

Multivariate Analysis Lecture 17: Factor Analysis

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Section 1

Motivating Example

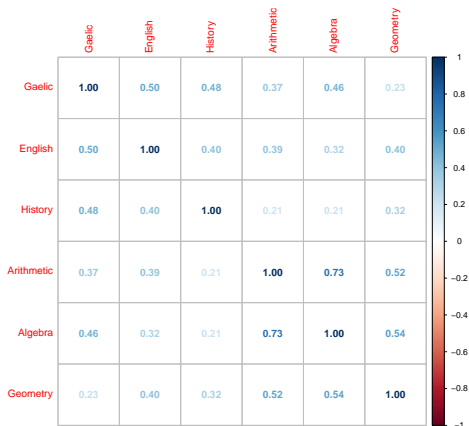
A Motivating Example: Exam Score

```
exam.cor=rbind(
c(1.000, 0.439, 0.410, 0.288, 0.329, 0.248),
c(0.439, 1.000, 0.351, 0.354, 0.320, 0.329),
c(0.410, 0.351, 1.000, 0.164, 0.190, 0.181),
c(0.288, 0.354, 0.164, 1.000, 0.595, 0.470),
c(0.329, 0.320, 0.190, 0.595, 1.000, 0.464),
c(0.248, 0.329, 0.181, 0.470, 0.464, 1.000))
exam.cov=diag(c(2, 3, 2, 3, 2, 2)) %*%exam.cor %*% diag(c(2, 3, 2, 3, 2, 2))

set.seed(2)
exam=round(mvrnorm(n=60, mu=rep(80,6), Sigma=exam.cov))
colnames(exam)=c("Gaelic",
"English",
"History",
"Arithmetic",
"Algebra",
"Geometry")
exam=data.frame(exam)
```

Pairwise Correlation

```
corrplot(cor(exam), method="number")
```



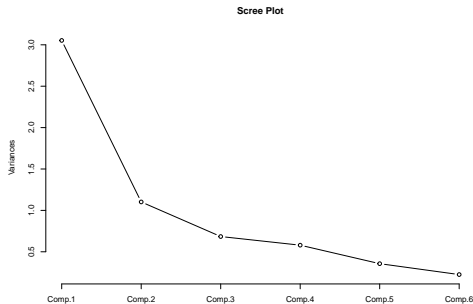
PCA

```
obj=princomp(exam, cor=TRUE)
obj$loadings
```

```
##
## Loadings:
##          Comp.1 Comp.2 Comp.3 Comp.4 Comp.5 Comp.6
## Gaelic      0.403  0.390  0.572      0.488  0.340
## English     0.398  0.322 -0.108 -0.801 -0.168 -0.238
## History     0.328  0.583 -0.324  0.552 -0.358 -0.122
## Arithmetic  0.448 -0.416  0.158      -0.581  0.514
## Algebra     0.453 -0.404  0.274  0.218      -0.708
## Geometry    0.408 -0.262 -0.676      0.511  0.217
##
##          Comp.1 Comp.2 Comp.3 Comp.4 Comp.5 Comp.6
## SS loadings  1.000  1.000  1.000  1.000  1.000  1.000
## Proportion Var 0.167  0.167  0.167  0.167  0.167  0.167
## Cumulative Var 0.167  0.333  0.500  0.667  0.833  1.000
```

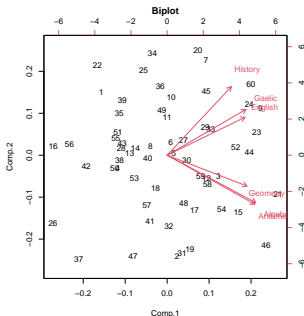
Visualize Eigenvalues

```
plot(obj, type="lines", main="Scree Plot")
```



Visualize First and Second PC (biplot)

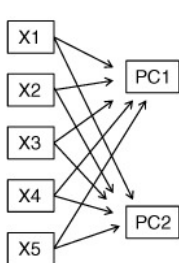
```
biplot(obj, main="Biplot")
```



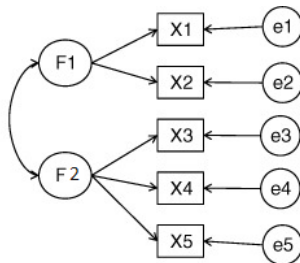
- How to interpret the results?

