# The Relationship Between Commodity Prices, Their Futures Trading Volumes, and Average Global Temperature

#### By York Wang

It is not implausible that the trading volumes of a commodity's futures contracts would be influenced by prior movements in the commodity's price itself. As futures contracts are utilized in hedging, it would make sense for commodity consumers to protect against potential upswings in commodity prices. Although there countless numbers of factors that can influence a commodity's prices and the trading volume of its futures, I wanted to see if there was a direct relationship between the two.

Specifically, I will be looking at agricultural commodities like cotton and wheat. In addition to examining the relationship between commodities prices and the trading volumes of their respective futures, I am also looking to see whether the global average temperature has any effect on the trading volumes of commodity futures, as one would expect the risk of climate change to be hedged against through futures contracts — which is why I've limited the commodities I'm looking at to agricultural ones.

The data I will be using will consist of:

- Monthly average global temperature deviations from the 1951-1980 mean
- Monthly trading volumes of agricultural commodities futures
- Monthly global prices for specified commodities

```
In [1]: # Packages needed to run

import pandas as pd
# DataFrame package

import matplotlib.pyplot as plt
# Plotting our data

import statsmodels.formula.api as sm
# Used for simple linear regressions

import seaborn as sns #statistical graphs
```

# I. Importing and Cleaning the Data

# i. Importing and Cleaning Global Temperature Data

I am importing global temperature data from Datahub.io, a site where certain datasets are already curated and cleaned up for public use.

For this data, I will need to isolate the proper data source and restrict the time frame to match that of my commodities data.

#### Out[2]:

	Source	Date	Mean
0	GCAG	2016-12-15	0.7895
1	GISTEMP	2016-12-15	0.8100
2	GCAG	2016-11-15	0.7504
3	GISTEMP	2016-11-15	0.9300
4	GCAG	2016-10-15	0.7292

```
In [3]: grouped data = raw temp.groupby("Source")
        # There is data from two different sources for each month
        # We try to utilize only data from "GISTEMP", which is a NASA source
        temp = grouped data.get group("GISTEMP")
        # We create a new dataframe with only "GISTEMP" data
        temp = temp.drop(temp.index[144:])
        # Since our historical data for futures trading volumes only goes back
        # to January 2005, we drop all data from before January 2005.
        temp = temp.reindex(index = temp.index[::-1])
        # We want to flip the data around so that it's in chronological order.
        temp = temp.drop("Source", axis = 1)
        # We no longer need the data source
        temp = temp.rename(columns = {"Mean": "MeanTemperatureDeviation"})
        df = temp
        # We'll use this as our main dataframe
        df = df.reset index()
        # Resetting our index so that it's no longer funky
        df = df.drop("index", axis = 1)
        # Dropping the old index
        df = df.reset index()
        # df.head()
        # Input to see if our DataFrame looks the way we want it to
```

# ii. Importing and Cleaning Futures Trading Volume Data

This data is pulled by hand from <u>Intercontinental Exchange (https://www.theice.com/marketdata/reports/8)</u>. I have cleaned the data up and uploaded it as a CSV to my Github.

```
In [4]: url2 = "https://raw.githubusercontent.com/yorktruewang/data-bootcamp/m
        aster/Final%20Project/Commodity%20Futures%20Volumes.csv"
        # Reading in our monthy futures trading volume historical data
        f v = pd.read csv(url2)
        f v.head()
        # Checking to see if the starting month is January 2005
        f v = f v.rename(columns = {"Unnamed: 0": "Date"})
        f v.tail()
        # Checking to see if the ending month is December 2016
        # We need to drop the last 16 rows of our dataframe to match the dates
        f v = f v.drop(f v.index[144:])
        # We should have the same number of rows, so we can use 144 as well
        f v = f v.drop(axis = 1, columns = "PULP")
        # Dropping "PULP" because its data ends somewhere within the time span
        # f v.tail()
        # Enter this command to see that our time span matches
In [5]: # Manually adding trading volumes to the centralized dataframe
        coffee = f v["COFFEE"]
        df["coffee"] = coffee
        sugar = f_v["SUGAR"]
        df["sugar"] = sugar
        cotton = f v["COTTON"]
        df["cotton"] = cotton
        fcoj = f v["FCOJ"]
        df["fcoj"] = fcoj
```

cocoa = f\_v["COCOA"]
df["cocoa"] = cocoa

mg = f\_v["METALS & GRAINS"]
df["grains proxy"] = mg

# iii. Importing and Cleaning Global Monthly Prices for Select Commodities

The commodities we've chosen are sugar, cocoa, cotton, and grains. We are limited in the number of commodities due to a lack of publicly-available monthly data on futures trading volumes and a lack of global monthly price data on other commodities (coffee, frozen concentrated orange juice).

