# MATH7501 Practical 11 (Week 12), Semester 1-2021

Topic: Probability Distributions

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## Pre-Tutorial Activity

- Students must have familiarised themselves with units 9 and 10 contents of the reading materials for MATH7501

## Resources

- Chapters 9 and 10: Computation of mean, variance, expectation, gradient decent method
- https://en.wikipedia.org/wiki/Rayleigh_distribution

## Section 1: The Rayleigh Distribution

In probability theory and statistics,
the **Rayleigh distribution** is a continuous probability distribution for nonnegative – valued random variables.

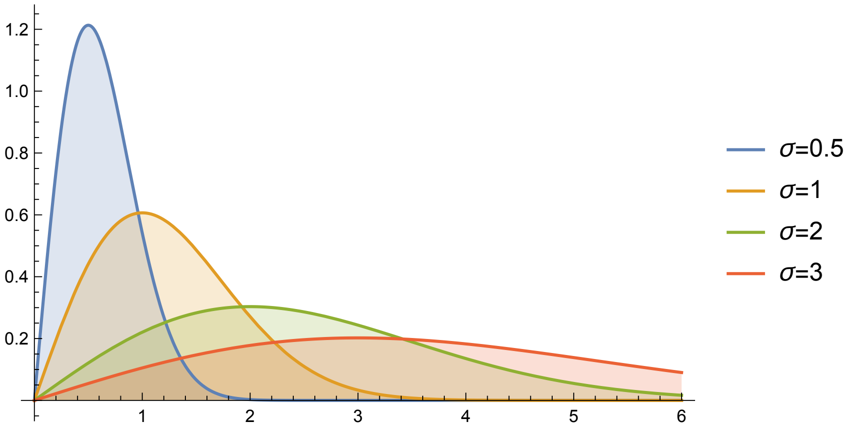
The notation X ~ Rayleigh ($\sigma$) means that the random variable X has a Rayleigh distribution with shape parameter $\sigma$. The probability density function (pdf) is :

$$f_X(x) = \begin{cases} \frac{x}{\sigma^2} e^{\frac{-x^2}{2\sigma^2}}, & x \geq 0, \ \sigma > 0 \\ 0, & x < 0 \end{cases}$$

Q1) Plot the pdf of X for $\sigma = 0.5, 1, 2, 3$

In[ ]:=
```
Plot[Table[PDF[RayleighDistribution[σ], x], {σ, {.5, 1, 2, 3}}] // Evaluate,
 {x, 0, 6}, Filling -> Axis, PlotRange → All,
 PlotLegends → {"σ=0.5", "σ=1", "σ=2", "σ=3"}]
```

Out[ ]=




Q2) Show that the cumulative distribution function (cdf) of X is

$$F_X(x) = \begin{cases} 1 - e^{\frac{-x^2}{2\sigma^2}}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

$$F_X(x) = P(X \leq x) = \int_0^x f_X(t)\,dt$$

$$= \int_0^x \frac{t}{\sigma^2} e^{\frac{-t^2}{2\sigma^2}}\,dt$$

Let $u = \frac{t^2}{2\sigma^2}$. Then $du = \frac{t}{\sigma^2}\,dt$,

and the limits of the integral are $u = 0$ and $u = \frac{x^2}{2\sigma^2}$. Thus,

substituting $dt = \frac{\sigma^2}{t}\,du$ in the above integration, we have,

$$F_X(x) = \int_0^{\frac{x^2}{2\sigma^2}} e^{-u}\,du$$

$$= -\left[e^{-u}\right]_{u=0}^{u=\frac{x^2}{2\sigma^2}}$$

$$= -\left\{e^{-\frac{x^2}{2\sigma^2}} - e^0\right\}$$

$$= \begin{cases} 1 - e^{-\frac{x^2}{2\sigma^2}}, & x \geq 0 \\ 0, & x < 0, \end{cases}$$

as required.

You can also use the Integrate[] function to compute the cdf as shown below.

