

# Aneuploidy can be an evolutionary diversion on the path to adaptation

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January 28, 2024

Classifications. Biological Sciences: Evolution, Genetics, Microbiology, Population Biology.

Keywords: whole-chromosome duplication, evolutionary model, adaptive evolution

**Abstract**

Aneuploidy is common in eukaryotes, often leading to decreased fitness. However, evidence from fungi and human tumor cells suggests that specific aneuploidies can be beneficial under stressful conditions and facilitate adaptation. In a previous evolutionary experiment with yeast, populations evolving under heat stress became aneuploid, only to later revert to euploidy after beneficial mutations accumulated. It was therefore suggested that aneuploidy is a “stepping stone” on the path to adaptation. Here, we test this hypothesis. We use Bayesian inference to fit an evolutionary model with both aneuploidy and mutation to the experimental results. We then predict the genotype frequency dynamics during the experiment, demonstrating that most of the evolved euploid population likely did not descend from aneuploid cells, but rather from the euploid wild-type population. Our model shows how the beneficial mutation supply—the product of population size and beneficial mutation rate—determines the evolutionary dynamics: with low supply, much of the evolved population descends from aneuploid cells; but with high supply, beneficial mutations are generated fast enough to outcompete aneuploidy due to its inherent fitness cost. Our results suggest that despite its potential fitness benefits under stress, aneuploidy can be an evolutionary “diversion” rather than a “stepping stone”: it can delay, rather than facilitate, the adaptation of the population, and cells that become aneuploid may leave less descendants compared to cells that remain diploid.

## 34 Introduction

Aneuploidy is an imbalance in the number of chromosomes in the cell: an incorrect karyotype.

36 Evidence suggests aneuploidy is very common in eukaryotes, e.g. animals<sup>52,41,3</sup>, and fungi<sup>44,77,48,66</sup>. Aneuploidy has been implicated in cancer formation, progression, and drug resistance<sup>6,54,52,51,25,37</sup>.

38 It is also common in protozoan pathogens of the *Leishmania* genus, a major global health concern<sup>39</sup>, and contributes to the emergence of drug resistance<sup>55</sup> and virulence<sup>40</sup> in fungal pathogens, which

40 are under-studied<sup>50</sup>, despite infecting a billion people per year, causing significant morbidity in >150 million and death in >1.5 million people per year<sup>55,50</sup>.

42 Experiments with human and mouse embryos found that most germ-line aneuploidies are lethal. Aneuploidies are also associated with developmental defects and lethality in other multicellular organisms<sup>58</sup>. For example, aneuploid mouse embryonic cells grow slower than euploid cells<sup>71</sup>. Similarly, in unicellular eukaryotes growing in benign conditions, aneuploidy usually leads to slower growth and

46 decreased overall fitness, in part due to proteotoxic stress due to increased expression, gene dosage imbalance, and hypo-osmotic-like stress<sup>42,69,44,58,53,29,76,70,72,49</sup>.

48 However, aneuploidy can be beneficial under stressful conditions due to the wide range of phenotypes it can produce, some of which are advantageous<sup>44,72</sup>. Indeed, in a survey of 1,011 yeast strains, 50 aneuploidy has been detected in about 19%<sup>45</sup>. Thus, aneuploidy can lead to rapid adaptation in unicellular eukaryotes<sup>18,68,23,47</sup>, as well as to rapid growth of somatic tumour cells<sup>54,60</sup>. For example,

52 aneuploidy in *Saccharomyces cerevisiae* facilitates adaptation to a variety of stressful conditions like heat and pH<sup>74</sup>, copper<sup>9,18</sup>, salt<sup>12,49</sup>, and nutrient limitation<sup>13,21,1</sup>, with similar results in *Candida*

54 *albicans*<sup>72</sup>. Importantly, aneuploidy can also lead to drug resistance in pathogenic fungi such as *C. albicans*<sup>57,56,17</sup> and *Cryptococcus neoformans*<sup>61</sup>, which cause candidiasis and meningoencephalitis,

56 respectively. Although we focus here on aneuploidy, a similar phenomena of adaptation via gene duplication or amplification has been observed in yeast<sup>36</sup>, bacteria<sup>64</sup>, and DNA viruses<sup>14</sup>.

58 Yona et al.<sup>74</sup> demonstrated experimentally the importance of aneuploidy in adaptive evolution. They evolved populations of *S. cerevisiae* under strong heat stress. The populations adapted to the heat stress

60 within 450 generations, and this adaptation was determined to be due a duplication of Chromosome III. Later on, after more than 1,500 generations, the populations reverted back to an euploid state, while

62 remaining adapted to the heat stress. Aneuploidy was therefore suggested to be a transient adaptive solution, because it can rapidly appear and take over the population under stressful conditions, and can

64 then be rapidly lost when the cost of aneuploidy outweighs its benefit—after the stress is removed, or after refined beneficial mutations appear and fix<sup>74</sup>. Furthermore, it has been suggested that aneuploidy

66 is an evolutionary “stepping stone” that facilitates future adaptation by genetic mutations, which require  
more time to evolve<sup>74,73</sup>.

68 Here, we test the hypothesis that aneuploidy is an evolutionary “stepping stone” that facilitates  
adaptive evolution by genetic mutations<sup>73</sup>. We develop an evolutionary genetic model and fit it to the  
70 experimental results of Yona et al.<sup>74</sup> to predict the genotype frequency dynamics in the experimental  
populations, thereby estimating the frequency of evolved euploid cells that descended from aneuploid  
72 cells. Our results show that although aneuploidy reached high frequencies in the experimental  
populations, the majority of cells in the evolved euploid population likely did not descend from  
74 aneuploid cells, but rather directly from wild-type euploid cells. These suggests that at the lineage  
level, aneuploidy may be an “evolutionary diversion”, rather than a “stepping stone”, on the path to  
76 adaptation.

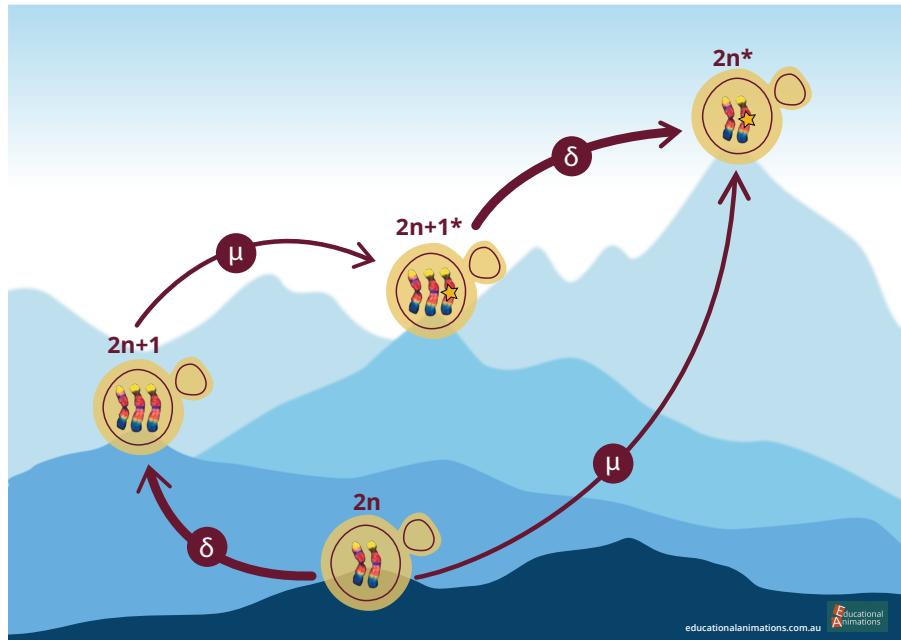
## Results

78 In the heat-stress experiment of Yona et al.<sup>74</sup>, four populations of *S. cerevisiae* evolved under 39 °C.  
Aneuploidy reached high frequency (>95%) in all four experimental repetitions in the first 450  
80 generations. Two of the repetitions, marked *H2* and *H4*, carried no large-scale duplications other  
than a Chromosome III trisomy. These two repetitions continued to evolve under the same conditions,  
82 wherein aneuploidy was eliminated by generation 1,700 and 2,350 in *H4* and *H2*, respectively.

**Evolutionary genetic model.** To explore the dynamics during the evolutionary experiments, we  
84 developed an evolutionary genetic model, fitted the model to empirical data, and used it to predict the  
genotype frequency dynamics, or specifically, the fraction of the evolved euploid population descended  
86 from aneuploid cells.

The model includes the effects of natural selection, genetic drift, aneuploidy, and mutation (i.e., other  
88 genetic variants), and follows a population of cells characterized by their genotype: euploid wild-  
type,  $2n$ , is the ancestral diploid genotype; euploid mutant,  $2n^*$ , has a diploid karyotype and a single  
90 beneficial mutation; aneuploid wild-type,  $2n + 1$ , has an extra chromosome due to a chromosome  
duplication event; and aneuploid mutant,  $2n + 1^*$ , has an extra chromosome (like  $2n + 1$ ) and a  
92 beneficial mutation (like  $2n^*$ ). Note that ‘mutation’ here refers to point mutations and other genetic  
variants unrelated to aneuploidy. Fitness values of the different genotypes are denoted by  $w_{2n}$ ,  $w_{2n^*}$ ,  
94  $w_{2n+1}$ , and  $w_{2n+1^*}$ , and the rate of mutation and aneuploidy are denoted by  $\mu$  and  $\delta$ , respectively. See  
Figure 1 for an illustration of the model.

96 We fitted this model to the experimental results<sup>74</sup>—time for fixation (frequency >95%) and for loss  
 98 (frequency <5%) of aneuploidy—using approximate Bayesian computation with sequential Monte Carlo  
 100 (ABC-SMC)<sup>62</sup>, thereby inferring the model parameters: rates of aneuploidy (i.e., mis-segregation,  
 non-disjunction) and mutation and the fitness of all genotypes. We then sampled posterior predictions  
 for the genotype frequency dynamics using the estimated parameter values and compared different  
 versions of the model to test additional hypotheses about the evolutionary process.



**Figure 1: Model Illustration.** There are four genotypes in our model: euploid wild-type,  $2n$ ; euploid mutant,  $2n^*$ ; aneuploid wild-type,  $2n + 1$ ; and aneuploid mutant,  $2n + 1^*$ . Overall there are two possible trajectories from  $2n$  to  $2n^*$ . Arrows denote transitions between genotypes, with transition rates  $\mu$  for the beneficial mutation rate and  $\delta$  for the aneuploidy rate. Elevation differences illustrate the expected, rather than the assumed, fitness differences between the genotypes.

102 **Estimated rates and fitness effects of aneuploidy and mutation.** We inferred the posterior distribution  
 104 of model parameters (Figure 2). We report parameter estimates using the MAP (maximum a  
 posteriori) and providing the 50% HDI (highest density interval) in square brackets. See Supplementary Material for sensitivity analysis.

106 The estimated beneficial mutation rate is  $\mu = 2.965 \cdot 10^{-6}$  [ $2.718 \cdot 10^{-7} - 3.589 \cdot 10^{-6}$ ] per genome  
 per generation (that is, roughly 3 out of  $10^6$  cell divisions produce a mutant cell with a fitness  
 108 advantage). From the literature, the mutation rate per base pair is roughly  $2 - 3 \cdot 10^{-10}$  (refs.<sup>78,38</sup>), but  
 it may be higher under heat stress, as several stresses<sup>22</sup>, including heat<sup>24</sup>, may cause hypermutation  
 110 in yeast. If we assume a 10-fold increase over the mutation rate reported in the literature, then the  
 estimated beneficial mutation rate can be explained by a genomic target size of 1,000 base pairs