Here, we pull the data from the IMF. Because it is all organized in the same way, we can write a function to quickly clean the data for us.

In addition, I've created individual DataFrames for each commodity so that the main DataFrame is not too populated.

```
In [6]: # Our global price indices are all from the IMF, and thus have the sam
        e file structure
        # Let's automate the process of cleaning the data
        WHEAT = "https://raw.githubusercontent.com/yorktruewang/data-bootcamp/m
        aster/Final%20Project/WHEAT%20INDEX.csv"
        COCOA = "https://raw.githubusercontent.com/yorktruewang/data-bootcamp/
        master/Final%20Project/COCOA%20INDEX.csv"
        SUGAR = "https://raw.githubusercontent.com/yorktruewang/data-bootcamp/
        master/Final%20Project/SUGAR%20INDEX.csv"
        COTTON = "https://raw.githubusercontent.com/yorktruewang/data-bootcamp
        /master/Final%20Project/COTTON%20INDEX.csv"
        def clean(url, name):
            placeholder = pd.read csv(url)
            # Import data
            placeholder.columns = ["Date", name]
            # Rename columns
            placeholder = placeholder.drop(placeholder.index[0:300])
            # Strip data before our range
            placeholder = placeholder.drop(placeholder.index[144:449])
            # Strip data after our range
            placeholder = placeholder.reset index()
            # Reset the index so we can add the values to df
            placeholder = placeholder.drop("index", axis = 1)
            # Drop the old index
            df[name] = placeholder[name]
            # Place into df
        # Manually clean the price indices
        clean(SUGAR, "sugarindex")
        clean(COTTON, "cottonindex")
        clean(COCOA, "cocoaindex")
        # Manually create new dataframes for each commodity
        sugardf = pd.DataFrame({"sugar":df["sugar"], "sugarindex": df["sugarin
        dex"]})
        cottondf = pd.DataFrame({"cotton":df["cotton"], "cottonindex": df["cot
        tonindex" | } )
        cocoadf = pd.DataFrame({"cocoa":df["cocoa"], "cocoaindex": df["cocoain
        dex"]})
```

Here we continue cleaning the data. This next section of code comprises of a function written to calculate the percentage change in commodity prices or futures trading volume from the previous month.

The other function written here is one that creates a column of "lagging" data in the DataFrame. In the case that markets aren't truly efficient, we take into consideration the possibility of futures contracts being bought one month after the commodity prices have begun to rise.

```
In [7]: def percentage calc(table, column, security):
            futures change = [0]
            for x in range(143):
                y = x + 1
                value = (table.iloc[y][column]/table.iloc[x][column])-1
                futures_change.append(value)
            table[security] = futures change
        percentage calc(sugardf, "sugar", "futures")
        percentage_calc(cottondf, "cotton", "futures")
        percentage calc(cocoadf, "cocoa", "futures")
        percentage_calc(sugardf, "sugarindex", "price")
        percentage calc(cottondf, "cottonindex", "price")
        percentage_calc(cocoadf, "cocoaindex", "price")
        def lagging_calc(table, column, security):
            futures change = [0,0]
            for x in range(142):
                y = x + 1
                value = (table.iloc[y][column]/table.iloc[x][column])-1
                futures change.append(value)
            table[security] = futures_change
        lagging_calc(sugardf, "sugar", "lagging_futures")
        lagging_calc(cottondf, "cotton", "lagging_futures")
        lagging calc(cocoadf, "cocoa", "lagging futures")
        lagging calc(sugardf, "sugarindex", "lagging price")
        lagging calc(cottondf, "cottonindex", "lagging_price")
        lagging_calc(cocoadf, "cocoaindex", "lagging_price")
        # cottondf.head()
        # Input code to see if a sample DataFrame that we created
```

# **II. Identifying Possible Correlations between Prices and Futures Trading Volume**

Refer to section iv. for discussion of the correlations results.