Introduction to Factor Analysis

```
knitr::include_graphics("img/PCA_FA.png")
```



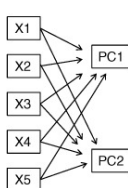
(a) Principal Components Model



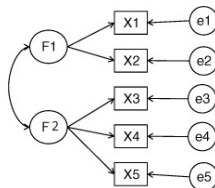
(b) Factor Analysis Model

<https://livebook.manning.com/book/r-in-action-second-edition/chapter-14/6>

PCA vs FA



(a) Principal Components Model



(b) Factor Analysis Model

- Both reduce dimensionality
- Both use linear combinations
- PCA leads to principal components, which are linear combinations of functions
- FA leads to factors (latent and unobserved)

Section 2

The FA Model

The Factor Model

- Consider a random vector $\mathbf{X} \in \mathbb{R}^p$
- Let $\boldsymbol{\mu} \in \mathbb{R}^p$ denote the population mean
- Let $F \in \mathbb{R}^m$ denote m factors

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix}, \boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{pmatrix}, \mathbf{F} = \begin{pmatrix} f_1 \\ f_2 \\ \vdots \\ f_m \end{pmatrix}$$

The Factor model

- X_j , which is observable, is assumed to be a linear function of the unobservable **common factors** f_1, \dots, f_m plus specific errors.

$$X_1 = \mu_1 + l_{11}f_1 + l_{12}f_2 + \dots + l_{1m}f_m + \epsilon_1$$

$$X_2 = \mu_2 + l_{21}f_1 + l_{22}f_2 + \dots + l_{2m}f_m + \epsilon_2$$

$$\vdots$$

$$X_p = \mu_p + l_{p1}f_1 + l_{p2}f_2 + \dots + l_{pm}f_m + \epsilon_p$$

where ϵ_j is called the specific factor for feature j .

- The means μ_1, \dots, μ_p are parameters
- The coefficients in the factor loading matrix are also parameters

The Factor model

- Let \mathbf{L} denote the $p \times m$ matrix of factor loadings

$$\mathbf{L} = \begin{pmatrix} l_{11} & l_{12} & \dots & l_{1m} \\ l_{21} & l_{22} & \dots & l_{2m} \\ \vdots & \vdots & & \vdots \\ l_{p1} & l_{p2} & \dots & l_{pm} \end{pmatrix}$$

- A compact expression of the factor model is

$$\mathbf{X}_{p \times 1} = \boldsymbol{\mu}_{p \times 1} + \mathbf{L}_{p \times m} \mathbf{F}_{m \times 1} + \boldsymbol{\epsilon}_{p \times 1}$$

Assumptions of FA

- \mathbf{F} and ϵ are uncorrelated
- The common factors are uncorrelated

$$\mathbb{E}(\mathbf{F}) = 0, \text{Cov}(\mathbf{F}) = \mathbf{I}$$

- The specific factors are uncorrelated

$$\mathbb{E}(\epsilon) = 0, \text{Cov}(\epsilon) = \Psi$$

where Ψ is a diagonal matrix with non-negative values, i.e.,

$$\Psi = \begin{pmatrix} \psi_1 & 0 & \dots & 0 \\ 0 & \psi_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \psi_p \end{pmatrix}$$

The Covariance

- By the factor model and its assumptions, we have

$$\begin{aligned}\Sigma &= \text{cov}(\mathbf{X}) \\ &= \text{cov}(\mathbf{LF} + \epsilon) \\ &= \mathbf{L} \text{Cov}(\mathbf{F}) \mathbf{L}^T + \Psi \\ &= \mathbf{LL}^T + \Psi\end{aligned}$$

The last step is due to our assumption that $\text{cov}(\mathbf{F}) = \mathbf{I}$

The Covariance

- For $i \neq j$, the covariance between X_i (feature i) and X_j (feature j) is

$$\sigma_{ij} = \text{cov}(X_i, X_j) = \sum_{k=1}^m l_{ik} l_{jk}$$

- The variance of X_i is

$$\sigma_{ii} = \sum_{k=1}^m l_{ik}^2 + \psi_i$$

Section 3

Communality

Communality and Specific Variance

- From last slide

$$\sigma_{ii} = \sum_{k=1}^m l_{ik}^2 + \psi_i$$

- We say that the variance of X_i is partitioned into communality and specific variance where
 - communality is defined as $h_i^2 = \sum_{k=1}^m l_{ik}^2$, which is the proportion of variance contributed by common factors
 - specific variance ψ_i , which is the specific variance of X_i