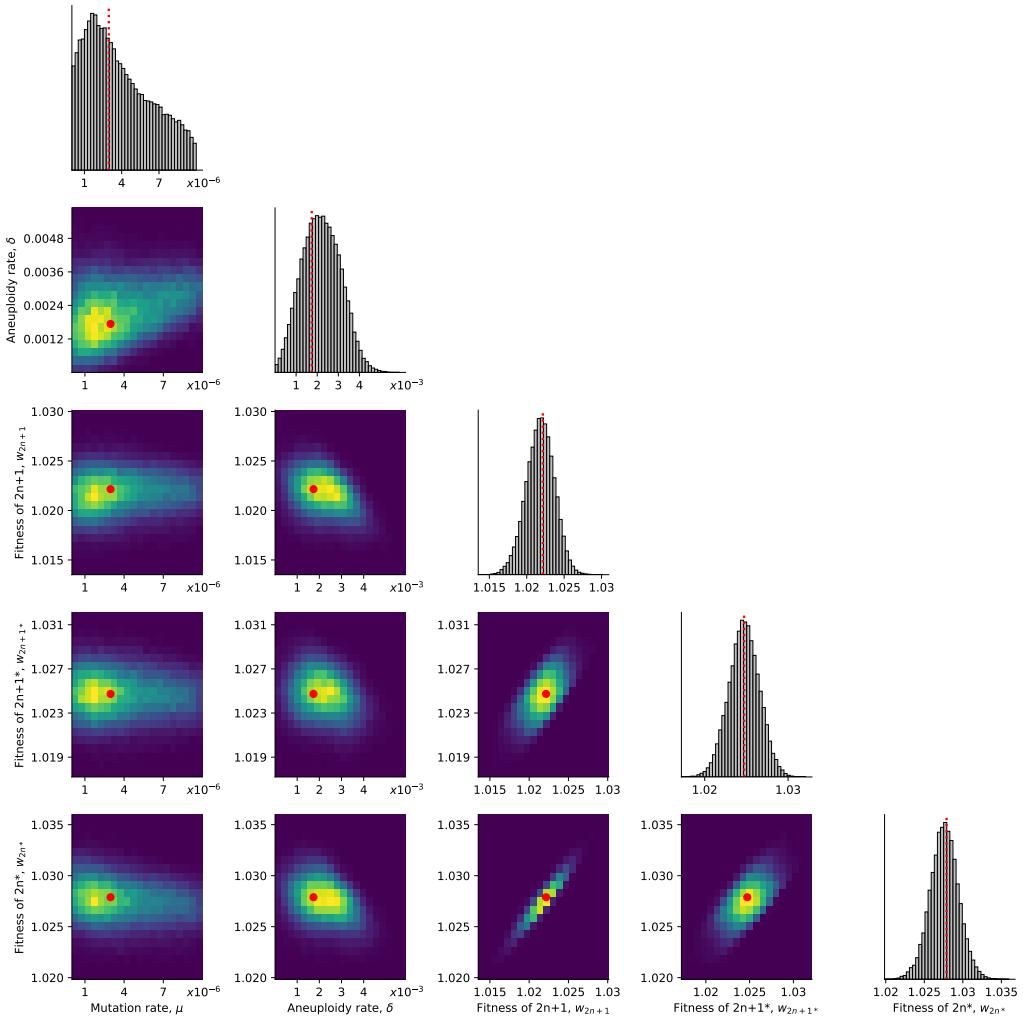
112 (that is, 1,000 base pairs across the genome in which a mutation would provide a fitness advantage):  
113  $3 \cdot 10^{-10} \times 10 \times 1,000 = 3 \cdot 10^{-6}$ . Supporting this, Jarolim et al.<sup>27</sup> found 279 genes that contributed  
114 to survival after a sudden shift from 30 °C to 50 °C, and Flynn et al.<sup>15</sup> used a deep mutational  
115 scan of a single protein, Hsp90, to find 465 amino-acid variants (out of 14,160) that significantly  
116 increased growth rate in 37 °C. Furthermore, Yona et al.<sup>74</sup> found at least 10 genes on Chromosome  
117 III that increased heat tolerance when over-expressed. Assuming that other chromosomes also have  
118 a similar number of heat-tolerance genes (and even more, as Chromosome III is one of the smallest  
119 chromosomes<sup>19</sup>), we estimate a total of 160 heat-tolerance genes in the genome. Indeed, mutations  
120 were found in 97 genes in an evolutionary experiment with yeast under heat stress<sup>24</sup>. Thus, to get a  
121 genomic target size of 1,000, it is enough that the average gene target size (number of base pairs in a  
122 gene in which a mutation is beneficial) is 6.25 base pairs. For example, Kohn and Anderson<sup>32</sup> found  
a target size of 11 in a proton exporter gene (*PMA1*) that contributes to high-salt adaptation.

124 The estimated rate of aneuploidy (i.e., mis-segregation, non-disjunction),  $\delta = 1.72 \cdot 10^{-3}$  [ $1.47 \cdot 10^{-3} -$   
125  $2.786 \cdot 10^{-3}$ ] is higher than in previous studies: for Chromosome III in diploid *S. cerevisiae*, Zhu  
126 et al.<sup>78</sup> estimated  $6.7 \cdot 10^{-6}$  chromosome gain events per generation, and Kumaran et al.<sup>35</sup> estimate  
127  $3.0 \cdot 10^{-5} - 4.3 \cdot 10^{-5}$  chromosome loss events per generation (95% confidence interval). However,  
128 this difference may be partly explained by an increased rate of aneuploidy during heat stress: heat  
shock can increase the rate of chromosome fragment loss by 2-3 orders of magnitude<sup>7</sup>.

130 The estimated fitness values are  $w_{2n+1} = 1.022$  [1.021 – 1.023],  $w_{2n+1^*} = 1.025$  [1.024 – 1.026],  
131  $w_{2n^*} = 1.028$  [1.026 – 1.029], all relative to the fitness of  $2n$ , which is set to  $w_{2n} = 1$ . If we allow for  
132 transitions (mutation, chromosome loss and gain) to less-fit genotypes (e.g.,  $2n^*$  to  $2n + 1^*$ ), then we  
infer similar but slightly different values, see Supplementary Material.

134 **Model comparison and goodness-of-fit.** To assess the fit of our model to the data, we use posterior  
135 predictive checks, in which we simulate the frequency dynamics using MAP parameter estimates and  
136 compare them to the data. Our model fits the data well:  $2n^*$  fixed in 63% of simulations by generation  
1,700 and in 100% of simulations by generation 2,350 (Figure 3).

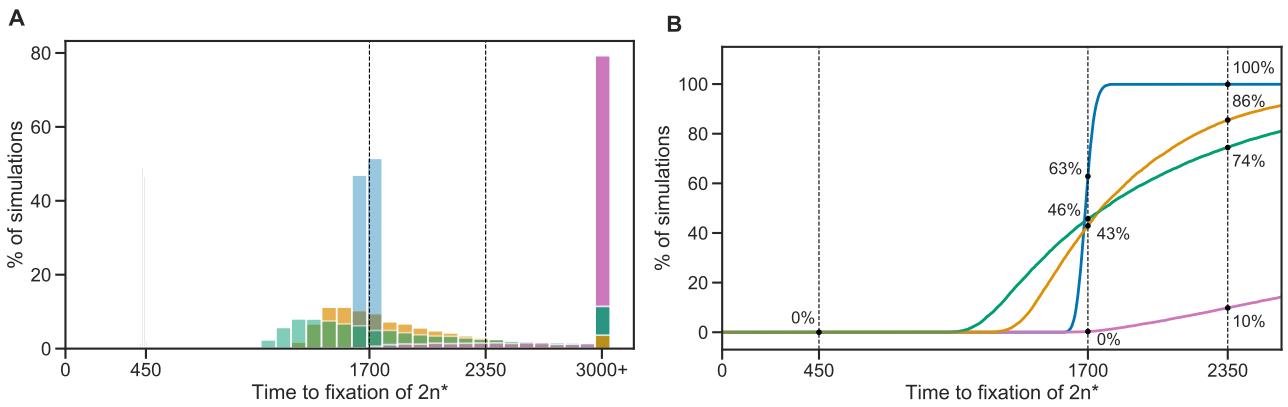
138 However, a model without aneuploidy (where the aneuploidy rate is fixed at zero,  $\delta = 0$ ), fails to  
139 explain the experimental observations (Figure 3). The estimated mutation rate without aneuploidy is  
140  $\mu = 7.98 \cdot 10^{-9}$  [ $7.906 \cdot 10^{-9} - 8.138 \cdot 10^{-9}$ ], much lower compared to a model with aneuploidy. The  
141 fitness of the mutant is also much lower at  $w_{2n^*} = 1.013$  [1.012 – 1.013]. This is because, without  
142 aneuploidy, a high mutation rate or fitness effect will lead to faster appearance and fixation of  $2n^*$  than  
in the experimental observations.



**Figure 2: Posterior distribution of model parameters.** On the diagonal, the marginal posterior distribution of each model parameter. Below the diagonal, the joint posterior distribution of pairs of model parameters (dark purple and bright yellow for low and high density, respectively). Red markers and orange lines for the joint MAP estimate (which may differ from the marginal MAP, as the marginal distribution integrates over all other parameters).

144 We also checked a model in which aneuploidy occurs but is adaptively neutral compared to the wild-type, that is,  $w_{2n+1} = w_{2n}$  and  $w_{2n+1}^* = w_{2n}^*$  but  $\delta > 0$ . This model fits the data better than the model  
 146 with no aneuploidy (in which  $\delta = 0$ ), but worse than a model with positive selection for aneuploidy, in which  $w_{2n} < w_{2n+1} < w_{2n+1}^* < w_{2n}^*$  (Figure 3).

148 **Model predictions of genotype frequency dynamics.** We simulated 50 replicate genotype frequency dynamics using the MAP estimate parameters. Figure 4A shows the simulated frequencies of  
 150 the four genotypes ( $2n$ ,  $2n + 1$ ,  $2n + 1^*$  and  $2n^*$ ), as well as the frequencies of  $2n^*$  cells that arose from either  $2n + 1$  cells via a sequences of mutation and chromosome loss events ( $2n_A^*$ ), or directly from  
 152  $2n$  cells via a mutation event ( $2n_M^*$ ). We find that  $2n + 1^*$  never reaches substantial frequency as it is



**Figure 3: Model fit with and without aneuploidy.** The distribution of time to fixation of  $2n^*$  (i.e., adaptation time) in 10,000 simulations using MAP parameters of the model with beneficial aneuploidy (blue;  $\delta > 0$ ,  $w_{2n} < w_{2n+1} < w_{2n+1^*} < w_{2n^*}$ ) compared to alternative models: a model with the same parameter values but without aneuploidy (gray,  $\delta = 0$ , concentrated at  $t = 450$ ); a model fitted to the data assuming no aneuploidy (green,  $\delta = 0$ ); a model fitted to the data assuming neutral aneuploidy (yellow,  $\delta > 0$ ,  $w_{2n+1} = w_{2n}$ ,  $w_{2n+1^*} = w_{2n^*}$ ); and a model with beneficial aneuploidy and an extended prior distribution (pink). In the experiment by Yona et al.<sup>74</sup>, one population lost aneuploidy by generation 1,700 and another by generation 2,350 (dashed lines) but not before generation 450. Thus, the blue distribution has a better fit compared to the other distributions (the gray distribution has a particularly poor fit). The MAP likelihood (eq. (4)) is 0.84, 0.78, 0.67, and 0.14 for the models represented by blue, yellow, green, and pink distributions, respectively. **(A)** Histogram of the time to fixation of  $2n^*$ . The last bin contains all values equal or greater than 3,000. **(B)** Cumulative distribution of the time to fixation.

quickly replaced by  $2n^*$  in a process similar to *stochastic tunneling*<sup>26,33</sup>.

- 154 To test the hypothesis that aneuploidy facilitates adaptation, we estimated  $F_A$ , the expected frequency  
 155 of  $2n^*$  that arose from  $2n + 1$ , computed as the average frequency of such  $2n_A^*$  cells at the end of  
 156 simulations using the MAP estimate parameters. Surprisingly, we observe that the majority of  $2n^*$   
 157 cells are  $2n_M^*$ , a product of a direct mutation in  $2n$  cells, rather than descending from  $2n + 1$  cells  
 158 ( $F_A^{MAP} = 0.106$ , average end point of 50 purple lines in Figure 4A). This is despite the fact that the  
 159  $2n + 1$  genotype reaches high frequencies in the population (at least 0.98, Figure 4A).
- 160 This result is not unique to the MAP parameter estimate. We simulated genotype frequency dynamics  
 161 using parameter samples from the posterior distribution, and computed the posterior distribution of  
 162  $F_A$  (Figure 4B). The posterior mode  $F_A$  was just 0.147 [0.0154-0.370 95% CI] and only in 489 of  
 163 100,000 posterior samples (0.489%)  $F_A$  was larger than 0.5 (see Supporting Material for results when  
 164 transitions to less-fit genotypes are allowed, such as  $2n^*$  to  $2n + 1^*$ ). Thus, if we sample a random

