1. Derive the following asymptotic moments associated with $m_3(x)$: mean, variance, first order auto-correlation. Furthermore, compute $\nabla_b g(b_0)$. Which moments are informative for estimating b?

Answer: We obtain the asymptotic moments as follows:

$$\mathbb{E}[x_t] = \mathbb{E}[\rho_0 x_{t-1} + \epsilon_t] = \rho_0 \, \mathbb{E}[x_{t-1}] + \mathbb{E}[\epsilon_t] = \rho_0 \, \mathbb{E}[\rho_0 x_{t-2} + \epsilon_{t-1}] = \rho_0^t \, \mathbb{E}[x_0] = 0.$$

$$\mathbb{E}[(x_t - \mathbb{E}[x_t])^2] = \mathbb{E}[x_t^2] = \mathbb{E}[(\rho_0 x_{t-1} + \epsilon_t)^2]$$

$$= \rho_0^2 \, \mathbb{E}[x_{t-1}^2] + 2\rho_0 \, \mathbb{E}[x_{t-1}\epsilon_t] + \mathbb{E}[\epsilon_t^2]$$

$$= \rho_0^2 \, \mathbb{E}[(\rho_0 x_{t-2} + \epsilon_{t-1})^2] + \sigma_0^2$$

$$= \rho_0^{2T} \, \mathbb{E}[x_0^2] + \sigma_0^2 \sum_{i=0}^{T-1} (\rho_0^2)^i \to \frac{\sigma_0^2}{1 - \rho_0^2} = \frac{4}{3}.$$

$$\mathbb{E}[(x_t - \mathbb{E}[x_t])(x_{t-1} - \mathbb{E}[x_{t-1}])] = \mathbb{E}[x_t x_{t-1}]$$

$$= \rho_0 \, \mathbb{E}[x_{t-1}^2] + \mathbb{E}[\epsilon_t x_{t-1}]$$

$$= \rho_0 \, \mathbb{E}[(\rho_0 x_{t-2} + \epsilon_{t-1})^2]$$

$$= \rho_0^3 \, \mathbb{E}[x_{t-2}^2] + \rho_0 \, \mathbb{E}[\epsilon_{t-1}^2]$$

$$= \rho_0^{2T-1} \, \mathbb{E}[x_0^2] + \rho_0 \sigma_0^2 \sum_{i=0}^{T-2} (\rho_0^2)^i \to \frac{\rho_0 \sigma_0^2}{1 - \rho_0^2} = \frac{2}{3}$$

Moreover, we have that $\nabla_b = \begin{pmatrix} \frac{\partial}{\partial \rho} \\ \frac{\partial}{\partial \sigma^2} \end{pmatrix}$. Evaluating the derivative of $g(\cdot)$ at $b = b_0$ gives us

$$\nabla_b g(b_0) = \begin{bmatrix} 0 & 0 \\ \frac{2\rho_0 \sigma_0^2}{(1-\rho_0)^2} & \frac{1}{1-\rho_0^2} \\ \frac{\sigma_0^2 (1+2\rho_0^2)}{(1-\rho_0^2)^2} & \frac{\rho_0}{1-\rho_0^2} \end{bmatrix} \xrightarrow{\text{true values}} \begin{bmatrix} 0 & 0 \\ 4 & \frac{4}{3} \\ \frac{8}{3} & \frac{2}{3} \end{bmatrix}$$

Both the variance and the first order correlation are informative for estimating *b*. This is because we can estimate the true parameters given the two moments as follows

$$\rho_0 = \frac{\mathbb{E}[(x_t - \mathbb{E}[x_t])(x_{t-1} - \mathbb{E}[x_{t-1}])]}{\mathbb{E}[(x_t - \mathbb{E}[x_t])^2]} \qquad \sigma^2 = \mathbb{E}[(x_t - \mathbb{E}[x_t])^2] - \frac{\mathbb{E}[(x_t - \mathbb{E}[x_t])(x_{t-1} - \mathbb{E}[x_{t-1}])]}{\mathbb{E}[(x_t - \mathbb{E}[x_t])^2]}$$