Example of Communality

```
obj=factanal(exam, factors=2)
L=obj$loadings[,1:2]
Psi=diag(obj$uniquenesses)
#communality
1-obj$uniquenesses
```

```
##      Gaelic      English      History Arithmetic      Algebra      Geometry
## 0.5217671 0.4459039 0.4655173 0.7016239 0.7759166 0.3925075
```

- For Geometry, the communality is 0.39 and the uniqueness (specific variance) is 0.61.

Section 4

Non-uniqueness

Non-uniqueness of Factor Loadings

- The factor loading coefficient is NOT unique.
- Suppose $\mathbf{X} = \boldsymbol{\mu} + \mathbf{LF} + \boldsymbol{\epsilon}$
- Consider any $m \times m$ orthogonal matrix Γ , which satisfies $\Gamma\Gamma^T = \Gamma^T\Gamma = \mathbf{I}$.
- Let $\tilde{\mathbf{L}} = \mathbf{L}\Gamma$ The model $\mathbf{X} = \boldsymbol{\mu} + \tilde{\mathbf{L}}\mathbf{F} + \boldsymbol{\epsilon}$

give the same Σ because

$$\text{cov}(\tilde{\mathbf{L}}\mathbf{F}) = \tilde{\mathbf{L}}\text{cov}(\mathbf{F})\tilde{\mathbf{L}}^T = \tilde{\mathbf{L}}\tilde{\mathbf{L}}^T = \mathbf{L}\Gamma\Gamma^T\mathbf{L}^T = \mathbf{L}\mathbf{L}^T = \text{cov}(\mathbf{L}\mathbf{F})$$

Non-uniqueness of Factor Loadings

```
#Estimated Sigma
L%*%t(L) + Psi
```

	Gaelic	English	History	Arithmetic	Algebra	Geometry
## Gaelic	1.0000011	0.4819908	0.4661743	0.3979476	0.4132344	0.3682496
## English	0.4819908	0.9999986	0.4363080	0.3514418	0.3645679	0.3308447
## History	0.4661743	0.4363080	0.9999990	0.2158592	0.2209179	0.2483943
## Arithmetic	0.3979476	0.3514418	0.2158592	1.0000009	0.7377925	0.5106265
## Algebra	0.4132344	0.3645679	0.2209179	0.7377925	1.0000003	0.5355603
## Geometry	0.3682496	0.3308447	0.2483943	0.5106265	0.5355603	0.9999993

Non-uniqueness of Factor Loadings

- Consider a rotation matrix R and define $\tilde{L} = LR$

```
theta=pi/6
R=matrix(c(cos(theta), sin(theta), -sin(theta), cos(theta)), 2,2)
L.tilde=L%*%R

L.tilde %*% t(L.tilde) + Psi
```

	Gaelic	English	History	Arithmetic	Algebra	Geometry
Gaelic	1.0000011	0.4819908	0.4661743	0.3979476	0.4132344	0.3682496
English	0.4819908	0.9999986	0.4363080	0.3514418	0.3645679	0.3308447
History	0.4661743	0.4363080	0.9999990	0.2158592	0.2209179	0.2483943
Arithmetic	0.3979476	0.3514418	0.2158592	1.0000009	0.7377925	0.5106265
Algebra	0.4132344	0.3645679	0.2209179	0.7377925	1.0000003	0.5355603
Geometry	0.3682496	0.3308447	0.2483943	0.5106265	0.5355603	0.9999993

Section 5

Factor Rotation

Subsection 1

Rotation for Better Interpretation

Rotation for Better Interpretation

- Interpretation of final results are easier for some choices of \mathbf{L} than others.
- We often rotate the factors to gain insights or for better interpretation
- This is one advantage of factor analysis
- In practice,
 - Step 1: fit a factor model by imposing conditions that lead to a unique solution
 - Step 2: the loading matrix \mathbf{L} is rotated (multiplied by an orthogonal matrix) in a way that gives a good interpretation of the data. Trial and error
- Well know criteria of rotation exist

