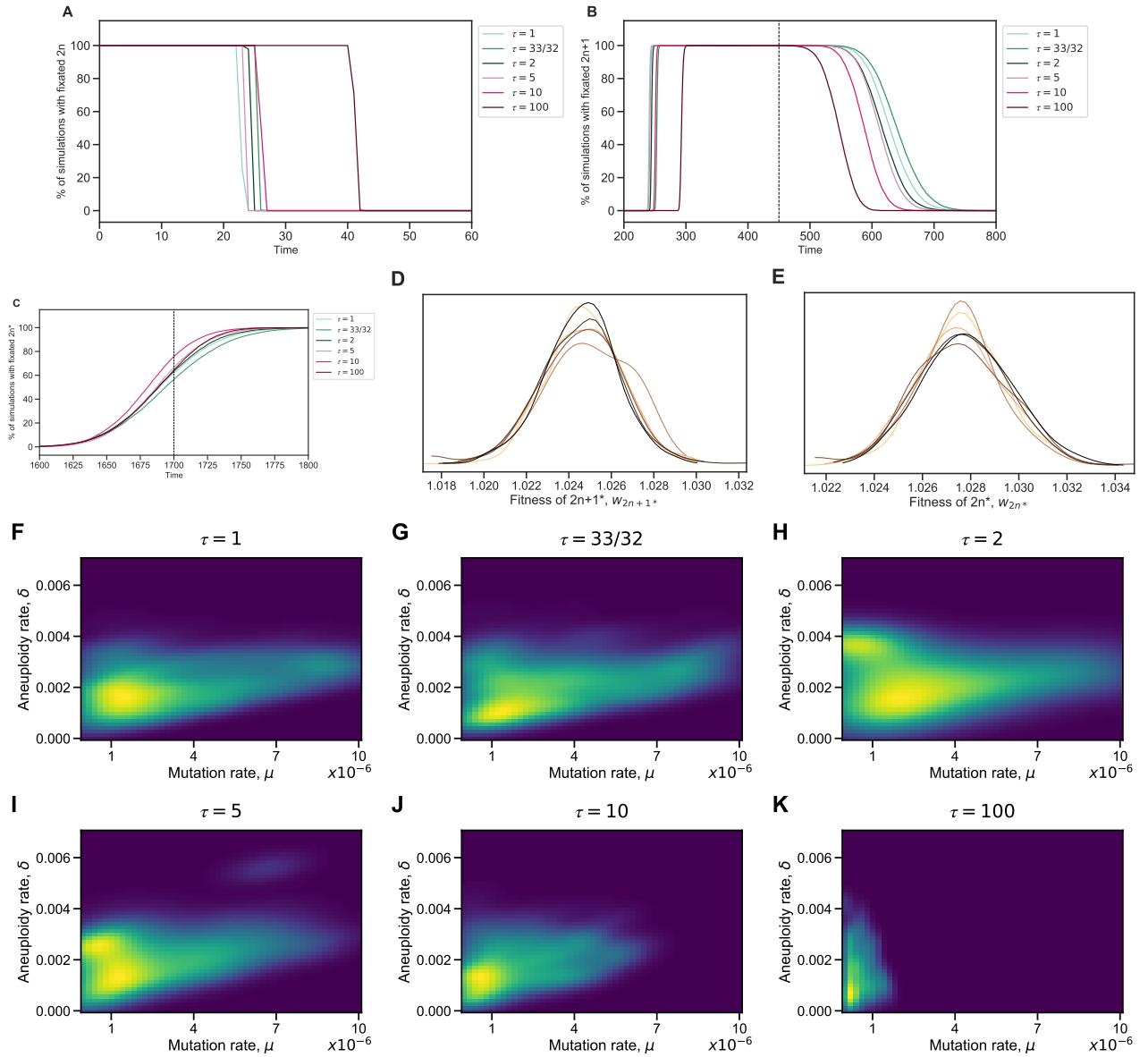
**Figure S2: Likelihood profiles.** Sensitivity of the approximate likelihood of the single-locus model,  $\mathcal{L}(\theta)$ , to changing a single parameter while the other parameters remain fixed at their MAP estimates. Dashed vertical line represents the MAP value. The prior distributions for the mutation rate and aneuploidy rate are  $\mu \sim U(10^{-9}, 10^{-5})$  and  $\delta \sim U(10^{-6}, 10^{-2})$ , respectively.



**Figure S3: Single-locus model inference convergence.** The ABC-SMC algorithm was used to infer the parameters of the single-locus model. **(A-E)** The approximate posterior distributions of model parameters at each iteration of the ABC-SMC algorithm demonstrates convergence, as the posterior did not significantly change after the first iteration,  $t = 1$ . **(F-I)** ABC-SMC measures of convergence. After iteration number 6, the acceptance threshold was  $\epsilon = 0.13$ , the acceptance rate was 0.018, the number of particles was 982, and the effective sample size ESS=651.



**Figure S4: Posterior distribution validation.** The posterior distribution of the single-locus model parameters is roughly the same regardless of the number of simulations (4-10,000 replicates) used to approximate the likelihood (eq. (4)).



**Figure S5: Single-locus model inference with elevated mutation rate in aneuploid cells.** (A-E) The inferred posterior distributions of model parameters for the single-locus model with different values of  $\tau$ , the fold-increase in mutation rate in aneuploid cells ( $2n+1$  and  $2n+1^*$ ). When the increase in mutation rate is high,  $\tau = 10$  and  $\tau = 100$ , the inferred mutation (A) and aneuploidy (B) rate tends to be lower. (F-K) The inferred joint posterior distribution of mutation rate ( $\mu$ ) and aneuploidy rate ( $\delta$ ) with different  $\tau$  values (dark purple and bright yellow for low and high density, respectively).