

IE 522 HW09

Q1

2 Points

For time series data, we often do a differencing and analyze changes (differences) instead of the original variables. This problem shows why. Let

$$X_i = 10 + t_i + \epsilon_i^X, \quad Y_{1i} = 20 + 2t_i + \epsilon_i^{Y_1}, \quad Y_{2i} = 20 + 2t_i + 5X_i + \epsilon_i^{Y_2},$$

where the error terms are i.i.d. with a $N(0, 1)$ distribution. The following generates 100 observations for X_i , Y_{1i} and Y_{2i} . Differencing leads to 99 observations for $\Delta X_i := X_{i+1} - X_i$, $\Delta Y_{1i} = Y_{1,i+1} - Y_{1i}$, $\Delta Y_{2i} = Y_{2,i+1} - Y_{2i}$. Assume $t_{i+1} - t_i = 1$.

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```
set.seed(1)
t=seq(1,100,by=1) #time
X =10+t+rnorm(100)
Y1 =20+2*t+rnorm(100)
Y2 =20+2*t+5*X+rnorm(100)
data = data.frame(X,Y1,Y2)
ddata=data.frame(apply(data,2,diff))
```

Q1.1

1 Point

When Y_1 and Y_2 are regressed on X , respectively, what R^2 's do you get? Other than a common time trend, are Y_1 and X related?

Call:

```
lm(formula = data$Y1 ~ data$X)
```

Multiple R-squared: 0.9988, Adjusted R-squared: 0.9988

Call:

```
lm(formula = data$Y2 ~ data$X)
```

Multiple R-squared: 0.9999, Adjusted R-squared: 0.9999



R_square for $Y1 \sim X$ = 0.9988

R_square for $Y2 \sim X$ = 0.9999

Without common time trend, X is $10 + \text{rnorm}$, and $Y1$ is $20 + \text{rnorm}$, which is independent (rnorm and rnorm is IID for each other).

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▼ Q1-1.R

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```
1 #Q1-1
2
3 #Data Preparation
4 set.seed(1)
5 t=seq(1,100,by=1) #time
6 X =10+t+rnorm(100)
7 Y1 =20+2*t+rnorm(100)
8 Y2 =20+2*t+5*X+rnorm(100)
9 data = data.frame(X,Y1,Y2)
10 ddata=data.frame(apply(data,2,diff))
11
12 #Regression
13 summary(lm(data$Y1 ~ data$X))
14 summary(lm(data$Y2 ~ data$X))
```

15

16

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1 Point

Write down the equations for ΔX_i , ΔY_{1i} , ΔY_{2i} . When ΔY_1 and ΔY_2 are regressed on ΔX , respectively, what R^2 's do you get? Why do the R^2 's change this way?

Call:

lm(formula = ddata\$Y1 ~ ddata\$X)

Multiple R-squared: 0.0001495, Adjusted R-squared: -0.01016

Call:

lm(formula = ddata\$Y2 ~ ddata\$X)

Multiple R-squared: 0.9465, Adjusted R-squared: 0.946

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There is linear correlation between Y_i and X , however not between ΔY_1 and ΔX , since the term that lefted (as the pdf shown) is iid for each other, so the R_{square} of $Y_1 \sim X$ fitting is reduced dramatically.

There is linear correlation between Y_2 and X , and as well between ΔY_2 and ΔX , since ΔY_2 contains the term ΔX in its equation. So the R_{square} of $Y_2 \sim X$ fitting is not change dramatically.

