

1. See the QuadLL2 function in the julia code attached below.
2. For $T_i = 1$, the probability can be written as follows

$$Pr\left(\eta_{i0} < \frac{-\alpha_0 - X_i\beta - Z_{it}\gamma}{\sigma_0}\right) \quad (1)$$

For $T_i = 2$, the probability is

$$Pr(\epsilon_{i0} < \alpha_0 + X_i\beta + Z_{it}\gamma, \epsilon_{i1} < -\alpha_1 - X_i\beta - Z_{it}\gamma) \quad (2)$$

The conditional probability can be written as follows

$$\begin{aligned} & Pr(\epsilon_{i1} < -\alpha_1 - X_i\beta - Z_{it}\gamma | \epsilon_{i0} < \alpha_0 + X_i\beta + Z_{it}\gamma) \\ &= Pr\left(\eta_{i0} + \rho\sigma_0\eta_{i1} < -\alpha_1 - X_i\beta - Z_{it}\gamma | \eta_{i0} < \frac{\alpha_0 + X_i\beta + Z_{it}\gamma}{\sigma_0}\right) \\ &= Pr\left(\eta_{i1} < \frac{-\alpha_1 - X_i\beta - Z_{it}\gamma - \eta_{i0}^*}{\rho\sigma_0} \middle| \eta_{i0} < \frac{\alpha_0 + X_i\beta + Z_{it}\gamma}{\sigma_0}\right) \end{aligned}$$

where η^* is a random variable from $\eta_{i0} < \frac{\alpha_0 + X_i\beta + Z_{it}\gamma}{\sigma_0}$. Then, the probability of $T_i = 2$ is

$$Pr\left(\eta_{i1} < \frac{-\alpha_1 - X_i\beta - Z_{it}\gamma - \eta_{i0}^*}{\rho\sigma_0} \middle| \eta_{i0} < \frac{\alpha_0 + X_i\beta + Z_{it}\gamma}{\sigma_0}\right) Pr\left(\eta_{i0} < \frac{\alpha_0 + X_i\beta + Z_{it}\gamma}{\sigma_0}\right) \quad (3)$$

Similarly, we can define probability in $T_i = 3, 4$. Based on these probabilities, we can define the log likelihood. The function is GHKLL2.

3. See the function AcceptRejectLL()
4. We got the following likelihoods:
 - Quadrature method:-40992
 - GHK method:-60572 (Changes somewhat each time the function is called)
 - Accept/Reject method:-75167 (Changes somewhat each time the function is called)

The calculated likelihood is different under each method, but the order (-1e5) is the same.

5. We got the following result.

$$\begin{aligned} \alpha_0 &= 3.12, \alpha_1 = 0.88, \alpha_2 = 2.29 \\ score_0 &= 0.00, ratespread = -0.25, largeloan = -0.79, mediumloan = -0.39 \\ irefinance &= -0.04, ager = -0.42, cltv = 0.34, dti = 0.67, cu = -0.43 \\ firstmort &= 0.59, iFHA = -0.05, openyear2 = -0.70, openyear3 = 0.07 \\ openyear4 &= 0.12, openyear5 = 0.02 \\ score0 &= 0.30, score1 = -0.16, score2 = 0.35, \rho = 0.57 \end{aligned}$$