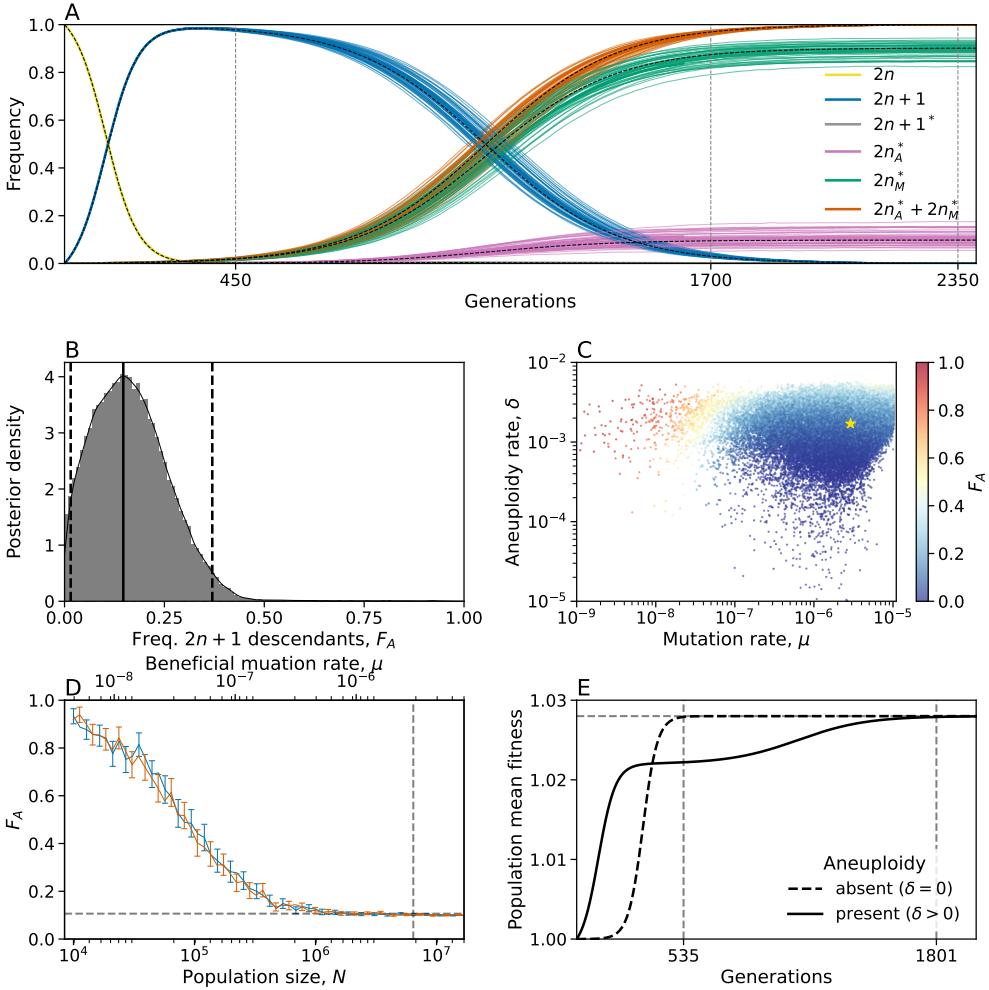
cell from the evolved  $2n^*$  population, it is more likely to have descended directly from an euploid cell than from an aneuploid cell. The probability of  $2n^*$  descending from  $2n + 1$  ( $F_A$ ) increases with the aneuploidy rate,  $\delta$ , and decreases with both the population size  $N$  and the mutation rate,  $\mu$  (Figure 4C,D). In some cases it can also be affected by the fitness parameters (Figure S10).

**Genetic instability in aneuploid cells.** It has been suggested that aneuploidy increases genomic instability: Sheltzer et al.<sup>59</sup> have demonstrated a fold increase of between 2.2 and 7.1 in the mutation rate of disomic yeast (rather than trisomic yeast, the focus of our analysis). Therefore, we inferred model parameters under the assumption that the mutation rate increases in aneuploid cells by a factor  $\tau = 1, 33/32$  (due to an additional chromosome), 2, 5, 10, or 100 (due to genetic instability). We found that the posterior distribution was similar for  $\tau = 1, 33/32, 2$ , and 5 (Figure S4). Furthermore, we computed the WAIC, a criterion for model selection (Methods). The WAIC values were similar for all  $\tau$  values (Table S1).

Assuming a strong increase of the mutation rate in aneuploid cells, i.e.  $\tau = 100$ , the inferred a mutation rate was  $\mu = 4.094 \cdot 10^{-7}$  [ $6.252 \cdot 10^{-8} - 6.046 \cdot 10^{-7}$ ]), and the inferred aneuploidy rate that was  $\delta = 0.744 \cdot 10^{-3}$  [ $0.506 \cdot 10^{-3} - 1.827 \cdot 10^{-3}$ ]. Compared to inference made assuming no effect of aneuploidy on the mutation rate, these rates were about 7-8-fold and 2-3-fold lower for  $\mu$  and  $\delta$ , respectively. Assuming  $\tau = 10$ , the inferred a mutation rate was only slightly lower compared to  $\tau = 1$  ( $\mu = 1.67 \cdot 10^{-6}$  [ $2.836 \cdot 10^{-8} - 2.245 \cdot 10^{-6}$ ]).

Therefore, we do not find evidence of an increase in mutation rate in aneuploid cells. This may be because, unless the increase is strong ( $\tau \geq 10$ ), it does not seem to affect our inference; or because Chromosome III is one of the smallest chromosomes<sup>19</sup>. We also checked the differences in genotype frequency dynamics for different  $\tau$  values. We observe  $\tau = 100$  could be distinguished if accurate data was available for the waiting time until the frequency of  $2n$  to decrease below 95% (Figure S5A) or for waiting time for the frequency of  $2n + 1$  to either reach or go below 95% (Figure S5B).

Similarly, we did not find evidence for an increase in the rate of chromosome loss in aneuploid cells<sup>59</sup>, probably due to lack of statistical power. Nevertheless, increasing the rate of chromosome loss (transitions from  $2n + 1^*$  to  $2n^*$ ) without increasing the rate of chromosome gain (transitions from  $2n$  to  $2n + 1$ ) increases  $F_A$  (Figure S11B), but not to the same extent as increasing the rate of chromosome gain (Figure S10). In contrast, increasing the mutation rate in aneuploid cells can have a marked effect on the dynamics: when using the MAP parameter estimates,  $F_A$  increases from 0.1 to 0.52 when the mutation rate in aneuploid cells increases 10-fold (Figure S11C).



**Figure 4: Predicted frequency of aneuploid-descended cells.** (A) Posterior predicted genotype frequencies over time, including the source of  $2n^*$ :  $2n_A^*$  arose from  $2n + 1$ , whereas  $2n_M^*$  arose directly from  $2n$ . Colored curves are 50 simulations using the MAP estimate parameters. Black dashed curves are the expected genotype frequencies without genetic drift (from a deterministic model). See Figure S9 for log-log scale, in which the sequence of events is easier to observe. (B) Posterior distribution of  $F_A$ , the expected frequency of  $2n^*$  cells descended from  $2n + 1$  cells, computed as the average frequency at the end of 100 simulations for 100,000 samples from the parameter posterior distribution. Solid and dashed lines show the mode and 95% CI. (C)  $F_A$  values (color coded) from panel B, with their corresponding mutation rate  $\mu$  on x-axis and aneuploidy rate  $\delta$  on the y-axis. Yellow star shows the MAP estimate. See also Figure S10. (D)  $F_A$  as a function of the population size ( $N$ , bottom x-axis) and the beneficial mutation rate ( $\mu$ , top x-axis) in posterior predictions with MAP parameters. Markers show  $F_A$  in 250 simulations per population size or mutation rate value. Error bars show mean  $F_A$  with 95% CI (bootstrap,  $n = 10,000$ ). Blue and red bars for varying population size and mutation rate, respectively. Vertical dashed line for population size in the experiment,  $6.425 \cdot 10^6$ , and the MAP mutation rate,  $2.965 \cdot 10^{-6}$ . Horizontal line for  $F_A^{MAP} = 0.106$ . (E) Population mean fitness in a model without drift using MAP estimate parameters. Solid lines for mean fitness with aneuploidy ( $\delta > 0$ ), where the population reaches adaptation (mean fitness at 99.99% of maximum value) at generation 1,802. Dashed lines for mean fitness without aneuploidy ( $\delta = 0$ ), where the population adapts much earlier, at generation 535.

196 **Discussion**

In a study on the role of chromosome duplication in adaptive evolution, Yona et al.<sup>74</sup> found that a  
198 Chromosome III trisomy was acquired by *S. cerevisiae* populations evolving under heat stress, only  
to be later replaced by euploid mutant cells that carry "refined" solutions to the stress. Additionally,  
200 such a replacement also occurred when they initiated evolutionary experiments with a population in  
which all cells carry a Chromosome III trisomy. They hypothesized that aneuploidy is a "useful yet  
202 short-lived intermediate that facilitates further adaptation", suggesting that the euploid mutant cells  
evolved by heat-resistance mutations in aneuploid cells followed by reversion of trisomy due to a  
204 chromosome loss event.

We developed an evolutionary genetic model of adaptive evolution by aneuploidy and mutation  
206 (Figure 1), fitted it to the experimental results of Yona et al.<sup>74</sup>, and used it to predict the genotype  
frequency dynamics. The model predicted that only about 10-15% of the evolved euploid population  
208 descended from aneuploid cells by acquiring a mutation and losing the extra chromosome—that is,  
the majority of the euploid population are not descended from aneuploid cells, but rather are direct  
210 descendants of the ancestral wild-type population (Figure 4).

This happens despite aneuploidy reaching a high frequency in the population (>95%). Conventional  
212 wisdom might suggest that once the aneuploid genotype  $2n + 1$  reaches high frequency, it will have a  
better chance at producing "refined" solutions via mutations, and its descendants will come to dominate  
214 the population: the frequency of  $2n_A^*$  (which arises from  $2n + 1^*$ ) will be higher than the frequency of  
 $2n_M^*$  (which arises directly from  $2n$ ).

216 So how does  $2n_M^*$  prevail? Initially, the supply rates of  $2n + 1$  and  $2n_M^*$  are  $N\delta \approx 11,000$  and  $N\mu \approx 19$ ,  
respectively (assuming MAP parameter estimates). Therefore, both genotypes are expected to appear  
218 immediately at the beginning of the experiment (Figure S9). However,  $2n + 1$  appears at a much  
higher frequency as  $\delta \gg \mu$  by 2-3 orders of magnitude. After they first appear,  $2n_M^*$  has higher  
220 fitness. But as long as the frequency of  $2n$  is high, the supply rate of  $2n + 1$  is higher than that of  
 $2n_M^*$ , again due to  $\delta \gg \mu$ . However, supply rates of both genotypes decrease with the frequency  
222 of  $2n$ . Therefore, when the latter decreases, mainly due to the increase in the frequency of  $2n + 1$ ,  
both supply rates diminish. At this stage, the higher fitness of  $2n_M^*$  comes into play and it starts to  
224 take over the population, which is mainly composed of  $2n + 1$ . For the aneuploid lineage to compete  
with the mutant lineage, it must produce  $2n_A^*$  via a mutation followed by chromosome loss. Although  
226 this is a stochastic process (due to drift), our results show that the time until  $2n_A^*$  reaches a frequency  
of 0.1% is roughly 450 generations, without much variation (intersection of purple lines and vertical

228 dashed line in Figure S9). However, by that time  $2n_M^*$  is already at a roughly 10-fold higher frequency  
229 (1.86%), and since both mutants have the same fitness, their relative frequency remains roughly the  
230 same until the end of the experiment.

232 **Predictions for small populations and low mutation rates.** We examined the effect of the popula-  
233 tion size,  $N$ , and the beneficial mutation rate,  $\mu$ , on the frequency of  $2n + 1$  descendants in the evolved  
234 population,  $F_A$ . We found that  $F_A$  is expected to decrease as the population size or mutation rate  
235 increase (Figure 4D), ranging from >90% when the population size is 10,000 or the mutation rate is  
236  $6 \cdot 10^{-9}$ , to about 10% when the population size is above 1,000,000 (less than the experimental popu-  
237 lation size, which was 6,425,000) or the mutation rate is above  $2 \cdot 10^{-6}$  (less than the inferred mutation  
238 rate, which is  $2.965 \cdot 10^{-6}$ ). Thus, our model provides a testable prediction: if the experiment was  
239 repeated under a lower population size (via stronger daily dilutions or in a smaller volume) or a lower  
240 mutation rate (via a non-mutagenic stress or stress with a smaller target size such as drug resistance),  
241 then the fraction of the population descending from aneuploid cells would be much higher.

242 **Aneuploidy delays rather than facilitates adaptation.** An additional interesting result of our  
243 study is that aneuploidy increases, rather than decreases, the adaptation time (Figure 4E). This  
244 happens despite the fact that the mean fitness initially increases faster in the presence of aneuploidy  
245 (Figure 4E). Aneuploidy increases adaptation time because once  $2n + 1$  is common, selection for the  
246 mutant strain ( $2n + 1^*$  or  $2n^*$ ) is weaker compared to when  $2n^*$  competes directly with  $2n$ . This is an  
interesting example of clonal interference<sup>20</sup> but between fast and slow mutational processes<sup>34</sup>.

