

# Multivariate Analysis Lecture 3: Random Vectors and Random Samples

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## Section 1

### Outline

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- ① Introduction
- ① Random variables and random samples
- ② Random vectors and random samples
- ③ Linear combinations of random vectors

## Section 2

### Introduction

# Random ???

$$\mathbf{X} = \begin{pmatrix} X_{11} & X_{12} & \cdots & X_{1j} & \cdots & X_{1p} \\ X_{21} & X_{22} & \cdots & X_{2j} & \cdots & X_{2p} \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ X_{i1} & X_{i2} & \cdots & X_{ij} & \cdots & X_{ip} \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ X_{n1} & X_{n2} & \cdots & X_{nj} & \cdots & X_{np} \end{pmatrix}$$

- $X_{ij}$ , e.g.,  $X_{12}$ , is a random variable.
- $(X_{i1}, \cdots, X_{ip})$ , is a random vector (row vector).
- $(X_{1j}, \cdots, X_{nj})^T$  is a random vector (column vector).
- $\mathbf{X}$  is a random matrix.

# Assumption

- A typical scenario for the  $n \times p$  matrix  $\mathbf{X}$  is that we have  $n$  independent and identically distributed (iid) rows
- Each row follows a  $p$ -dimensional multivariate distribution (say, multivariate normal)

# A Univariate Random Sample

- Consider a specific column, say the  $j$ th column.

$$\mathbf{X} = \begin{pmatrix} X_{11} & X_{12} & \cdots & X_{1j} & \cdots & X_{1p} \\ X_{21} & X_{22} & \cdots & X_{2j} & \cdots & X_{2p} \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ X_{i1} & X_{i2} & \cdots & X_{ij} & \cdots & X_{ip} \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ X_{n1} & X_{n2} & \cdots & X_{nj} & \cdots & X_{np} \end{pmatrix}$$

- $X_{1j}, X_{2j}, \dots, X_{nj}$  are independent and identically (iid) distributed random variables.
- We say  $X_{1j}, X_{2j}, \dots, X_{nj}$  is a random sample from a univariate distribution, the distribution of the  $j$ th variable/feature.
- In other words, each column is a random sample from a univariate distribution.

# A Multivariate Random Sample

$$\mathbf{X} = \begin{pmatrix} X_{11} & X_{12} & \cdots & X_{1j} & \cdots & X_{1p} \\ X_{21} & X_{22} & \cdots & X_{2j} & \cdots & X_{2p} \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ \color{red}{X_{i1}} & \color{red}{X_{i2}} & \cdots & \color{red}{X_{ij}} & \cdots & \color{red}{X_{ip}} \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ X_{n1} & X_{n2} & \cdots & X_{nj} & \cdots & X_{np} \end{pmatrix}$$

- $\mathbf{X}$  has  $n$  rows and  $p$  columns. Each row is from a  $p$ -dimensional multivariate distribution.
- The rows are independent and identically distributed random vectors.
- We say the  $n$  rows are a random sample from a  $p$ -dimensional multivariate distribution.



## Section 3

### Review: Univariate R.V.

## Subsection 1

# Random Variables

# What Is a Random Variable?

- A random variable is a numerical quantity that takes on different values with certain probabilities.
- e.g., a normal distributed random variable takes values between  $-\infty$  to  $\infty$ .
- It represents the outcome of a random event or experiment.
- e.g., the BMI of a randomly chosen adult living in Canada
- Random variables can be discrete or continuous.

# The Mean of a Random Variable

- The mean of a random variable  $X$  measures its central tendency, often denoted by  $\mu$  or  $E(X)$ .
- It is the expected value of the random variable, weighted by the probabilities of each possible outcome:
  - Continuous:  $\mu = E(X) \stackrel{\text{def}}{=} \int_{-\infty}^{\infty} xf(x)dx$
  - Discrete:  $\mu = E(X) \stackrel{\text{def}}{=} \sum_{i=1} x_i p_i$
- $E(aX + b) = aE(X) + b$ , where  $X$  is random and  $a$  and  $b$  are fixed.

# Variance of a Random Variable

- The variance of a random variable is a measure of how spread out its values are around the mean.
- It represents the expected value of the squared deviation of the random variable from its mean.  $\sigma^2 \stackrel{\text{def}}{=} E[(X - \mu)^2]$ , specifically,
  - Continuous:  $\sigma^2 \stackrel{\text{def}}{=} \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$
  - Discrete:  $\sigma^2 \stackrel{\text{def}}{=} \sum_{i=1} (x_i - \mu)^2 p_i$
- $\sigma$ , the square root of the variance, is called the standard deviation (SD) of  $X$ .

