Stress-induced mutagenesis and the evolution of complex adaptations

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# Summary (max 200)

# Keywords (3-6)

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# Short title

Stress-induced mutagenesis and complex adaptations

# Introduction

Stress-induced mutagenesis, the phenomenon in which stressed or maladapted individuals increase their mutation rate, has been demonstrated in numerous species, both prokaryote and eukaryote [1–3]. The phenomenon is considered by many to have a meaningful impact on *evolvability* - the capacity of individuals and populations to adapt to environmental changes [REFS]. In a previous work we showed that stress-induced mutagenesis is favored by natural selection over constant rate mutagenesis in asexual populations and that it increases the mean fitness of populations due to the increased generation of beneficial mutations in maladapted individuals [4]. Here, we focus instead on the effect of stress-induced mutagenesis on the evolution of complex traits.

Complex traits, coded by multiple genes, present an open evolutionary question, first described by Sewall Wright in 1931 [5]: *if different alleles are separately deleterious but jointly advantageous, how can a population evolve from one co-adapted gene complex to a better one?* Or, in terms of adaptive landscapes [6], how can a population cross a fitness valley and shift from one adaptive peak to a higher one?

Wright suggested the "shifting-balance theory of evolution" [7], which is based on the division of the population into small sub-populations and relies on genetic drift and migration as complementary processes to mutation and selection. This solution is valid [8–10] but seems to be limited to specific parameter ranges [11–14]. As a result, there is a disagreement if the "shifting-balance theory" is an important process in evolution [15,16].

Here, we analyze a population genetic model of an asexual population undergoing an adaptive peak shift in a rugged adaptive landscape. We derive analytical expressions that suggest that stress-induced mutagenesis increases the population adaptation rate and show the results of stochastic simulations that validate our analytic expressions. We then discuss possible predictions of our model and how it relates to the literature.

# Material and methods

## Analytical model

We consider two loci with alleles *A/a* and *B/b* and a population at a mutation-selection balance (MSB) in an environment in which *ab* is the wildtype with a fitness value of *1*, single mutants (*Ab* and *aB*) have a fitness value of *1-s*, with *s* as the selection coefficient, and double mutants (*AB)* have a fitness value of *(1-s)2*. This corresponds to a multiplicative fitness function *(1-s)m* where *m* is the number of deleterious mutations the individual has.

Mutation from *a* to *A* and from *b* to *B* occurs with a probability *µ* at reproduction (we neglect back-mutation) - *µ* is therefore the site-specific beneficial mutation rate. In addition, the number of new deleterious mutations that occur across the genome at reproduction follows a Poisson distribution with an average *U* - the genomic deleterious mutation rate.

Individuals with stress-induced mutegenesis and a fitness lower than 1 increase both mutation rates *τ*-fold.

We are interested in the capacity of the population to adapt to an environmental change in which the fitness of the double mutant *AB* changes from *(1-s)2*  to *1+sH*, where *H* scales the advantage of *AB* in comparison with the disadvantage of single mutants. Figure 1a presents an illustration of the adaptive landscape of the analytical model.

