

## **COEN281 -- Introduction to Pattern Recognition and Data Mining**

### **Lecture 5: Linear Discriminant Functions**

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## **Syllabus**

Week 1	Introduction; R (Ch.1)
Week 2	Bayesian Decision Theory (Ch.2; DHS: 2.1-2.6, 2.9) Parameter Estimation (DHS: 3.1-3.4)
Week 3	<b>Linear Discriminant Functions</b> (Ch.3&4; DHS: 3.8.2, 5.1-5.8) Regularization (Ch.6; SE: Ch.3)
Week 4	Neural Networks (DHS: 6.1-6.6, 6.8); Deep Learning
Week 5	Support Vector Machines (Ch.9)
Week 6	Decision Trees (Ch. 8.1; DHS: 8.3; Ch 2 SE)
Week 7	Ensemble Methods (Ch. 8.2; SE: Ch 4, 5)
Week 8	Clustering (Ch. 10; DHS: 10.6, 10.7) Clustering (DHS: 10.9); How many clusters are there? (DHS: 10.10)
Week 9	Non-metric: Association Rules Collaborative Filtering
Week 10	Text Retrieval; Other topics

## Overview

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- Introduction
  - Approaches to building classifiers
  - Geometry of linear discriminant functions
- Linear separable case – Perceptron criteria
- Other methods
  - Linear Discriminant Analysis (**LDA**)
    - Restricted Gaussian classifier (see Lecture 2)
  - Linear Regression – Least Squares (**LS**) solution
  - Fisher's geometric view of LDA
  - **Logistic Regression**

## Introduction

### Building Classifiers

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- 1) *Class-conditional* ("generative") approach
  - $p_{\theta_j}(x|\omega_j)$  are modeled explicitly;  $\hat{\theta}_j$  are estimated via ML
  - Combined with estimates of  $P(\omega_j)$  are inverted via Bayes rule to arrive at  $P(\omega_j|x)$
- 2) *Regression* approach
  - $P(\omega_j|x)$  are modeled explicitly
  - e.g., **Logistic regression**
- 3) *Discriminative* approach
  - Try to model the decision boundary directly – i.e., a mapping from inputs  $x$  to one of the classes
  - Assume we know the form for the discriminant functions  $g_i(x)$

## Introduction

### Building Classifiers (2)

- Classification is an easier problem than density estimation (Vapnik)
  - Why use density estimation as an intermediate step?
  - Remember likelihood ratio:

$$\frac{p(\mathbf{x} | \omega_1)}{p(\mathbf{x} | \omega_2)} > \frac{\lambda_{12} - \lambda_{22}}{\lambda_{21} - \lambda_{11}} \times \frac{P(\omega_2)}{P(\omega_1)}$$

$\Rightarrow$  we only need to know if  $\frac{P(\omega_i) \cdot p(\mathbf{x} | \omega_i)}{P(\omega_j) \cdot p(\mathbf{x} | \omega_j)} > 1$

- i.e., only ratios matter!

## Introduction

### Linear Discriminant Functions

- Definition
  - Just a linear combination of the measurements of  $x$  written as  $g_{\theta}(x) = \boldsymbol{\theta}^t \mathbf{x} + \theta_0$
  - $\boldsymbol{\theta}$  is the “weight”<sup>T</sup> (or coefficient) vector of the model
  - $\theta_0$  the “bias” or “threshold” weight
- Optimal if underlying distributions are “cooperative”
  - Gaussians with  $\Sigma_i = \sigma^2 I$  or  $\Sigma_i = \Sigma$  (**LDA** - see Lecture 2)
  - Simplicity makes them attractive for initial, trial classifiers
  - Can be generalized to be linear in some given set of functions  $\varphi(x)$