- 2. Simulate a series of "true" data of length T=200 using (1). We will use this to compute $M_T(x)$.
- 3. Set H = 10 and simulate H vectors of length T = 200 random variables e_t from N(0,1). We will use this to compute $M_{TH}(y(b))$. Store these vectors. You will use the same vector of random variables throughout the entire exercise. Since this exercise requires you to estimate σ^2 , you want to change the variance of e_t during the estimation. You can simply use σ_{e_t} when the variance is σ^2 .
- 4. We will start by estimating the l=2 vector b for the just identified case where m_2 uses mean and variance. Given what you found in part (i), do you think there will be a problem? Of course, in general we would not know whether this case would be a problem, so hopefully the standard error of the estimate of b as well as the J test will tell us something. Let's see.

<u>Answer:</u> The code for completing this question, as well as the other questions for this problem set, is included in the appendix. From the results of part (1), we know that the mean will be uninformative about the true parameter

values, and so in this case only the variance will be useful in identifying the parameter values. One problem that we've notice here and elsewhere is that occasionally the random draws selected by the code will produce issues which make it so that a solution cannot be found. In this problem and for the remainder of the problem set, then, we set a specific random seed value that avoid this issue and produce nice results. This value is 9193, as shown in the code.

(a) $\hat{b}_{TH}^1 = [0.7178, 0.855].$ Figure 1 displays a graph of the objective function.

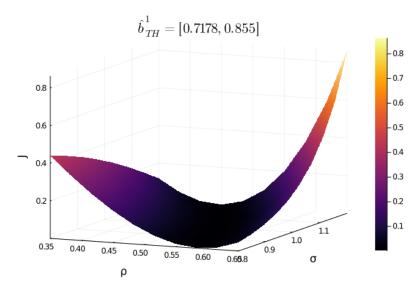


Figure 1: Objective Function for Exercise 4

(b) $\hat{b}_{TH}^2 = [0.7179, 0.8549]$ after the Newey-West correction. Figure 2 displays a graph of the objective function.

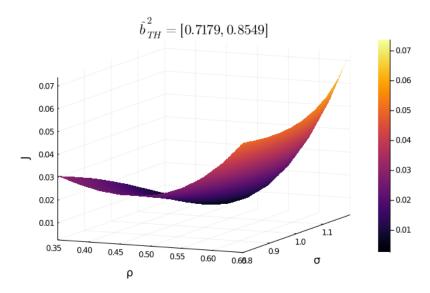


Figure 2: Objective Function for Exercise 4 with the Newey-West Correction

(c)
$$\nabla_b g_T(\hat{b}_{TH}^2) = \begin{bmatrix} 0.22 & 0.06 \\ 3.66 & 3.11 \end{bmatrix}$$

The variance-covariance matrix of \hat{b}_{TH}^2 is given by

$$\begin{bmatrix} 0.01 & -4.07e - 3 \\ -4.07e - 3 & 0.01 \end{bmatrix}$$

The standard errors are given by

$$\begin{bmatrix} 0.08 \\ 0.11 \end{bmatrix}$$

For local identification, it is useful to have the elements of $\nabla_b g_T(\hat{b}_{TH}^2)$ to be relatively large. This is because if the distance between the modeled moments and the data moments is changing rapidly around the optimal parameter choices, we can be more confident in the precision with which we've identified the local optimal parameter values.

- (d) The J-test value is 1.17e 6.
- 5. Next we estimating the l=2 vector b for the just identified case where m_2 uses the variance and autocorrelation. Given what you found in part (i), do you now think there will be a problem? If not, hopefully the standard error of the estimate of b as well as the J test will tell us something. Let's see. For this case, perform steps (a)-(d) above.

Answer: Given our results from part (1), we know that the variance and first order correlation are both informative for estimating *b*. Therefore, we do not anticipate a major problem here.

(a) $\hat{b}_{TH}^1 = [0.4408, 1.0751]$. Figure 3 displays a graph of the objective function.