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▼ Q1-2.R


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```
1 #Q1-2
2
3 #Data Preparation
4 set.seed(1)
5 t=seq(1,100,by=1) #time
6 X =10+t+rnorm(100)
7 Y1 =20+2*t+rnorm(100)
```

```

8 Y2 =20+2*t+5*X+rnorm(100)
9 data = data.frame(X,Y1,Y2)
10 ddata=data.frame(apply(data,2,diff))
11
12 #Regression
13 summary(lm(ddata$Y1 ~ ddata$X))
14 summary(lm(ddata$Y2 ~ ddata$X))
15

```

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Q2

8 Points

The dataset USMacroG in the AER package contains quarterly macroeconomic variables from 1950 to 2000. Click [HERE](#) for details. We want to study the relationship between consumption (dependent variable) and other variables (independent variables). In all of the regression fitting below, it's assumed that the constant regressor is included, significance level is 1% for tests, and the regression models are normally distributed.

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```

library(AER)
data("USMacroG")

dim(USMacroG)

USMacroG=na.omit(USMacroG)
dim(USMacroG)

Macro=as.data.frame(apply(USMacroG,2,diff))
dim(Macro)

```

```
sum(is.na(USMacroG))  
sum(is.na(Macro))
```

USMacroG contains the original variables before differencing, Macro contains the variables after differencing (that is, quarterly changes).

Q2.1

1 Point

What's the R^2 when the consumption (USMacroG[, "consumption"]) is regressed on the cpi (USMacroG[, "cpi"]) (the variables here are before differencing)? What's the R^2 when Macro\$consumption is regressed on Macro\$cpi (the variables here are after differencing, that is, they are quarterly changes)? What could have caused the big difference in the values of the R^2 's? Does the cpi quarterly change seem to be useful in explaining the consumption quarterly change?

Call:

```
lm(formula = USMacroG[, "consumption"] ~ USMacroG[, "cpi"])
```

Multiple R-squared: 0.9593, Adjusted R-squared: 0.9591

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Call:

```
lm(formula = Macro$consumption ~ Macro$cpi)
```

Multiple R-squared: 0.003662, Adjusted R-squared: -0.001319

It is possible that by differentiating, the correlations(the Xs in $Y = BX + \text{error}$ model) are being reduced, so that only residual error is left, which is random and not correlated.

From the R^2 achieved, we do not conclude that cpi quarterly change is useful in explaining the consumption quarterly change.



