

1 Principal Component Analysis(PCA)

1.1 Settings

Let $\mathbf{X} = \{\mathbf{x}_i\}_{i=1}^n \in \mathbb{R}^p$ be a sample of n observations of p dimensional vectors. The first principal component of this sample is a real variable transformed from X

$$\mathbf{z}_1 = \mathbf{X}\mathbf{a}_1^T$$

where the vector $\mathbf{a}_1 = (a_{11}, a_{21}, \dots, a_{p1}) \in \mathbb{R}^p$ is chosen so that $\|\mathbf{a}_1\|_2 = 1$ and $\text{var}[\mathbf{z}_1]$ is maximized.

The k^{th} principal component of the sample is

$$\mathbf{z}_k = \mathbf{X}\mathbf{a}_k^T, \quad k = 2, \dots, p$$

where the vector $\mathbf{a}_k = (a_{1k}, a_{2k}, \dots, a_{pk}) \in \mathbb{R}^p$ is chosen so that $\|\mathbf{a}_k\|_2 = 1$, $\text{var}[\mathbf{z}_k]$ is maximized and $\text{cov}[\mathbf{z}_k, \mathbf{z}_l] = 0$ for $k > l \geq 1$.

Geometrically, at first step, PCA finds a unit vector $\mathbf{a}_1 \in \mathbb{R}^p$ that the projections of sample points on it has maximum variance. \mathbf{z}_1 is actually a variable of projection length of sample points on \mathbf{a}_1 . Secondly, PCA finds the second component \mathbf{z}_2 in the same logic but with further restriction that \mathbf{z}_2 has no relation with \mathbf{z}_1 , i.e. $\text{cov}[\mathbf{z}_2, \mathbf{z}_1] = 0$.

1.2 Computations

To find \mathbf{a}_1 , first note that

$$\begin{aligned} \text{var}[\mathbf{z}_1] &= \mathbb{E}[\mathbf{z}_1^2] - \mathbb{E}[\mathbf{z}_1]^2 \\ &= \sum_{i,j=1}^p a_{i1}a_{j1}\mathbb{E}[X_iX_j] - \sum_{i,j=1}^p a_{i1}a_{j1}\mathbb{E}[X_i]\mathbb{E}[X_j] \\ &= \sum_{i,j=1}^p a_{i1}a_{j1}\text{cov}[X_i, X_j] \\ &= (a_{11}, a_{21}, \dots, a_{p1}) \begin{pmatrix} \text{cov}[X_1, X_1] & \text{cov}[X_1, X_2] & \dots & \text{cov}[X_1, X_p] \\ \text{cov}[X_2, X_1] & \text{cov}[X_2, X_2] & \dots & \text{cov}[X_2, X_p] \\ \vdots & \vdots & \ddots & \vdots \\ \text{cov}[X_p, X_1] & \text{cov}[X_p, X_2] & \dots & \text{cov}[X_p, X_p] \end{pmatrix} \begin{pmatrix} a_{11} \\ a_{21} \\ \vdots \\ a_{p1} \end{pmatrix} \\ &= \mathbf{a}_1 \Sigma \mathbf{a}_1^T \end{aligned}$$

where Σ is the covariance matrix of X . The problem becomes

$$\max(\mathbf{a}_1 \Sigma \mathbf{a}_1^T) \quad \text{s.t.} \quad \|\mathbf{a}_1\|_2 = 1$$

Let λ be Lagrange multiplier and let $g(\mathbf{a}_1) = \mathbf{a}_1 \Sigma \mathbf{a}_1^T - \lambda(\mathbf{a}_1 \mathbf{a}_1^T - 1)$. Then we need

$$\nabla g(\mathbf{a}_1) = \mathbf{a}_1 \Sigma^T - \lambda \mathbf{a}_1 = (\Sigma \mathbf{a}_1^T - \lambda \mathbf{a}_1^T)^T = 0$$

Therefore, \mathbf{a}_1^T is the eigenvector of Σ corresponding to eigenvalue $\lambda \equiv \lambda_1$. Furthermore, we have maximized $\text{var}[\mathbf{z}_1] = \mathbf{a}_1 \Sigma \mathbf{a}_1^T = \lambda_1$, this means λ_1 is the largest eigenvalue of Σ .

For the second component, note that $\text{cov}[\mathbf{z}_2, \mathbf{z}_1] = \mathbf{a}_2 \Sigma \mathbf{a}_1^T = \mathbf{a}_2 \lambda_1 \mathbf{a}_1^T = 0$. For another multipliers ϕ and λ and another $g(\mathbf{a}_2) = \mathbf{a}_2 \Sigma \mathbf{a}_2^T - \lambda(\mathbf{a}_2 \mathbf{a}_2^T - 1) - \phi \lambda_1 \mathbf{a}_2 \mathbf{a}_1^T$, we need

$$\nabla g(\mathbf{a}_2) = (\Sigma \mathbf{a}_2^T - \lambda \mathbf{a}_2^T - \phi \lambda_1 \mathbf{a}_1^T)^T = 0$$

Multiply \mathbf{a}_1 both side, we get

$$\begin{aligned} & \mathbf{a}_1 \Sigma \mathbf{a}_2^T - \mathbf{a}_1 \lambda \mathbf{a}_2^T - \phi \lambda_1 \mathbf{a}_1 \mathbf{a}_1^T \\ &= (\mathbf{a}_2 \Sigma \mathbf{a}_1)^T - \lambda (\mathbf{a}_2 \mathbf{a}_1)^T - \phi \lambda_1 = 0 \end{aligned}$$

So ϕ must be zero. Hence

$$\Sigma \mathbf{a}_2^T = \lambda \mathbf{a}_2^T$$

\mathbf{a}_2^T is also an eigenvector and has $\lambda \equiv \lambda_2$ as its eigenvalue. And, again, $var[\mathbf{z}_2] = \mathbf{a}_2 \Sigma \mathbf{a}_2^T = \lambda_2$ is the second largest eigenvalue.

In general,

$$var[\mathbf{z}_k] = \mathbf{a}_k \Sigma \mathbf{a}_k^T = \lambda_k$$

The k^{th} largest eigenvalue of Σ is the variance of the k^{th} PC.