Appendix

The first codefile named "runfile.jl" runs the code.

```

#==
    This file conducts the analyses for JF's PS2
==#

using StatFiles, DataFrames, Optim, BenchmarkTools, Latexify, CSV
# We can use BenchmarkTools for better precision. Just need to replace time
# with btime. The runtime of the overall code get longer as btime runs the
# code multiple times to reduce noise

#include("../PS2b/JuliaCode/functions.jl")
include("../functions.jl")

## load the mortgage data and sparse grid weights as a DataFrames (and
## convert weights to matrices
df = DataFrame(StatFiles.load("PS2b/data/Mortgage_performance_data.dta"))
w1 = DataFrame(CSV.File("PS2b/data/KPU_d1_l20.csv")) |> Matrix
w2 = DataFrame(CSV.File("PS2b/data/KPU_d2_l20.csv")) |> Matrix

# Use this if you are loading data from the root folder.
df = DataFrame(StatFiles.load("../data/Mortgage_performance_data.dta"))
w1 = DataFrame(CSV.File("../data/KPU_d1_l20.csv")) |> Matrix
w2 = DataFrame(CSV.File("../data/KPU_d2_l20.csv")) |> Matrix

#df = DataFrame(CSV.File("C:/Users/ryana/OneDrive/Documents/School/PhD Economics/Research/GitHub/ECON-899/PS1b/

## Separate data into independent variable matrices X and Z and
## dependent variable vector Y
X = df[:, [:score_0, :rate_spread, :i_large_loan, :i_medium_loan,
           :i_refinance, :age_r, :cltv, :dti, :cu,
           :first_mort_r, :i_FHA, :i_open_year2,
           :i_open_year3, :i_open_year4, :i_open_year5]] |> Matrix;

Z = df[:, [:score_0, :score_1, :score_2]] |> Matrix;

Y = df[:, :duration]; #|> Matrix

for name in names(df)
    println(name)
end
## 1. Evaluate log-likelihood using the quadrature method
println("See QuadLL() function")

## 2. Evaluate simulated log-likelihood function using GHK
println("see GHKLL() function")

## 3. Evaluate simulated log-likelihood function using accept/reject
println("see AcceptRejectLL() function")

## 4. Compare predicted choice probabilities for each of the above methods
θ₀ = [0, -1, -1, 0 * ones(size(X, 2), 1), 0.3 * ones(size(Z, 2), 1), 0.5]

ll_quad=QuadLL2(Y, X, Z, w1, w2, θ₀)
ll_ghk=GHKLL2(Y, X, Z, θ₀)
ll_ar=AcceptRejectLL(Y, X, Z, θ₀)

## 5. Maximize quadrature log-likelihood function using BFGS
θ₀ = vcat([0, -1, -1], 0 * ones(size(X, 2), 1), 0.3 * ones(size(Z, 2), 1), [0.5])

θ = optimize(t -> -QuadLL2(Y, X, Z, w1, w2,
```

```

[t[1],t[2],t[3],t[4:(3+size(X,2))],
 t[(4+size(X,2)):(3+size(X,2)+size(Z,2))],
 t[(4+size(X,2)+size(Z,2))]],  $\theta_0$ ,
 method = BFGS(), f_tol = 1e-5, g_tol = 1e-5).minimizer