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▼ Q2-1.R

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```
1 #Q2-1
2
3 #Data Preparation
4 library(AER)
5 data("USMacroG")
6 dim(USMacroG)
7 USMacroG=na.omit(USMacroG)
8 dim(USMacroG)
9 Macro=as.data.frame(apply(USMacroG,2,diff))
10 dim(Macro)
11
12 sum(is.na(USMacroG))
13 sum(is.na(Macro))
14
15 #Regression for USMacroG
16 summary(lm(USMacroG[, "consumption"] ~ USMacroG[, "cpi"]))
17
18 #Regression for Macro
19 summary(lm(Macro$consumption ~ Macro$cpi))
20
21
```

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Q2.2

1 Point

From now on, we work with quarterly changes only. Consumption refers to quarterly change in consumption, gdp refers to quarterly change in gdp, etc. Construct a scatter plot for consumption and gdp. Does it show some linear relationship between the two variables? Let consumption (dependent variable) be regressed on gdp. Give the equation for the estimated regression line, the estimate for σ and its degrees of freedom.

It shows on linear correlation between consumption and gdp, yee it is still scattered.

Call:

`lm(formula = Macro$consumption ~ Macro$gdp)`

Residuals:

Min	1Q	Median	3Q	Max
-56.713	-9.064	-2.733	9.491	56.189

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	11.06909	1.72775	6.407	1.04e-09 ***
Macro\$gdp	0.39623	0.03038	13.042	< 2e-16 ***

Significance levels: 0. '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

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Residual standard error: 18.33 on 200 degrees of freedom

Multiple R-squared: 0.4596, Adjusted R-squared: 0.4569



F-statistic: 170.1 on 1 and 200 DF, p-value: < 2.2e-16

The estimated regression line is $\text{consumption_hat} = 11.06909 + 0.39623 \cdot \text{gdp_hat}$, since it is estimated regression line, we use estimated value, consumption_hat and dgp_hat.

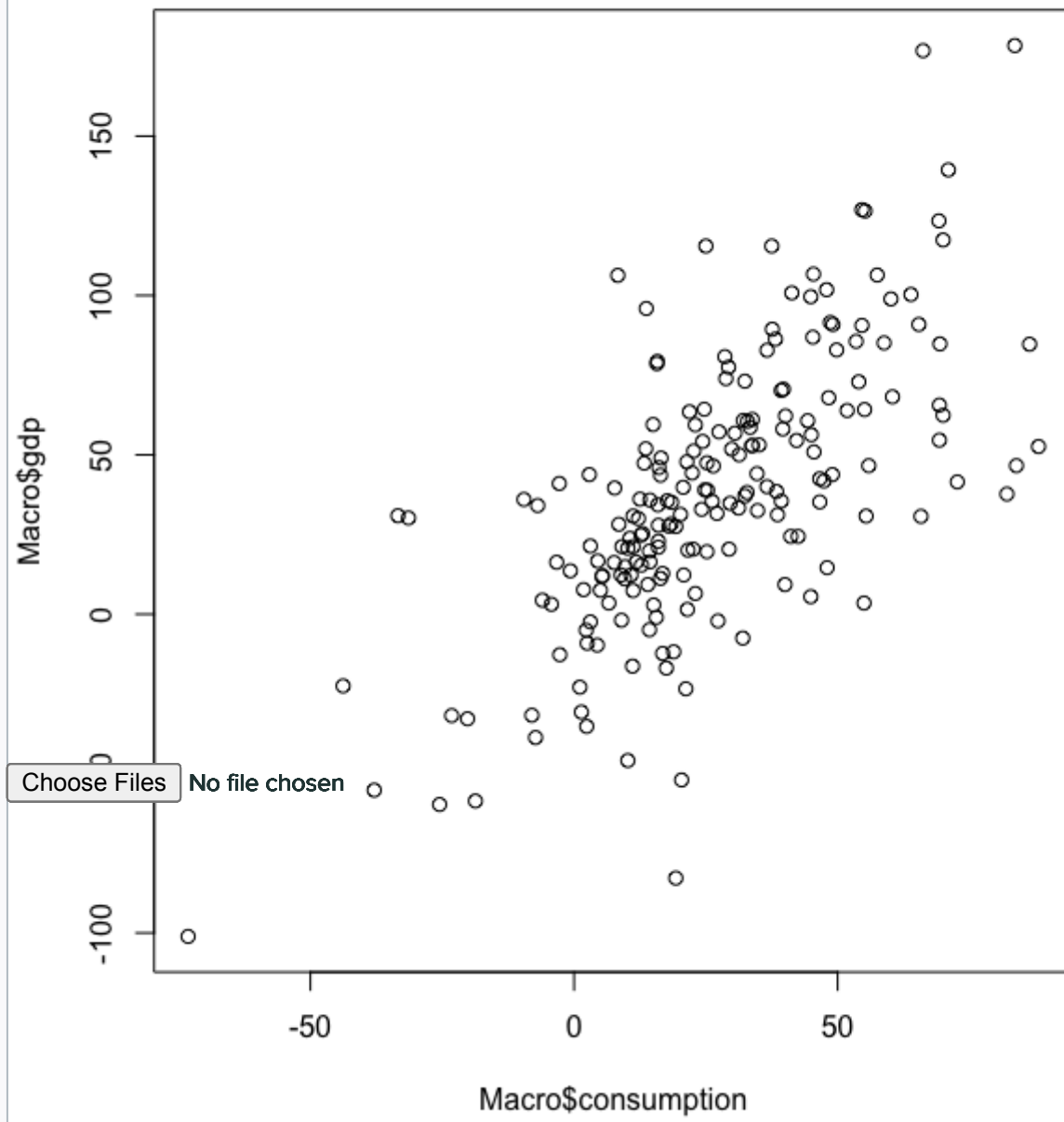
The estimation for sigma is 18.33, and the degree of freedom is 200.

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▼ Q2-2_Scatter_Plot.png

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```
1 #Q2-2
2
3 #Data Preparation
4 library(AER)
5 data("USMacroG")
6 dim(USMacroG)
7 USMacroG=na.omit(USMacroG)
8 dim(USMacroG)
9 Macro=as.data.frame(apply(USMacroG,2,diff))
10 dim(Macro)
11 sum(is.na(USMacroG))
12 sum(is.na(Macro))
13
14 #Scatter plot (consumption, gdp)
15 plot(Macro$consumption, Macro$gdp)
16
17 #Regression (consumption, gdp)
18 fit2_2 = lm(Macro$consumption ~ Macro$gdp)
19 summary(fit2_2)
20
21
22
23
```