# Properties of Variance

- The variance is a non-negative quantity.
- The variance of a constant is 0:  $\text{Var}(c) = 0$ , where  $c$  is a constant.
- The variance is affected by changes in the scale of the random variable but not by a shift in locations:  
 $\text{Var}(aX + b) = a^2 \text{Var}(X)$ , where  $a$  is a constant.
- The variance of a sum of **independent** random variables is the sum of their individual variances:  
 $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$ , provided that  $X$  and  $Y$  are independent. More general, if  $X_1, \dots, X_n$  are mutually independent, then  $\text{Var}(\sum_{i=1}^n X_i) = \sum_{i=1}^n \text{Var}(X_i)$ .

## Subsection 2

### A Random Sample of Random Variables

## Random Samples (from Simple Random Sampling)

- In a simple random sample, each member of the population is selected independently and with equal probability.
- Obtaining a truly random sample can often be challenging.

Reasons:

- it may be difficult or impossible to obtain a complete list of all members of the population of interest.
  - it may be costly or time-consuming to sample from the entire population.
  - there may be practical constraints on the sampling process, such as geographic distance, language barriers, or legal restrictions.
  - certain subgroups of the population may be underrepresented or difficult to reach, leading to potential biases in the sample.
- Nevertheless, we assume the samples are simple random samples for theoretical derivations



# Sample Mean and Variance from a Simple Random Samples

- Let  $(X_1, \dots, X_n)$  be a simple random sample from a distribution with mean  $\mu$  and variance  $\sigma^2$ . The notation we will use is

$$X_1, \dots, X_n \stackrel{iid}{\sim} (\mu, \sigma^2)$$

- Summary Statistics and their Expectations:
  - The sample mean  $\bar{X}$  is defined as  $\bar{X} \stackrel{def}{=} \frac{1}{n} \sum_{i=1}^n X_i$ .
  - $\bar{X}$  is unbiased for  $\mu$ , i.e.,  $E(\bar{X}) = \mu$ .  $Var(\bar{X}) = \sigma^2/n$ .
  - The sample variance  $S^2 \stackrel{def}{=} \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$ .
  - $S^2$  is unbiased for  $\sigma^2$ , i.e.,  $E(S^2) = \sigma^2$ .

# Sample Mean is Unbiased

- The proof of unbiasedness follows from the linearity of the expected value operator:

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \sum_{i=1}^n \mu = \mu$$

- The unbiasedness of the sample mean is a fundamental property of statistical estimation.

# The Variance of the Sample Mean

$$\begin{aligned} \text{Var}(\bar{X}) &= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n X_i\right) \\ &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{\sigma^2}{n} \end{aligned}$$

- The variability of the sample means decreases as the sample size increases.
- The result is important for the design of experiments and surveys. E.g., what is a minimum sample size to achieve a desired level of precision?

# Sample Variance is Unbiased

- The proof of unbiasedness follows from the properties of the variance operator and the linearity of the expected value operator:

$$\begin{aligned} E(S^2) &= \frac{1}{n-1} \sum_{i=1}^n E[(X_i - \bar{X})^2] = \frac{1}{n-1} \sum_{i=1}^n E[(X_i - \mu + \mu - \bar{X})^2] \\ &= \frac{1}{n-1} \sum_{i=1}^n E[(X_i - \mu)^2 + 2(X_i - \mu)(\mu - \bar{X}) + (\mu - \bar{X})^2] \\ &= \frac{1}{n-1} [n\sigma^2 - 2nE[(\mu - \bar{X})^2] + nE[(\mu - \bar{X})^2]] \\ &= \frac{1}{n-1} (n-1)\sigma^2 = \sigma^2 \end{aligned}$$

## Section 4

# Multivariate R.V.

## Subsection 1

### Random Vectors

# Notations for Random **Vectors**

- A random vector is a vector whose elements are random variables. e.g.,

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix}$$

where each  $X_i$  is a random variable

# The Expectation of A Random Vector

- Let  $E(\mathbf{X})$  denote the mean vector of  $\mathbf{X}_{p \times 1}$ . We have

$$\boldsymbol{\mu} = E(\mathbf{X}) \stackrel{\text{def}}{=} \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_p \end{pmatrix},$$

where  $\mu_i = E(X_i), i = 1, \dots, p$ .