248 **Rate and fitness effect of aneuploidy and mutation.** We inferred the rates of aneuploidy and  
mutation and their effects on fitness. We estimate that the aneuploidy rate (i.e., number of chromosome  
gains per generation) is  $1.7 \cdot 10^{-3}$ , higher than a previous estimate of  $6.7 \cdot 10^{-6}$  (ref. <sup>77</sup>). This may  
250 be due to genetic instability caused by heat stress<sup>7</sup>, but we note that there is a general scarcity of  
empirical data on aneuploidy rates. In addition, we did not find evidence for increased mutation rates  
252 in aneuploid cells. Previous empirical studies have suggested that genetic instability (e.g., elevated  
mutation rates) in aneuploid cells is due to stress associated with the aneuploid state<sup>5,8,75,25</sup>. However,  
254 in the experiment of Yona et al.<sup>74</sup>, both the wild-type and the aneuploid were under heat stress, which  
may explain why we didn't find evidence for an increased mutation rate specifically in aneuploid  
256 cells.

258 **Effect of ploidy.** The evolutionary dynamics may change in haploid yeast, in which aneuploidy  
results in a second, rather than third, chromosome copy. For example, it has been demonstrated that

drug resistance mainly evolves via recessive mutations and aneuploidy in haploid yeast<sup>63</sup>, whereas in  
260 diploid yeast it evolves via dominant mutations, aneuploidy, and gene/segmental duplications<sup>4</sup>. Thus,  
the second chromosome copy of disomic yeast may facilitate further adaptation via duplications,  
262 rearrangements, and increased mutational tolerance<sup>2</sup>, while decreasing the chance for adaptation via  
recessive mutations. Future models and experiments can consider how ploidy and other genomic  
264 contexts affect the role of aneuploidy in adaptive evolution.

**Conclusions.** Here, we tested the hypothesis that aneuploid cells are an evolutionary “stepping  
266 stone”, or adaptive intermediate, between wild-type euploid cells and mutant euploid cells in the  
evolutionary experiment of Yona et al.<sup>73</sup>. Our results suggest that, although it seems the population  
268 goes from euploid to aneuploid and back, this is not the case at the individual level. We estimate  
that only about 10-15% of the euploid cells descended from aneuploid cells, whereas the rest are  
270 direct descendants of the wild-type euploid cells. Thus, aneuploidy can delay, rather than accelerate,  
adaptation, and cells that become aneuploid may leave less descendants than cells that remain euploid.  
272 This surprising result reinforces the importance of mathematical models when interpreting evolutionary  
dynamics. Moreover, our study emphasizes the unintuitive outcomes of clonal interference  
274 between mechanisms for generation of variation that differ in their rate of formation and distribution  
of fitness effects, including mutation, copy number variation, horizontal gene transfer, and epigenetic  
276 modifications.

## Models and Methods

278 **Evolutionary genetic model.** We model the evolution of a population of cells using a Wright-Fisher  
model<sup>43</sup>, assuming a constant effective population size  $N$ , non-overlapping generations, and including  
280 the effects of natural selection, genetic drift, aneuploidy, and mutation. We focus on beneficial ge-  
netic modifications, neglecting the effects of deleterious and neutral mutations or karyotypic changes.  
282 The model allows for a single aneuploid karyotype (e.g., Chromosome III duplication) and a single  
mutation to accumulate in the genotype. Thus, the model follows four genotypes (Figure 1): euploid  
284 wild-type,  $2n$ , the initial genotype; euploid mutant,  $2n^*$ , with the standard karyotype and a single  
beneficial mutation; aneuploid wild-type,  $2n + 1$ , with an extra chromosome, i.e., following chro-  
286 mosome duplication; and aneuploid mutant,  $2n + 1^*$ , with an extra chromosome and a beneficial  
mutation.  
288 Transitions between the genotypes occur as follows (Figure 1): Beneficial mutations from  $2n$  to  $2n^*$   
and from  $2n+1$  to  $2n+1^*$  occur with probability  $\mu$ , the mutation rate. We neglect back-mutations (i.e.,

290 from  $2n^*$  to  $2n$  and from  $2n+1^*$  to  $2n+1$ ). Aneuploidy is formed by chromosome mis-segregation,  
 so that cells transition from  $2n$  to  $2n+1$  and from  $2n+1^*$  to  $2n^*$  with probability  $\delta$ , the aneuploidy  
 292 rate. That is, we assume chromosomes are gained and lost at the same rate, and we neglect events that  
 form a less-fit genotype (i.e.,  $2n+1$  to  $2n$  and  $2n^*$  to  $2n+1^*$ ). A model that assumes an increased  
 294 rate of chromosome loss in aneuploid cells (as in Sheltzer et al.<sup>59</sup>) did not perform well, probably due  
 to lack of statistical power, and was abandoned.

296 In the experiment by Yona et al.<sup>74</sup>, the population was grown every day from  $1.6 \cdot 10^6$  cells until  
 reaching stationary phase and then diluted 1:120. Thus, we set the population size to  $N = 6.425 \cdot 10^6$ ,  
 298 the harmonic mean of  $\{2^k \cdot 1.6 \cdot 10^6\}_{k=0}^7$ <sup>11</sup>. The initial population has  $N$  cells with genotype  $2n$ .  
 The effect of natural selection on the frequency  $f_i$  of genotype  $i = 2n, 2n+1, 2n+1^*$ , or  $2n^*$  is given  
 300 by

$$f_i^s = \frac{f_i w_i}{\bar{w}}, \quad (1)$$

302 where  $w_i$  is the fitness of genotype  $i$  and  $\bar{w} = \sum_j f_j w_j$  is the population mean fitness. The effect of  
 mutation and aneuploidy on genotype frequencies is given by

$$\begin{aligned} f_{2n}^m &= (1 - \delta - \mu) f_{2n}^s, \\ f_{2n+1}^m &= \delta f_{2n}^s + (1 - \mu) f_{2n+1}^s, \\ f_{2n+1^*}^m &= \mu f_{2n+1}^s + (1 - \delta) f_{2n+1^*}^s, \\ f_{2n^*}^m &= \mu f_{2n}^s + \delta f_{2n+1^*}^s + f_{2n^*}^s. \end{aligned} \quad (2)$$

Finally, random genetic drift is modeled using a multinomial distribution<sup>43</sup>,

$$306 \quad \mathbf{f}' \sim \frac{1}{N} \cdot \text{Mult}(N, \mathbf{f}^m), \quad (3)$$

where  $\mathbf{f}^m = (f_{2n}^m, f_{2n+1}^m, f_{2n+1^*}^m, f_{2n^*}^m)$  are the frequencies of the genotypes after mutation and  
 308 aneuploidy,  $\mathbf{f}'$  are the genotype frequencies in the next generation, and  $\text{Mult}(N, \mathbf{f})$  is a multinomial  
 distribution parameterized by the population size  $N$  and the genotype frequencies  $\mathbf{f}$ . Overall, the change  
 310 in genotype frequencies from one generation to the next is given by the transformation  $f_i \rightarrow f'_i$ .

**Empirical data for model inference.** We use the results of evolutionary experiments reported by  
 312 Yona et al.<sup>74</sup>. In their heat-stress experiment, four populations of *S. cerevisiae* evolved under 39 °C.  
 Aneuploidy fixed (frequency >95%) in all four population in the first 450 generations. Hereafter,  
 314 fixation or elimination of a genotype by generation  $t$  means that more than 95% or less than 5% of the  
 population carry the genotype at generation  $t$ , and possibly earlier. In the original analysis of Yona  
 316 et al.<sup>74</sup>, samples were routinely extracted from the evolving populations and tested for indication of

heat-shock tolerance. The first generation in which such indication was found was generation 200.

318 Therefore, we determine that aneuploidy did not reach high frequency before generation 200. The experiment continued with two populations, in which aneuploidy was eliminated by generation 1,700  
320 and 2,350 while still under the same conditions of elevated heat (39 °C).

322 **Likelihood function.** Because our model, just like the Wright-Fisher model, is non-linear and stochastic, computing the distribution of fixation time  $T(g)$  of genotype  $g$  for use in the likelihood function is intractable (it is even hard to use a diffusion-equation approximation due to the model having  
324 multiple genotypes, rather than just two). We overcome this problem by approximating the likelihood using simulations. We simulate 1,000 experiments per parameter vector  $\theta = (\mu, \delta, s, b, c)$ , resulting  
326 in a set of simulated observations  $\tilde{\mathbf{X}} = \{\tilde{X}_i\}_{i=1}^{1000}$ . We then compute the approximate likelihood,

$$\begin{aligned} \mathcal{L}(\theta) = P^4(200 \leq T(2n+1) \leq 450) \cdot & \left[ 1 - \right. \\ & P_{\tilde{\mathbf{X}}}^4(\{T(2n^*) < 1700\} \mid 200 \leq T(2n+1) \leq 450) - \\ & P_{\tilde{\mathbf{X}}}^4(\{1700 < T(2n^*) < 2350\} \mid 200 \leq T(2n+1) \leq 450) + \\ & \left. P_{\tilde{\mathbf{X}}}^4(\{T(2n^*) < 1700\} \wedge \{1700 < T(2n^*) < 2350\} \mid 200 \leq T(2n+1) \leq 450) \right], \end{aligned} \quad (4)$$

328 where  $\{ \dots \}$  is the "logical not" operator,  $P^4(\dots)$  is the 4th power of  $P(\dots)$ , and all probabilities  
330  $P_{\tilde{\mathbf{X}}}(\dots)$  are approximated from the results of the simulations  $\tilde{\mathbf{X}}$ . For example,  $P_{\tilde{\mathbf{X}}}(\{T(2n^*) < 1700\} \mid$   
332  $200 \leq T(2n+1) \leq 450)$  is approximated by taking simulations in which  $2n+1$  fixed (reached >95%)  
334 before generation 450 but not before generation 200, and computing the fraction of such simulations  
in which  $2n^*$  did not fix by generation 1,700, and hence aneuploidy did not extinct (reach <5%)  
before generation 1,700. Figure S1 compares results with less and more simulated experiments,  
demonstrating that 1,000 simulations are likely sufficient.

For a model without aneuploidy (that is, when the aneuploidy rate is fixed at zero,  $\delta = 0$ ), we disregard  
336 the increased expression in Chromosome III and the growth advantage measured in generation 450,  
and focus on the growth advantage measured in later generations, presumably due to a beneficial  
338 mutation. Therefore, the likelihood is approximated by

$$\begin{aligned} \mathcal{L}_!(\theta) = 1 - P_{\tilde{\mathbf{X}}}^4(\{T(2n^*) < 1700\}) - \\ P_{\tilde{\mathbf{X}}}^4(\{1700 < T(2n^*) < 2350\}) + \\ P_{\tilde{\mathbf{X}}}^4(\{T(2n^*) < 1700\} \wedge \{1700 < T(2n^*) < 2350\}). \end{aligned} \quad (5)$$

340 **Parameter inference.** To infer model parameters, we use approximate Bayesian computation with a sequential Monte-Carlo scheme, or ABC-SMC<sup>62</sup>, implemented in the pyABC Python package<sup>31</sup>

342 `pyabc.readthedocs.io`. This approach uses numerical stochastic simulations of the model to infer a  
343 posterior distribution over the model parameters. It is a method of likelihood-free, simulation-based  
344 inference<sup>10</sup>, that is, for estimating a posterior distribution when a likelihood function cannot be  
345 directly computed. It is therefore suitable in our case, in which the likelihood function can only be  
346 approximated from simulations, and cannot be directly computed.

The ABC-SMC algorithm employs sequential importance sampling over multiple iterations<sup>67,30,65</sup>. In  
348 iteration  $t$  of the algorithm, a set of parameter vectors,  $\{\theta_{i,t}\}_{i=1}^{n_t}$ , also called *particles*, are constructed  
349 in the following way. A proposal particle,  $\theta^*$ , is sampled from a proposal distribution, and is either  
350 accepted or rejected, until  $n_t$  particles are accepted. The number of particles,  $n_t$ , is adapted at every  
351 iteration  $t$  using the adaptive population strategy<sup>31</sup> `pyabc.readthedocs.io`. For  $t = 0$ , the proposal  
352 particle is sampled from the prior distribution,  $p(\theta)$ . For  $t > 0$ , the proposal particle is sampled from  
353 the particles accepted in the previous iteration,  $\{\theta_{i,t-1}\}_{i=1}^{n_{t-1}}$ , each with a probability relative to its weight  
354  $W_{t-1}(\theta_{i,t-1})$  (see below). The proposal particle is then perturbed using a kernel perturbation kernel,  
355  $K_t(\theta^* | \theta)$  where  $\theta$  is the sample from the previous iteration. Then, a set of synthetic observations  
356  $\tilde{\mathbf{X}}^*$  is simulated, and the proposal particle  $\theta^*$  is accepted if its approximate likelihood (eq. (4)) is high  
357 enough,  $\mathcal{L}(\theta^*) > 1 - \epsilon_t$  (or more commonly, if  $1 - \mathcal{L}(\theta^*) < \epsilon_t$ ), where  $\epsilon_t > 0$  is the *acceptance*  
358 *threshold*, as higher values of  $\epsilon_t$  allow more particles to be accepted. The acceptance threshold  $\epsilon_t$  is  
359 chosen as the median of the  $1 - \mathcal{L}(\theta)$  of the particles accepted in the previous iteration,  $t - 1$ , and  
360  $\epsilon_0 = 0.01$ . For each accepted particle  $\theta_{i,t}$  a weight  $W_t(\theta_{i,t})$  is assigned: for  $t = 0$ ,  $W_0(\theta_{i,0}) = 1$ ,  
361 and for  $t > 0$ ,  $W_t(\theta_{i,t}) = p(\theta_{i,t}) / \sum_{i=1}^{n_{t-1}} W_{t-1}(\theta_{i,t-1}) K_t(\theta_{i,t}, \theta_{i,t-1})$ , where  $p(\theta)$  is the prior density of  
362  $\theta$  and  $K_t(\theta', \theta)$  is the probability of a perturbation from  $\theta$  to  $\theta'$ .  $K_t(\theta' | \theta)$  is a multivariate normal  
363 distribution, fitted at iteration  $t$  to the particles from the previous iteration,  $\{\theta_{i,t-1}\}_{i=1}^{n_{t-1}}$ , and their  
364 weights,  $\{W(\theta_{i,t-1})\}_{i=1}^{n_{t-1}}$ .