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Q2.3

1 Point

Let β_{gdp} be the coefficient for gdp and $\hat{\beta}_{gdp}$ the corresponding OLS estimator. What's the estimated standard error for $\hat{\beta}_{gdp}$? Construct a 99% confidence interval for β_{gdp} . Is there strong enough statistical evidence that $\beta_{gdp} \neq 0$?

Call:

```
lm(formula = Macro$consumption ~ Macro$gdp)
```

Residuals:

Min	1Q	Median	3Q	Max
-56.713	-9.064	-2.733	9.491	56.189

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	11.06909	1.72775	6.407	1.04e-09 ***
Macro\$gdp	0.39623	0.03038	13.042	< 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 18.33 on 200 degrees of freedom

Multiple R-squared: 0.4596, Adjusted R-squared: 0.4569

Choose Files No file chosen 201 file chosen 200 DF, p-value: < 2.2e-16

The estimated standard error of Beta_hat_gdp is 0.03038.

The 99% CI of Beta_gdp is [0.317222725828501, 0.475237274171499]


Yes, since the p-value of Beta_gdp is smaller than 2e-16, which is smaller enough for H0:Beta_gdp = 0 to be rejected. Thus, Beta_gdp is unlikely to be zero.

▼ Q2-3.R

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```
1 #Q2-3
2
3 #Data Preparation
4 library(AER)
5 data("USMacroG")
6 dim(USMacroG)
7 USMacroG=na.omit(USMacroG)
8 dim(USMacroG)
9 Macro=as.data.frame(apply(USMacroG,2,diff))
10 dim(Macro)
11 sum(is.na(USMacroG))
12 sum(is.na(Macro))
13
14
15 #Regression (consumption, gdp)
16 fit2_3 = lm(Macro$consumption ~ Macro$gdp)
17 summary(fit2_3)
18
19 #99% CI of Beta_gdp
20 Beta_gdp_hat = 0.39623
21 print(paste(Beta_gdp_hat - qt(p = 0.995, df = 200) * 0.03038 , Beta_gdp_hat +
22           qt(p = 0.995, df = 200) * 0.03038))
```

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Q2.4

1 Point

What's the SST , SSR , SSE , R^2 of the above simple linear regression model? What's the percentage of the variation in consumption that could be explained by gdp in this model.

```
> anova(fit2_4)
```

Analysis of Variance Table

Response: Macro\$consumption

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Macro\$gdp	1	57139	57139	170.1	< 2.2e-16 ***
Residuals	200	67185	336		

Signif. codes:

0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

SSR = 57139

SSE = 67186

SST = SSR + SSE = 124134

The percentage of the variation in consumption that can be explained by gdp is 45.95975%

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▼ Q2-4.R


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```
1 #Q2-4
2
3 #Data Preparation
4 library(AER)
5 data("USMacroG")
6 dim(USMacroG)
7 USMacroG=na.omit(USMacroG)
8 dim(USMacroG)
9 Macro=as.data.frame(apply(USMacroG,2,diff))
```

```

10 dim(Macro)
11 sum(is.na(USMacroG))
12 sum(is.na(Macro))
13
14 #Regression
15 fit2_4 = lm(Macro$consumption ~ Macro$gdp)
16 summary(fit2_4)
17
18 #Defination
19 anova(fit2_4)
20 57139/(57139+67185)
21