- Suppose  $X_1, \dots, X_n \stackrel{iid}{\sim} (\mu, \sigma^2)$ , and  $\mathbf{X} = (X_1, \dots, X_n)^T$ . What is  $E(\mathbf{X})$ ?

$$E(\mathbf{X}) = E\left[\begin{pmatrix} X_1 \\ \vdots \\ X_n \end{pmatrix}\right] = \begin{pmatrix} \mu \\ \vdots \\ \mu \end{pmatrix} = \mu \mathbf{1}$$



# The Variance-Covariance of A Random Vector

- The variance-covariance matrix of a random vector  $\mathbf{X}$  is a square matrix that summarizes the variability and dependence among its components.
- It is denoted by the symbol  $\text{Var}(\mathbf{X})$ ,  $\text{Cov}(\mathbf{X})$ , or  $\Sigma$  and is given by:

$$\Sigma \stackrel{\text{def}}{=} E[(\mathbf{X} - \boldsymbol{\mu})(\mathbf{X} - \boldsymbol{\mu})^T]$$
$$= \begin{bmatrix} \text{Var}(X_1) & \text{Cov}(X_1, X_2) & \cdots & \text{Cov}(X_1, X_p) \\ \text{Cov}(X_2, X_1) & \text{Var}(X_2) & \cdots & \text{Cov}(X_2, X_p) \\ \vdots & \vdots & \ddots & \vdots \\ \text{Cov}(X_p, X_1) & \text{Cov}(X_n, X_2) & \cdots & \text{Var}(X_p) \end{bmatrix}$$

# The Variance-Covariance of A Random Vector

- Alternative notations

$$\text{Var}(\mathbf{X}) = \Sigma = (\sigma_{ij}) \stackrel{\text{def}}{=} \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1p} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p1} & \sigma_{p2} & \cdots & \sigma_p^2 \end{bmatrix}$$

- Remarks

- The covariance between two components measures how much they vary together, and it can be positive, negative, or zero.
- $\Sigma$  is a symmetric matrix because  $\sigma_{ij} = \text{Cov}(X_i, X_j) = \sigma_{ji}$ .
- The diagonal elements of  $\Sigma$  represent the variances of the components of the random vector:  
 $\sigma_i^2 = \text{Var}(X_i) = \text{Cov}(X_i, X_i)$ .

# Correlation Matrix

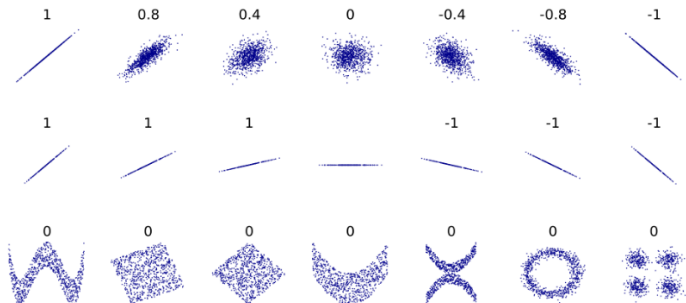
- A correlation matrix is a table showing correlation coefficients between different variables.
- The correlation coefficient measures the strength and direction of the linear relationship between two variables.

$$\text{Corr}(X_i, X_j) = \frac{\text{Cov}(X_i, X_j)}{\sqrt{\text{Var}(X_i)}\sqrt{\text{Var}(X_j)}} = \frac{\sigma_{ij}}{\sigma_i\sigma_j}$$

- The correlation coefficient ranges from -1 to 1 (by Cauchy-Schwarz inequality).
- Values close to -1 indicating a strong negative linear relationship, values close to 1 indicating a strong positive linear relationship, and values close to 0 indicating no linear relationship.

# Examples of Correlations

- Correlations are about linear relationship!