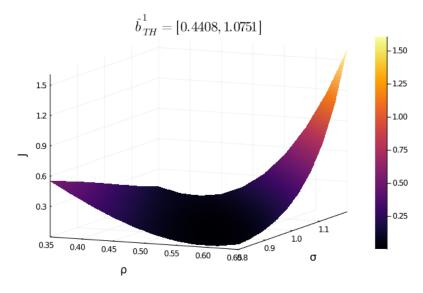


Figure 3: Objective Function for Exercise 5

(b) $\hat{b}_{TH}^2 = [0.4409, 1.075]$ after the Newey-West correction. Figure 4 displays a graph of the objective function.

(c)
$$\nabla_b g_T(\hat{b}_{TH}^2) = \begin{bmatrix} 1.44 & 2.78 \\ 1.83 & 1.22 \end{bmatrix}$$

The variance-covariance matrix of \hat{b}_{TH}^2 is given by

$$\begin{bmatrix} 7.43e - 5 & -1.02e - 3 \\ -1.02e - 3 & 0.01 \end{bmatrix}$$

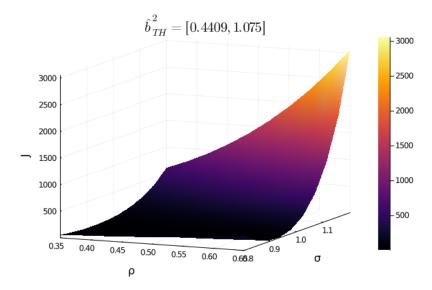


Figure 4: Objective Function for Exercise 5 with the Newey-West Correction

The standard errors are given by

$$\begin{bmatrix} 0.01 \\ 0.12 \end{bmatrix}$$

For local identification, it is useful to have the elements of $\nabla_b g_T(\hat{b}_{TH}^2)$ to be relatively large. This is because if the distance between the modeled moments and the data moments is changing rapidly around the optimal parameter choices, we can be more confident in the precision with which we've identified the local optimal parameter values.

- (d) The J-test value is 2.29e 7.
- 6. Next, we will consider the overidentified case where m₃ uses the mean, variance and autocorrelation. Let's see. For this case, perform steps (a)-(d) above. Furthermore, bootstrap the finite sample distribution of the estimators using the following algorithm:
 - i. Draw ϵ_t and e_t^h from N(0,1) for $t=1,2,\ldots,T$ and $h=1,2,\ldots,H$. Compute $(\hat{b}_{TH}^1),\hat{b}_{TH}^2$ as described.
 - ii. Repeat (e) using another seed.

Answer:

- (a) $\hat{b}_{TH}^1 = [0.4414, 1.0748]$. Figure 5 displays a graph of the objective function.
- (b) $\hat{b}_{TH}^2 = [0.3894, 0.4101]$ after the Newey-West correction. Figure 6 displays a graph of the objective function.

(c)
$$\nabla_b g_T(\hat{b}_{TH}^2) = \begin{bmatrix} 0.02 & 0.03 \\ 0.14 & 0.92 \\ 0.24 & 0.34 \end{bmatrix}$$

The variance-covariance matrix of \hat{b}_{TH}^2 is given by

$$\begin{bmatrix} 3.58e - 3 & 0.02 \\ 0.02 & 0.07 \end{bmatrix}$$

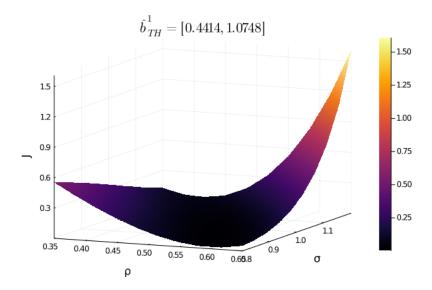


Figure 5: Objective Function for Exercise 6

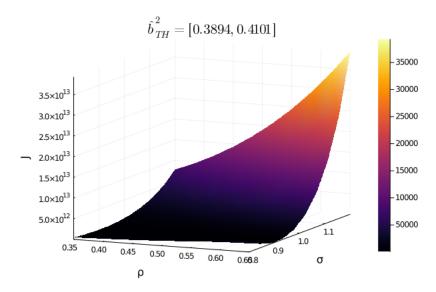


Figure 6: Objective Function for Exercise 6 with the Newey-West Correction

The standard errors are given by

$$\begin{bmatrix} 0.06 \\ 0.27 \end{bmatrix}$$

For local identification, it is useful to have the elements of $\nabla_b g_T(\hat{b}_{TH}^2)$ to be relatively large. This is because if the distance between the modeled moments and the data moments is changing rapidly around the optimal parameter choices, we can be more confident in the precision with which we've identified the local optimal parameter values.