```

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Q2.5

1 Point

Let n be the sample size of the above regression model. Regress consumption[1:($n-1$)] on gdp[1:($n-1$)] on consumption using gdp[n]. What's the 95% prediction interval? Compare with the actual value consumption[n]: what's the prediction error?

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Y_star_hat = 28.31972

95% prediction interval is [-7.890101, 64.52954]

The true consumption[202] is 49, and the predicted consumption is 28.31972 with 95% PI [-7.890101, 64.52954].

The prediction error is 20.68028.



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Q2.6

1 Point

Suppose you regress consumption on tbill, inflation and cpi. What's the null hypothesis of the F-test? Is the null hypothesis rejected? What does it mean?

The null hypo of F-test (H_0) is H_0 : None of regressor is useful : $B_{\text{tbill}} = B_{\text{inflation}} = B_{\text{cpi}} = 0$.

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Call:

`lm(formula = consumption ~ tbill + inflation + cpi)`

Residuals:

Min	1Q	Median	3Q	Max
-81.723	-15.179	-3.492	16.394	59.113

Coefficients:

Estimate	Std. Error	t value	Pr(> t)
----------	------------	---------	----------


```
(Intercept) 25.75793  2.61828  9.838 <2e-16 ***
tbill      4.30044  2.44569  1.758  0.0802 .
inflation  0.82269  0.68256  1.205  0.2295
cpi        0.09873  0.87997  0.112  0.9108
---
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 24.65 on 198 degrees of freedom



Multiple R-squared: 0.03264, Adjusted R-squared: 0.01799

F-statistic: 2.227 on 3 and 198 DF, p-value: 0.08627

Under the confidence level of 95%, since $p = 0.08627 > \alpha = 0.05$, null hypo H_0 is not rejected, means that we maintain our hypo that none of the regressor is useful.


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▼ Q2-6.R

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```
1 #Q2-6
2 #Data Preparation
3 library(AER)
4 data("USMacroG")
5 MacroG=na.omit(USMacroG)
6 USMacroG=na.omit(USMacroG)
7 dim(USMacroG)
8 Macro=as.data.frame(apply(USMacroG,2,diff))
9 dim(Macro)
10 sum(is.na(USMacroG))
11 sum(is.na(Macro))
12
13 #Regression
14 fit2_6 = lm(Macro$consumption ~ Macro$tbill + Macro$inflation + Macro$cpi)
15 summary(fit2_6)
16
```

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Q2.7

1 Point

When consumption is regressed on m1 only, is m1 significant (i.e., is $H_0 : \beta_{m1} = 0$ rejected)? Does adding a second order term improve the fitting? Construct a scatter plot of consumption (y axis) and m1 (x axis), add the estimated regression line resulting from the simple linear regression model, and the quadratic curve resulting from the fitting with a second order term in addition to m1 itself.

Call:

```
lm(formula = Macro$consumption ~ Macro$m1)
```

Residuals:

Min	1Q	Median	3Q	Max
-97.31	-13.58	-1.63	15.70	70.84

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Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	23.3945	1.9608	11.931	< 2e-16 ***
Macro\$m1	0.5526	0.1948	2.837	0.00503 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 24.45 on 200 degrees of freedom

Multiple R-squared: 0.03868, Adjusted R-squared: 0.03387

F-statistic: 8.046 on 1 and 200 DF, p-value: 0.005029

Since $p = 0.00503$ (simple regression so F-test = t-test), so H_0 is rejected, which means that m_1 is significant under 95% confidence level.