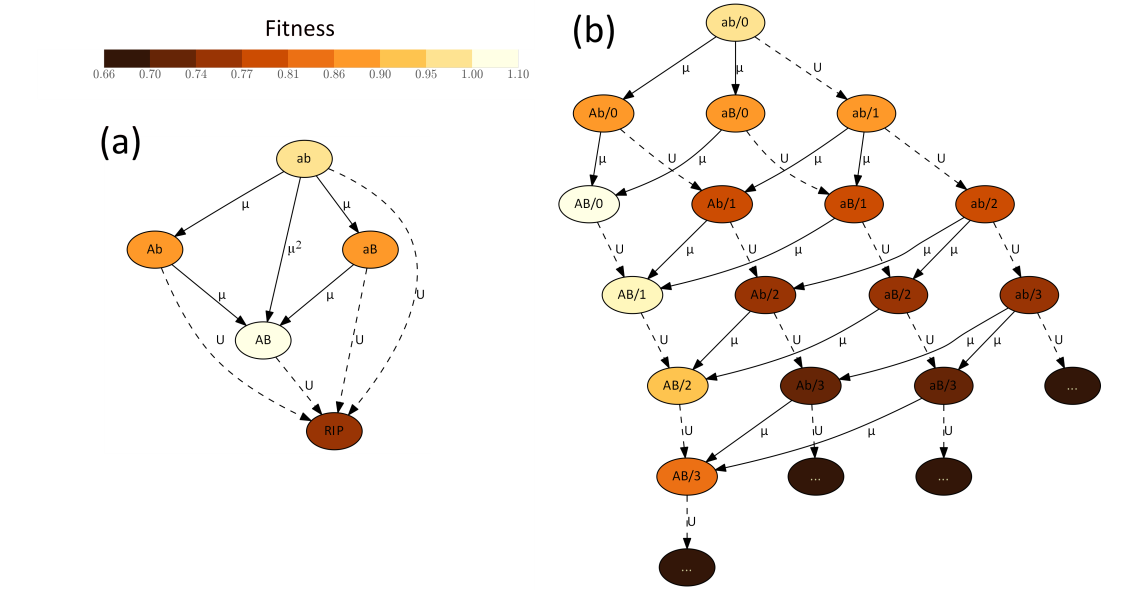


Figure 1 – Adaptive landscape illustration. Each node represent a specific genotype. Node labels specify the alleles at the *A/*a and *B/*b loci and the number of additional deleterious alleles after the forward slash (only in (b) and only as much as 3 deleterious mutations are shown to keep the figure simple). Arrows define the direction of mutation (solid for site-specific beneficial mutations with rate µ, dashed for genomic deleterious mutations with rate *U*). Node colour indicates the fitness of a genotype, from pale brown for optimal fitness to dark brown for lower fitness, see the colourbar. Parameters used: *s=0.01*, *H=2*. (a) In the analytical model individuals with deleterious mutations are considered evolutionary dead-ends (*RIP*) anddo not contribute to adaptation.(b) In the stochastic model we make no such assumption. In this panel only single mutations are shown for simplicity.

At the MSB, the frequency of wildtype (*ab*) individuals is , the frequency of single mutants (*Ab* and *aB* combined) is and the frequency of double mutants (*AB*) is . Therefore, there are several constraints on the parameter range:

1. The above MSB approximations are only valid when *µ/s<1* or *µ<s*.
2. If *N*(*µ/s)2>1* then there are double mutants in the population at the MSB and therefore adaptation will be rapid and will not require new mutations.
3. If *Nµ/s<1* then there are no single mutants at the MSB and double mutatns must be generated by a double mutation in a wildtype individual. Therefore, increasing the mutation rate of individuals with fitness below 1 will have a much smaller effect than if single mutants were abundant.
4. If we assume that individuals that accumulated deleterious mutations are "evolutionary dead-ends" (figure 1a) and cannot be the origin of adaptation, then the fraction of such individuals must be small - *U/s<1*. This replaces the first constraint above because *µ* is much smaller than *U*.

To summarize the above constraints:

|  |  |
| --- | --- |
|  | (1) |
|  | (2) |

Table 1 has estimation on the parameter values in *Escherischia coli*. Taking the conservative estimations, we get this constraint on the population size - *105 ≤ N ≤ 10*7.

In the following derivations we assume both of the above constraints.

Table 1 – Model parameters and estimated values for bacteria

|  |  |  |  |
| --- | --- | --- | --- |
| Symbol | Name | Estimate | References |
| *s* | Selection coefficient | 0.001-0.03 | [17,18] |
| *H* | Double mutant advantage | 1-10 | [18] |
| *U* | Genomic deleterious mutation rate | 0.0004-0.003 | [19,20] |
| *µ* | Site-specific beneficial mutation rate | U/5000 | [18] |
| *τ* | Fold-increase in mutation rate | 1-100 | [21,22] |
| *N* | Population size | 104-1010 |  |

## Stochastic model

To validate our analytic approximations, we developed a Wright-Fisher simulation with mutation, selection and random genetic drift. The main differences between the analytical and the stochastic models are: (i) The simulations incorporate genetic drift by randomly sampling each generation from the previous one using a multinomial distribution. (ii) Individuals with deleterious mutations are not "evolutionary dead-ends" - individuals are allowed to accumulate up to 25 deleterious mutations (figure 1b). (iii) Simulations start with an *ab* mutation-free population. After reaching MSB, the environment is changed so that *AB* is advantageous. Therefore, the stochastic model assumes nothing about the distribution of deleterious mutations at the MSB.