# i. Sugar

```
In [8]: fig, (ax1, ax2) = plt.subplots(1, 2, sharey = True, figsize = (12,5))
    fig.suptitle("Movements in Sugar's Price and Futures Trading Volume")

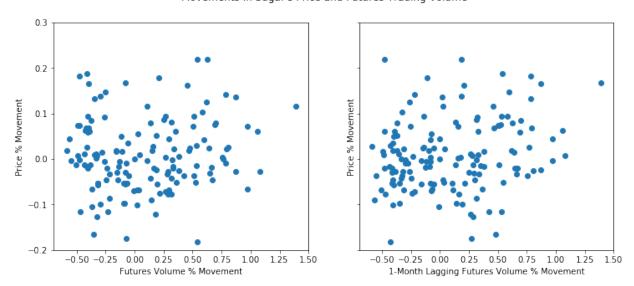
ax1.scatter(sugardf["futures"], sugardf["price"])

ax1.set_xlabel('Futures Volume % Movement')
    ax1.set_ylabel('Price % Movement')
    ax1.set_ylim(-0.2,0.3)

ax2.scatter(sugardf["lagging_futures"], sugardf["price"])

ax2.set_xlabel('1-Month Lagging Futures Volume % Movement')
    ax2.set_ylabel('Price % Movement')
```

Movements in Sugar's Price and Futures Trading Volume



#### a. Regression of Price Movements on Futures Trading Volume Movements

```
In [9]: mod = sm.ols('price ~ futures', data=sugardf)
    res = mod.fit()
    print(res.summary())
```

	OLS Regression Results						
=======================================	:=======		======		=======		
Dep. Variable:		price	R-sq	uared:			
0.011		_	_				
Model:		OLS	Adj.	R-squared:			
0.004							
Method:		Least Squares	F-sta	atistic:			
1.634	m	15 Mars 2010	Dwoh	/E atatiatia).			
Date: 0.203	Tue	e, 15 May 2018	PLOD	(F-statistic):			
Time:		00:11:20	I₁oα−1	Likelihood:			
163.15		00122120					
No. Observatio	ns:	144	AIC:				
-322.3							
Df Residuals:		142	BIC:				
-316.4		_					
Df Model:		1					
Covariance Typ	e: 	nonrobust					
========							
	coef	std err	t	P> t	[0.025		
0.975]					-		
	0.0065	0.007	0 070	0 224	0 007		
0.020	0.0065	0.007	0.970	0.334	-0.007		
	0.0193	0.015	1.278	0.203	-0.011		
0.049	0.0193	0.013	1.270	0.203	0.011		
	=======		======		=======		
=======							
Omnibus:		4.234	Durb	in-Watson:			
1.378							
<pre>Prob(Omnibus): 3.874</pre>		0.120	Jarqi	ue-Bera (JB):			
Skew:		0.396	Proh	(.TR) •			
Skew: 0.396 Prob(JB): 0.144							
Kurtosis: 3.130 Cond. No.							
2.33							
=========	=======		======		=======		
=======							

#### Warnings:

b. Regression of Price Movements on 1-Month Lagging Futures Trading Volume Movements

```
In [10]: mod = sm.ols('futures ~ lagging_price', data=sugardf)
    res = mod.fit()
    print(res.summary())
```

OLS Regression Results					
=======					
Dep. Variable: 0.024		futures	R-squared	:	
Model:		OLS	Adj. R-sq	uared:	
0.017					
Method: 3.463	L€	east Squares	F-statist	ic:	
Date:	Tue,	15 May 2018	Prob (F-s	tatistic):	
0.0648	·	-	·	ŕ	
Time: -82.244		00:11:20	Log-Likel	ihood:	
No. Observations	:	144	AIC:		
168.5					
Df Residuals: 174.4		142	BIC:		
Df Model:		1			
Covariance Type:		nonrobust			
	=======		=======	=======	======
	coef	std err	t	P> t	[0.02
5 0.975]					
Intercept	0.0975	0.036	2.695	0.008	0.02
6 0.169					
<pre>lagging_price 1 0.053</pre>	-0.8590	0.462	-1.861	0.065	-1.77
========	=======	-=======	=======	=======	======
Omnibus:		7.746	Durbin-Wa	tson:	
2.881 Prob(Omnibus):		0.021	Jarque-Be	ra (JR)•	
6.908		0.021	ourque be	1a (5b):	
Skew:		0.462	Prob(JB):		
0.0316 Kurtosis:		2.454	Cond. No.		
12.8		2.131	cona. No.		
		-========	=======	=======	======
========					

#### Warnings:

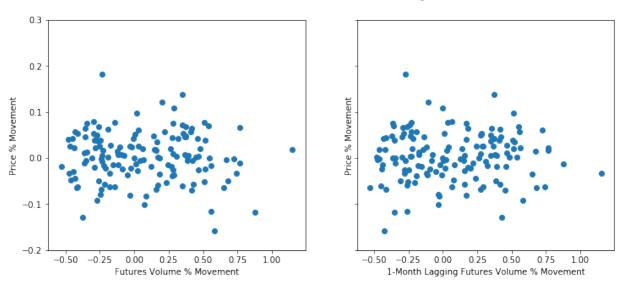
# ii. Cocoa

```
In [11]: fig, (ax1, ax2) = plt.subplots(1, 2, sharey = True, figsize = (12,5))
fig.suptitle("Movements in Cocoa's Price and Futures Trading Volume")

ax1.scatter(cocoadf["futures"], cocoadf["price"])
ax1.set_xlabel('Futures Volume % Movement')
ax1.set_ylabel('Price % Movement')
ax1.set_ylim(-0.2,0.3)

ax2.scatter(cocoadf["lagging_futures"], cocoadf["price"])
ax2.set_xlabel('1-Month Lagging Futures Volume % Movement')
ax2.set_ylabel('Price % Movement')
plt.show()
```