Acceptance is determined according to the approximate likelihood (eq. (4)), which has a maximum  
366 value of  $\mathcal{L}_{max} = 0.875$  (giving a minimal value of  $\epsilon_{min} = 0.125$ ). We terminated the inference  
367 iterations when the change in  $\epsilon$  value from one iteration to the next was small. With our standard prior  
368 and model, we reached  $\epsilon = 0.13$  (or  $\mathcal{L} = 0.87$ ) after six iterations, with  $n_6 = 982$  accepted parameter  
369 vectors and effective sample size ESS=651 (Figure S2). Running the inference algorithm with different  
370 initialization seeds and less or more simulations for approximating the likelihood produced similar  
371 posterior distributions (Figure S1).

372 After producing a set of weighted particles from the the posterior distribution using the above ABC-  
373 SMC algorithm, we approximate the posterior using kernel density estimation (KDE) with Gaussian  
374 kernels. We truncate the estimated posterior to avoid positive posterior density for values with zero

prior density. The MAP (maximum a posteriori) estimate is computed as the the maximum of the  
376 estimated joint posterior density. We then draw 5,000,000 samples from the posterior distribution  
to compute the HDI (highest density interval) and draw 50,000 samples to visualize the posterior  
378 distribution with histograms.

**Model comparison.** We examine several versions of our evolutionary models, e.g. without aneuploidy or with increased mutation rate in aneuploid cells, as well as several different prior distributions  
380 (see below). To compare these, we plot posterior predictions: for each model we execute 10,000  
382 simulations using the MAP parameter estimates and plot the distributions of time to fixation of  $2n^*$ ,  
one of key properties of the model likelihood. These plots visualize the fit of each model to the  
384 data. Also, for similar models we plot the marginal and joint posterior distributions of the parameters;  
if these are similar, we consider the models interchangeable. We validate this by comparing HDI  
386 (highest density interval) of posterior distributions.

Where posterior plots are very similar and the number of parameters is the same, we use WAIC, or  
388 the widely applicable information criterion<sup>16</sup>, defined as

$$WAIC(\theta) = -2 \log \mathbb{E}[\mathcal{L}(\theta)] + 2\mathbb{V}[\log \mathcal{L}(\theta)] \quad (6)$$

390 where  $\theta$  is a parameter vector, and  $\mathbb{E}[\cdot]$  and  $\mathbb{V}[\cdot]$  are the expectation and variance taken over the  
posterior distribution, which in practice are approximated using 50,000 samples from the posterior  
392 KDE. We validated that upon resampling WAIC values do not significantly change and that differences  
in WAIC between models are preserved. WAIC values are scaled as a deviance measure: lower values  
394 imply higher predictive accuracy.

**Prior distributions.** We used informative prior distributions for  $w_{2n+1}$ ,  $w_{2n+1^*}$  and  $w_{2n^*}$  (we set  
396  $w_{2n} = 1$ ), which we estimated from growth curves data from mono-culture growth experiments  
previously reported by Yona et al.<sup>74</sup>, Figs. 3C, 4A, and S2. We used Curveball, a method  
398 for predicting results of competition experiments from growth curve data<sup>46</sup> [curveball.yoavram.com](http://curveball.yoavram.com).  
Briefly, Curveball takes growth curves of two strains growing separately in mono-culture and predicts  
400 how they would grow in a mixed culture, that is, it predicts the results of a competition assay. From these  
predictions, relative fitness values can be computed. Because Curveball uses a maximum-likelihood  
402 approach to estimate model parameters, we were able to estimate a distribution of relative fitness  
values to be used as a prior distribution by sampling 10,000 samples from a truncated multivariate  
404 normal distribution defined by the maximum-likelihood covariance matrix (Figure S3).

We used growth curves of  $2n$  and  $2n + 1$  in 39 °C to estimate an informative prior distribution for

406  $w_{2n+1}$  (Figure S3D, assuming  $w_{2n} = 1$ ). In this prior distribution, we used the same prior for  $w_{2n+1^*}$   
408 and  $w_{2n^*}$ . To increase computational efficiency, we also assumed  $w_{2n^*} > w_{2n+1^*} > w_{2n+1} > w_{2n}$ ;  
410 running the inference without this assumption produced similar results. See *supporting material* for  
an extended informative prior distribution that uses growth curves of  $2n^*$  and  $2n+1$  growing in 39 °C;  
this prior distribution proved to be less useful.

As a control, we tested an uninformative uniform prior with  $U(1, 6)$ , for (i) all  $w_{2n+1}$ ,  $w_{2n+1^*}$ ,  $w_{2n^*}$ ,  
412 or (ii) only for  $w_{2n+1^*}$ ,  $w_{2n^*}$ , using the above informative prior for  $w_{2n+1}$ . In these cases the inference  
algorithm failed to converge.

414 For the mutation rate,  $\mu$ , and aneuploidy rate,  $\delta$ , we used uninformative uniform priors,  $\mu \sim U(10^{-9}, 10^{-5})$  and  $\delta \sim U(10^{-6}, 10^{-2})$ . A wider mutation rate prior,  $\mu \sim U(10^{-9}, 10^{-3})$ , produced  
416 similar results.

## Acknowledgements

418 We thank Lilach Hadany, Judith Berman, David Gresham, Shay Covo, Martin Kupiec, Uri Obolski, Daniel  
Weissman, and Tal Simon for discussions and comments. This work was supported in part by the Israel Science  
420 Foundation (ISF, YR 552/19), the US–Israel Binational Science Foundation (BSF, YR 2021276), Minerva  
Center for Live Emulation of Evolution in the Lab (YR, YP), Minerva Stiftung short-term research grant (MP),  
422 the Burroughs Wellcome Fund (AS), and the Jane Coffin Childs Memorial Fund (PVZ).

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## 626 Supplementary Material

### Supplementary Analysis

628 **Sensitivity analysis.** Changing a single parameter while keeping the rest fixed at the MAP estimate produces a worse fit to the data (Figure S6). Furthermore, we fitted models with a mutation rate  
630 fixed at  $\mu = 10^{-5}$ ,  $10^{-6}$  and  $10^{-7}$ . We inferred similar parameters estimates for the model with  
632  $\mu = 10^{-6}$  compared to the model with a free  $\mu$  parameter, in which the inferred mutation rate is  
634  $\mu \approx 3 \cdot 10^{-6}$ . Inference assuming  $\mu = 10^{-5}$  or  $\mu = 10^{-7}$  produced similar estimates except that the estimated aneuploidy rate,  $\delta$ , was higher, and assuming  $\mu = 10^{-7}$ , the estimated fitness of  $2n + 1$  was lower (Figure S7).

636 **Extended informative prior distribution.** In an extended informative prior distribution, we used additional growth curves of  $2n^*$  (*refined* strain from Yona et al.<sup>74</sup>) and  $2n + 1$  in 39 °C to estimate  
638  $w_{2n^*}/w_{2n+1}$  (Figure S3H). The same distribution was used for  $w_{2n^*}/w_{2n+1^*}$ . Thus, our main informative prior uses a single prior distribution for fitness values of  $2n + 1$ ,  $2n + 1^*$ , and  $2n^*$ , whereas the extended informative prior uses one distribution for  $2n + 1$ , and another distribution for both  $2n + 1^*$   
640 and  $2n^*$ .

We estimated the parameters under this extended informative prior. Inference took much longer  
642 to run but the posterior distribution seemed to converge, as it did not change much in the final iterations. The posterior predictive plot shows that inference with this extended prior produces a  
644 posterior distribution that fails to explain the empirical observations (pink in Figure 3). However, the inferred posterior distribution is considerably narrower (compare Figures 2 and S8) and therefore  
646 parameter estimates are less variable. The estimated mutation rate was much lower compared to the main informative prior, with  $\mu = 2.474 \cdot 10^{-9}$  [ $2.423 \cdot 10^{-9} - 2.612 \cdot 10^{-9}$ ]. Other parameter  
648 estimates are:  $\delta = 2.705 \cdot 10^{-3}$  [ $2.094 \cdot 10^{-3} - 3.094 \cdot 10^{-3}$ ],  $w_{2n+1} = 1.022$  [ $1.021 - 1.024$ ],  
650  $w_{2n+1^*} = 1.052$  [ $1.05 - 1.054$ ],  $w_{2n^*} = 1.053$  [ $1.051 - 1.055$ ], the latter two being much higher  
652 compare to the main informative prior. Notably, the mode of the posterior ratio  $w_{2n^*}/w_{2n+1} = 1.0009$  is much lower than the mode of the prior ratio of 1.033 (Figure S3H) and closer to the ratio of 1 that we assume in the main informative prior. Together with the posterior predictive results, we conclude  
654 that the main informative prior is preferable over the extended informative prior.

654 **Model with transitions to less-fit genotypes** We also estimated the parameters of a version of the model that includes transitions (mutation, chromosome loss and gain) to less-fit genotypes (e.g.,  $2n^*$

656 to  $2n + 1^*$ ),

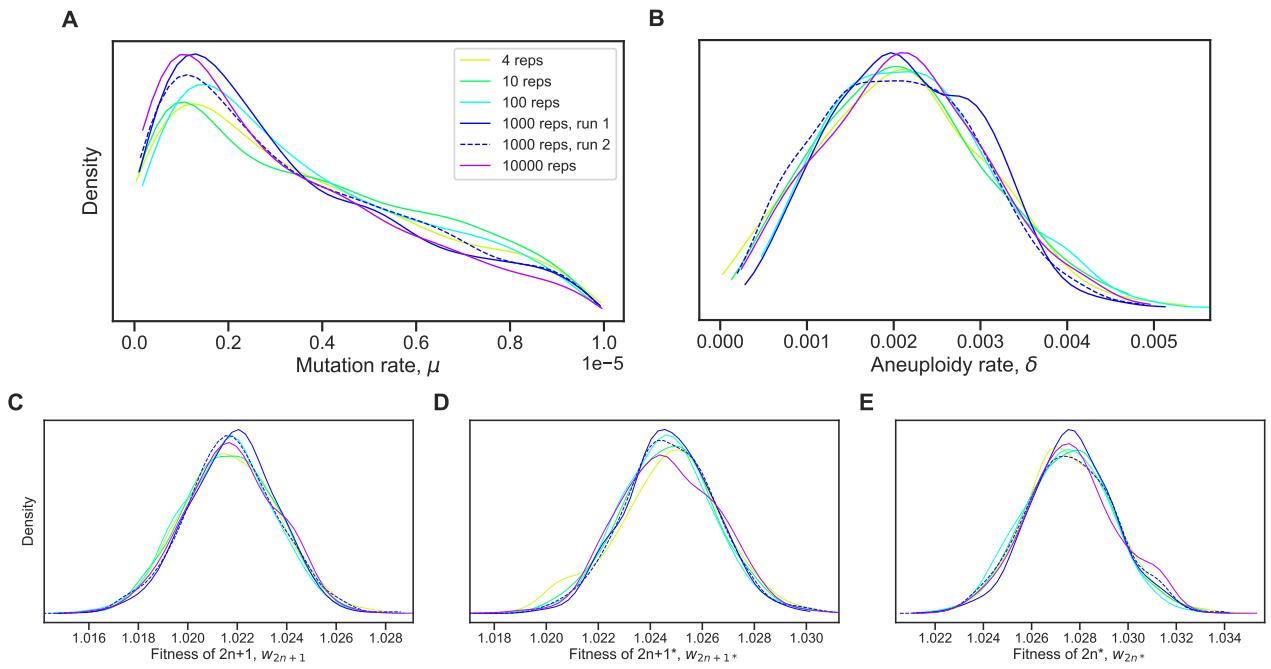
$$\begin{aligned}
 f_{2n}^m &= (1 - \delta - \mu)f_{2n}^s + \delta f_{2n+1}^s + \mu f_{2n^*}^s , \\
 f_{2n+1}^m &= \delta f_{2n}^s + (1 - \delta - \mu)f_{2n+1}^s + \mu f_{2n+1^*}^s , \\
 f_{2n+1^*}^m &= \mu f_{2n+1}^s + (1 - \delta - \mu)f_{2n+1^*}^s + \delta f_{2n^*}^s , \\
 f_{2n^*}^m &= \mu f_{2n}^s + \delta f_{2n+1^*}^s + (1 - \delta - \mu)f_{2n^*}^s .
 \end{aligned} \tag{7}$$