# Results

The adaptation process is divided to two distinct processes: (i) the appearance of the double mutant *AB*; and (ii) its subsequent fixation. The appearance of the double mutant mainly depends on mutation and therefore takes much longer than fixation, which depends on the interplay between selection and genetic drift. In the following sections we provide analytic expressions for the appearance and fixation probabilities of the double mutant and the expected total waiting time for adaptation. In addition, we compare these analytic approximations to simulation results.

First-order approximations are marked by a star (\*). The complete derivations of all the expressions are given in the *Electronic Supporting Material*.

## Appearance of a double mutant

In a population at an MSB without double mutants *AB*, the probability *q* that a random newborn is a double mutant and neglecting individuals with deleterious mutations is:

|  |  |
| --- | --- |
|  |  |

However, if mutation is stress-induced, then the mutation rate of single mutants is increased *τ*-fold and the appearance probability is:

|  |  |
| --- | --- |
|  |  |

Deriving with respect to τ, we find that as long as , increasing *τ* increases the appearance probability of the double mutant.

## Fixation of a double mutation

Assuming that the advantage of the double mutant is considerable (for example, *H>1)* and that the population size is large (constraint (2) ensures that), a double mutant has two possible fates after its appearance: fixation or extinction. We derive the fixation probability *ρ* of the double mutant following Eshel [23]. Assuming that fitness is measured by the number of progeny which is Poisson distributed we find:

which is a classic result in population genetics [23].

Because stress-induced mutagenesis can have an effect on the population mean fitness [4], the fixation probability with stress-induced mutagenesis is:

Comparing the righthand side of the last two equations, and because and , this demonstrates that , that is, stress-induced mutagenesis increases the fixation probability of the double mutant. This is because maladapted individuals accumulate more deleterious mutations then adapted individuals, producing a wider fitness distribution and higher relative fitness for adapted individuals. However, the derivative of with respect to *τ* is , so increasing *τ* only has a small effect on .

## Adaptation rate

The adaptation time can be approximated by the number of generation for appearance of a double mutant *AB* that will to fixation. This number is geometrically distributed with probability *1/Nqρ*, where *N* is the population size and *q* and *ρ* are the probabilities shown in the previous sections. The expected adaptation time without and with stress-induced mutagenesis is therefore:

We compared these expressions to find an approximate sufficient condition for stress-induced mutagenesis to decrease the adaptation time:

That is, the mutation rate of maladapted individuals must be larger than that of well-adapted individuals but lower than one mutation per genome per generation. Two estimates of the genomic deleterious mutation rate *U* in *E. coli* are 0.003 and 0.0004 (see Table 1), which sets the upper limit on *τ* to be between 333 and 2,500. This is much higher than the values of *τ* in the literature *- 1<τ<100* (see Table 1).

## Simulation results

Figure 2 shows the effect os stress-induced mutagenesis on complex adaptation. The approximations (in dashed blue), fit the simulation results (in black circles) very well.

The starting point of all the lines is at *τ=1* which represents populations without stress-induced mutagenesis. Hence, both the approximations and the simulation results agree that stress-induced mutagenesis increases the adaptation rate, and that this effect increases with *τ*.



Figure 2 – Stress-induced mutagenesis reduces the adaptation time. Analytical approximations in dashed blue and simulation results in black circles. (a) Fold-change in waiting time for appearance of a double mutant with stress-induced mutagenesis compared to without. Note that a 10-fold increase in mutation rate results in a ~10-fold decrease in waiting time. (b) Fold-change in the number of double mutant appearances before fixation. The increase in fixation probability (12) is to small to be seen on this scale. (c) The waiting time for adaptation – appearance and fixation of a double mutant – in generations, as a function of the fold-increase in mutation rate in stressed individuals. Note how a minor increase in mutation rate results in a considerable decreae in adaptation time. Parameters used: selection coefficient *s=0.05*, double mutant advantage *H=2*, genomic mutation rate *U=0.0004*, locus specific mutation rate *µ=U/5000*, population size *N=106*.