Call:

`lm(formula = Macro$consumption ~ Macro$m1 + I(Macro$m1^2))`

Residuals:

Min	1Q	Median	3Q	Max
-96.549	-12.536	-2.095	15.933	62.358

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	23.71325	1.91247	12.399	< 2e-16 ***
Macro\$m1	-0.33071	0.32011	-1.033	0.302804
I(Macro\$m1^2)	0.03899	0.01138	3.426	0.000743 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Choose Files and error: 23.81 on 199 degrees of freedom

Multiple R-squared: 0.09223, Adjusted R-squared: 0.08311

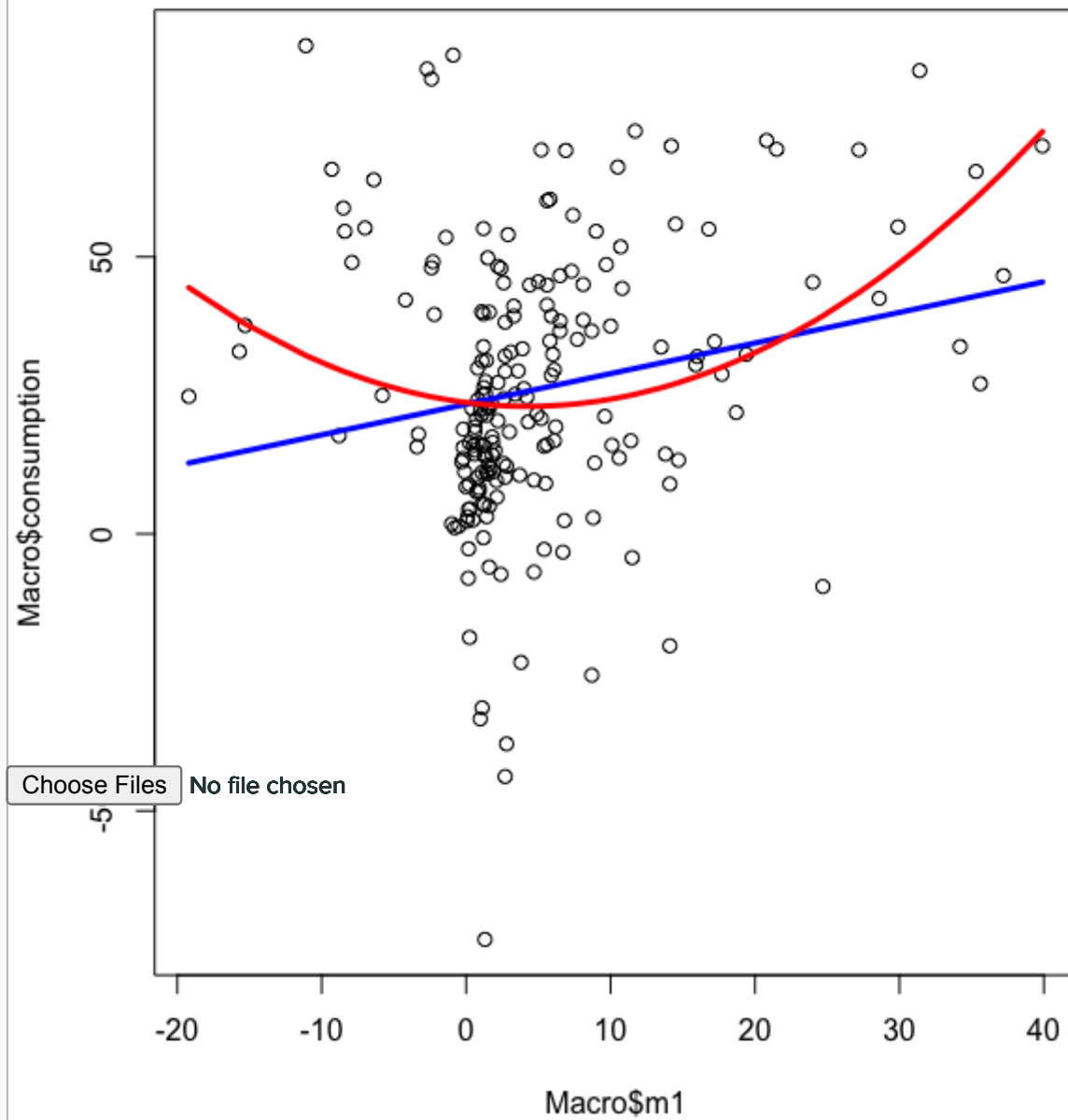
F-statistic: 10.11 on 2 and 199 DF, p-value: 6.586e-05