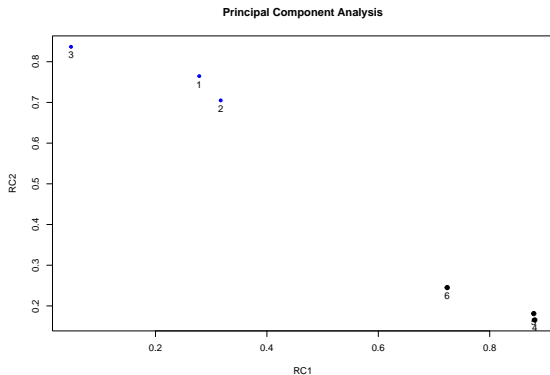
Factor Rotation

```
principal(exam, nfactors=2)
```

```
## Principal Components Analysis
## Call: principal(r = exam, nfactors = 2)
## Standardized loadings (pattern matrix) based upon correlation matrix
##          RC1  RC2   h2   u2 com
## Gaelic      0.28 0.76 0.66 0.34 1.3
## English      0.32 0.70 0.60 0.40 1.4
## History      0.05 0.84 0.70 0.30 1.0
## Arithmetic   0.88 0.17 0.80 0.20 1.1
## Algebra      0.88 0.18 0.81 0.19 1.1
## Geometry     0.72 0.25 0.58 0.42 1.2
##
##
##          RC1  RC2
## SS loadings      2.25 1.90
## Proportion Var    0.38 0.32
## Cumulative Var    0.38 0.69
## Proportion Explained 0.54 0.46
## Cumulative Proportion 0.54 1.00
##
## Mean item complexity = 1.2
## Test of the hypothesis that 2 components are sufficient.
##
## The root mean square of the residuals (RMSR) is 0.11
## with the empirical chi square 22.04 with prob < 2e-04
##
## Fit based upon off diagonal values = 0.93
```

Factor Rotation

```
plot(principal(exam, nfactors=2))
```



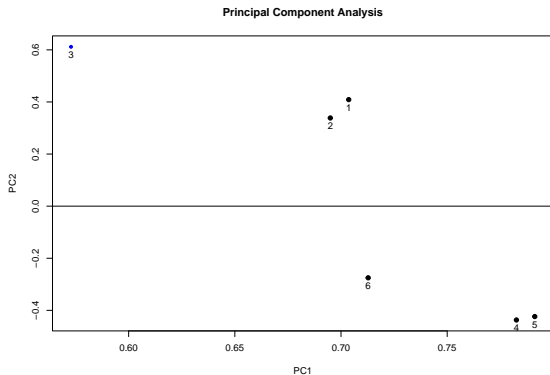
Factor Rotation

```
principal(exam,nfactors=2, rotate="none")
```

```
## Principal Components Analysis
## Call: principal(r = exam, nfactors = 2, rotate = "none")
## Standardized loadings (pattern matrix) based upon correlation matrix
##          PC1  PC2  h2  u2 com
## Gaelic    0.70  0.41 0.66 0.34 1.6
## English    0.69  0.34 0.60 0.40 1.4
## History    0.57  0.61 0.70 0.30 2.0
## Arithmetic 0.78 -0.44 0.80 0.20 1.6
## Algebra    0.79 -0.42 0.81 0.19 1.5
## Geometry   0.71 -0.28 0.58 0.42 1.3
##
##          PC1  PC2
## SS loadings      3.05 1.10
## Proportion Var    0.51 0.18
## Cumulative Var    0.51 0.69
## Proportion Explained 0.73 0.27
## Cumulative Proportion 0.73 1.00
##
## Mean item complexity = 1.6
## Test of the hypothesis that 2 components are sufficient.
##
## The root mean square of the residuals (RMSR) is 0.11
## with the empirical chi square 22.04 with prob < 2e-04
##
## Fit based upon off diagonal values = 0.93
```