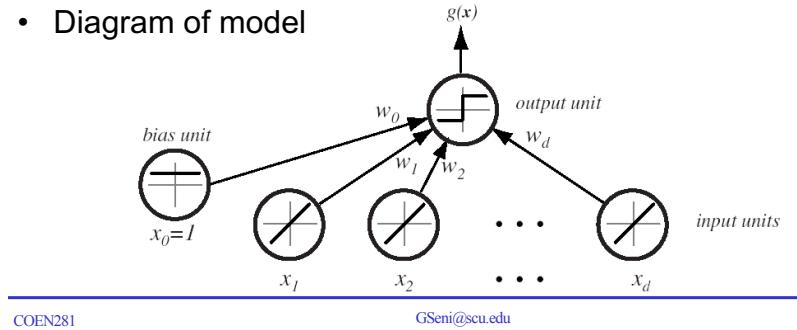
<sup>T</sup>sometimes we write  $w$  to refer to  $\theta$

## Introduction

### Linear Discriminant Functions (2)

- Decision rule - two-class case
  - Decide  $\omega_1$  if  $g(\mathbf{x}) > 0$  and  $\omega_2$  if  $g(\mathbf{x}) < 0$   
i.e., assign  $\mathbf{x}$  to  $\omega_1$  if  $\mathbf{w}'\mathbf{x}$  exceeds threshold  $-w_0$
  - If  $g(\mathbf{x}) = 0$  assignment is undefined – i.e., can go either way

- Diagram of model



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## Introduction

### Linear Discriminant Functions (3)

- Homogeneous form

$$g(\mathbf{x}) = w_0 + \sum_{i=1}^d w_i x_i = \sum_{i=0}^d w_i x_i \quad \text{where } x_0 = 1$$

- Augmented weight & feature vector

$$\mathbf{a} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_d \end{bmatrix} \quad \mathbf{x} = \begin{bmatrix} 1 \\ x_1 \\ \vdots \\ x_d \end{bmatrix}$$

- We write  $g(\mathbf{x}) = \mathbf{a}'\mathbf{x}$

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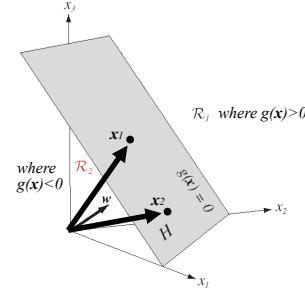
## Introduction

### Decision Surface

- Equation  $g(\mathbf{x})=0$  defines surface that separates points assigned to the category  $\omega_1$  from points assigned to the category  $\omega_2$ 
  - $g(\mathbf{x})$  linear  $\Rightarrow$  surface is a *hyperplane*  $H$
  - Consider  $\mathbf{x}_1$  and  $\mathbf{x}_2$  both on the decision surface:

$$\begin{aligned} \mathbf{w}'\mathbf{x}_1 + w_0 &= \mathbf{w}'\mathbf{x}_2 + w_0 \\ \text{or } \mathbf{w}'(\mathbf{x}_1 - \mathbf{x}_2) &= 0 \\ \Rightarrow \mathbf{w} &\text{ is normal to any vector lying in the hyperplane} \end{aligned}$$

- Orientation of  $H$  is determined by  $\mathbf{w}$



## Introduction

### Decision Surface (2)

- $g(\mathbf{x}) \propto$  distance from  $\mathbf{x}$  to  $H$

– Express  $\mathbf{x}$  as  $\mathbf{x} = \mathbf{x}_p + r \frac{\mathbf{w}}{\|\mathbf{w}\|}$

because  $g(\mathbf{x}_p) = 0$

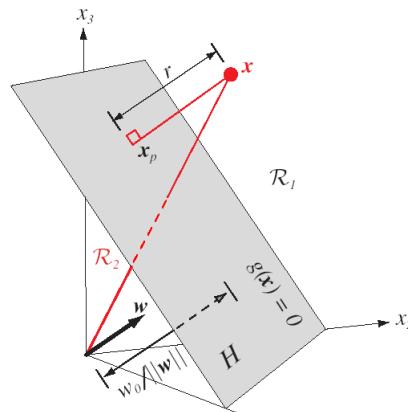
$$g(\mathbf{x}) = \mathbf{w}'\mathbf{x} + w_0 = g(\mathbf{x}_p) + r \frac{\mathbf{w}'\mathbf{w}}{\|\mathbf{w}\|}$$