In[39]:=
```
f[x_] := x/σ^2 E^(-x^2/(2 σ^2)) (*Rayleigh probability density function*)

Integrate[f[x], {x, 0, u}] (*cdf*)
```

Out[ ]= $1 - e^{-\frac{u^2}{2\sigma^2}}$

Q3) Show that $f_X(x)$ is a valid probability density function by showing that the integral over $[0, \infty)$ is 1

$$\int_0^\infty f_X\ (t)\ dt\ =\ \int_0^\infty \frac{x}{\sigma^2}\ e^{\frac{-x^2}{2\sigma^2}}\ dx$$

$$= \text{limit}_{y\to\infty} \int_0^y \frac{x}{\sigma^2}\ e^{\frac{-x^2}{2\sigma^2}}\ dx$$

$= \text{limit}_{y\to\infty} F_X\ (y)$, by the definition of the cdf

$$= \text{limit}_{y\to\infty} \left(1 - e^{-\frac{x^2}{2\sigma^2}}\right)$$

$= 1 - 0$

$= 1$, as required.

You can use the Integrate[] function to compute this integration exactly.

```
In[65]:= Integrate[f[x], {x, 0, ∞}, Assumptions → σ > 0]
Out[•]= 1
```

You can also use NIntegrate[] function to derive a numerical approximation to $\int_0^\infty \frac{x}{\sigma^2}\ e^{\frac{-x^2}{2\sigma^2}}\ dx$, for a given value of $\sigma$ (say for example, $\sigma = 1$).

```
NIntegrate[x E^(-x^2/2), {x, 0, ∞}]
Out[•]= 1.
```

**Q4)** Show that the mean of X ~ Rayleigh ($\sigma$) $= \sigma\sqrt{\frac{\pi}{2}}$

$$\mu_X\ =\ E\ (X)\ =\ \int_0^\infty x f_X\ (x)\ dx$$

$$= \int_0^\infty x\ \frac{x}{\sigma^2}\ e^{\frac{-x^2}{2\sigma^2}}\ dx$$

Let $u = \frac{x^2}{2\sigma^2}$. Then $du = \frac{x}{\sigma^2}\ dx$,

and the limits of the integral are still $u = 0$ and $u = \infty$. Thus,

substituting $dx = \frac{\sigma^2}{x}\ du$ and using $x = \sqrt{(2\sigma^2 u)}$ in the above integration we have,

$$\mu_X\ =\ \int_0^\infty \sqrt{(2\sigma^2 u)}\ e^{-u}\ du$$

$$= \sigma\sqrt{2} \int_0^\infty \sqrt{u}\ e^{-u}\ du$$

$$= \sigma\sqrt{2} \int_0^\infty u^{3/2-1}\ e^{-u}\ du$$

$= \sigma\sqrt{2}\,\Gamma\left(\frac{3}{2}\right)$, $\Gamma\ (x)$ is the Gamma function defined as : $\Gamma\ (x)\ =\ \int_0^\infty t^{x-1}\ e^{-t}\ dt$

$$= \sigma\sqrt{2}\,\Gamma\left(\frac{1}{2} + 1\right)$$

$= \sigma\sqrt{2} \times \frac{1}{2}\,\Gamma\left(\frac{1}{2}\right)$, using the property of Gamma function, $\Gamma\ (x+1)\ = x\Gamma\ (x)$

$= \sigma\sqrt{2} \times \frac{1}{2}\sqrt{\pi}$, using the property of Gamma function, $\Gamma\left(\frac{1}{2}\right)\ = \sqrt{\pi}$

$$= \sigma\sqrt{\left(\frac{\pi}{2}\right)}$$

You can use the Integrate[] function to check your derivation of the mean as follows.

In[•]:=
```
Integrate[x f[x], {x, 0, ∞}, Assumptions → σ > 0] (*mean (μ) *)
```
Out[•]= $\sqrt{\frac{\pi}{2}}\ \sigma$

In[•]:=
```
(*using numerical approximation to compute the mean when σ =1*)
```

In[•]:=
```
NIntegrate[x x E^(-x^2/2), {x, 0, ∞}]
```
Out[•]= 1.25331

```
δ = 0.01
```
Out[•]= 0.01