# Discussion

We studied the effect of stress-induced mutagenesis on adaptation in rugged adaptive landscapes. Our approximations, validated by simulation results, show that stress-induced mutagenesis can considerably reduce the adaptation time by increasing the mutation rate of single mutants, as well as increasing the fixation probability of double mutants.

An interesting comparison can be made to constitutive mutagenesis, in which all individuals increase their mutation rate all the time. Although constitutive mutagenesis can help to reduce the adaptation time by increasing the generation of beneficial mutations, its long-term effect is largely detrimental [24] and it is easily outcompeted by stress-induced mutagenesis [4].

Our work provides a theoretical and formal basis to the conjecture that stress-induced mutagenesis facilitates adaptation and increases the evolvability of populations, as has been suggested by many authors before [REF]. The next step would be to experimentaly verify our results. This can be done, for example, in a bacterial system such as *E. coli*, in which one can interfere with the regulation of hypermutation by stress [REF- Cirz?]. If, for example, frameshift mutations are introduced to the lacI and lacZ genes and the population is allowed to evolve in a lactose rich environment, one could measure the average adaptation time with and without stress-induced mutagenesis and compare it to our analytical approximations. Validation of our model in an experimental setting will provide strong evidence of the contribution of stress-induced mutagenesis to adaptive evolution.

# Acknowledgments

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# Data accessability

The simulations source code and the source code used to analyse the data and generate the figures are available as a *git* repository on GitHub at <https://github.com/yoavram/ruggedsim.git>. The codeis written in Python (<http://www.python.org>) using NumPy (<http://www.numpy.org>) and SciPy (<http://www.scipy.org>).

# References

1 Galhardo, R. S., Hastings, P. J. & Rosenberg, S. M. 2007 Mutation as a stress response and the regulation of evolvability. *Critical reviews in biochemistry and molecular biology* **42**, 399–435. (doi:10.1080/10409230701648502)

2 Sharp, N. P. & Agrawal, A. F. 2012 Evidence for elevated mutation rates in low-quality genotypes. *Proceedings of the National Academy of Sciences of the United States of America* **109**, 6142–6. (doi:10.1073/pnas.1118918109)

3 MacLean, R. C., Torres-Barceló, C. & Moxon, R. 2013 Evaluating evolutionary models of stress-induced mutagenesis in bacteria. *Nature Reviews Genetics* **14**, 221–227. (doi:10.1038/nrg3415)

4 Ram, Y. & Hadany, L. 2012 The evolution of stress-induced hypermutation in asexual populations. *Evolution* **66**, 2315–28. (doi:10.1111/j.1558-5646.2012.01576.x)

5 Wright, S. 1931 Evolution in Mendelian Populations. *Genetics* **16**, 97–159.

6 Gavrilets, S. 2004 *Fitness Landscapes and the Origin of Species (MPB-41) (Monographs in Population Biology)*. Princeton University Press.

7 Wright, S. 1988 Surfaces of selective value revisited. *American Naturalist* **131**, 115–123.

8 Crow, J. F., Engels, W. R. & Denniston, C. 1990 Phase Three of Wright’s Shifting-Balance Theory. *Evolution* **44**, 233–247.

9 Wade, M. & Goodnight, C. 1991 Wright’s shifting balance theory: an experimental study. *Science* **253**, 1015–1018. (doi:10.1126/science.1887214)

10 Peck, S. L., Ellner, S. P. & Gould, F. 2000 Varying migration and deme size and the feasibility of the shifting balance. *Evolution; international journal of organic evolution* **54**, 324–7.

11 Moore, F. B. G. & Tonsor, S. J. 1994 A Simulation of Wright Shifting-Balance Process - Migration and the Three Phases. *Evolution* **48**, 69–80.