# Correlation Matrix

$$\mathbf{R} = \begin{bmatrix} 1 & \text{Corr}(X_1, X_2) & \cdots & \text{Corr}(X_1, X_p) \\ \text{Corr}(X_2, X_1) & 1 & \cdots & \text{Corr}(X_2, X_p) \\ \vdots & \vdots & \ddots & \vdots \\ \text{Corr}(X_p, X_1) & \text{Corr}(X_p, X_2) & \cdots & 1 \end{bmatrix}$$

-  $\rho_{ij} \stackrel{\text{def}}{=} \text{Corr}(X_i, X_j)$

- The diagonal  $\rho_{ii}$  of the correlation matrix shows the correlation of each variable with itself, which is always equal to **1**
- The matrix is symmetric since the correlation between X and Y is the same as the correlation between Y and X:  $\rho_{ij} = \rho_{ji}$ .
- Correlation matrix can help identify variables that are correlated.

# Covariance Matrix of Two Random Vectors

- The covariance matrix of two random vectors  $\mathbf{X} = (X_1, \dots, X_p)^T$  and  $\mathbf{Y} = (Y_1, \dots, Y_q)^T$  is a  $p \times q$  matrix defined as

$$\begin{aligned}\mathbf{Cov}(\mathbf{X}, \mathbf{Y}) &\stackrel{\text{def}}{=} E[(\mathbf{X} - \boldsymbol{\mu}_X)(\mathbf{Y} - \boldsymbol{\mu}_Y)^T] \\ &= \begin{bmatrix} \text{Cov}(X_1, Y_1) & \cdots & \text{Cov}(X_1, Y_q) \\ \vdots & \ddots & \vdots \\ \text{Cov}(X_p, Y_1) & \cdots & \text{Cov}(X_p, Y_q) \end{bmatrix}\end{aligned}$$

- Each element of the matrix is the covariance between two corresponding elements of the vectors.

# Covariance Matrix of Two Random Vectors

- E.g.,

$$\mathbf{X}_{2 \times 1} = \begin{pmatrix} X_1 & X_2 \end{pmatrix}^T, \mathbf{Y}_{3 \times 1} = \begin{pmatrix} Y_1 & Y_2 & Y_3 \end{pmatrix}^T$$

$$\begin{aligned} \text{Cov}(\mathbf{X}, \mathbf{Y}) &= E[(\mathbf{X} - E(\mathbf{X}))(\mathbf{Y} - E(\mathbf{Y}))^T] \\ &= E\left[\begin{pmatrix} X_1 - \mu_{x1} \\ X_2 - \mu_{x2} \end{pmatrix} (Y_1 - \mu_{y1}, Y_2 - \mu_{y2}, Y_3 - \mu_{y3})\right] \\ &= \begin{bmatrix} E[(X_1 - \mu_{x1})(Y_1 - \mu_{y1})] & E[(X_1 - \mu_{x1})(Y_2 - \mu_{y2})] & E[(X_1 - \mu_{x1})(Y_3 - \mu_{y3})] \\ E[(X_2 - \mu_{x2})(Y_1 - \mu_{y1})] & E[(X_2 - \mu_{x2})(Y_2 - \mu_{y2})] & E[(X_2 - \mu_{x2})(Y_3 - \mu_{y3})] \end{bmatrix} \\ &= \begin{bmatrix} \text{Cov}(X_1, Y_1) & \text{Cov}(X_1, Y_2) & \text{Cov}(X_1, Y_3) \\ \text{Cov}(X_2, Y_1) & \text{Cov}(X_2, Y_2) & \text{Cov}(X_2, Y_3) \end{bmatrix} \end{aligned}$$

- Note:  $\text{Cov}(\mathbf{X}, \mathbf{Y}) = [\text{Cov}(\mathbf{Y}, \mathbf{X})]^T$

## Subsection 2

### A Random Sample of Random Vectors



# Notations about A Random Sample of Random Vectors

- Consider a random sample from a **multivariate** distribution with mean vector  $\boldsymbol{\mu}_{p \times 1}$  and covariance  $\boldsymbol{\Sigma}_{p \times p}$
- A random sample of random vectors is a collection of  $n$  independent and identically distributed random vectors, denoted as  $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n$ .
- The random sample of random vectors is denoted by

$$\mathbf{X}_{n \times p} = \begin{pmatrix} \mathbf{X}_1^T \\ \vdots \\ \mathbf{X}_n^T \end{pmatrix}$$