In[•]:=
```
Total[Table[x x E^(-x^2/2) δ, {x, 0, 10, δ}]]
```
Out[•]= 1.25331

**Q5)** Show that the variance of $X\sim$ Rayleigh $(\sigma)$ $=\sigma^2\left(\frac{4-\pi}{2}\right)$

From Week 11 prctical, we know that $\sigma^2{}_X = \text{var}\,(X) = E[X^2] - \mu_X{}^2$, where $E[X^2]$ is the second moment of X and $\mu_X$ is the mean of X, which we computed in Q4. Thus, it remains to find an expression for $E[X^2]$ . We will require the integration bi - parts formula $\int u\,dv = uv - \int v\,du$ for this calculation.

$$E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x)\,dx$$
$$= \int_0^{\infty} x^2 \frac{x}{\sigma^2} e^{\frac{-x^2}{2\sigma^2}}\,dx$$
$$= \left[x^2 \int_0^{\infty} \frac{x}{\sigma^2} e^{\frac{-x^2}{2\sigma^2}}\,dx\right]_{x=0}^{x=\infty} - \int_0^{\infty} 2\,x \left(\int_0^{\infty} \frac{x}{\sigma^2} e^{\frac{-x^2}{2\sigma^2}}\,dx\right)dx,$$

by using $u = x^2$ and $dv = \frac{x}{\sigma^2} e^{\frac{-x^2}{2\sigma^2}}$,

$$= \left[-x^2 e^{\frac{-x^2}{2\sigma^2}}\right]_{x=0}^{x=\infty} + \int_0^{\infty} 2\,x\,e^{\frac{-x^2}{2\sigma^2}}\,dx$$

To evaluate the integral, let $u = \frac{x^2}{2\sigma^2}$. Then $du = \frac{x}{\sigma^2}dx$, and the limits of the integral are still $u = 0$ and $u = \infty$. Thus, substituting $dx = \frac{\sigma^2}{x}du$ we have

$$E[X^2] = 0 + 2\sigma^2\left[\int_0^{\infty} e^{-u}\,du\right]$$
$$= -2\sigma^2\left[e^{-u}\right]_{u=0}^{u=\infty}$$
$$= -2\sigma^2[0 - 1]$$
$$= 2\sigma^2$$

Thus, $\sigma^2{}_X = \text{var}\,(X) = E[X^2] - \mu_X{}^2$
$$= 2\sigma^2 - \left(\sigma\sqrt{\left(\frac{\pi}{2}\right)}\right)^2$$
$$= \sigma^2\left(\frac{4-\pi}{2}\right) \text{ as required.}$$

You can use the Integrate[] function to check your work.

```
Integrate[x^2 f[x], {x, 0, ∞}, Assumptions → σ > 0]
(*compute second moment (E[X^2]*)
```

Out[•]= $2\,\sigma^2$

```
(*using numerical approximation to compute the second moment when σ =1*)
```

In[•]:= $\text{NIntegrate}\left[x^2\, x\, E^{-\frac{x^2}{2}}, \{x, 0, \infty\}\right]$

Out[•]= 2.

In[•]:= $\text{Total}\left[\text{Table}\left[x^2\, x\, E^{-\frac{x^2}{2}}\, \delta, \{x, 0, 10, \delta\}\right]\right]$

Out[•]= 2.

In[•]:= (*variance*)

$$E[X^2] - \mu^2 = 2\,\sigma^2 - \left(\sqrt{\frac{\pi}{2}}\,\sigma\right)^2$$

Out[•]= $2\,\sigma^2 - \frac{\pi\,\sigma^2}{2}$

**Q6)** Find the median of X.

Note that the median of X is the number M such that,

$$\int_0^M f_X\,(x)\,dx = \frac{1}{2}$$

Note that the left hand side is the cdf of X evaluated at M. Thus we have

$$F_X\,(M) = \frac{1}{2}$$

$$1 - e^{-\frac{M^2}{2\sigma^2}} = \frac{1}{2}$$

$$e^{-\frac{M^2}{2\sigma^2}} = \frac{1}{2}$$

$$-\frac{M^2}{2\,\sigma^2} = \ln\left(\frac{1}{2}\right)$$

$$M^2 = -2\,\sigma^2 \ln\left(\frac{1}{2}\right)$$

$M = \sigma\sqrt{\left(-2\ln\left(\frac{1}{2}\right)\right)}$ as the median.