```

The second codefile named "functions.jl" contains the relevant functions.

```

==
This file defines functions used in JF's PS2
==#
using Optim, Distributions, Parameters, LinearAlgebra

# structure of model parameters
mutable struct ModelParameters
     $\alpha_0$ ::Float64
     $\alpha_1$ ::Float64
     $\alpha_2$ ::Float64
     $\beta$ ::Array{Float64}
     $\gamma$ ::Array{Float64}
     $\rho$ ::Float64
end # parameters struct

# Calculate log-likelihood using quadrature method
function QuadLL2(Y, X, Z, W1, W2,  $\theta$ )

    u = W1[:, 1]; w = W1[:, 2]
     $\mu_0$  = W2[:, 1];  $\mu_1$  = W2[:, 2];  $\omega$  = W2[:, 3]

    param = ModelParameters( $\theta$ [1],  $\theta$ [2],  $\theta$ [3],  $\theta$ [4],  $\theta$ [5],  $\theta$ [6])
    @unpack  $\alpha_0$ ,  $\alpha_1$ ,  $\alpha_2$ ,  $\beta$ ,  $\gamma$ ,  $\rho$  = param

    # Calculate  $\sigma_0$  and  $\sigma_0^2$ 
     $\sigma_0^2$  = 1/(1- $\rho$ )^2
     $\sigma_0$  = 1/(1- $\rho$ )

    tmp =  $\alpha_0$  .+ X* $\beta$  .+ Z* $\gamma$ ;
    mp = zeros(size(X,1), size(u,1));

    for i in 1:size(X,1) # For each observation, get range based on domain (0,1) at t = 0
        mp[i,:] = log.(u') .+ tmp[i]
    end

    tmp =  $\alpha_1$  .+ X* $\beta$  .+ Z* $\gamma$ 
    mp1 = zeros(size(X,1), size(u,1));

    for i in 1:size(X,1) # For each observation, get range based on domain (0,1) at t = 1
        mp1[i,:] = log.(u') .+ tmp[i]
    end

    mdp = ones(size(mp,1), size(mp,2));

    for i in 1:size(X,1) # For each observation, get Jacobian 1/u
        mdp[i,:] = mdp[i,:] ./u
    end

    L1 = cdf.(Normal(), (- $\alpha_0$  .- X* $\beta$  .- Z* $\gamma$ )./ $\sigma_0$ )

    density = pdf.(Normal(), mp./ $\sigma_0$ )./ $\sigma_0$ 

    L2 = (cdf.(Normal(), -  $\alpha_1$  .- X* $\beta$  .- Z* $\gamma$  .-  $\rho$  .* mp) .* density .* mdp) * w

    density = pdf.(Normal(), mp1 -  $\rho$ *mp) .* pdf.(Normal(), mp./ $\sigma_0$ ) ./  $\sigma_0$ 

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L3 = (cdf.(Normal(), -  $\alpha_2$  .- X* $\beta$  .- Z* $\gamma$  .-  $\rho$  .* m $\rho$ 1) .* density .* md $\rho$  .* md $\rho$ ) *
w

L4 = 1 .- L1 .- L2 .- L3

ll = 0

for i = 1:size(Y, 1)

    if Y[i] == 1

        ## If the likelihood becomes minus, I evaluate this value as 1e-10.
        if L1[i] < 0
            L1[i]=1e-10
        else
            end

        ll = ll + log(L1[i])

    elseif Y[i] == 2

        if L2[i] < 0
            L2[i]=1e-10
        else
            end

        ll = ll + log(L2[i])

    elseif Y[i] == 3

        if L3[i] < 0
            L3[i]=1e-10
        else
            end

        ll = ll + log(L3[i])

    elseif Y[i] == 4

        if L4[i] < 0
            L4[i]=1e-10
        else
            end

        ll = ll + log(L4[i])

    end

end # for i

return(ll)

end

# Calculate log-likelihood using quadrature method
function QuadLL(Y, X, Z, W1, W2,  $\theta$ )
    # separate weights and nodes from W1 and W2
    u = W1[:, 1]; w = W1[:, 2]
     $\mu_0$  = W2[:, 1];  $\mu_1$  = W2[:, 2];  $\omega$  = W2[:, 3]

    # unpack model parameters
    param = ModelParameters( $\theta$ [1],  $\theta$ [2],  $\theta$ [3],  $\theta$ [4],  $\theta$ [5],  $\theta$ [6])