(d) The J-test value is 27.19.

Bootstrapping: Figures 7 and 8 display the bootstrapped densities of parameter estimates for \hat{b}_{TH}^1 and \hat{b}_{TH}^2 , respectively. The density of \hat{b}_{TH}^2 is calculated using only those random seeds which allowed for a solution to be found, as shown in the code.

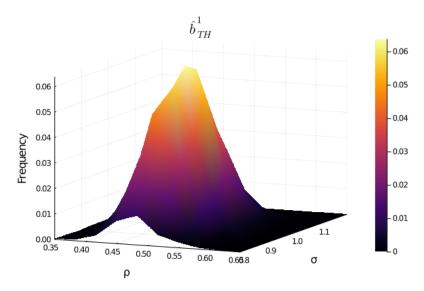


Figure 7: Bootstrapped Density of Estimates for \hat{b}_{TH}^1

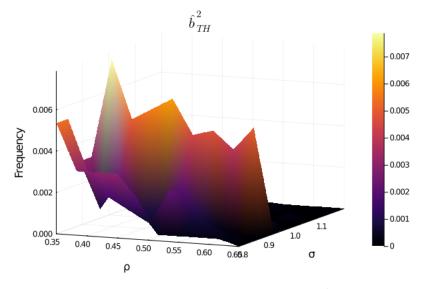


Figure 8: Bootstrapped Density of Estimates for \hat{b}_{TH}^2

Appendix

The first codefile named "RunCode.jl" runs the code.

```
# theme(:juno)
# default(fontfamily="Computer Modern", framestyle=:box)
# Load the Code
   include("./Code.jl");

#Run The Code
   #All exercises are done at once like this since they are supposed to all use
        #the same e draw
   #Note that the results are highly dependent upon the epsilon draw,
   #estpecially for exercise 6. Oftentimes standard errors cannot be
   #calculated. Setting UseRandomSeed to false gives a consistent draw where this
   #is not the case. Refresh Julia after you do this to reset the seed somewhere
   #more truly random
   StepsAThroughD(UseRandomSeed=false)
   #StepsAThroughD(UseRandomSeed=true)
```