12 Gavrilets, S. 1996 On phase three of the shifting-balance theory. *Evolution* **50**, 1034–1041. (doi:10.2307/2410644)

13 Phillips, P. C. 1996 Waiting for a compensatory mutation: phase zero of the shifting-balance process. *Genetical Research* **67**, 271–283.

14 Coyne, J. A., Barton, N. H. & Turelli, M. 1997 Perspective: A Critique of Sewall Wright’s Shifting Balance Theory of Evolution. *Evolution* **51**, 643. (doi:10.2307/2411143)

15 Coyne, J. A., Barton, N. H. & Turelli, M. 2000 Is Wright’s shifting balance process important in evolution? *Evolution* **54**, 306–317. (doi:10.1111/j.0014-3820.2000.tb00033.x)

16 Whitlock, M. C. & Phillips, P. C. 2000 The exquisite corpse: a shifting view of the shifting balance. *Trends in Ecology and Evolution* **15**, 347–348.

17 Kibota, T. T. & Lynch, M. 1996 Estimate of the genomic mutation rate deleterious to overall fitness in E. coli. *Nature* **381**, 694–6. (doi:10.1038/381694a0)

18 Gordo, I., Perfeito, L. & Sousa, A. 2011 Fitness effects of mutations in bacteria. *Journal of molecular microbiology and biotechnology* **21**, 20–35. (doi:10.1159/000332747)

19 Drake, J. W., Charlesworth, B., Charlesworth, D. & Crow, J. F. 1998 Rates of spontaneous mutation. *Genetics* **148**, 1667–86.

20 Wielgoss, S., Barrick, J. E., Tenaillon, O., Cruveiller, S., Chane-Woon-Ming, B., Médigue, C., Lenski, R. E. & Schneider, D. 2011 Mutation Rate Inferred From Synonymous Substitutions in a Long-Term Evolution Experiment With Escherichia coli. *G3: Genes, Genomes, Genetics* **1**, 183. (doi:10.1534/g3.111.000406)

21 Bjedov, I., Tenaillon, O., Gérard, B., Souza, V., Denamur, E., Radman, M., Taddei, F. & Matic, I. 2003 Stress-induced mutagenesis in bacteria. *Science* **300**, 1404–9. (doi:10.1126/science.1082240)

22 Hall, L. M. C. & Henderson-Begg, S. K. 2006 Hypermutable bacteria isolated from humans--a critical analysis. *Microbiology (Reading, England)* **152**, 2505–14. (doi:10.1099/mic.0.29079-0)

23 Eshel, I. 1981 On the survival probability of a slightly advantageous mutant gene with a general distribution of progeny size—a branching process model. *Journal of mathematical biology* **12**, 355–362.

24 Taddei, F., Radman, M., Maynard Smith, J., Toupance, B., Gouyon, P.-H. & Godelle, B. 1997 Role of mutator alleles in adaptive evolution. *Nature* **387**, 700–2. (doi:10.1038/42696)

# Electronic supporting material

## Appearance of a double mutant

Because there are no double mutants (*AB*) at the time of the environmental change, double mutants can appear either via a double mutation in a wildtype individual, *ab*, or via a single mutation in a single mutant, *Ab* or *aB* (figure 1a). At the MSB the number of deleterious mutations per individual follows a Poisson distribution [@Haigh1978]. Therefore, the frequencies of mutation-free wildtype *ab* and single mutants *aB* and *Ab* are and . The probability *q* that a random newborn is a double mutant, given there are no double mutants and neglecting individuals with deleterious mutations is:

|  |  |
| --- | --- |
|  | (3) |

If mutation is stress-induced, then the mutation rate of single mutants is increased *τ*-fold and the appearance probability is:

|  |  |
| --- | --- |
|  | (4) |

Note that stress-induction increases the transition from single mutants to other types, but does not significantly change the MSB frequency of single mutants, because this frequency is mainly determined by the mutation rate of the wildtype which does not hypermutate.