#### Movements in Cocoa's Price and Futures Trading Volume



#### a. Regression of Price Movements on Futures Trading Volume Movements

```
In [12]: mod = sm.ols('price ~ futures', data=cocoadf)
  res = mod.fit()
  print(res.summary())
```

#### OLS Regression Results

\_\_\_\_\_

========						
Dep. Variable	•	pri	ce	R-sq	uared:	
0.002						
Model:		01	LS	Adj.	R-squared:	
-0.005						
Method:		Least Square	es	F-st	atistic:	
0.3419	_	15 00	1.0	_ ,	·	
Date:	Tue	e, 15 May 20	18	Prob	(F-statistic):	
0.560 Time:		00:11:	2.0	T 0 0 1	Likelihood:	
218.49		00:11:	20	rog-1	Likelinood:	
No. Observation	ong•	1.	44	AIC:		
-433.0	JIIS •	1.	77	AIC.		
Df Residuals:		1.	42	BIC:		
-427.0		_				
Df Model:			1			
Covariance Ty	pe:	nonrobu	st			
==========			====:	=====		
=======						
	coef	std err		t	P> t	[0.025
0.975]						
Intercept	0 0046	0 005	1	025	0.307	-0.004
0.014	0.0040	0.003	1	• 023	0.507	-0.004
	-0.0073	0.013	-0	.585	0.560	-0.032
0.017		0.020	·			0000
			====	=====		
=======						
Omnibus:		2.4	63	Durb	in-Watson:	
1.470						
Prob(Omnibus)	:	0.29	92	Jarqı	ue-Bera (JB):	
2.346						
Skew:		-0.0	42	Prob	(JB):	
0.309		2 6	2.0	a 1	NT -	
<pre>Kurtosis: 2.83</pre>		3.6	<b>2</b> U	Cond	• NO •	
	=====	=====				
========						

#### Warnings:

# b. Regression of Price Movements on 1-Month Lagging Futures Trading Volume Movements

<sup>[1]</sup> Standard Errors assume that the covariance matrix of the errors is correctly specified.

```
In [13]: mod = sm.ols('futures ~ lagging_price', data=cocoadf)
    res = mod.fit()
    print(res.summary())
```

OLS Regression Results					
=======					
Dep. Variable: 0.008		futures	R-squared	:	
Model:		OLS	Adj. R-squared:		
0.001			, ,		
Method:	Le	east Squares	F-statist	ic:	
1.107 Date:	<b>Т</b> 11 <b>0</b>	15 May 2018	Prob (F-s	tatistic).	
0.295	140,	15 11dy 2010	1100 (1 5	caciscie).	
Time:		00:11:21	Log-Likel	ihood:	
-54.770		1 4 4	n T C		
No. Observations: 113.5		144	AIC:		
Df Residuals:		142	BIC:		
119.5					
Df Model:		1			
Covariance Type:	=======		=======	=======	======
========					
5 0.975]	coef	std err	t	P> t	[0.02
Intercept	0.0655	0.030	2.197	0.030	0.00
7 0.124 lagging price	0.5933	0.564	1.052	0.295	-0.52
1 1.708	0.3733	0.304	1.032	0.233	0.32
=======================================	=======		=======	=======	======
Omnibus:		6.459	Durbin-Wa	tson:	
Prob(Omnibus):		0.040	Jarque-Be	ra (JB):	
5.166			-	, ,	
Skew:		0.360	Prob(JB):		
0.0756 Kurtosis:		2.415	Cond. No.		
19.0		2.413	CO114. 140.		
=======================================	=======		=======	=======	======
=======					