    @unpack  $\alpha_0$ ,  $\alpha_1$ ,  $\alpha_2$ ,  $\beta$ ,  $\gamma$ ,  $\rho$  = param

    # Calculate  $\sigma_0$  and  $\sigma_0^2$ 
     $\sigma_0^2$  = 1/(1- $\rho$ )^2
     $\sigma_0$  = 1/(1- $\rho$ )

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# map integral bounds to [0, 1] Upper bound (-Infinity,  $\alpha + X\beta + Z\gamma$ )
#  $\rho(u) = \ln(u) + \alpha + X\beta + Z\gamma$ 

b0 = (a, x, z) -> log.(a) .+ ( $\alpha_0 + \text{dot}(x, \beta) + \text{dot}(z, \gamma)$ )
b1 = (a, x, z) -> log.(a) .+ ( $\alpha_1 + \text{dot}(x, \beta) + \text{dot}(z, \gamma)$ )

# per-observation likelihood:

L1 = (x, z) -> log(cdf(Normal(), (- $\alpha_0 - \text{dot}(x, \beta) - \text{dot}(z, \gamma)$ )/ $\sigma_0$ ))
# I think this should just be  $\sigma_0$  not  $\sigma_0^2$ 

# I think the location of the () about  $\sigma_0$  wrong.
#L2 = (x, z) -> log(sum(w.*
#   (cdf.(Normal(), (- $\rho$ )*b0(u, x, z) .- ( $\alpha_1 + \text{dot}(x, \beta) + \text{dot}(z, \gamma)$ ))/ $\sigma_0$ )).*
#   pdf.(Normal(), b0(u./ $\sigma_0$ , x, z)) ./ u))

function L2(x, z)
    out=0
    try
        out=log(sum(w.*
            (cdf.(Normal(), - $\alpha_1 - \text{dot}(x, \beta) - \text{dot}(z, \gamma) - \rho*b_0(u, x, z))$ )).* # Function
            (pdf.(Normal(), b0( $\sigma_0$  .* u, x, z)/ $\sigma_0$ )/ $\sigma_0$ )).* # Density
            (1 ./u))) # Jacobian

    catch #Sometimes parameters will be tried that make the above try to take the log
        #of a negative number. This is an attempted fix for that which just returns something
        #awful in that case so that it won't be picked
        out=log(1e-5)
        print("Attempted a point that does not work.")
    end
    return out
end

function L3(x, z)
    out=0
    try
        out=log(sum( $\omega$ .*(cdf.(Normal(), (- $\rho$ )*b1( $\mu_1$ , x, z) .- ( $\alpha_2 + \text{dot}(x, \beta)$ 
            +  $\text{dot}(z, \gamma)$ ))/ $\sigma_0$ )).*pdf.(Normal(), b0( $\mu_0$ , x, z)/ $\sigma_0$ )).*pdf.(Normal(), b1( $\mu_1$ , x, z) .-
             $\rho*b_0(\mu_0, x, z)) ./ (\mu_0 .* \mu_1)$ ))

    catch
        out=log(1e-5)
        print("Attempted a point that does not work.")
    end
    return out
end

function L4(x, z)
    out=0
    try
        out=log(sum( $\omega$ .*(cdf.(Normal(), -( $\rho$ )*b1( $\mu_1$ , x, z) .+ ( $\alpha_2 + \text{dot}(x, \beta) +$ 
             $\text{dot}(z, \gamma)$ ))/ $\sigma_0$ )).*pdf.(Normal(), b1( $\mu_1$ ./ $\sigma_0$ , x, z)).*pdf.(Normal(), b1( $\mu_1$ , x, z) .-
             $\rho*b_0(\mu_0, x, z)) ./ (\mu_0 .* \mu_1)$ ))

    catch
        out=log(1e-5)
        print("Attempted a point that does not work.")
    end
    return out
end

# calculate the log-likelihood for all observations
ll = 0
for i = 1:size(Y, 1)
    if Y[i] == 1
        ll = ll + L1(X[i, :], Z[i, :])
    elseif Y[i] == 2

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        ll = ll + L2(X[i, :], Z[i, :])
    elseif Y[i] == 3
        ll = ll + L3(X[i, :], Z[i, :])
    elseif Y[i] == 4
        ll = ll + L4(X[i, :], Z[i, :])
    end
end # for i

return ll
end # quadrature log-likelihood function

# Calculate log-likelihood using quadrature method
function GHKLL(Y, X, Z, θ; sims = 100)