$$= r\|\mathbf{w}\|$$

$$\Rightarrow r = \frac{g(\mathbf{x})}{\|\mathbf{w}\|}$$

– Also,  $d(0, H) = w_0 / \|\mathbf{w}\|$

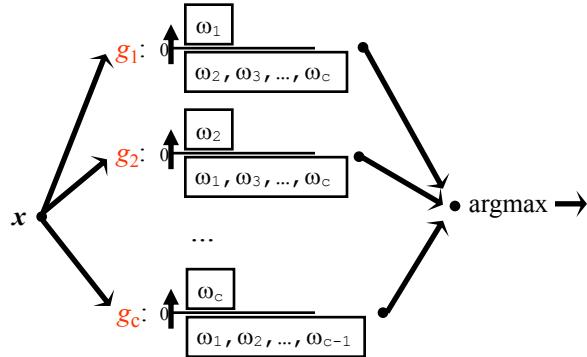
- Location of  $H$  is determined by  $w_0$



## Introduction

### Multiclass Case

- One per class decomposition (*linear machine*)
  - i.e.,  $C$  discriminant functions
  - $\omega_i$  vs.  $\neg\omega_i$



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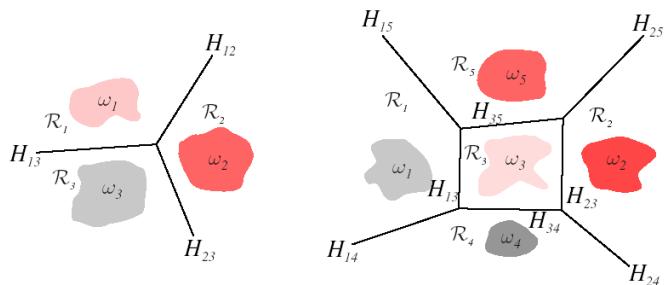
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## Introduction

### Multiclass Case (2)

- Decision boundaries



- $H_{ij}$  defined by  $g_i(\mathbf{x}) = g_j(\mathbf{x})$
- Number of  $H_{ij}$  is often fewer than  $c(c-1)/2$
- Decision regions are convex and singly connected
  - Most suitable when  $p(\mathbf{x}|\omega_j)$  is unimodal
    - Many exceptions!

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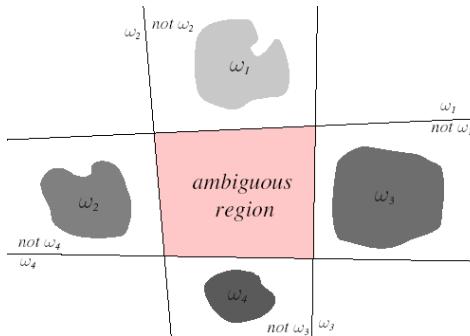
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## Introduction

### Multiclass Case (3)

- Without  $\text{argmax}$ , ambiguous class assignments can arise



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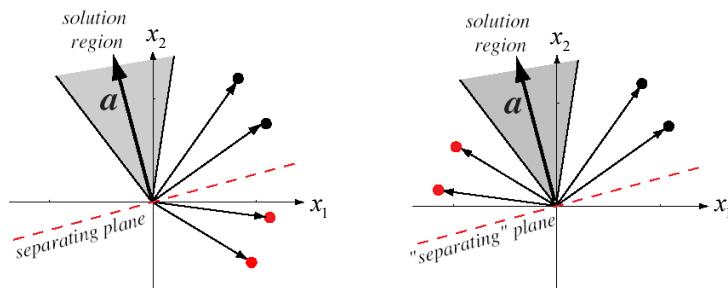
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## Linear Separable Case

### Perceptron

- Simplifying normalization
    - Replace  $\omega_2$  samples by their negatives
- ⇒ Find  $a$  such that  $a'x > 0$  for all samples



- Note that  $a$  is not unique!