#### Warnings:

# iii. Cotton

```
In [14]: fig, (ax1, ax2) = plt.subplots(1, 2, sharey = True, figsize = (12,5))
    fig.suptitle("Movements in Cotton's Price and Futures Trading Volume")

ax1.scatter(cottondf["futures"], cottondf["price"])

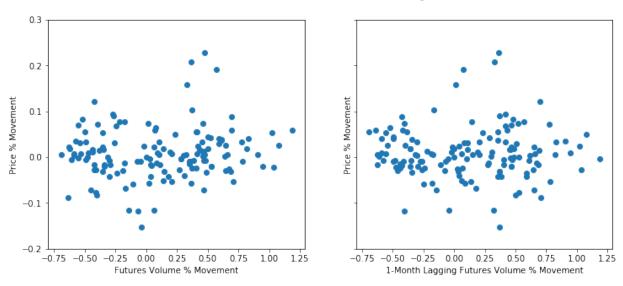
ax1.set_xlabel('Futures Volume % Movement')
    ax1.set_ylabel('Price % Movement')
    ax1.set_ylim(-0.2,0.3)

ax2.scatter(cottondf["lagging_futures"], cottondf["price"])

ax2.set_xlabel('1-Month Lagging Futures Volume % Movement')
    ax2.set_ylabel('Price % Movement')

plt.show()
```

#### Movements in Cotton's Price and Futures Trading Volume



#### a. Regression of Price Movements on Futures Trading Volume Movements

```
In [15]: mod = sm.ols('price ~ futures', data=cottondf)
  res = mod.fit()
  print(res.summary())
```

#### OLS Regression Results

=========	=======	==========		:=====================================	=======
========					
Dep. Variable:		price	R-sqı	uared:	
0.021		-	-		
Model:		OLS	Adj.	R-squared:	
0.014			,	-	
Method:		Least Squares	F-sta	atistic:	
3.041		-			
Date:	Tue	e, 15 May 2018	Prob	(F-statistic):	
0.0834		· -		,	
Time:		00:11:21	Log-l	Likelihood:	
196.93			_		
No. Observatio	ns:	144	AIC:		
-389.9					
Df Residuals:		142	BIC:		
-383.9					
Df Model:		1			
Covariance Typ	e <b>:</b>	nonrobust			
=========	=======		======		=======
=======					
	coef	std err	t	P> t	[0.025
0.975]					
=	0.0029	0.005	0.550	0.583	-0.008
0.013					
	0.0199	0.011	1.744	0.083	-0.003
0.042					
=========	=======		======		=======
========					
Omnibus:		17.955	Durb	in-Watson:	
1.138					
Prob(Omnibus):		0.000	Jarqı	ue-Bera (JB):	
75.716					
Skew:		-0.084	Prob	(JB):	
3.62e-17					
Kurtosis:		6.548	Cond	. No.	
2.23					
=========	=======	=========	======		=======
=======					

#### Warnings:

# b. Regression of Price Movements on 1-Month Lagging Futures Trading Volume Movements

<sup>[1]</sup> Standard Errors assume that the covariance matrix of the errors is correctly specified.

```
In [16]: mod = sm.ols('futures ~ lagging_price', data=cottondf)
    res = mod.fit()
    print(res.summary())
```

OLS Regression Results					
=======================================	=======	=======	========	:=======	:======
Dep. Variable: 0.001		futures	R-squared:		
Model: -0.006		OLS	Adj. R-squ	ared:	
Method: 0.1709	Le	east Squares	F-statisti	.c:	
Date: 0.680	Tue,	15 May 2018	Prob (F-st	atistic):	
Time: -90.653		00:11:22	Log-Likeli	hood:	
No. Observations:	:	144	AIC:		
Df Residuals:		142	BIC:		
Df Model: Covariance Type:		1 nonrobust			
				:=======	======
5 0.975]	coef	std err	t	P> t	[0.02
Intercept 1 0.182	0.1069	0.038	2.796	0.006	0.03
<pre>lagging_price 2     0.956</pre>	-0.2529	0.612	-0.413	0.680	-1.46
=======================================	======	=======	========	:=======	======
Omnibus: 3.086		21.403	Durbin-Wat	son:	
Prob(Omnibus): 5.988		0.000	Jarque-Ber	a (JB):	
Skew:		0.077	Prob(JB):		
0.0501 Kurtosis: 16.1		2.013	Cond. No.		
=======================================		========	=======		

#### Warnings:

#### iv. Discussion

So the results are pretty disappointing, given that we had hoped to see correlations between the price and trading volume movements of commodities and their futures contrats. For each commodity, we graphed two scatterplots.

- The first scatterplot (left-hand side) of each commodity's price/trading volume movements looked for positive linear correlation between the present price movements and present trading volume movements. This works under the assumption that markets are perfectly efficient, and that market participants will instantly begin purchasing (selling) futures contracts to hedge when prices begin to move upward (downward).
- The second scatterplot (right-hand side) of each commodity's price/trading volume movements looked for positive linear correlation between present price movements and 1 month-lagged trading volume movements. We thought that, in the case that markets are not perfectly efficient (as they are not in the real world), market participants would need a fair amount of time to recognize when the prices began moving. We used a lag period of 1 month because we felt that 30 days was enough time for participants to notice price changes and decide whether to engage in buying/selling contracts.