The above expressions can be simplified by using first-order approximations. Starting with Eq. (3) for populations without stress-induced mutation:

The last step assumes that *2s* is much larger than *s2* and *sU* is much larger than *2µ.* Rearranging the last expression gives us

|  |  |
| --- | --- |
|  | (5) |

For a population with stress-induced mutation the first-order approximation is based on the full expression in Eq. (4):

The last approximation assumes that *Us* is smaller than *U* and that *τU* is much larger than *µ/s.* Now,

The last approximation assumed that *2τ>s* and 2*τ2>1*, because *τ>1* and probably even *τ≥10*. Rearranging the last expression gives us the first order approximation for populations with stress-induced mutation:

|  |  |
| --- | --- |
|  | (6) |

Note that by setting and because , is consistent with .

## Fixation probability with stress-induced mutation

Assuming that the advantage of the double mutant is considerable (for example, *H>1)* and that the population size is large (constraint (2) ensures that), a double mutant has two possible fates after its appearance: fixation or extinction. Following Eshel [@Eshel1981] the fixation probability *ρ* of the double mutant is:

|  |  |
| --- | --- |
|  | (7) |

where *α* is the fitness of the double mutant relative to the population mean fitness and assuming that fitness is measured by the number of progeny which is Poisson distributed:

|  |  |
| --- | --- |
|  | (8) |

Here, we only take the fraction of progeny that do not have deleterious mutations - . This factor cannot be ignored because there is variation in mutation rates in the population.

At this stage, double mutants are still very rate, so we can use the population mean fitness at the MSB. Without stress-induced mutagenesis, this evaluates to [@Kimura1966]. Therefore, and

|  |  |
| --- | --- |
|  | (9) |

Assuming that is small we get

|  |  |
| --- | --- |
|  | (10) |

However, as we have shown before [@Ram2012], the mean fitness of a population with stress-induced mutagenesis can be different from because of mutation rate variation. The mean fitness with stress-induced mutagenesis can be calculated by divding the population to the mutation-free fraction which has fitness *1* and the rest of the population with a fraction of . Within the the latter fraction, individuals have at least one deleterious mutation and additional mutations are Poisson distributed with expectation , because these individuals are hypermutating. Therefore the mean fitness of this fraction is . Taken together, the mean fitness of a population with stress-induced mutagenesis is:

|  |  |
| --- | --- |
|  | (11) |

Pluging in the population mean fitness to the relative fitness of the double mutant we get

Pluging that in the fixation probability gives the final result:

|  |  |
| --- | --- |
|  | (12) |

This fixation probability can be further simplified by first-order approximations:

|  |  |
| --- | --- |
|  | (13) |

## Adaptation time

From the probability *q* that a random newborn is a double mutant we can derive the probability that some double mutants appear in the next generation: . The constraint (2) guarantees that *Nq* is very small and therefore this probability can be approximated by *Nq*.

Once a double mutant appears it has a probability *𝜌* to go to fixation.

The time for adaptation *T* can be approximated by the waiting time for a double mutant which will go to fixation *Tw*. This is true as long as fixation is much faster than mutation (guaranteed by *µ*2*<2* which is a weaker constraint than that given by Eq. (1)). *Tw* follows a geometric distribution with probability *Nq𝜌* and therefore the expected time for adaptation can be approximated by:

|  |  |
| --- | --- |
|  | (14) |

Without stress-induced mutation, we plug in Eqs. (10) and (5) and get:

|  |  |
| --- | --- |
|  | (15) |

With stress-induced mutation we plug in Eqs. (6) and (13) and get:

|  |  |
| --- | --- |
|  | (16) |

Comparing these two expressions, we can write the adaptation rate (the inverse of the expected adaptation time) with stress-induced mutagenesis as a function of the adaptation rate without it:

Now, because the second term is positive, if then we can infer that the rate with stress-induced mutagenesis is faster than without. This condition can ve rewritten:

Using the quadrate formula this translates to:

Because *U* is very small, *1-2U* is well apprxomated by *1*, and *1/U2* is much larger than *2*, the RHS of this inequality can be approximated by 2 and we get the following condition:

Therefore, an approximate sufficient condition for stress-induced mutagenesis to decrease the adaptation time is:

|  |  |
| --- | --- |
|  | (17) |