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## Linear Separable Case

### Perceptron (2)

- Score function

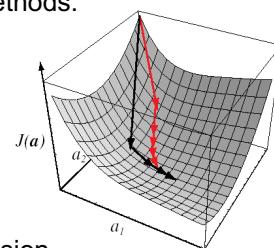
- A scalar function  $J(\mathbf{a})$  that is minimized if  $\mathbf{a}$  is a solution vector
- Allows use of *Gradient Descent* (search) methods:

$$\mathbf{a}(k+1) = \mathbf{a}(k) - \eta(k) \nabla J(\mathbf{a}) \quad \text{or}$$

$$\mathbf{a}(k+1) = \mathbf{a}(k) - \mathbf{H}^{-1} \nabla J(\mathbf{a}) \quad (\text{Newton})$$

- Idea 1:  $J(\mathbf{a})$  is # of misclassified samples
- Idea 2:  $J_p(\mathbf{a})$  is  $\propto$  to sum of distances to decision boundary

$$J_p(\mathbf{a}) = \sum_{\mathbf{x} \in \text{Miss}(\mathbf{a})} (-\mathbf{a}^T \mathbf{x}) \quad \text{where } \text{Miss}(\mathbf{a}) \text{ is misclassified set}$$



## Linear Separable Case

### Perceptron (3)

- Fixed-increment, single-sample

```
k ← 0
do {
    k ← k+1
    if ( $x^k$  is misclassified by  $\mathbf{a}$ ) {
         $\mathbf{a} ← \mathbf{a} + \underline{\mathbf{x}^k} \rightarrow -\nabla J_p$ 
    }
} until (all patterns are properly classified)
```

- Convergence Theorem – Perceptron algorithm is guaranteed to find a solution if samples are linearly separable
- In nonseparable case, error-correcting algorithm produces an infinite sequence  $\mathbf{a}(k) \Rightarrow$  limited applicability

## Linear Regression

### Least Squares Error

- Class encoding:  $y = 1 \text{ if } \text{class}(\mathbf{x}) = \omega_1, \text{ else } y = -1$
- Model:  $\hat{y} = \hat{w}_0 + \sum_{j=1}^p x_j \hat{w}_j = \mathbf{x}' \hat{\mathbf{w}}$
- Score function:  $J_s(\mathbf{w}) = \text{RSS}(\mathbf{w}) = \sum_{i=1}^n (y_i - \mathbf{x}'_i \mathbf{w})^2$ 
  - Rationale - minimizing the size of the error vector  $\mathbf{e} = \mathbf{Y} - \hat{\mathbf{Y}}$  (see Lecture 4)
- Note that  $\mathbf{X}$  is rectangular and  $\mathbf{w}$  is overdetermined
  - $\mathbf{Y} = \mathbf{X}\mathbf{w}$  ordinarily has no exact solution
- $J_s(\mathbf{w})$  is quadratic – we can look for a single global minimum  
 $(\nabla J_s = 0)$

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## Linear Regression

### Least Squares Error (2)

- Closed-form solution

$$\begin{aligned}\nabla J_s &= -\sum_{i=1}^n 2(y_i - \mathbf{x}'_i \mathbf{w}) \mathbf{x}_i = -2\mathbf{X}'(\mathbf{Y} - \mathbf{X}\mathbf{w}) \\ \nabla J_s &= 0 \quad \Rightarrow \quad \mathbf{X}'\mathbf{Y} = \mathbf{X}'\mathbf{X}\mathbf{w} \\ \hat{\mathbf{w}} &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \boxed{\mathbf{X}^+\mathbf{Y}}\end{aligned}$$

- A more general definition of the *pseudoinverse* always exists:  $\mathbf{X}^+ \equiv \lim_{\varepsilon \rightarrow 0} (\mathbf{X}'\mathbf{X} + \varepsilon \mathbf{I})^{-1}\mathbf{X}'$
- We expect to obtain a useful discriminant in both the separable and the nonseparable cases
  - When  $c$  is large, sensitive to “masking” problem (Hastie)