The two regressions calculated for both combinations of variables for each commodity were incredibly uncorrelated. The R-squared values yieled attest to this. From this, we can say that the relationship between 1) present prices and present futures trading volumes, and 2) present prices and 1 month-lagged futures trading volumes is statistically insignificant.

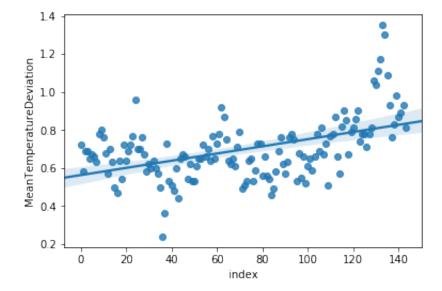
# III. Commodity Price Relationships with Average Global Temperature Deviations

Here, we try to see if there is any relationship between average global temperatures and futures trading volumes. Although actual crop harvests are difficult to project, it may be easier to hedge based on changes in temperature. In this case, our temperature dataset acts as a proxy for the overall climate change.

# i. Mean Temperature Deviation Growth Over Time

```
In [17]: sns.regplot(x="index", y="MeanTemperatureDeviation", data=df)
```

Out[17]: <matplotlib.axes.\_subplots.AxesSubplot at 0x11618fcf8>



```
In [18]: mod = sm.ols('MeanTemperatureDeviation ~ index', data=df)
    res = mod.fit()
    print(res.summary())
```

#### OLS Regression Results

=============	======= ==				
Dep. Variable:	Mean	Temperature	Deviation	R-squared:	
0.235		_		_	
Model:			OLS	Adj. R-squar	red:
0.229					
Method:		Least	Squares	F-statistic	:
43.53		. 15		<b>5</b> 1 (5 )	
Date: 7.69e-10		Tue, 15	May 2018	Prob (F-stat	tistic):
7.09e-10 Time:			00:11:23	Log-Likeliho	ood•
76.928			00.11.25	под-піксіїї	30 <b>u</b> •
No. Observatio	ns:		144	AIC:	
-149.9					
Df Residuals:			142	BIC:	
-143.9					
Df Model:			1		
Covariance Typ	e:	r	nonrobust		
=======================================	=======	========		:=======	=======
	coef	std err	+	P> t	10 025
0.975]	COCI	Sta CII	C	17   0	[0.023
Intercept	0.5624	0.024	23.749	0.000	0.516
0.609					
	0.0019	0.000	6.598	0.000	0.001
0.002					
=======================================	=======	=======			=======
Omnibus:		17.1	l21 Durbi	n-Watson:	
0.656		_,,,			
Prob(Omnibus):		0.0	000 Jarqu	ue-Bera (JB):	
25.305			-	, ,	
Skew:		0.6	35 Prob(	JB):	
3.20e-06					
Kurtosis:		4.6	Cond.	No.	
165.					
=========	======	========		========	=======
=======					

\_\_\_\_\_

#### Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

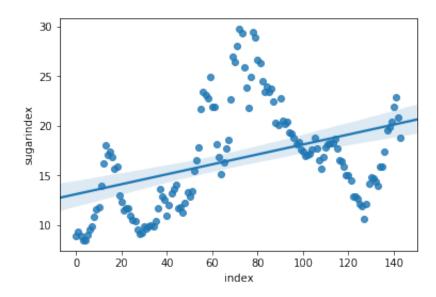
Looking at the scatterplot, we can see a weak positive linear relationship between the x-axis (Number of months since January 2005) and the y-axis (Mean Temperature Deviations from the 1950-1980 mean).

We ran a regression to check the strength of teh correlation. The R-squared values show weak positive correlation, confirming our visual observations.

# ii. Commodity Prices over Time

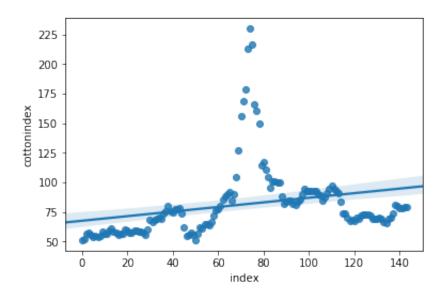
In [19]: sns.regplot(x="index", y="sugarindex", data=df)