The second codefile named "Code.jl" contains the relevant functions.

```
using Parameters, Optim, Distributions, LinearAlgebra, Plots, LaTeXStrings,
     Random
@with_kw mutable struct Primitives
    T ::Int64
H ::Int64
                                #Time Series Length
                               #Number of Simulations
                                = .5

ho0
               ::Float64
    \sigma0
             ::Float64
     \times 0
              ::Float64
                                        = 0
                                       = 10
= collect(range(0.35, length = gpoints, stop = 0.65))
    gpoints :: Int64
     \rhogrid ::Array{Float64}
     \sigmagrid
             ::Array{Float64} = collect(range(0.8, length = gpoints, stop = 1.2))
              ::Int64
end # Primitives
@with_kw mutable struct Results
    \overline{\text{bHat1TH}} ::Array{Float64} = [0,0]
     bHat2TH
                       ::Array{Float64} = [0,0]
                      ::Array{Float64} = zeros(5000,2,2) #Iteration, bhat1 or 2, then each parameter
     b_dist
                      ::Array{Float64}
                                            =100
                      ::Float64
                       ::Array{Float64} = [0, 0] #The true data
     td
end # Primitives
function GetMoments (\rho, \sigma, T, H)
    Random.seed! (100)
     \epsilon = zeros(T, H)
     for hi=1:H
          \epsilon [:, hi]=rand(Normal(0,\sigma), T)
     x_t = zeros(T+1, H)
     for t=2:(T+1)
         \mathbf{x}_t \left[ \mathtt{t}, : \right] \; = \; \rho_0 \; \mathrel{.*} \; \mathbf{x}_t \left[ \mathtt{t-1}, : \right] \; \mathrel{.+} \; \epsilon \left[ \mathtt{t-1}, : \right]
    x_t_1 = x_t[1:end-1,:]
     x_t = x_t [2:end,:]
    \texttt{barx} = \texttt{mean}(\texttt{x}_t, \texttt{dims}{=}1)
    m1 = mean(mean(x_t, dims=1), dims = 2)[1]
     \begin{array}{l} \text{m2 = mean(mean((x_t .- mean(x_t, \text{dims} - 2)[1] \\ m3 = mean((x_t .- \text{barx}) .* (x_t - 1 .- \text{barx}))[1] \\ \end{array} 
    m = [m1, m2, m3]
    return m
end
=#
```

```
function TrueData(prim)
         Qunpack T, \sigma0, x0, \rho0 = prim \epsilon = rand(Normal(0, sqrt(\sigma0)), T)
          xt=zeros(T+1)
          xt[1]=x0
          for t=2:(T+1)
                  xt[t] = \rho 0 * xt[t-1] + \epsilon[t-1]
          end
          xt=xt.[2:end]
          return xt
end
function eDrawsForModel(prim;URS=false)
          if ~URS
                    Random.seed!(9193) #543
          end
         @unpack H, T=prim
          e=zeros(T,H)
          for hi=1:H
                  e[:,hi]=rand(Normal(0,1),T)
          end
          return e
end
function ModelData(prim, e, b)
          @unpack T, \sigma0, x0, \rho0, H = prim
          #if b[2]<0
                      #Here J is set to return a very high value so this will not be picked
                      return zeros(T,H)
          #else
                    yt=zeros(T+1,H)
                    for hi=1:H
                              for t=2:(T+1)
                                         # yt[t,hi]=b[1]*yt[t-1,hi]+sqrt(b[2])*e[t-1,hi]
                                        yt[t,hi]=b[1]*yt[t-1,hi]+b[2]*e[t-1,hi] # Removed sqrt if sigma is being used
                    yt=yt[2:end,:]
                    return yt
end
function FindM2_MeanVar(data)
          M=zeros(2)
          if size(data, 2) > 1 #We are dealing with the Model Data
                    Hmeans=sum(data, dims=1)./size(data, 1)
                    M[1]=sum(Hmeans)/size(data,2)
                    #data.-Hmeans does the subtraction for each row
                    Hvars=sum((data.-Hmeans).^2,dims=1)./size(data,1)
                    M[2]=sum(Hvars)/size(data,2)
          else
                   M[1]=sum(data)/length(data)
                    M[2] = sum((data.-M[1]).^2)/length(data)
         end
          return M
end
function m_MeanVar(x,ind, mean,prim)
          return [x[ind] (x[ind]-mean)^2]
end
function FindM2 VarCoVar(data)
         M=zeros(2)
          if size(data, 2) > 1 #We are dealing with the Model Data