    # unpack model parameters
    param = ModelParameters(θ[1], θ[2], θ[3], θ[4], θ[5], θ[6])
    @unpack α₀, α₁, α₂, β, γ, ρ = param
    σ₀ = 1/(1-ρ)

    ll = 0
    for i=1:size(Y, 1)
        ll_i = 1
        ε_draws = zeros(sims, 3)
        if Y[i] > 1 # Need to draw from a distribution which won't make the borrower repay in period 1
            ε_draws[:, 1] = rand.(truncated(Normal(0, σ₀), -Inf, -α₀ - dot(X[i, :],
β) - dot(Z[i, :], γ)), sims)
            elseif Y[i] == 1 # Find the probability that this draw would have occurred
                ll_i = 1 - cdf(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ))/σ₀)
            end
            if Y[i] > 2 # Need to draw from a distribution which won't make the borrower repay in period 2
                ε_draws[:, 2] = [rand(truncated(Normal(0, σ₀), -Inf, -α₀ - dot(X[i, :],
β) - dot(Z[i, :], γ) - ρ*ε_draws[si, 1])) for si = 1:sims]
                elseif Y[i] == 2 # Find the probability that this draw would have occurred
                    ll_i = (1/sims)*cdf(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ))/σ₀)*
                        sum(1 .- cdf.(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ))
.- ρ*ε_draws[:, 1]))
                end
                if Y[i] > 3 # Need to draw from a distribution which won't make the borrower repay in period 3
                    ε_draws[:, 3] = [rand(truncated(Normal(0, σ₀), -Inf, -α₀ - dot(X[i, :],
β) - dot(Z[i, :], γ) - (ρ²)*ε_draws[si, 1] - ρ*ε_draws[si, 2])) for si = 1:sims]
                    # Find the probability that Y[i]==4 would have occurred
                    ll_i = (1/sims)*cdf(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ))/σ₀)*
                        sum((cdf.(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ)) .-
ρ*ε_draws[:, 1])).*
                        cdf.(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ)) .- (ρ²)*ε_draws[:, 1] .-
ρ*ε_draws[:, 2]))
                    elseif Y[i] == 3 # Find the probability that this draw would have occurred
                        ll_i = (1/sims)*cdf(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ))/σ₀)*
                            sum( (cdf.(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ)) .-
ρ*ε_draws[:, 1])).*
                            (1 .- cdf.(Normal(), (-α₀ - dot(X[i, :], β) - dot(Z[i, :], γ)) .-
(ρ²)*ε_draws[:, 1] .- ρ*ε_draws[:, 2])) )
                        end
                    end
                    ll += log(ll_i)
                end
            return ll
        end # GHK log-likelihood function

function GHKLL2(Y, X, Z, θ; sims = 100)
    #ref: http://fmwww.bc.edu/repec/bocode/g/GHK_note.pdf
    param = ModelParameters(θ[1], θ[2], θ[3], θ[4], θ[5], θ[6])
    @unpack α₀, α₁, α₂, β, γ, ρ = param
    σ₀ = 1/(1-ρ)

    ll_store = 0

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    for i=1:size(Y, 1)

        if Y[i] == 1
            ll = cdf(Normal(0, 1), (-alpha_0 - dot(X[i, :], beta) - dot(Z[i, :], gamma))/sigma_0)
        elseif Y[i] == 2
            eta_0 = rand.(truncated(Normal(0, 1), -Inf, (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))/sigma_0), 100)