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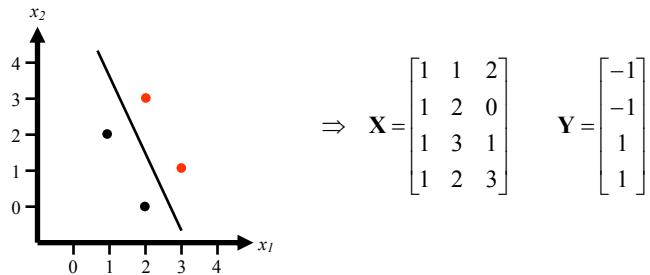
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## Linear Regression

### Minimum Squared Error (3)

- Example



- In R: `X.plus <- solve(t(X) %*% X) %*% t(X)`

$$\mathbf{X}^+ = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' = \begin{bmatrix} 5/4 & 13/12 & -3/4 & -7/12 \\ -1/2 & -1/6 & 1/2 & 1/6 \\ 0 & -1/3 & 0 & 1/3 \end{bmatrix} \Rightarrow \mathbf{X}^+\mathbf{Y} = \mathbf{w} = \begin{bmatrix} -11/3 \\ 4/3 \\ 2/3 \end{bmatrix}$$

$$\Rightarrow g(\mathbf{x}) = \mathbf{w}'\mathbf{x} = -\frac{11}{3} + \frac{4}{3}x_1 + \frac{2}{3}x_2$$

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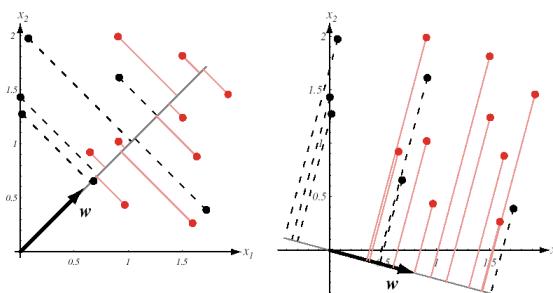
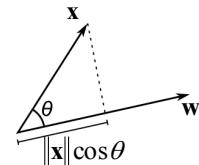
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## Fisher Linear Discriminant

### Low-Dimensional Projection

- Geometric interpretation of dot product
  - Length of the projection of  $\mathbf{x}$  onto the (unit) vector  $\mathbf{w}$
$$\mathbf{w}'\mathbf{x} = \|\mathbf{w}\| \|\mathbf{x}\| \cos \theta$$
- Searching for the  $\mathbf{w}$  that best separates the projected data



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## Fisher Linear Discriminant

### Low-Dimensional Projection (2)

- Criterion function
  - Idea 1: use the distance between the projected sample means
 
$$|\tilde{m}_1 - \tilde{m}_2| = |\mathbf{w}'(\mathbf{m}_1 - \mathbf{m}_2)| \quad \text{where } \mathbf{m}_i = \frac{1}{n_i} \sum_{\mathbf{x} \in D_i} \mathbf{x}$$
    - Dependent on  $\|\mathbf{w}\|$ ... could be made arbitrarily large
  - Idea 2: maximize ratio of between-class scatter (as above) to within-class scatter
 
$$J_F(\mathbf{w}) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{S}_1^2 + \tilde{S}_2^2} \quad \text{where } \tilde{S}_i^2 = \sum_{\mathbf{x} \in D_i} (\mathbf{w}' \mathbf{x} - \mathbf{w}' \mathbf{m}_i)^2$$
    - Clearly,  $(1/n)(\tilde{S}_1^2 + \tilde{S}_2^2)$  is an estimate of the variance of the pooled data

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## Fisher Linear Discriminant