The 2nd_order model is better since its Adjusted R_square is better than the 1st order model.

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```

1 #Q2-7
2 #Data Preparation
3 library(AER)
4 data("USMacroG")
5 dim(USMacroG)
6 USMacroG=na.omit(USMacroG)
7 dim(USMacroG)
8 Macro=as.data.frame(apply(USMacroG,2,diff))
9 dim(Macro)
10 sum(is.na(USMacroG))
11 sum(is.na(Macro))
12
13 #Regression
14 fit2_7 =lm(Macro$consumption ~ Macro$m1)
15 summary(fit2_7)
16
17 #Regression 2nd order
18 fit2_7.2 =lm(Macro$consumption ~ Macro$m1 + I(Macro$m1^2))
19 summary(fit2_7.2)
20
21 #Plotting
22 plot(Macro$m1, Macro$consumption)
23 lines(Macro$m1, fit2_7$fitted.values, lwd=3, col = 'blue')
24 lines(sort(Macro$m1), fit2_7.2$fitted.values[order(Macro$m1)], lwd = 3, col =

```

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Q2.8

1 Point

Regress consumption on all other 11 variables. What's the adjusted R^2 now? Is m1 significant? From results in #7 and #8, what can you say about whether m1 is useful in explaining consumption?

Call:

```
lm(formula = Macro$consumption ~ ., data = Macro)
```

Residuals:

Min	1Q	Median	3Q	Max
-34.287	-6.922	-1.108	6.922	38.723

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-3.14859	3.07949	-1.022	0.30787
gdp	0.77061	0.04636	16.624	< 2e-16 ***
invest	-0.74133	0.05781	-12.822	< 2e-16 ***
government	-0.58702	0.08414	-6.976	4.9e-11 ***
dpi	0.10500	0.03550	2.958	0.00349 **
cpi	-0.39117	0.46705	-0.838	0.40334
m1	0.14953	0.11079	1.350	0.17873
unemp	-3.82193	3.05721	-1.250	0.21279
population	10.22284	3.99745	2.557	0.01133 *
inflation	-16.05374	245.71204	-0.065	0.94798
interest	-16.77342	245.72689	-0.068	0.94565

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 12.71 on 190 degrees of freedom

Multiple R-squared: 0.7533, Adjusted R-squared: 0.739

F-statistic: 52.74 on 11 and 190 DF, p-value: < 2.2e-16

Multiple R-squared: 0.7533

The p-value of m1 is 0.17873, which we consider it as not significant.

With only use m1 as the regressor, m1 can be the significant However, with other regressors , m1;s significance seemed to be dimmed, use makes m1 not significant in explaining the consumption anymore.




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▼ Q2-8.R

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```
1 #Q2-8
2 #Data Preparation
3 library(AER)
4 data("USMacroG")
5 dim(USMacroG)
6 USMacroG=na.omit(USMacroG)
7 dim(USMacroG)
8 Macro=as.data.frame(apply(USMacroG,2,diff))
9 dim(Macro)
10 sum(is.na(USMacroG))
11 sum(is.na(Macro))
12
13 #Regression
14 fit2_8 = lm(Macro$consumption ~., Macro)
15 summary(fit2_8)
16
```

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