658 The inferred values are slightly different. The estimated mutation rate,  $\mu = 1.036 \cdot 10^{-7}$  [ $8.01 \cdot 10^{-8} - 1.339 \cdot 10^{-7}$ ], corresponds to a mutation target size of  $\sim 300 - 500$ , assuming the mutation  
 660 rate per base pair is roughly  $2 \cdot 10^{-10}$  (ref.<sup>78</sup>) or  $3.3 \cdot 10^{-10}$  (ref.<sup>38</sup>). The estimated aneuploidy  
 662 rate,  $\delta = 2.358 \cdot 10^{-4}$  [ $1.766 \cdot 10^{-4} - 2.837 \cdot 10^{-4}$ ] is 5-35-fold higher than in previous studies:  
 664 for Chromosome III in diploid *S. cerevisiae*, Zhu et al.<sup>78</sup> estimated  $6.7 \cdot 10^{-6}$  chromosome gain  
 events per generation, and Kumaran et al.<sup>35</sup> estimate  $3.0 - 4.3 \cdot 10^{-5}$  chromosome loss events per  
 666 generation (95% confidence interval). The estimated fitness values are  $w_{2n+1} = 1.024$  [ $1.023 - 1.025$ ],  
 $w_{2n+1^*} = 1.025$  [ $1.024 - 1.026$ ],  $w_{2n^*} = 1.032$  [ $1.031 - 1.033$ ], all relative to the fitness of  $2n$ , which

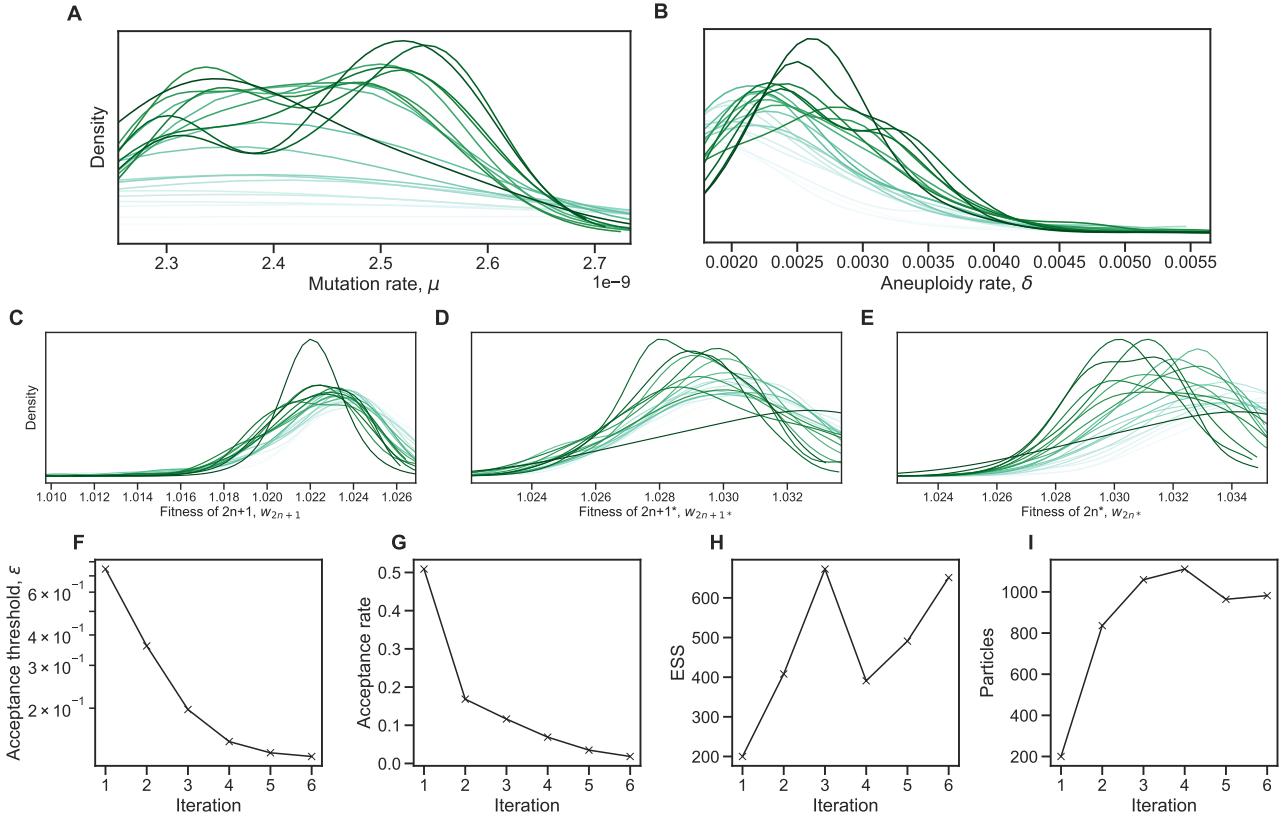
is set to  $w_{2n} = 1$ .

We simulated genotype frequency dynamics using parameter samples from the posterior distribution,  
 668 and computed the posterior distribution of  $F_A$ . The mean  $F_A$  in this case is just 0.0189 [0.0004 -  
 670 0.1214 95% CI], lower than without the transitions to less-fit genotypes. Here,  $F_A$  is the sum of  
 frequencies of both  $2n_A^*$  and  $2n + 1^*$ , which reaches a frequency of 0.0007. Out of 100,000 posterior  
 samples, none had  $F_A$  above 0.05 (i.e., 5% of the population).

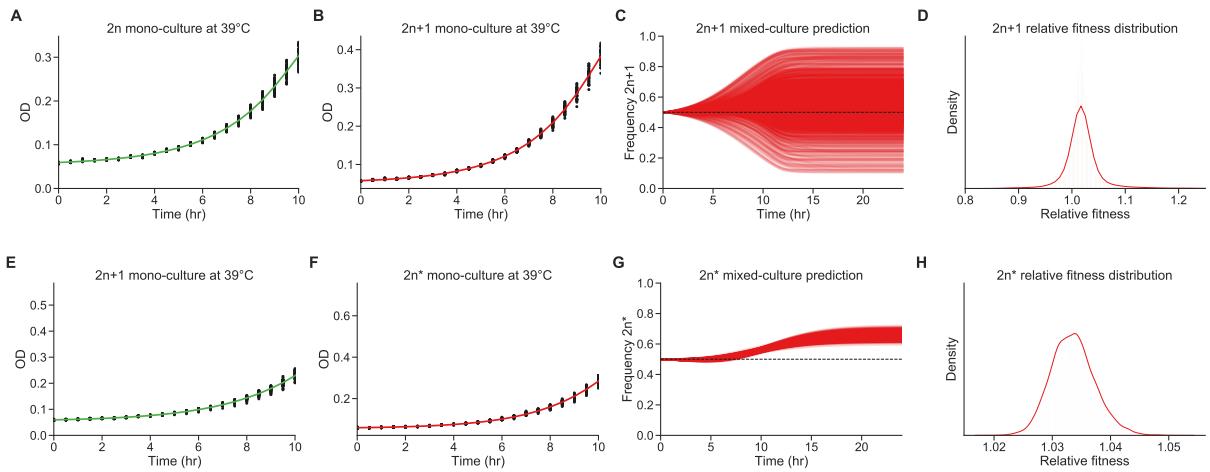
672 **Supplementary Figures & Tables**



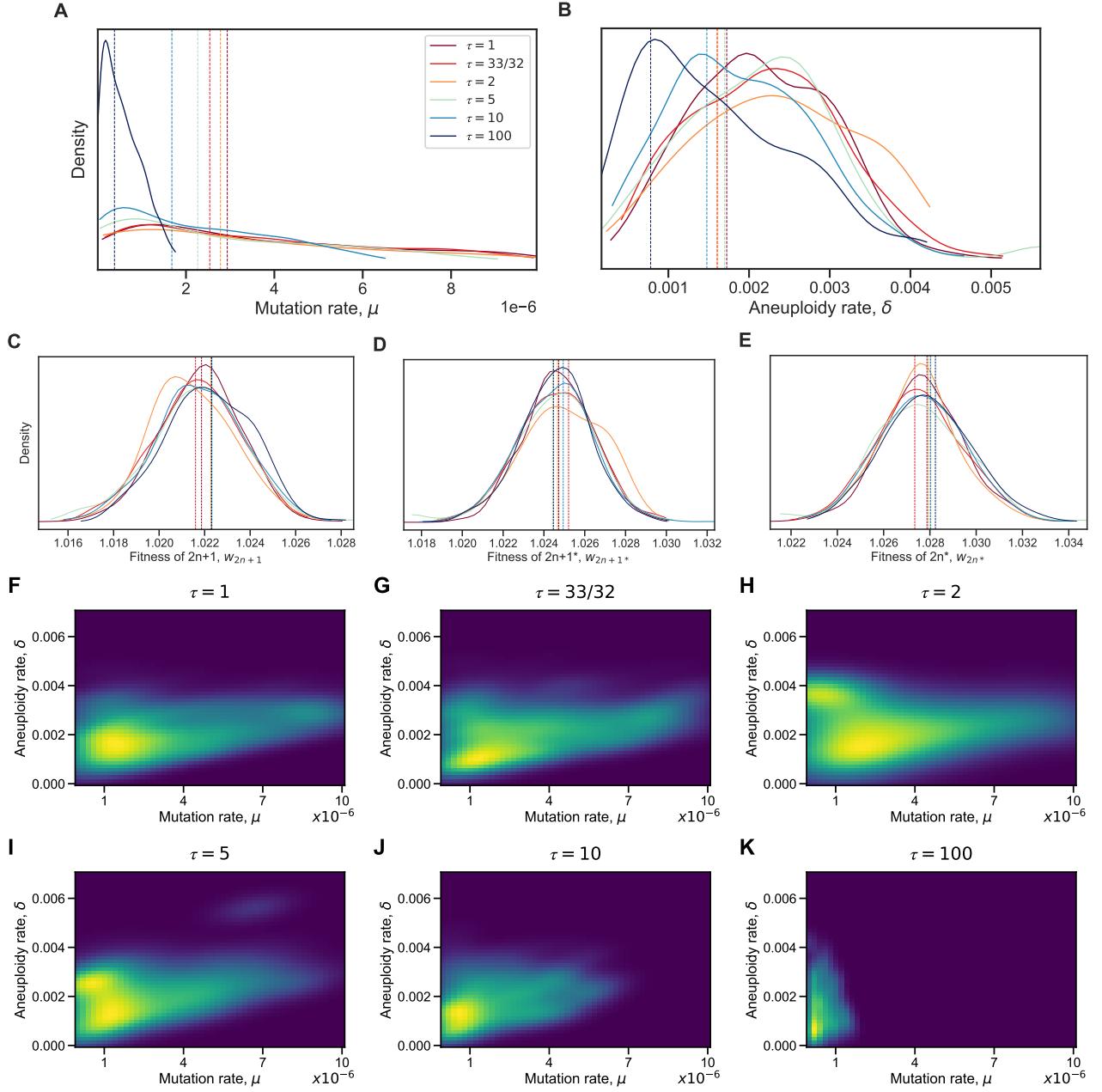
**Figure S1: Posterior distribution validation.** The posterior distribution of model parameters is roughly the same regardless of the number of simulations (4-10,000 replicates) used to approximate the likelihood (eq. (4)).