- Each random vector  $\mathbf{X}_i$  is of dimension  $p$  and can be represented as:

$$\mathbf{X}_i = (X_{i1}, X_{i2}, \dots, X_{ip})^T$$

# Sample Mean Vector $\bar{\mathbf{X}}_{p \times 1}$

- The sample mean vector, denoted as  $\bar{\mathbf{X}}$ , is a random vector of dimension  $p$ , defined as:

$$\bar{\mathbf{X}} = \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i$$

- It is unbiased for the population mean vector  $\boldsymbol{\mu}$  because

$$E[\bar{\mathbf{X}}] = E\left[\frac{1}{n} \sum_{i=1}^n \mathbf{X}_i\right] = \frac{1}{n} \sum_{i=1}^n E[\mathbf{X}_i] = \frac{1}{n} \sum_{i=1}^n \boldsymbol{\mu} = \boldsymbol{\mu}$$

- The sample mean vector  $\bar{\mathbf{X}}$  is often used to estimate the population mean vector  $\boldsymbol{\mu}$ .

# The Covariance of the Sample Mean Vector

- The sample mean vector, denoted as  $\bar{\mathbf{X}}$ , is a random vector of dimension  $p$ . We can also compute its covariance matrix
- Because  $(\mathbf{X}_1, \dots, \mathbf{X}_n)$  are iid ,

$$\text{Cov}(\bar{\mathbf{X}}) = \text{Cov}\left(\frac{1}{n} \sum_{i=1}^n \mathbf{X}_i\right) = \frac{1}{n^2} \sum_{i=1}^n \text{Cov}(\mathbf{X}_i) = \frac{1}{n} \boldsymbol{\Sigma}$$

- Similar to the population mean vector, the population covariance  $\boldsymbol{\Sigma}$  is typically unknown. If we have a random sample, we can estimate it - the sample covariance matrix.S

# Sample Covariance Matrix $\mathbf{S}_{p \times p}$

- The sample covariance matrix, denoted as  $\mathbf{S}$ , is a  $p \times p$  symmetric matrix, defined as:

$$\mathbf{S} = \frac{1}{n-1} \sum_{i=1}^n (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})^T$$

- Next, we show that the sample covariance matrix  $\mathbf{S}$  is an unbiased estimator of  $\mathbf{\Sigma}$ :

$$\mathbb{E}[\mathbf{S}] = \mathbf{\Sigma}$$

# The Sample Covariance Matrix is Unbiased: Lemmas

- Lemma 1:  $E(\mathbf{X}_i \mathbf{X}_i^T) = \mu \mu^T + \text{Cov}(\mathbf{X}_i) = \mu \mu^T + \Sigma$ . -Proof.  
By the definition of Cov, we have

$$\begin{aligned}\Sigma &= E[(\mathbf{X}_i - \mu)(\mathbf{X}_i - \mu)^T] \\ &= E[\mathbf{X}_i \mathbf{X}_i^T - \mu \mathbf{X}_i^T - \mathbf{X}_i \mu^T + \mu \mu^T] \\ &= E[\mathbf{X}_i \mathbf{X}_i^T] - \mu E[\mathbf{X}_i^T] - E[\mathbf{X}_i] \mu^T + \mu \mu^T \\ &= E[\mathbf{X}_i \mathbf{X}_i^T] - \mu \mu^T - \mu \mu^T + \mu \mu^T \\ &= E[\mathbf{X}_i \mathbf{X}_i^T] - \mu \mu^T\end{aligned}$$

As a result,  $E[\mathbf{X}_i \mathbf{X}_i^T] = \mu \mu^T + \Sigma$ .

- Similarly, we have Lemma 2:

$$E(\bar{\mathbf{X}} \bar{\mathbf{X}}^T) = \mu \mu^T + \text{Cov}(\bar{\mathbf{X}}) = \mu \mu^T + \frac{1}{n} \Sigma$$

# The Sample Covariance Matrix is Unbiased: Proof

- Proof: Expand the product:

$$\begin{aligned}\mathbf{S} &= \frac{1}{n-1} \sum_{i=1}^n (\mathbf{x}_i \mathbf{x}_i^T - \mathbf{x}_i \bar{\mathbf{x}}^T - \bar{\mathbf{x}} \mathbf{x}_i^T + \bar{\mathbf{x}} \bar{\mathbf{x}}^T) \\ &= \frac{1}{n-1} \left[ \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i^T - n \bar{\mathbf{x}} \bar{\mathbf{x}}^T - n \bar{\mathbf{x}} \bar{\mathbf{x}}^T + n \bar{\mathbf{x}} \bar{\mathbf{x}}^T \right] \\ &= \frac{1}{n-1} \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i^T - \frac{n}{n-1} \bar{\mathbf{x}} \bar{\mathbf{x}}^T\end{aligned}$$