            ll = mean(cdf(Normal(0, 1), (-alpha_1 - dot(X[i, :], beta) - dot(Z[i, :], gamma)
                .- eta_0)./(rho .* sigma_0)) .*
                cdf(Normal(0, 1), (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))./(sigma_0)))
        elseif Y[i] == 3
            eta_0 = rand.(truncated(Normal(0, 1), -Inf, (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))/sigma_0), 100)
            tmp = ((alpha_1 + dot(X[i, :], beta) + dot(Z[i, :], gamma) .- eta_0)./sigma_0)
            eta_1 = zeros(100)
            for i in 1:100
                eta_1[i] = (rand.(truncated(Normal(0, 1), -Inf, (tmp[i])[1]), 1))[1]
            end
            ll = mean(cdf(Normal(0, 1), (-alpha_2 - dot(X[1, :], beta) - dot(Z[i, :], gamma)
                .- eta_0 .- rho.* eta_1)./((rho.^2) .* sigma_0)) .*
                cdf(Normal(0, 1), (alpha_1 + dot(X[i, :], beta) + dot(Z[i, :], gamma) .- eta_1)./(rho
                .* sigma_0)) .*
                cdf(Normal(0, 1), (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))./(sigma_0)))
        else
            ll1 = cdf(Normal(0, 1), (-alpha_0 - dot(X[i, :], beta) - dot(Z[i, :], gamma))/sigma_0)
            eta_0 = rand.(truncated(Normal(0, 1), -Inf, (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))/sigma_0), 100)
            ll2 = mean(cdf(Normal(0, 1), (-alpha_1 - dot(X[i, :], beta) - dot(Z[i, :], gamma)
                .- eta_0)./(rho .* sigma_0)) .*
                cdf(Normal(0, 1), (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))./(sigma_0)))
            eta_0 = rand.(truncated(Normal(0, 1), -Inf, (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))/sigma_0), 100)
            tmp = ((alpha_1 + dot(X[i, :], beta) + dot(Z[i, :], gamma) .- eta_0)./sigma_0)
            eta_1 = zeros(100)
            for j in 1:100
                eta_1[j] = (rand.(truncated(Normal(0, 1), -Inf, (tmp[j])[1]), 1))[1]
            end
            ll3 = mean(cdf(Normal(0, 1), (-alpha_2 - dot(X[i, :], beta) - dot(Z[i, :], gamma)
                .- eta_0 .- rho.* eta_1)./((rho.^2) .* sigma_0)) .*
                cdf(Normal(0, 1), (alpha_1 + dot(X[i, :], beta) + dot(Z[i, :], gamma) .- eta_1)./(rho
                .* sigma_0)) .*
                cdf(Normal(0, 1), (alpha_0 + dot(X[i, :], beta) + dot(Z[i, :], gamma))./(sigma_0)))

            ll = 1 - ll1 - ll2 - ll3

        end

        ll_store += log(ll)

    end

    return(ll_store)

end

# Calculate log-likelihood using accept-reject method
function AcceptRejectLL(Y, X, Z, theta; sims = 100, k = maximum(Y))

    # unpack model parameters
    param = ModelParameters(theta[1], theta[2], theta[3], theta[4], theta[5], theta[6])
    @unpack alpha_0, alpha_1, alpha_2, beta, gamma, rho = param

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 $\sigma_0 = 1 / (1 - \rho)$ 

l1 = 0 # initialize log-likelihood

# Define index functions for each outcome
I1 = (x, z,  $\epsilon$ ) ->  $\epsilon$  < -( $\alpha_0$  .+ x *  $\beta$  .+ z *  $\gamma$ )
I2 = (x, z,  $\epsilon$ ) -> ( $\epsilon[:, 1]$  < -( $\alpha_0$  .+ x *  $\beta$  .+ z *  $\gamma$ )) .& ( $\epsilon[:, 2]$  < -( $\alpha_1$  .+
x *  $\beta$  .+ z *  $\gamma$ ) .-  $\rho$  *  $\epsilon[:, 1]$ )