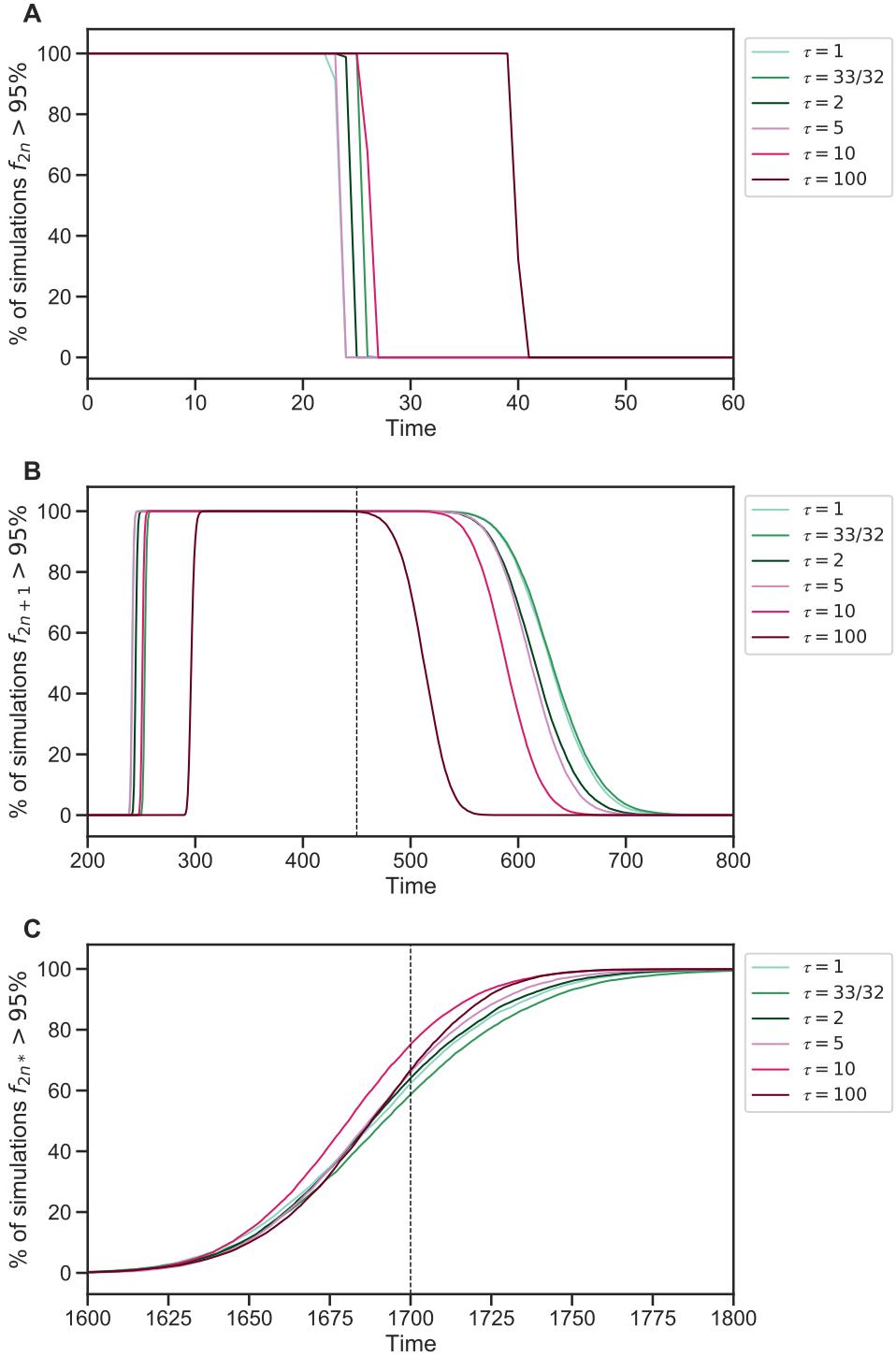
**Figure S2: Inference convergence.** The ABC-SMC algorithm was used to infer the model parameters. **(A-E)** The approximate posterior distributions of model parameters at each iteration of the ABC-SMC algorithm demonstrates convergence, as the posterior did not significantly change after the first iteration,  $t = 1$ . **(F-I)** ABC-SMC measures of convergence. After iteration number 6, the acceptance threshold was  $\epsilon = 0.13$  (i.e.,  $\mathcal{L} = 0.87$ , eq. (4)), the acceptance rate was 0.018, the number of particles was 982, and the effective sample size ESS=651.



**Figure S3: Fitness estimation from growth curves.** **(A-D)** Fitness estimation from growth curves of  $2n$  and  $2n+1$  at  $39^{\circ}\text{C}$ .  $w_{2n+1}/w_{2n}=1.024$  (95% CI: 0.959 - 1.115). Curveball **(E-H)** Fitness estimation from growth curves of  $2n+1$  and  $2n^*$  at  $39^{\circ}\text{C}$ .  $w_{2n^*}/w_{2n+1}=1.033$  (95% CI: 1.027 - 1.041). Growth curves previously described in Yona et al.<sup>74</sup>, Figs. 3C, 4A, and S2. Fitness estimated from growth curves using Curveball, a method for predicting results of competition experiments from growth curve data<sup>46</sup> [curveball.yoavram.com](http://curveball.yoavram.com). See *Models and Methods, Prior distributions* for more details. **(A,B;E,F)** Mono-culture growth curve data (markers) and best-fit growth models (lines). **(C,G)** The mixed-culture prediction for the strains from A,B and E,F respectively, 6,375 generated curves. **(D,H)** The relative fitness distribution for  $2n+1$  relative to  $2n$  (panel D) and  $2n^*$  relative to  $2n+1$  (panel H). Figures generated by Curveball.



**Figure S4: Model with elevated mutation rate in aneuploid cells.** (A-E) The inferred posterior distributions for models with different values of  $\tau$ , the fold-increase in mutation rate in aneuploid cells ( $2n + 1$  and  $2n + 1^*$ ). Vertical dashed lines represent the MAP (maximum a posteriori) of each distribution. When the increase in mutation rate is high,  $\tau = 10$  and  $\tau = 100$ , the inferred mutation (A) and aneuploidy (B) rates tend to be lower. (F-K) The inferred joint posterior distribution of mutation rate ( $\mu$ ) and aneuploidy rate ( $\delta$ ) with different  $\tau$  values (dark purple and bright yellow for low and high density, respectively).

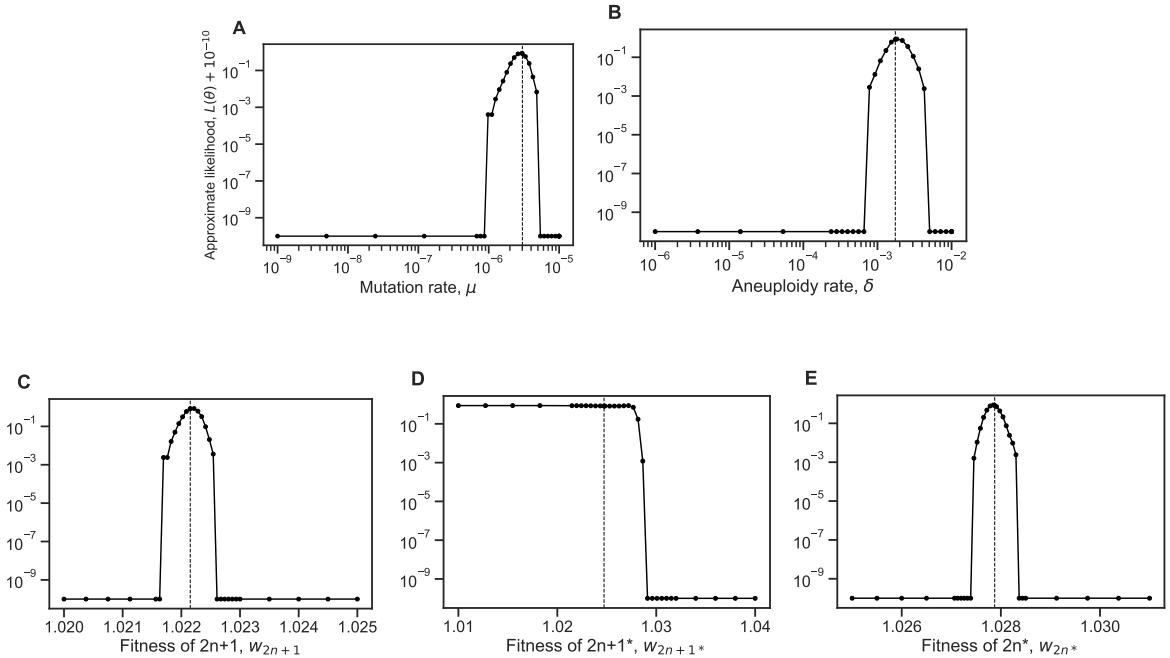


**Figure S5: Genotype fixations for models with increased genetic instability.** We estimated the parameters for different models, each assuming a different value of  $\tau$ , the fold-increase in mutation rate in aneuploid cells. We then generated 10,000 simulations using the MAP estimate of each model and evaluated the fraction of simulations in which the frequency of genotype  $2n$  (**A**),  $2n + 1$  (**B**), and  $2n^*$  (**C**) is above 95% (y-axis) at each generation (x-axis). Note that  $2n + 1^*$  did not fix. We can see that  $\tau = 100$  can be distinguished if the waiting time for  $f_{2n} < 95\%$  is known (panel A) or if the waiting time for  $f_{2n+1} > 95\%$  or  $f_{2n+1} < 95\%$  is known (panel B). It is harder to distinguish between  $1 \leq \tau \leq 10$ .

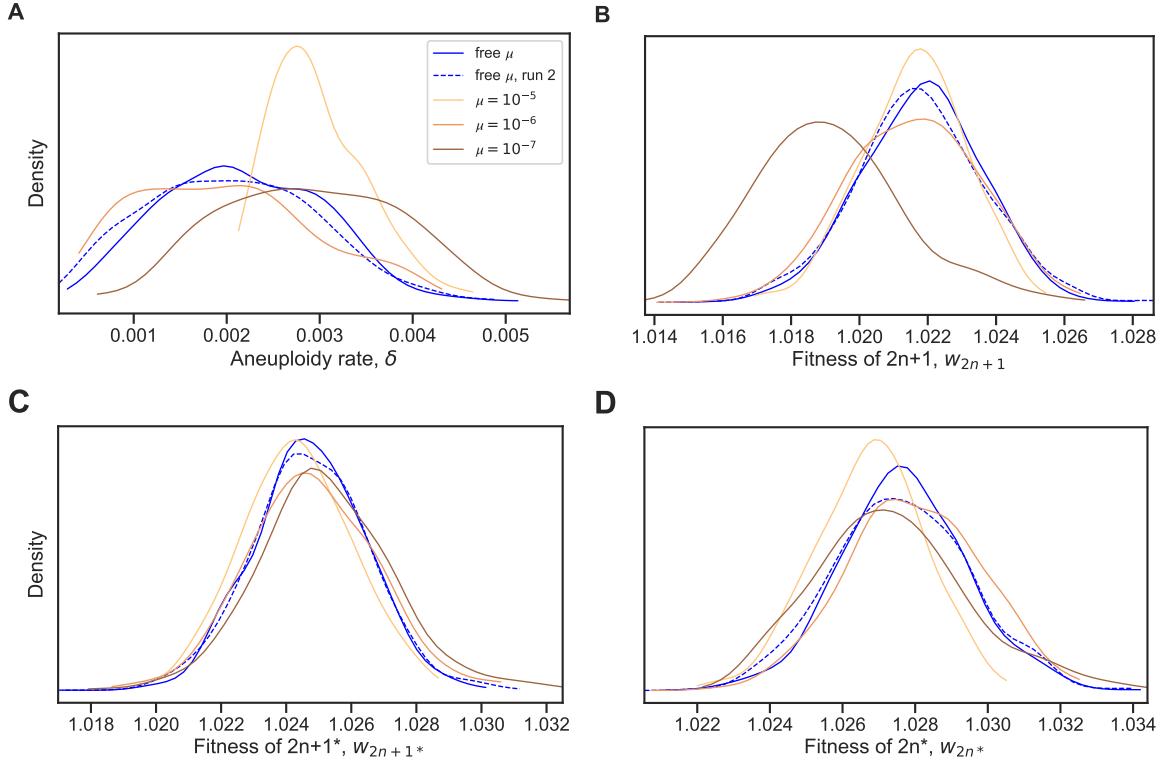
**Table S1: WAIC values for different  $\tau$  values.** Differences of less than 6 are considered of weak significance<sup>28</sup>.