# The Sample Covariance Matrix is Unbiased: Proof (continued)

- Taking the expected value:

$$\begin{aligned}\mathbb{E}[\mathbf{S}] &= \mathbb{E}\left[\frac{1}{n-1} \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i^T - \frac{n}{n-1} \bar{\mathbf{x}} \bar{\mathbf{x}}^T\right] \\ &= \frac{1}{n-1} \sum_{i=1}^n \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T] - \frac{n}{n-1} \mathbb{E}[\bar{\mathbf{x}} \bar{\mathbf{x}}^T] \\ &= \frac{1}{n-1} \sum_{i=1}^n (\boldsymbol{\Sigma} + \boldsymbol{\mu} \boldsymbol{\mu}^T) - \frac{n}{n-1} \left(\frac{1}{n} \boldsymbol{\Sigma} + \boldsymbol{\mu} \boldsymbol{\mu}^T\right) \\ &= \frac{n}{n-1} \boldsymbol{\Sigma} + \frac{n}{n-1} \boldsymbol{\mu} \boldsymbol{\mu}^T - \frac{1}{n-1} \boldsymbol{\Sigma} - \frac{n}{n-1} \boldsymbol{\mu} \boldsymbol{\mu}^T \\ &= \boldsymbol{\Sigma}\end{aligned}$$

- Therefore, the sample covariance matrix is unbiased

# Examples: The Iris Setosa Data

- The iris data consists of three random samples, one for each species. Consider the setosa sample.
- It is a random sample (let's assume it) of size 50.
- The data matrix has  $n = 50$  rows and  $p = 4$  columns



# The Data Matrix of Iris Setosa

```
setosa=as.matrix(iris[iris$Species=="setosa", 1:4])  
dim(setosa)
```

```
## [1] 50  4
```

```
head(setosa)
```

```
##      Sepal.Length Sepal.Width Petal.Length Petal.Width  
## 1           5.1           3.5           1.4           0.2  
## 2           4.9           3.0           1.4           0.2  
## 3           4.7           3.2           1.3           0.2  
## 4           4.6           3.1           1.5           0.2  
## 5           5.0           3.6           1.4           0.2  
## 6           5.4           3.9           1.7           0.4
```

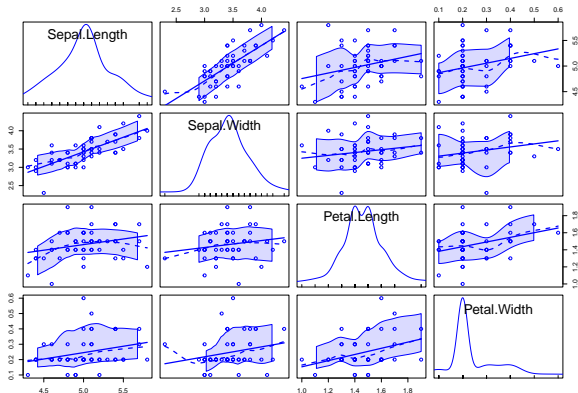
# The Sample Mean of Iris Setosa

```
sample.meanvec=matrix(colMeans(setosa), 4, 1)
rownames(sample.meanvec)=colnames(setosa)
colnames(sample.meanvec)="mean"
sample.meanvec
```

```
##              mean
## Sepal.Length 5.006
## Sepal.Width  3.428
## Petal.Length 1.462
## Petal.Width  0.246
```

# Pairwise Scatter Plot of the Features of Iris Setosa

```
scatterplotMatrix(setosa)
```



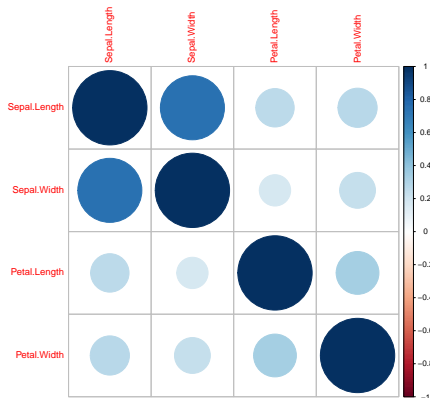
# The Sample Covariance Matrix of Iris Setosa

```
sample.cov=cov(setosa)  
round(sample.cov,2)
```

| ##              | Sepal.Length | Sepal.Width | Petal.Length | Petal.Width |
|-----------------|--------------|-------------|--------------|-------------|
| ## Sepal.Length | 0.12         | 0.10        | 0.02         | 0.01        |
| ## Sepal.Width  | 0.10         | 0.14        | 0.01         | 0.01        |
| ## Petal.Length | 0.02         | 0.01        | 0.03         | 0.01        |
| ## Petal.Width  | 0.01         | 0.01        | 0.01         | 0.01        |