### Low-Dimensional Projection (3)

- $\mathbf{w}$  that optimizes  $J_F()$  can be shown to be

$$\mathbf{w} = \mathbf{S}_w^{-1}(\mathbf{m}_1 - \mathbf{m}_2) \quad \text{where } \mathbf{S}_w = \mathbf{S}_1 + \mathbf{S}_2$$

$$\mathbf{S}_i = \sum_{\mathbf{x} \in D_i} (\mathbf{x} - \mathbf{m}_i)(\mathbf{x} - \mathbf{m}_i)^t$$

–  $\mathbf{S}_w$  is called the within class scatter matrix

- Connection to LDA --  $p(\mathbf{x} | \omega_i) \sim N(\boldsymbol{\mu}_i, \boldsymbol{\Sigma})$

$$\begin{aligned} g(\mathbf{x}) &= g_i(\mathbf{x}) - g_j(\mathbf{x}) \\ &= (\mathbf{w}_i^t \mathbf{x} + w_{i0}) - (\mathbf{w}_j^t \mathbf{x} + w_{j0}) \\ &= \mathbf{x}^t \underbrace{\boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_i - \boldsymbol{\mu}_j)}_{\text{since } \mathbf{w}_i = \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i} + (w_{i0} - w_{j0}) \end{aligned} \quad \text{(see Lecture 2)}$$

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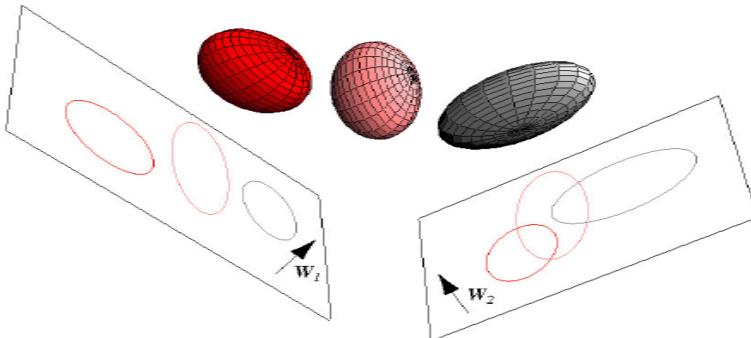
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## Fisher Linear Discriminant

### Low-Dimensional Projection (4)

- For the  $c$ -class problem,  $c-1$  functions are required

– Projection is from a  $d$  to a  $(c-1)$  dimensional space ( $d > c$ )



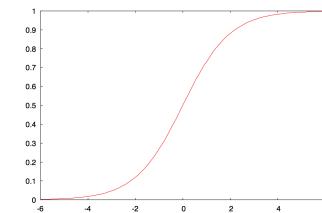
– Sacrifice performance for the advantage of lower-dimensional space

## Logistic Regression

### Modeling Posteriors

- Model form:  $P(\omega_1 | \mathbf{x}) = \phi(\beta_0 + \beta' \mathbf{x})$  where  $\phi$  is the “logistic” function

$$\phi(z) = \frac{e^z}{1+e^z} = \frac{1}{1+e^{-z}}$$



– Two-class case:  $P(\omega_2 | \mathbf{x}) = 1 - P(\omega_1 | \mathbf{x}) = \frac{1}{1+e^{\beta_0+\beta' \mathbf{x}}}$

- Log of “odds ratio” is linear

$$\log \frac{P(\omega_1 | \mathbf{x})}{P(\omega_2 | \mathbf{x})} = \beta_0 + \beta' \mathbf{x} \quad \Rightarrow \text{decision boundaries are linear}$$

## Logistic Regression

### Fitting Model

- $\phi'$  is given by:

$$\phi'(z) = \frac{e^{-z}}{(1+e^{-z})^2} = \frac{e^{-z}}{1+e^{-z}} \frac{1}{1+e^{-z}} = \frac{1}{1+e^z} \frac{e^z}{1+e^z} = \phi(z)(1-\phi(z))$$