Out[19]: <matplotlib.axes.\_subplots.AxesSubplot at 0x1161c6048>



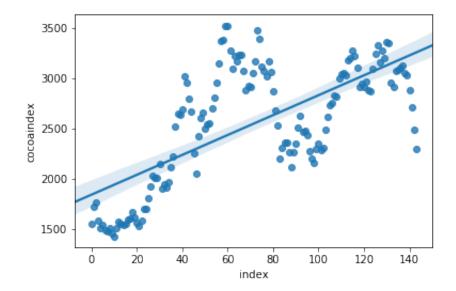
In [20]: sns.regplot(x="index", y="cottonindex", data=df)

Out[20]: <matplotlib.axes.\_subplots.AxesSubplot at 0x116047390>



In [21]: sns.regplot(x="index", y="cocoaindex", data=df)

Out[21]: <matplotlib.axes.\_subplots.AxesSubplot at 0x116038f28>



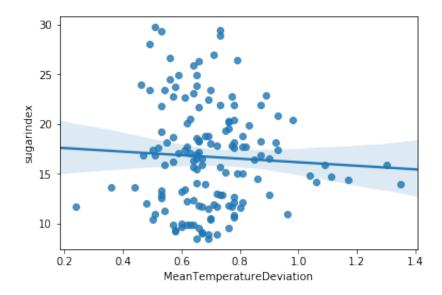
Although we can clearly see the positive correlation between our commodity prices and the number of months since January 2005, the prices are extremely cyclical, and these cycles span many months or years. Interestingly, we can see the commodity prices spike between 36 and 84 months since January 2005, which translates to 2008-2012 — the years of the Great Recession.

Having visually confirmed the connection between commodity prices and time since January 2005, we then try to identify any correlation between these commodity prices and the degree of deviation from the mean temperature between 1950 and 1980.

# iii. Relationship between Commodity Prices and Mean Temperature Deviation

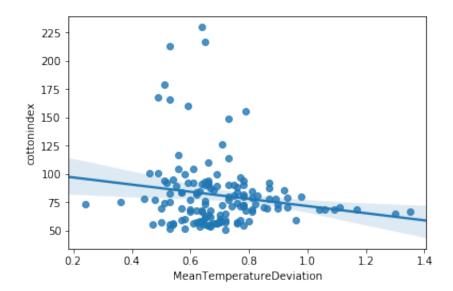
In [22]: sns.regplot(x="MeanTemperatureDeviation", y="sugarindex", data=df)

Out[22]: <matplotlib.axes.\_subplots.AxesSubplot at 0x1164ea9b0>



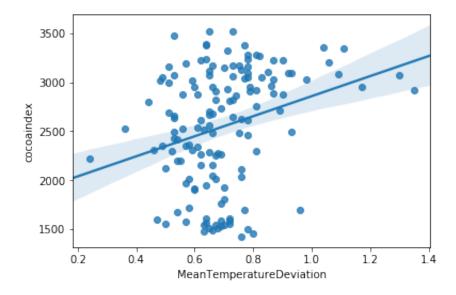
In [23]: sns.regplot(x="MeanTemperatureDeviation", y="cottonindex", data=df)

Out[23]: <matplotlib.axes.\_subplots.AxesSubplot at 0x116000208>



```
In [24]: sns.regplot(x="MeanTemperatureDeviation", y="cocoaindex", data=df)
```

Out[24]: <matplotlib.axes.\_subplots.AxesSubplot at 0x10c81deb8>



### iv. Discussion

As we can see, the connection between Mean Temperature Deviation and global commodity prices is largely zero or leaning negative. We recognize that commodity prices are affected by far more factors than simply the weather, however, and thus view these results as insignificant due to the sheer amount of noise that can affect the data.

However, we also decided to run a regression to confirm the lack of a relationship between commodity prices and Mean Temperature Deviation.

# v. Further Regression

```
In [25]: mod = sm.ols('sugarindex ~ MeanTemperatureDeviation', data=df)
    res = mod.fit()
    print(res.summary())
```

#### OLS Regression Results

=======================================		=======		======
Dep. Variable:	sugarindex	R-square	d:	
0.003				
Model:	OLS	Adj. R-s	quared:	
-0.004 Method:	Least Squares	E statis	+ia.	
0.4338	neast squares	r-statis	CIC.	
	e, 15 May 2018	Prob (F-	statistic):	
0.511	,		,	
Time:	00:11:25	Log-Like	lihood:	
-442.67				
No. Observations:	144	AIC:		
889.3	1.40	DIG		
Df Residuals: 895.3	142	BIC:		
Df Model:	1			
Covariance Type:	nonrobust			
=======================================		=======	========	======
	=			
	coef	std err	t	P> t
[0.025 0.975]				
Intercept	17.9515	1.940	9.254	0.000
14.117 21.786	2,,,,,,,		31201	
MeanTemperatureDeviation	n -1.7845	2.709	-0.659	0.511
-7.140 3.571				
=======================================		=======	========	======
Omnibus:	6 071	Durbin-W	at aon e	
0.071	0.9/1	Dulbin-w	atson:	
Prob(Omnibus):	0.031	Jarque-R	era (JB):	
6.213	0.031	cardae p	(	
Skew:	0.434	Prob(JB)	:	
0.0448		, ,		
Kurtosis:	2.468	Cond. No	•	
9.22				
		=======	========	======