Factor Rotation

```
plot(principal(exam,nfactors=2, rotate="none"))
```



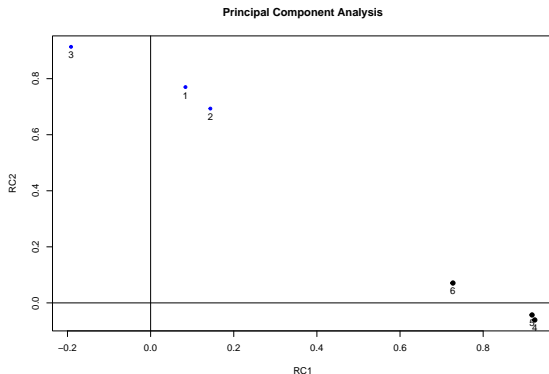
Factor Rotation

```
principal(exam,nfactors=2, rotate="promax")
```

```
## Principal Components Analysis
## Call: principal(r = exam, nfactors = 2, rotate = "promax")
## Standardized loadings (pattern matrix) based upon correlation matrix
##          RC1  RC2  h2  u2 com
## Gaelic      0.08  0.77 0.66 0.34 1.0
## English      0.14  0.69 0.60 0.40 1.1
## History     -0.19  0.91 0.70 0.30 1.1
## Arithmetic   0.92 -0.06 0.80 0.20 1.0
## Algebra      0.92 -0.04 0.81 0.19 1.0
## Geometry     0.73  0.07 0.58 0.42 1.0
##
##          RC1  RC2
## SS loadings      2.26 1.89
## Proportion Var    0.38 0.32
## Cumulative Var    0.38 0.69
## Proportion Explained 0.54 0.46
## Cumulative Proportion 0.54 1.00
##
## With component correlations of
##          RC1  RC2
## RC1 1.00 0.48
## RC2 0.48 1.00
##
## Mean item complexity = 1
## Test of the hypothesis that 2 components are sufficient.
##
```

Factor Rotation

```
plot(principal(exam,nfactors=2, rotate="promax"))
```



Section 6

Computation

Subsection 1

Method 1: Use PCA

Method 1: Use PCA

- By the spectral decomposition of Σ we have

$$\Sigma = \Gamma \Lambda \Gamma^T$$

where $\Gamma = (\gamma_1, \dots, \gamma_p)$ is an orthogonal matrix and $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$ be the diagonal matrix of eigenvalues.

- The spectral decomposition can be rewritten to

$$\Sigma = \sum_{i=1}^p \lambda_i \gamma_i \gamma_i^T = \sum_{i=1}^p (\sqrt{\lambda_i} \gamma_i)(\sqrt{\lambda_i} \gamma_i)^T$$

Subsection 2

Method 1: Use PCA

Method 1: Use PCA

- Suppose that $\lambda_m, \lambda_{m+1}, \dots, \lambda_p$ are small. Then

$$\mathbf{\Sigma} \approx \sum_{i=1}^m (\sqrt{\lambda_i} \gamma_i) (\sqrt{\lambda_i} \gamma_i)^T$$

- Let $\mathbf{L} = (\sqrt{\lambda_1} \gamma_1, \dots, \sqrt{\lambda_m} \gamma_m)$
- Let $\mathbf{\Psi} = \mathbf{\Sigma} - \mathbf{L} \mathbf{L}^T$

Subsection 3

Method 2: MLE

Method 2: MLE

- We impose multivariate normality on the common and specific factors

$$\mathbf{f} \sim N(\mathbf{0}, \mathbf{I}), \epsilon \sim N(\mathbf{0}, \Psi)$$

- The log-likelihood is

$$l(\mu, \mathbf{L}, \Psi) = -\frac{np}{2} \log 2\pi - \frac{n}{2} \log |\mathbf{L}\mathbf{L}^T + \Psi| - \frac{1}{2} \sum_{i=1}^n (\mathbf{X}_i - \mu)^T (\mathbf{L}\mathbf{L}^T + \Psi)^{-1} (\mathbf{X}_i - \mu)$$

where $\mathbf{X}_i \in \mathbb{R}^p$ denotes the i observation (not the i th feature).