Model	WAIC
$\tau = 1$	-9
$\tau = 33/32$	-9
$\tau = 2$	-8
$\tau = 5$	-12
$\tau = 10$	-9
$\tau = 100$	-12

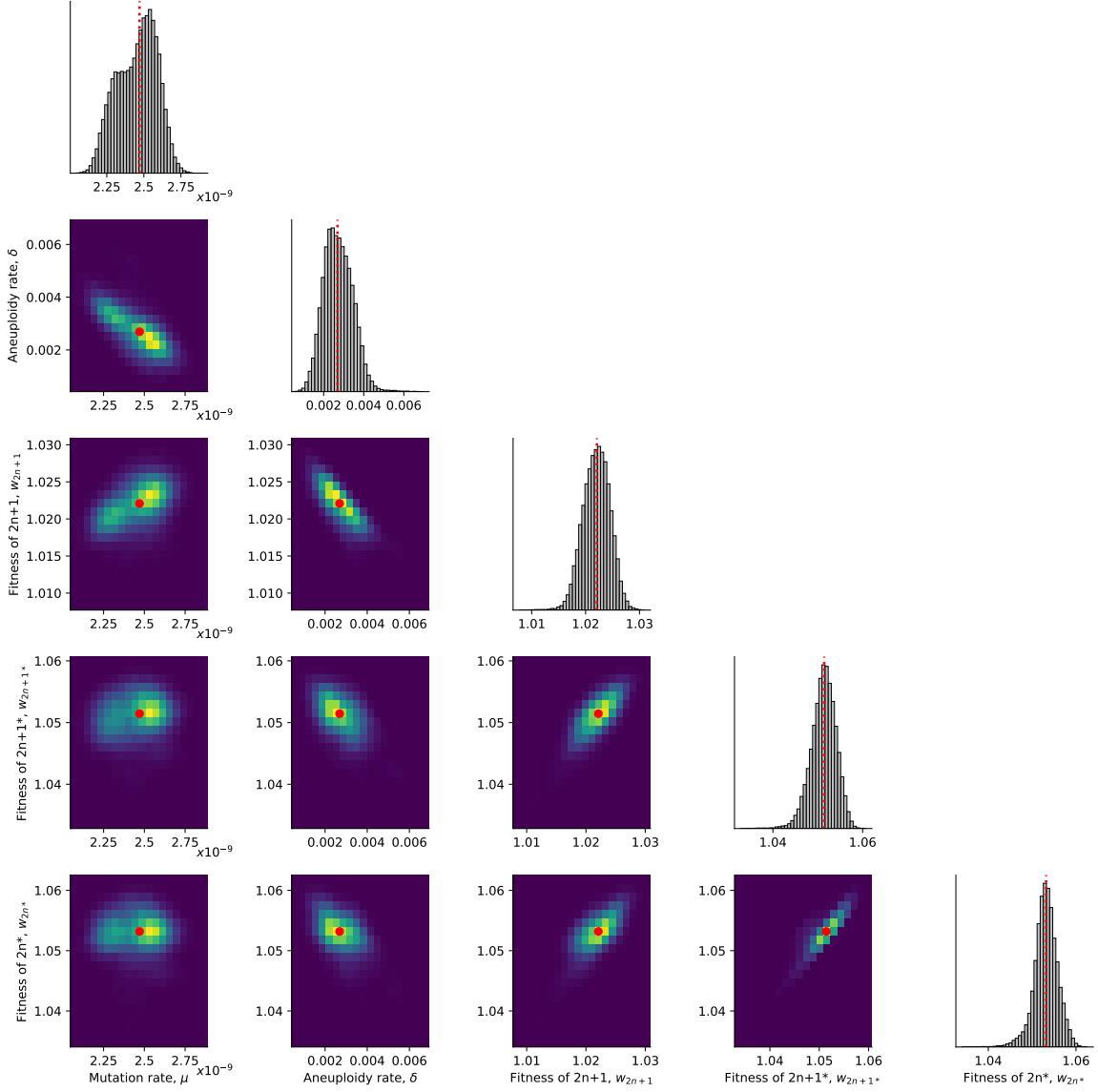
WAIC defined in eq. (6).



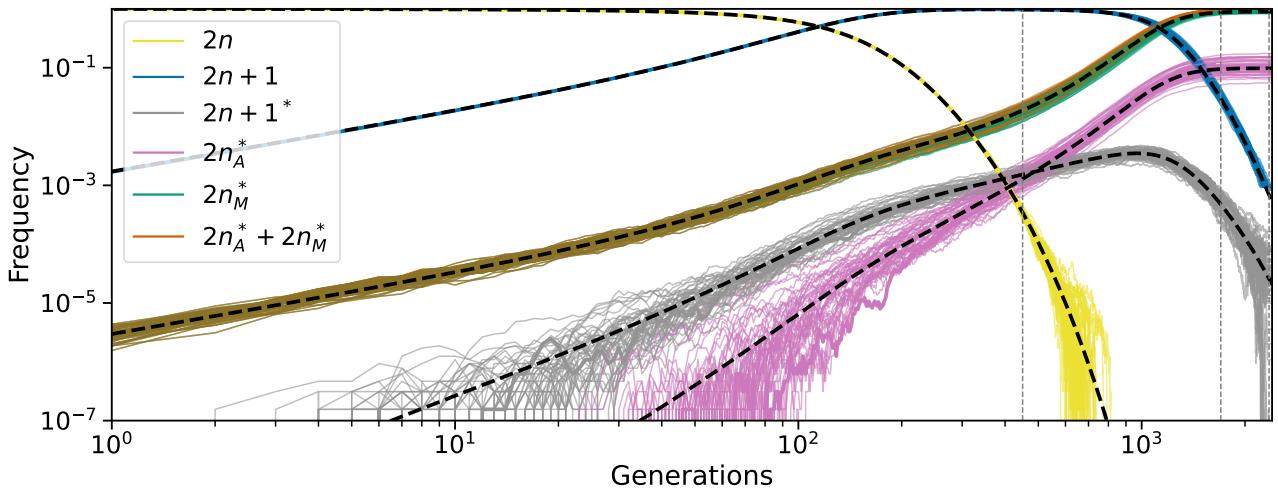
**Figure S6: Likelihood profiles.** Sensitivity of the model approximate likelihood,  $\mathcal{L}(\theta)$ , to changing a single parameter while the other parameters remain fixed at their MAP estimates. Dashed vertical line represents the MAP value. The prior distributions for the mutation rate and aneuploidy rate are  $\mu \sim U(10^{-9}, 10^{-5})$  and  $\delta \sim U(10^{-6}, 10^{-2})$ , respectively.



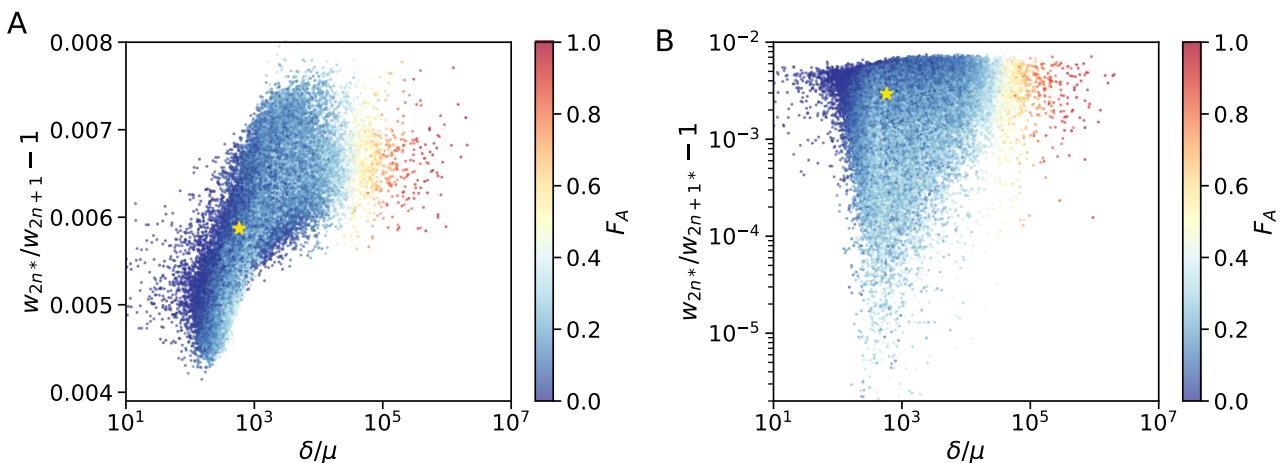
**Figure S7: Model with fixed mutation rate. (A-D)** The inferred posterior distributions for models with free and fixed mutation rate,  $\mu$ . The MAP (maximum a posteriori) and 50% HDI (highest density interval) for each model are: **free  $\mu$ , run 1:**  $\delta = 1.720 \cdot 10^{-3}$  [ $1.470 \cdot 10^{-3} - 2.786 \cdot 10^{-3}$ ],  $w_{2n+1} = 1.022$  [1.021 – 1.023],  $w_{2n+1^*} = 1.025$  [1.024 – 1.026],  $w_{2n^*} = 1.028$  [1.026 – 1.029]; **free  $\mu$ , run 2:**  $\delta = 2.129 \cdot 10^{-3}$  [ $1.334 \cdot 10^{-3} - 2.695 \cdot 10^{-3}$ ],  $w_{2n+1} = 1.022$  [1.02 – 1.023],  $w_{2n+1^*} = 1.025$  [1.023 – 1.026],  $w_{2n^*} = 1.028$  [1.026 – 1.029];  **$\mu = 10^{-5}$ :**  $\delta = 2.903 \cdot 10^{-3}$  [ $2.399 \cdot 10^{-3} - 3.156 \cdot 10^{-3}$ ],  $w_{2n+1} = 1.022$  [1.021 – 1.023],  $w_{2n+1^*} = 1.024$  [1.023 – 1.025],  $w_{2n^*} = 1.027$  [1.026 – 1.028];  **$\mu = 10^{-6}$ :**  $\delta = 1.917 \cdot 10^{-3}$  [ $9.624 \cdot 10^{-4} - 2.447 \cdot 10^{-3}$ ],  $w_{2n+1} = 1.022$  [1.02 – 1.023],  $w_{2n+1^*} = 1.025$  [1.023 – 1.026],  $w_{2n^*} = 1.028$  [1.027 – 1.029];  **$\mu = 10^{-7}$ :**  $\delta = 2.901 \cdot 10^{-3}$  [ $2.139 \cdot 10^{-3} - 3.671 \cdot 10^{-3}$ ],  $w_{2n+1} = 1.019$  [1.017 – 1.02],  $w_{2n+1^*} = 1.025$  [1.024 – 1.026],  $w_{2n^*} = 1.027$  [1.026 – 1.029].



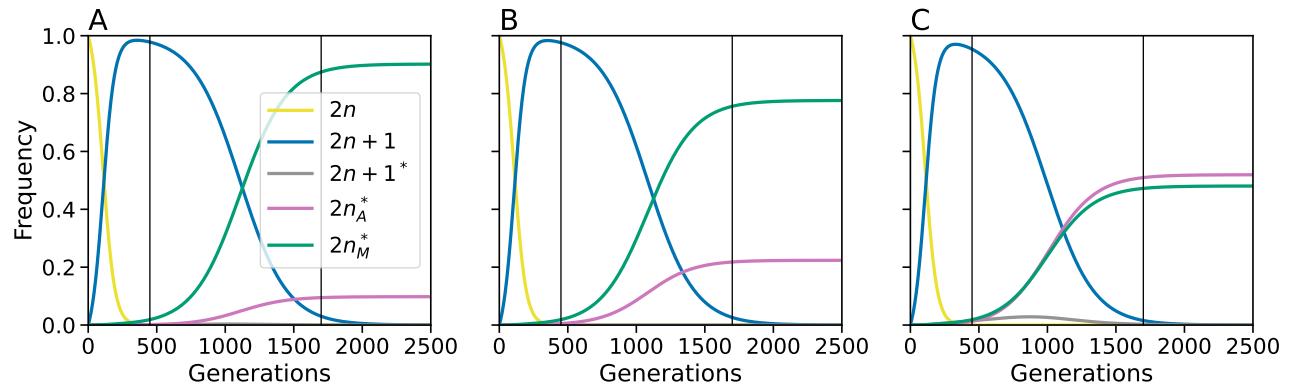
**Figure S8: Posterior distribution of parameters inferred with the extended prior distribution.** On the diagonal, the inferred posterior distribution of each model parameter. Below the diagonal, the inferred joint posterior distribution of pairs of model parameters (dark purple and bright yellow for low and high density, respectively). Red markers and orange lines for the joint MAP estimate (which may differ from the marginal MAP, as the marginal distribution integrates over all other parameters).



**Figure S9: Posterior predicted genotype frequencies in log-log scale.** Frequency dynamics of the different genotypes with MAP parameter estimates, same as Figure 4A, but in log-log scale. Black dashed curves for a deterministic model without genetic drift. Clearly, appearance of  $2n + 1$  and  $2n_M^*$  is deterministic. Appearance of  $2n + 1^*$ , and therefore  $2n_A^*$ , is stochastic, however, the frequency dynamics are deterministic above a frequency of roughly 0.001. Note that the  $2n_M^*$  and the  $2n_A^* + 2n_M^*$  lines are overlapping for much of their trajectories.



**Figure S10: Posterior distribution of  $F_A$ .** (A,B)  $F_A$  values (color coded) as in Figure 4 for different parameter choices on the x- and y-axes. White star denotes the MAP estimate.



**Figure S11: Effect of genomic instability on genotype frequencies.** Genotype frequencies in the deterministic model without drift and (A) with MAP parameter estimates; (B) with 100-fold increase in rate of chromosome loss (transition from  $2n+1^*$  to  $2n^*$ ); or (C) with 10-fold increase in mutation rate in aneuploid cells (transition from  $2n+1$  to  $2n+1^*$ ). Corresponding  $F_A$  values (purple line at generation 2,500) are 0.098, 0.223, and 0.519, respectively.