# The Sample Correlation Matrix of Iris Setosa

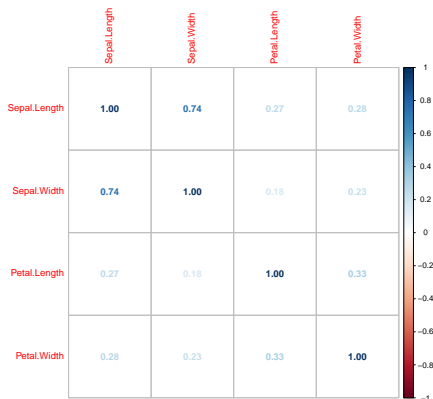
```
sample.corr=cor(setosa)  
corrplot(sample.corr)
```



## A Random Sample of Random Vectors

## The Sample Correlation Matrix of Iris Setosa

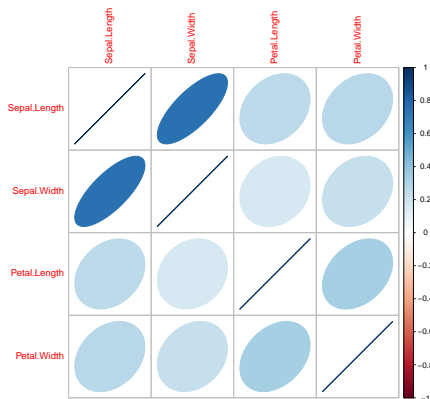
```
corrplot(sample.corr, method="number")
```



## A Random Sample of Random Vectors

## The Sample Correlation Matrix of Iris Setosa

```
corrplot(sample.corr, method="ellipse")
```



# Sample Covariate Matrix as a Quadratic Form

$$(n-1)S = \sum (X_i - \bar{X})(X_i - \bar{X})^T = (X_1 - \bar{X} \quad \cdots \quad X_n - \bar{X}) \begin{pmatrix} (X_1 - \bar{X})^T \\ \vdots \\ (X_n - \bar{X})^T \end{pmatrix}$$

Note that

$$\begin{pmatrix} (X_1 - \bar{X})^T \\ \vdots \\ (X_n - \bar{X})^T \end{pmatrix} = \begin{pmatrix} X_1 - \bar{X} \\ \vdots \\ X_n - \bar{X} \end{pmatrix}^T = \mathbf{C}\mathbf{X}$$

where  $\mathbf{C}$  is the centering matrix defined in assignment 1, i.e.,  $\mathbf{C} = \mathbf{I}_n - \frac{1}{n}\mathbf{1}\mathbf{1}^T$ . In addition, it can be verified that  $\mathbf{C}^T\mathbf{C} = \mathbf{C}$ .

Therefore,

$$(n-1)S = (\mathbf{C}\mathbf{X})^T\mathbf{C}\mathbf{X} = \mathbf{X}^T\mathbf{C}\mathbf{X}$$



## Section 5

# Linear Combinations

## Subsection 1

## Definition

# Definition of a Linear Combination of a Random Vector

- Let  $\mathbf{X}$  be a  $p$ -dimensional random vector with mean vector  $\boldsymbol{\mu}$  and covariance matrix  $\boldsymbol{\Sigma}$ .
- Consider a linear combination of the form:

$$Y = \mathbf{a}^T \mathbf{X}$$

where  $\mathbf{a}$  is a  $p$ -dimensional constant vector.

- E.g.,  $\mathbf{X} = (X_1, X_2, X_3)^T$ ,  $\mathbf{a} = (1/3, 1/3, 1/3)^T$ . Then

$$Y = \mathbf{a}^T \mathbf{X} = \frac{1}{3}(X_1 + X_2 + X_3)$$

## Subsection 2

Mean

# Mean of $Y = \mathbf{a}^T \mathbf{X}$

- The mean of  $Y$  can be expressed as:

$$\begin{aligned} E(Y) &= E(\mathbf{a}^T \mathbf{X}) \\ &= \mathbf{a}^T E(\mathbf{X}) \\ &= \mathbf{a}^T \boldsymbol{\mu} \end{aligned}$$

- Intuitively, the mean of  $Y$  is a weighted average of the components of  $\mathbf{X}$ , with weights given by the corresponding components of  $\mathbf{a}$ .