The Number of Common Factors

- The $p \times p$ covariance matrix Σ is symmetric. As a result, there are $\frac{p(p+1)}{2}$ parameters.
- A factor model imposes a structure on Σ
- For a FA model with m common factors
- A FA model a small number common factors, i.e., when m is small, the model uses fewer parameters
 - the model is more parsimonious
 - the model might not be adequate if m is too small
- One can test whether an m is large enough

Choose m of Factors Computed using PCA

- Cumulative variance explained is should be reasonably large, such as $>80\%$
- Look for elbow from the scree plot

A Goodness of Fit Test for the Adequacy of the Number of Common Factors

```
knitr::include_graphics("img/TestNumberCommonFactors.png")
```

$$H_0: \underset{(p \times p)}{\Sigma} = \underset{(p \times m)}{\mathbf{L}} \underset{(m \times p)}{\mathbf{L}'} + \underset{(p \times p)}{\Psi}$$

$$-2 \ln \Lambda = -2 \ln \left[\frac{\text{maximized likelihood under } H_0}{\text{maximized likelihood}} \right]$$

$$= -2 \ln \left(\frac{|\hat{\Sigma}|}{|\mathbf{S}_n|} \right)^{-n/2} + n [\text{tr}(\hat{\Sigma}^{-1} \mathbf{S}_n) - p]$$

It can be shown that $\text{tr}(\hat{\Sigma}^{-1} \mathbf{S}_n) - p = 0$

$$= n \ln \left(\frac{|\hat{\Sigma}|}{|\mathbf{S}_n|} \right)$$

Test for the Adequacy of the Number of Common Factors

- The number of parameters for covariance in the full model is

$$\frac{p(p+1)}{2}$$

- The number of parameters for covariance in the reduced model is

$$mp + p - \frac{m(m-1)}{2}$$

Note: $-\frac{m(m-1)}{2}$ is due to the nonuniqueness of \mathbf{L} .

- The difference is

$$\begin{aligned} df &= \frac{p(p+1)}{2} - \left[mp + p - \frac{m(m-1)}{2} \right] \\ &= \frac{1}{2}[(p-m)^2 - p - m] \end{aligned}$$

Method 2: MLE

Test for the Adequacy of the Number of Common Factors

- The result indicates that 1 factor might be inadequate as the p-value is small.

```
factanal(exam, factors=1)
```

```
##
## Call:
## factanal(x = exam, factors = 1)
##
## Uniquenesses:
##      Gaelic      English      History Arithmetic      Algebra      Geometry
##      0.714      0.752      0.876      0.315      0.289      0.593
##
## Loadings:
##              Factor1
## Gaelic      0.535
## English     0.498
## History     0.353
## Arithmetic  0.828
## Algebra     0.843
## Geometry    0.638
##
##              Factor1
## SS loadings    2.462
## Proportion Var 0.410
##
```

Method 2: MLE

Test for the Adequacy of the Number of Common Factors

- The result indicates that 2 factors is adequate because the fit is not substantially from the full model.

```
factanal(exam, factors=2)
```

```
##
## Call:
## factanal(x = exam, factors = 2)
##
## Uniquenesses:
##      Gaelic      English      History      Arithmetic      Algebra      Geometry
##      0.478       0.554       0.534       0.298       0.224       0.607
##
## Loadings:
##              Factor1 Factor2
## Gaelic      0.323   0.646
## English     0.275   0.608
## History     0.000   0.676
## Arithmetic  0.811   0.211
## Algebra     0.855   0.213
## Geometry    0.554   0.294
##
##              Factor1 Factor2
## SS loadings  1.882   1.421
## Proportion Var 0.314   0.237
## Cumulative Var 0.314   0.551
##
```

Section 7

Confirmatory FA

Exploratory or Confirmatory Factor Analysis

- Exploratory Factor Analysis (EFA)
- The FA approach we have discussed is exploratory in nature.
- In fact, we can perform EFA and identify latent factors by using only correlations, not the data
- The purpose of EFA is to explore the possible underlying structure that can explain the observed correlations
- EFA is used when researchers do not have a specific idea about the underlying structure of data
- EFA is hypothesis-generating

Exploratory or Confirmatory Factor Analysis

- Confirmatory Factor Analysis (CFA) is used when a researcher has specific hypotheses or theories about the factor structure of their data.
- It is a “theory-driven” approach.
- In CFA, the researcher specifies the number of factors and which variables load onto which factors.
- CFA is typically used in later stages of research to test or confirm the factor structure suggested by EFA
- CFA is hypothesis testing. A pre-specified model is required

Exploratory or Confirmatory Factor Analysis

- Use EFA when:
 - You are unsure about the underlying structure.
 - You aim to uncover complex patterns.
 - You need to form hypotheses and develop theory.
- Use CFA when:
 - You have a predetermined theory or model.
 - You aim to test the hypothesis about the factor structure.
 - You need to confirm or disconfirm theories.