- Log-likelihood (two-class case)

$$l(\beta) = \sum_{i=1}^n b_i \ln P(\mathbf{x}_i; \beta) + (1-b_i) \ln(1-P(\mathbf{x}_i; \beta))$$

$$b_i = \begin{cases} 1 & x \in \omega_1 \\ 0 & \text{otherwise} \end{cases}$$

$$\partial l / \partial \beta_r = \sum_{i=1}^n \left( \frac{b_i}{P_i} - \frac{1-b_i}{1-P_i} \right) \phi'(\beta^t \mathbf{x}_i) x_{ir}$$

$$\partial l / \partial \beta = \sum_{i=1}^n \left( \frac{b_i}{P_i} - \frac{1-b_i}{1-P_i} \right) P_i(1-P_i) \mathbf{x}_i = \sum_{i=1}^n [b_i(1-P_i) - P_i(1-b_i)] \mathbf{x}_i$$

$$= \sum_{i=1}^n (b_i - P_i) \mathbf{x}_i = \mathbf{X}^t (\mathbf{b} - \mathbf{P})$$

## Logistic Regression

### Fitting Model (2)

- Differentiating again to obtain the Hessian:

$$\partial^2 l / \partial \beta_s \partial \beta_r = \sum_{i=1}^n \partial \beta_s (b_i - P_i) x_{ir} = -\sum_{i=1}^n \phi'(\beta^t \mathbf{x}_i) x_{ir} x_{is} = -\sum_{i=1}^n P_i(1-P_i) x_{ir} x_{is}$$

$$\mathbf{H} = -\mathbf{X}^t \mathbf{W} \mathbf{X} \quad \text{where } \mathbf{W} = \begin{pmatrix} P_1(1-P_1) & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & P_n(1-P_n) \end{pmatrix}$$

- Newton steps is:

$$\begin{aligned} \boldsymbol{\beta}(k+1) &= \boldsymbol{\beta}(k) - \mathbf{H}^{-1} \nabla J(\boldsymbol{\beta}) \\ &= \boldsymbol{\beta}(k) + [\mathbf{X}^t \mathbf{W} \mathbf{X}]^{-1} \mathbf{X}^t (\mathbf{b} - \mathbf{P}) \end{aligned}$$

## Logistic Regression

### Comparison to LDA

- We had  $g(\mathbf{x}) = g_i(\mathbf{x}) - g_j(\mathbf{x}) = (\mathbf{w}_i^t \mathbf{x} + w_{i0}) - (\mathbf{w}_j^t \mathbf{x} + w_{j0})$   
 $= \mathbf{x}^t \Sigma^{-1} (\mu_i - \mu_j) + (w_{i0} - w_{j0})$  since  $\mathbf{w}_i = \Sigma^{-1} \mu_i$   
 $= \alpha_0 + \alpha^t \mathbf{x}$
- Simply note that  $g(\mathbf{x}) = \log \frac{P(\omega_i | \mathbf{x})}{P(\omega_j | \mathbf{x})}$ 
  - LR's  $\beta$  computed directly not via  $\mu_i, \mu_j, \Sigma$ 
    - i.e., optimizing different criteria
  - LR holds also for some non-normal densities... it only needs the ratio to be of the logistic type
  - If  $x_i$  are normal, then LDA is 30% more efficient

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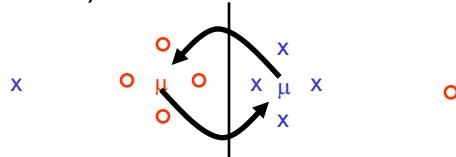
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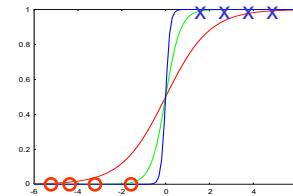
## Logistic Regression

### Comparison to LDA (2)

- If  $x_i$  are not normal, then LDA can be much worse (e.g., extreme outliers)



- LR can be degenerate on separable data
  - Numerical issues when  $\|\beta\| = \infty$
- In general, LR is a safer, more robust bet, but often similar results



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