## Subsection 3

### Variance

# Variance of $Y$

- The variance of  $Y$  can be expressed as:

$$\begin{aligned}\text{Var}(Y) &= \text{Var}(\mathbf{a}^T \mathbf{X}) \\ &= E[(\mathbf{a}^T \mathbf{X} - \mathbf{a}^T \boldsymbol{\mu})^2] \\ &= E[(\mathbf{a}^T \mathbf{X} - \mathbf{a}^T \boldsymbol{\mu})(\mathbf{a}^T \mathbf{X} - \mathbf{a}^T \boldsymbol{\mu})^T] \\ &= E[\mathbf{a}^T (\mathbf{X} - \boldsymbol{\mu})(\mathbf{X} - \boldsymbol{\mu})^T \mathbf{a}] \\ &= \mathbf{a}^T \boldsymbol{\Sigma} \mathbf{a}\end{aligned}$$

# Variance of $Y$ and Quadratic Forms

- The variance of  $Y$  depends on the covariance structure of  $\mathbf{X}$ , as well as the weights given by  $\mathbf{a}$ .
- We call forms like  $\mathbf{a}^T \boldsymbol{\Sigma} \mathbf{a}$  as quadratic forms.
- Note, we can also write the variance of  $Y$  as:

$$\mathbf{a}^T \boldsymbol{\Sigma} \mathbf{a}$$

$$= \sum_i \sum_j \sigma_{ij} a_i a_j$$

$$= \sigma_{11} a_1^2 + \sigma_{12} a_1 a_2 + \cdots + \sigma_{1p} a_1 a_p + \sigma_{21} a_2 a_1 + \cdots + \sigma_{p1} a_p a_1 + \cdots$$

$$= \sigma_{11} a_1^2 + \sigma_{22} a_2^2 + \cdots + \sigma_{pp} a_p^2 + 2(\sigma_{12} a_1 a_2 + \sigma_{13} a_1 a_3 + \cdots + \sigma_{p-1,p} a_{p-1} a_p)$$

- A quadratic form is a polynomial of degree 2 in the components of a vector, and it can be expressed as:

$$Q(\mathbf{X}) = \mathbf{X}^T \mathbf{A} \mathbf{X}$$



## Subsection 4

### Example

# Linear Combinations of Iris Setosa Features

- Recall that for the iris setosa,  $\mathbf{X}$  is  $50 \times 4$ .
- Consider a linear combination of the features  $Y = \mathbf{X}b$ , where

$$b = \begin{pmatrix} 1/4 \\ 1/4 \\ 1/4 \\ 1/4 \end{pmatrix}$$

- $Yb$  is a  $50 \times 1$  vector, with the  $i$ th row be the average of the four features of the  $i$ th iris setosa flower. To see this

## Example

## Linear Combinations of Iris Setosa Features

$$Y = \mathbf{X}b = \begin{pmatrix} X_1^T \\ \vdots \\ X_n^T \end{pmatrix} b = \begin{pmatrix} X_1^T b \\ \vdots \\ X_n^T b \end{pmatrix} = \begin{pmatrix} \frac{x_{11} + x_{12} + x_{13} + x_{14}}{4} \\ \vdots \\ \frac{x_{n1} + x_{n2} + x_{n3} + x_{n4}}{4} \end{pmatrix}$$

## Example

## Linear Combinations of Iris Setosa Features: sample mean

```
b=matrix(1/4, 4, 1)
Y=setosa%*%b
#sample mean of Y: the following two results are the same
mean(Y)
```

```
## [1] 2.5355
```

```
t(b)%*%sample.meanvec
```

```
##          mean
## [1,] 2.5355
```

## Example

# Linear Combinations of Iris Setosa Features: sample variance

```
#sample variance of Y: the following two results are the same  
var(Y)
```

```
##                [,1]  
## [1,] 0.03844617
```

```
t(b)%*%cov(setosa)%*%b
```

```
##                [,1]  
## [1,] 0.03844617
```