Q7) The quantile function of the distribution, $q\,(u)$ for $u \in [0, 1)$, is defined as follows: For each u, we should have,

$$\int_0^{q\,(u)} f_X\,(x)\,dx = u$$

a) Find an expression for $q\,(u)$

By noting that the left -
hand side of the equation is the cdf of X evaluated at $q\,(u)$, we have

$$F_X(q(u)) = u$$
$$1 - e^{-\frac{q(u)^2}{2\sigma^2}} = u$$
$$e^{-\frac{q(u)^2}{2\sigma^2}} = 1 - u$$
$$-\frac{q(u)^2}{2\sigma^2} = \ln(1 - u)$$
$$q(u)^2 = -2\sigma^2 \ln(1 - u)$$
$$q(u) = \sigma\sqrt{(-2\ln(1 - u))}$$

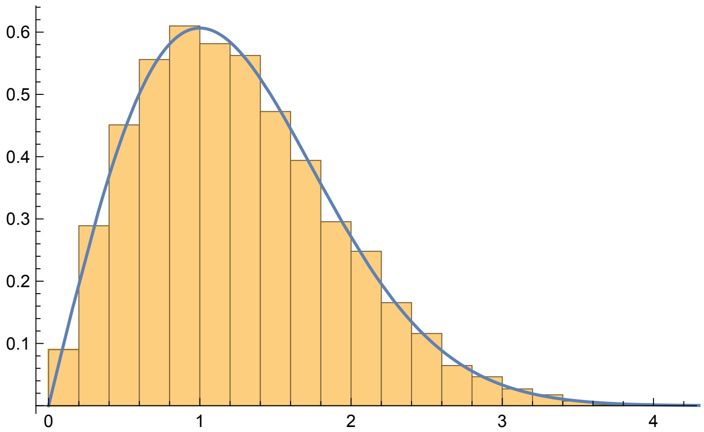
b) Say that X = q (U) with U as Uniformly distributed on [0, 1]. Then, X has a Rayleigh distribution. Show this empirically for $\sigma = 1$, by generating $10^4$ uniform random variables on [0, 1]

```
(*Generate a Rayleigh random Variable from Uniform[0,1] random variable*)
```

In[189]:=
```
U = RandomReal[1, 10000];
X = Sqrt[-2 Log[1 - U]];
```

In[194]:=
```
Show[Histogram[X, Automatic, "PDF"],
 Plot[PDF[RayleighDistribution[1], x],
  {x, 0, 8}, PlotRange → All]]
```

Out[194]=




The plot shows histogram of the data X along with the pdf of X ~ RayleighDistribution (1).

## Section 2: Simple Linear Regression Problem

Consider the simple linear regression problem with data points $(x_1, y_1)$, ..., $(x_n, y_n)$. The aim is to seek $\beta_0$ and $\beta_1$ to fit the line,

$$y = \beta_0 + \beta_1 x,$$

by minimizing the loss function

$$L(\beta_0\ \beta_1) = \sum_{i=1}^{n} (y_i - (\beta_0 + \beta_1 x_i))^2 .$$

Here, $\beta_0$ and $\beta_1$ are the slope and the intercept, respectively, of the line of best fit.

Q1) Compute an expression for the gradient of $L(\beta_0\ \beta_1)$

**Since the loss function has two variables, we need to use partial derivatives here.**

$$\begin{aligned}\frac{\partial L(\beta_0\, \beta_1)}{\partial \beta_1} &= \frac{\partial}{\partial \beta_1}\sum_{i=1}^{n}(y_i - (\beta_0 + \beta_1 x_i))^2 \\ &= \sum_{i=1}^{n}\frac{\partial}{\partial \beta_1}(y_i - (\beta_0 + \beta_1 x_i))^2 \\ &= -2\sum_{i=1}^{n} x_i\,(y_i - (\beta_0 + \beta_1 x_i))\end{aligned}$$

$$\begin{aligned}\frac{\partial L(\beta_0\, \beta_1)}{\partial \beta_0} &= \frac{\partial}{\partial \beta_0}\sum_{i=1}^{n}(y_i - (\beta_0 + \beta_1 x_i))^2 \\ &= \sum_{i=1}^{n}\frac{\partial}{\partial \beta_0}(y_i - (\beta_0 + \beta_1 x_i))^2 \\ &= -2\sum_{i=1}^{n}(y_i - (\beta_0 + \beta_1 x_i))\end{aligned}$$