                    Hmeans=sum(data,dims=1)./size(data,1)
                    #data.-Hmeans does the subtraction for each row
                    Hvars=sum((data.-Hmeans).^2,dims=1)./size(data,1)
                    M[1]=sum(Hvars)/size(data,2)
                    \texttt{HCovars} = \texttt{sum} ((\texttt{data[2:end,:]}.-\texttt{Hmeans}). * (\texttt{data[1:(end-1),:]}.-\texttt{Hmeans}), \texttt{dims=1}). / (\texttt{size}(\texttt{data,1})-\texttt{1}) = \texttt{dimpersion} = \texttt{dimpers
                    M[2] = sum (HCovars) / size (data, 2)
          else #We are dealing with the true data
                    Mean=sum(data)/length(data)
                    M[1] = sum((data.-Mean).^2)/length(data)
                    \texttt{M[2]} = \texttt{sum((data[2:end].-Mean).*(data[1:(end-1)].-Mean))/(length(data)-1)}
          end
          return M
end
function m_VarCovar(x,ind,mean,prim)
          Qunpack x0 = prim

if ind -1 == 0
                    return [(x[ind]-mean)^2
                                                                                              (x[ind]-mean)*(x0-mean)]
          else
                    return [(x[ind]-mean)^2
                                                                                              (x[ind]-mean)*(x[ind-1]-mean)]
          end
end
function FindM3(data)
```

```
M=zeros(3)
    if size(data, 2) > 1 #We are dealing with the Model Data
        Hmeans=sum(data,dims=1)./size(data,1)
        M[1]=sum(Hmeans)/size(data,2)
         #data.-Hmeans does the subtraction for each row
        Hvars=sum((data.-Hmeans).^2, dims=1)./size(data,1)
        M[2]=sum(Hvars)/size(data,2)
        HCovars=sum((data[2:end,:].-Hmeans).*(data[1:(end-1),:].-Hmeans),dims=1)./(size(data,1)-1)
        M[3] = sum(HCovars)/size(data, 2)
    else #We are dealing with the true data
        Mean=sum(data)/length(data)
        M[1] = Mean;
        M[2] = sum((data.-Mean).^2)/length(data)
        \texttt{M[3]} = \texttt{sum((data[2:end].-Mean).*(data[1:(end-1)].-Mean))/(length(data)-1)}
    end
    return M
end
function m3(x,ind,mean,prim)
    Qunpack x0 = prim

if ind -1 == 0
        return [x[ind]
                            (x[ind]-mean)^2
                                                   (x[ind]-mean)*(x0-mean)] #We can replace this by x0
    else
        return [x[ind]
                            (x[ind]-mean)^2
                                                   (x[ind]-mean)*(x[ind-1]-mean)]
    end
end
function J(g,W,b) #The objective function
    \mbox{if } \mbox{b[2]<0} #We shouldn't allow there to be a negative variance term
        return 1e25
        return transpose(q) *W*q
    end
end
function GraphAndFindbHat(W,prim,res,FindM,Exercise; NeweyWest=false,Graph=true)
    Qunpack 
hogrid,\sigmagrid,gpoints=prim
    @unpack e, td = res
    Jgrid=zeros(gpoints, gpoints)
    for \rho i=1:gpoints, \sigma i=1:gpoints
        md=ModelData(prim,e, [\rhogrid[\rhoi] \sigmagrid[\sigmai]])
        {\tt Jgrid[\rho i,\sigma i]=J(FindM(td).-FindM(md),W,[\rho grid[\rho i]\ \sigma grid[\sigma i]])}
    Solution=optimize(b->J(FindM(td).-FindM(ModelData(prim,e,b)),W,b),[.3 1.2], NelderMead())
    # Solution=optimize(b->JFindM(td).-FindM(ModelData(prim,e,b)),W,b), # [.3 1.2],[\rhogrid[1]],[\rhogrid[end]] \sigmagrid[end], NelderMead()) #I want to restrict the parameter s
    bHat=Solution.minimizer
    if Graph
      println("The minimizer is ", bHat)
    if NeweyWest
        plot(\rhogrid,\sigmagrid,Jgrid, st=:surface,
            title=L"\hat{b}^{2}_{TH}=[%(round(bHat[1],digits=4)), %(round(bHat[2],digits=4))]", xlabel =
"ρ",
             vlabel = "\sigma", zlabel ="J")
        # savefig("PS7\\Figures\\Exercise$(Exercise)NeweyWestCorrection.png")
        savefig("PS7/Figures/Exercise$(Exercise) NeweyWestCorrection.png")
    else
        plot(\rhogrid,\sigmagrid,Jgrid, st=:surface,
            title=L"\hat{b}^{1}_{TH}=[%$(round(bHat[1],digits=4)) , %$(round(bHat[2],digits=4)) ]", xlabel =