# Warnings:

========

In [26]: mod = sm.ols('cottonindex ~ MeanTemperatureDeviation', data=df)
 res = mod.fit()
 print(res.summary())

OLS Regression Results				
=======				
Dep. Variable:	cottonindex	R-squared	d:	
0.027				
Model:	OLS	Adj. R-sc	quared:	
0.021				
	east Squares	F-statist	cic:	
4.006				
-	15 May 2018	Prob (F-s	statistic):	
0.0472				
Time:	00:11:25	Log-Likel	Lihood:	
-696.33				
No. Observations:	144	AIC:		
1397.				
Df Residuals:	142	BIC:		
1403.	_			
Df Model:	1			
Covariance Type:	nonrobust			
	=========	=======	=======	======
			<b>.</b>	ם או ב
10 025 0 0751	coei	std err	t	P> t
[0.025 0.975]				
Intercept	103.3649	11.293	9.153	0.000
81.041 125.689	10010019	111230	J. 130	0.000
MeanTemperatureDeviation	-31.5683	15.772	-2.002	0.047
-62.746 -0.391	0113000	130,72	21002	0.017
=======================================	=========	========		======
=======				
Omnibus:	94.195	Durbin-Wa	atson:	
0.085				
<pre>Prob(Omnibus):</pre>	0.000	Jarque-Be	era (JB):	
486.892		-	,	
Skew:	2.463	Prob(JB):	1	
1.87e-106		. ,		
Kurtosis:	10.542	Cond. No.	•	
9.22				
				======
========				

#### Warnings:

In [27]: mod = sm.ols('cocoaindex ~ MeanTemperatureDeviation', data=df)
 res = mod.fit()
 print(res.summary())

OLS Regression Results				
=======================================		=======	======	======
Dep. Variable:	cocoaindex	R-squared:		
0.078				
Model:	OLS	Adj. R-squ	ared:	
0.071	T			
	Least Squares	F-statisti	c:	
11.95 Date: Tue	, 15 May 2018	Drob (F st	atistis).	
0.000721	, 13 May 2016	Prob (F-st	aciscic):	
Time:	00:11:26	Log-Likeli	hood:	
-1118.4	***************************************	_09		
No. Observations:	144	AIC:		
2241.				
Df Residuals:	142	BIC:		
2247.				
Df Model:	1			
Covariance Type:	nonrobust			
	=========	========	=======	======
		std err	t	P> t
[0.025 0.975]	0001	bed ell	C	1,   0
Intercept	1834.9008	211.774	8.664	0.000
1416.263 2253.539				
MeanTemperatureDeviation	1022.5108	295.764	3.457	0.001
437.840 1607.181				
=======================================	=========	========	=======	======
Omnibus:	13.256	Durbin-Wat	son•	
0.107	13.230	Darbin wat	5011.	
Prob(Omnibus):	0.001	Jarque-Ber	a (JB):	
8.216		1	- (- ,	
Skew:	-0.434	Prob(JB):		
0.0164				
Kurtosis:	2.215	Cond. No.		
9.22				
=======================================	=========	========	=======	======
=======				

#### Warnings:

As predicted, the correlations we observe are next to useless. Perhaps completely useless. The R-squared values, once again, are spectacularly low, and imply that the correlation is statistically insignificant.

# IV. Findings and Concluding Remarks

From our data analysis, we can see that there is very little relationship between commodity prices, commodity futures trading volumes, and average global temperature deviations from the 1950-1980 mean (our proxy for quantifying climate change).

Even when considering the correlations between these variables in different scenarios, statistical analysis reveals that such correlations are not significant, and that they definitely do not fully explain changes in each other. Again, this makes sense because there are countless factors and risks that can affect these variables, whether the factors are political, social, economic, or environmental in nature. This is also due in part to the fact that the geographic location in which agricultural activity occurs is guite literally the entire world.

Perhaps if we had more focused data, such as commodity prices in the United States and futures trading volumes of American agricultural firms, we would be able to reach more productive conclusions. However, such granular data is nigh-impossible to obtain with the budget of an undergraduate, assuming such data even exists and is organized for analytical use.

In short, our data was likely too noisy due to its wide coverage of the world, and we found no statistically significant relationships between any of our factors, even when adjusting for market inefficiencies.