## Introduction to Statistical Machine Learning Homework 1

Yota Toyama

October 20, 2016

1.

$$\frac{\partial}{\partial \mathbf{w}} R(\mathbf{w}) = 0$$

$$\frac{\partial}{\partial \mathbf{w}} E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x})^2 \right] = 0$$

$$E_{p(\mathbf{x},y)} \left[ 2(y - \mathbf{w}^T \mathbf{x})(-\mathbf{x}) \right] = 0$$

$$E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) \mathbf{x} \right] = 0$$

$$\mathbf{a}^T E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) \mathbf{x} \right] = 0$$

$$E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) \mathbf{a}^T \mathbf{x} \right] = 0$$

2.

$$E_{p(\mathbf{x},y)} \left[ y - \mathbf{w}^T \mathbf{x} \right] = 0$$

$$E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) E_{p(\mathbf{x})} \left[ \mathbf{x} \right] \right] = 0$$

$$E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) \mathbf{x} \right] + E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) E_{p(\mathbf{x})} \left[ \mathbf{x} \right] \right] = 0$$

$$E_{p(\mathbf{x},y)} \left[ (y - \mathbf{w}^T \mathbf{x}) (\mathbf{x} - E_{p(\mathbf{x})} \left[ \mathbf{x} \right]) \right] = 0$$

 $\therefore$  the correration between data and prediction errors is 0.

3.

$$\hat{\mathbf{w}} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{i=1}^{N} (y_i - \mathbf{w}^T \mathbf{x}_i)^2$$
$$= (X^T X)^{-1} X^T y$$

Let  $C \in \mathbb{R}^{(d+1) \times (d+1)}$  be a diagonal matrix s.t.  $\tilde{X} = XC$ 

$$\begin{split} \hat{\tilde{\mathbf{w}}} &= \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{i=1}^{N} (y_i - \mathbf{w}^T \tilde{\mathbf{x}}_i)^2 \\ &= (\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T y \\ &= ((XC)^T X C)^{-1} (XC)^T y \\ &= (CX^T X C)^{-1} CX^T y \\ &= C^{-1} (X^T X)^{-1} C^{-1} CX^T y \\ &= C^{-1} (X^T X)^{-1} X^T y \\ \tilde{X} \hat{\tilde{\mathbf{w}}} &= XCC^{-1} (X^T X)^{-1} X^T y \\ &= X(X^T X)^{-1} X^T y \\ &= X \hat{\mathbf{w}} \text{ as required} \end{split}$$

4.

$$\hat{\sigma^2} = \underset{\sigma^2}{\operatorname{argmax}} \sum_{i=1}^{N} \log p(y_i | \mathbf{x}_i; \mathbf{w}, \sigma)$$

$$= \underset{\sigma^2}{\operatorname{argmax}} - \frac{1}{2\sigma^2} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2 - N \log \sigma \sqrt{2\pi}$$

$$= \underset{\sigma^2}{\operatorname{argmin}} \frac{1}{\sigma^2} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2 + N \log 2\pi \sigma^2$$

$$= \underset{\sigma^2}{\operatorname{argmin}} \frac{1}{\sigma^2} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2 + N \log \sigma^2$$

$$\frac{\partial}{\partial \sigma^2} \frac{1}{\sigma^2} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2 + N \log \sigma^2 = 0$$

$$-\frac{1}{\sigma^4} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2 + \frac{N}{\sigma^2} = 0$$

$$\sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2 - N\sigma^2 = 0$$

$$\hat{\sigma^2} = \frac{1}{N} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2$$

$$\therefore \hat{\sigma^2} = \frac{1}{N} \sum_{i=1}^{N} (y_i - f(\mathbf{x}_i; \mathbf{w}))^2$$

5.

Let 
$$\beta_{y \cdot y} = \begin{cases} 1 & \text{if } \hat{y} \leq y \\ \alpha & \text{otherwise} \end{cases}$$

$$l_{\alpha}(\hat{y}, y) = \begin{cases} (\hat{y} - y)^2 & \text{if } \hat{y} \leq y \\ \alpha(\hat{y} - y)^2 & \text{otherwise} \end{cases}$$

$$= \beta_{y \cdot y} (\hat{y} - y)^2$$

$$\frac{\partial}{\partial \mathbf{w}} L_{\alpha} = \frac{\partial}{\partial \mathbf{w}} \sum_{i=1}^{N} l_{\alpha}(\hat{y}, y)$$

$$= \frac{\partial}{\partial \mathbf{w}} \frac{1}{N} \sum_{i=1}^{N} \beta_{y \cdot y} (\hat{y}_i - y_i)^2$$

$$= \frac{\partial}{\partial \mathbf{w}} \frac{1}{N} \sum_{i=1}^{N} \beta_{y \cdot y} (\mathbf{w}^T \mathbf{x}_i - y_i)^2$$

$$= \frac{1}{N} \sum_{i=1}^{N} 2\beta_{y \cdot y} (\mathbf{w}^T \mathbf{x}_i - y_i) \mathbf{x}_i$$