            ylabel = "\sigma", zlabel ="J")
        # savefig("PS7\\Figures\\Exercise$(Exercise).png")
        savefig("PS7/Figures/Exercise$(Exercise).png")
    end
    end #if Graph statement
    return bHat
function NeweyWest(prim::Primitives,simdata::Array{Float64},
        m::Function, MTH::Array{Float64})
    @unpack iT,H,T = prim \#Can either use the simulation-H specific mean
        {\tt Hmeans=sum}\,({\tt simdata},{\tt dims=1})\,./{\tt size}\,({\tt simdata},1)
    #Or can Try using the overall mean:
    \texttt{\#Hmeans=ones(size(simdata,1))*(sum(simdata)/(T*H))}\\ \texttt{\#Defining }\Gamma \texttt{ function}
        function \Gamma jTH(j)
             out=zeros(length(MTH),length(MTH))
             for hi=1:H, ti=(j+1):T
                 \# Necessary to add one twice above because one moment includes a lag ( I removed the twice addition
                  # changed the m function to consider cases)
                  out+=(m(simdata[:,hi],ti,Hmeans[hi],prim).-MTH)*
                    transpose(m(simdata[:,hi],ti-j,Hmeans[hi],prim).-MTH)
```

```
end
              return (1/(T*H)) *out
         end
    #Finding STH
         SyTH=ΓjTH(0)
         for ji=1:iT
              SyTH+=(1-ji/(iT+1))*(\Gamma jTH(ji)+transpose(\Gamma jTH(ji)))
         end
         STH = (1+1/H) *SVTH
     #Returning WStan
         return inv(STH)
end
function Find∇g(res,prim, FindM; s=1e-15)
     @unpack e, bHat2TH = res
     \#\partial g\partial \rho = -(\texttt{FindM}(\texttt{ModelData}(\texttt{prim}, \texttt{e}, \texttt{bHat2TH})). - \texttt{FindM}(\texttt{ModelData}(\texttt{prim}, \texttt{e}, \texttt{bHat2TH} - [\texttt{s} \ 0])))./\texttt{s}
     \#\partial g\partial \sigma = -(\texttt{FindM}\,(\texttt{ModelData}\,(\texttt{prim},\texttt{e},\texttt{bHat2TH})\,)\,. - \texttt{FindM}\,(\texttt{ModelData}\,(\texttt{prim},\texttt{e},\texttt{bHat2TH}-[0\ \texttt{s}]))\,)\,./s
     #Alternate Derivative calculation for (hopefully) improved accuracy from
     #https://en.wikipedia.org/wiki/Numerical_differentiation#
     \partial g \partial \rho = \text{(FindM(ModelData(prim,e,bHat2TH+[s 0])).-FindM(ModelData(prim,e,bHat2TH-[s 0])))./(2*s)}
     \partial g \partial \sigma = (\text{FindM}(\text{ModelData}(\text{prim}, e, b\text{Hat2TH} + [0 s])). - \text{FindM}(\text{ModelData}(\text{prim}, e, b\text{Hat2TH} - [0 s]))). / (2*s))
     return [\partial g \partial \rho \ \partial g \partial \sigma]
end
function StepsAThroughD(;T=200,H=10,UseRandomSeed=false)
    prim=Primitives (T=T, H=H)
     res=Results(e=eDrawsForModel(prim,URS=UseRandomSeed))
     res.td=TrueData(prim)
     for Exercise=4:6
         if Exercise==4
              FindM=FindM2_MeanVar
              m=m_MeanVar
          elseif Exercise==5
              FindM=FindM2_VarCoVar
              m=m_VarCovar
          elseif Exercise==6
              FindM=FindM3
              m=m3
          #Function for parts a and b
          #Part a: Graph in three Dimensions
              res.bHat1TH=GraphAndFindbHat(I,prim,res,FindM,Exercise)
              print("
                         Results for Exercise $(Exercise)
              println("The estimate of b using W = I is ", res.bHat1TH,".")
               # \nabla g1 = Find \nabla g (res, prim, Find M)
                      print("∇g= \n\n")
                      display(∇q1)
               # StdErrorsbHat1TH = sqrt.(diag((1/prim.T)*inv(transpose(\nabla g1)*I*\nabla g1))
               # print("\n The Standard errors are given by \n\n")
                     display(StdErrorsbHat1TH)
          #Part b: Use NeweyWest to update your guess of bHat
              md_bHat1TH=ModelData(prim, res.e, res.bHat1TH)
              WStar=NeweyWest(prim, ModelData(prim, res.e, md_bHat1TH),
                       m, FindM(md bHat1TH))
               res.bHat2TH=GraphAndFindbHat(WStar,prim,res,FindM,Exercise,
                   NeweyWest=true)
              println("The estimate of b using Wstar is ", res.bHat2TH,".")
         #Part c
              \nabla g = \text{Find} \nabla g \text{ (res, prim, FindM)}