I3 = (x, z,  $\epsilon$ ) -> ( $\epsilon[:, 1]$  < -( $\alpha_0$  .+ x *  $\beta$  .+ z *  $\gamma$ )) .& ( $\epsilon[:, 2]$  < -( $\alpha_1$  .+
x *  $\beta$  .+ z *  $\gamma$ ) .-  $\rho$  *  $\epsilon[:, 1]$ ) .& (
 $\epsilon[:, 3]$  < -( $\alpha_1$  .+ x *  $\beta$  .+ z *  $\gamma$ ) .- ( $\rho^2$ ) *  $\epsilon[:, 1]$  .-  $\rho$  *  $\epsilon[:, 2]$ )
I4 = (x, z,  $\epsilon$ ) -> ( $\epsilon[:, 1]$  < -( $\alpha_0$  .+ x *  $\beta$  .+ z *  $\gamma$ )) .& ( $\epsilon[:, 2]$  < -( $\alpha_1$  .+
x *  $\beta$  .+ z *  $\gamma$ ) .-  $\rho$  *  $\epsilon[:, 1]$ ) .& (
 $\epsilon[:, 3]$  <  $\alpha_1$  .+ x *  $\beta$  .+ z *  $\gamma$  .- ( $\rho^2$ ) *  $\epsilon[:, 1]$  .-  $\rho$  *  $\epsilon[:, 2]$ )

# Calculate log-likelihood for Y = 1 observations
x, z = repeat(X[Y==1, :], inner = [sims, 1]), repeat(Z[Y==1, :], inner = [sims, 1])
 $\epsilon$  = rand.(Normal(0,  $\sigma_0$ ), size(x, 1))
for i = 1:sum(Y == 1)
    ind = ((i-1)*sims+1):(i*sims)
    if sum(I1(x[ind, :], z[ind, :],  $\epsilon$ [ind, :]))!=0
        l1 += log(1/sum(I1(x[ind, :], z[ind, :],  $\epsilon$ [ind, :])))
        #l1 += log(sum(I1(x[ind, :], z[ind, :],  $\epsilon$ [ind, :])) / sims)
    else
        l1+=log(1/sims) #Act as though it happened at least once to avoid -infinity
    end
end

# Calculate log-likelihood for Y = 2 observations
x, z = repeat(X[Y==2, :], inner = [sims, 1]), repeat(Z[Y==2, :], inner = [sims, 1])
 $\epsilon$  = [rand.(Normal(0,  $\sigma_0$ ), size(x, 1)) rand.(Normal(), size(x, 1))]
for i = 1:sum(Y == 2)
    ind = ((i-1)*sims+1):(i*sims)
    if sum(I2(x[ind, :], z[ind, :],  $\epsilon$ [ind, :]))!=0
        l1 += log(1/sum(I2(x[ind, :], z[ind, :],  $\epsilon$ [ind, :])))
        #l1 += log(sum(I2(x[ind, :], z[ind, :],  $\epsilon$ [ind, :])) / sims)
    else
        l1+=log(1/sims) #Act as though it happened at least once to avoid -infinity
    end
end

# Calculate log-likelihood for Y = 3 observations
x, z = repeat(X[Y==3, :], inner = [sims, 1]), repeat(Z[Y==3, :], inner = [sims, 1])
 $\epsilon$  = [rand.(Normal(0,  $\sigma_0$ ), size(x, 1)) rand.(Normal(), size(x, 1)) rand.(Normal(), size(x, 1))]
for i = 1:sum(Y == 3)
    ind = ((i-1)*sims+1):(i*sims)
    if sum(I3(x[ind, :], z[ind, :],  $\epsilon$ [ind, :]))!=0
        #l1 += log(1/sum(I3(x[ind, :], z[ind, :],  $\epsilon$ [ind, :]))
        l1 += log(sum(I3(x[ind, :], z[ind, :],  $\epsilon$ [ind, :])) / sims)
    else
        l1+=log(1/sims) #Act as though it happened at least once to avoid -infinity
    end
end

# Calculate log-likelihood for Y = 4 observations
x, z = repeat(X[Y==4, :], inner = [sims, 1]), repeat(Z[Y==4, :], inner = [sims, 1])
 $\epsilon$  = [rand.(Normal(0,  $\sigma_0$ ), size(x, 1)) rand.(Normal(), size(x, 1)) rand.(Normal(), size(x, 1))]
for i = 1:sum(Y == 4)
    ind = ((i-1)*sims+1):(i*sims)
    if sum(I4(x[ind, :], z[ind, :],  $\epsilon$ [ind, :]))!=0
        #l1 += log(1/sum(I4(x[ind, :], z[ind, :],  $\epsilon$ [ind, :]))
        l1 += log(sum(I4(x[ind, :], z[ind, :],  $\epsilon$ [ind, :])) / sims)
    else
        l1+=log(1/sims) #Act like it happened at least once to avoid -infinity
    end
end
return l1
end # accept-reject log-likelihood function

```