                   print("\nabla g = \n\n")
                   display(∇g)
               \label{eq:VarCovarbHat2TH=(1/prim.T)*inv(transpose($\nabla g$)*WStar*$\nabla g$)} \\
                    print("\n The variance-covariance matrix for bHat2TH is given by \n\
                    display(VarCovarbHat2TH)
              try
                    StdErrorsbHat2TH=sqrt.(diag(VarCovarbHat2TH))
                         print("\n The Standard errors are given by \n\n")
                         display(StdErrorsbHat2TH)
              catch
                   print("The Standard Errors cannot be found for this epsilon draw")
              end
          #Part d: Computing the Value of the J test
               res.JTest=prim.T*(prim.H/(1+prim.H))*
                   J(FindM(res.td).-FindM(ModelData(prim,res.e,res.bHat2TH)),
                         WStar, res.bHat2TH)
```

```
println("\n The J-Test is $(res.JTest)")
          #Bootstrapping for Exercise 6
         if Exercise==6
              println("\n\n Beginning Bootstrapping")
              Qunpack \rho \mathrm{grid}, \sigma \mathrm{grid}, \mathrm{gpoints=prim}
              Density=zeros(2, gpoints, gpoints)
              FailedDraws=0
              for iter=1:size(res.b_dist,1)
                   res=Results(e=eDrawsForModel(prim,URS=true))
                    res.td=TrueData(prim)
                    #bHat1TH
                        res.b_dist[iter,1,:]=GraphAndFindbHat(I,prim,res,FindM,Exercise, Graph=false)
                    #bHat2TH
                        md_bHat1TH=ModelData(prim, res.e, res.b_dist[iter, 1,:])
                        try
                             WStar=NeweyWest(prim, ModelData(prim, res.e, md_bHat1TH),
                                       m, FindM(md_bHat1TH))
                             res.b_dist[iter,2,:]=GraphAndFindbHat(WStar,prim,res,FindM,Exercise,
                                  NeweyWest=true, Graph=false)
                        catch
                             res.b_dist[iter,2,:]=[-Inf, -Inf]
                             \texttt{FailedDraws+=} \mathbf{1}
                        end
                   for b_type=1:2
                         \textbf{if} \ \text{res.b\_dist[iter,b\_type,1]} <= \rho \\  \text{grid[1]} \ | | 
                             \texttt{res.b\_dist[iter,b\_type,1]} >= \rho \texttt{grid[gpoints]} \ \mid \mid
                                  res.b_dist[iter,b_type,2]>= \sigmagrid[gpoints] ||
                                  res.b_dist[iter,b_type,2] <= \sigmagrid[1]
                                  #Out of range, do nothing
                             for \rhoi=1:gpoints,\sigmai=1:gpoints
                                   \textbf{if} \ (\rho \texttt{grid}[\rho \texttt{i}+1] \texttt{>=} \texttt{res.b\_dist}[\texttt{iter,b\_type,1}] \texttt{>=} \rho \texttt{grid}[\rho \texttt{i}] \ \& \& \\
                                       \sigmagrid[\sigmai+1]>=res.b_dist[iter,b_type,2]>=\sigmagrid[\sigmai])
                                       Density[b_type,\rhoi,\sigmai]+=1
                                       break
                                  end
                             end
                        end
                   end
                    if iter % 250 ==0
                        print("\n Iteration $(iter) of Bootstrapping. Cumulative failed draws=$(FailedDraws)")
              end #End bootstrapping with iter
              Density[1,:,:] = Density[1,:,:]./size(res.b_dist,1) #Divide by the number of iterations
              Density[2,:,:] = Density[2,:,:]./(size(res.b_dist,1) - FailedDraws)
              for b_type=1:2
                   plot(ρgrid,σgrid,Density[b_type,:,:], st=:surface,
title=L"\hat{b}^{%$(b_type)}_{TH}", xlabel = "ρ",
                        ylabel = "\sigma", zlabel = "Frequency")
                   savefig("PS7/Figures/Exercise$(Exercise)Bootstrapping$(b_type).png")
              end
         end
    end #Exercise Loop
end #End Function StepsAThroughD
```