6.842 Randomness and Computation February 23, 2022

Homework 1

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1. *Collaborators and sources:* none.

Proof. We construct an approximation scheme $\mathcal{B}$ for f as follows: On input (x, ε, δ), run $\mathcal{A}(x, \varepsilon)$ independently for $k := \lceil 12 \log(1/\delta) \rceil$ times with outputs $y_1, \ldots, y_k$, and output a median of $y_1, \ldots, y_k$.

Let $t_{\mathcal{A}}(x, \varepsilon)$ be the running time of $\mathcal{A}$ on input (x, ε). Then $\mathcal{B}$ runs in $O(kt_{\mathcal{A}}(x, \varepsilon)) = O(\log(1/\delta) t_{\mathcal{A}}(x, \varepsilon))$. Since $\mathcal{A}$ runs in time polynomial in $1/\varepsilon$ and $|x|$, then $\mathcal{B}$ runs in time polynomial in $1/\varepsilon$, $|x|$ and $\log(1/\delta)$.

By the definition of medians, if more than half of $y_1, \ldots, y_k$ fall in $[f(x)/(1+\varepsilon), f(x)(1+\varepsilon)]$, then $\mathcal{B}(x, \varepsilon, \delta) \in [f(x)/(1+\varepsilon), f(x)(1+\varepsilon)]$. Let $X_1, \ldots, X_k \in \{0, 1\}$ be random variables so that $X_i = 1$ with probability $p := \mathbb{P}[\mathcal{A}(x, \varepsilon) \notin [f(x)/(1+\varepsilon), f(x)(1+\varepsilon)]] \leq 1 - 3/4 = 1/4$. Then $\sum_{i=1}^k \mathbb{E}[X_i] = kp \leq k/4$. Therefore,

$$
\begin{aligned}
&\mathbb{P}\left[\mathcal{B}(x, \varepsilon, \delta) \notin \left[\frac{f(x)}{1+\varepsilon}, f(x)(1+\varepsilon)\right]\right] \\
&\leq \mathbb{P}\left[\text{at least half of } y_1, \ldots, y_k \text{ do not fall in } \left[\frac{f(x)}{1+\varepsilon}, f(x)(1+\varepsilon)\right]\right] \\
&= \mathbb{P}\left[\sum_{i=1}^k X_i \geq \frac{k}{2}\right] \\
&= \mathbb{P}\left[\sum_{i=1}^k X_i \geq (1+1) \cdot \frac{k}{4}\right] \\
&\leq e^{-\frac{k/4}{3}} && \text{(Chernoff bound)} \\
&= e^{-\frac{\lceil 12 \log \frac{1}{\delta} \rceil}{12}} \leq e^{-\frac{12 \log \frac{1}{\delta}}{12}} = \delta.
\end{aligned}
$$

Therefore,

$$
\mathbb{P}\left[\mathcal{B}(x, \varepsilon, \delta) \in \left[\frac{f(x)}{1+\varepsilon}, f(x)(1+\varepsilon)\right]\right] = 1 - \mathbb{P}\left[\mathcal{B}(x, \varepsilon, \delta) \notin \left[\frac{f(x)}{1+\varepsilon}, f(x)(1+\varepsilon)\right]\right] \geq 1 - \delta.
$$

This completes the proof. $\square$

2. *Collaborators and sources:* none.

Proof. Suppose $\binom{m}{t}2^{1-\binom{t}{2}} < 1$. To prove $R(t) > m$, it suffices to show that there exists a 2-edge-coloring of K_m such that for all $S \subset V(K_m)$ of size t, $E(K_m[S])$ is not monochromatic. We randomly color the edges of K_m red or blue, independently and equiprobably. For each $S \subset V(K_m)$ of size t, there are exactly $2^{\binom{t}{2}}$ two-colorings of $E(K_m[S])$, amongst which two are monochromatic colorings (all red and all blue), so

$$\mathbb{P}\left[E\left(K_m[S]\right) \text{ is monochromatic}\right] = \frac{2}{2^{\binom{t}{2}}} = 2^{1-\binom{t}{2}}.$$

By the union bound,

$$\begin{aligned}
&\mathbb{P}\left[\exists S \subset V\left(K_m\right), |S| = t, E\left(K_m[S]\right) \text{ is monochromatic}\right] \\
\leq &\sum_{\substack{S \subset V(K_m) \\ |S|=t}} \mathbb{P}\left[E\left(K_m[S]\right) \text{ is monochromatic}\right] \\
= &\binom{m}{t} 2^{1-\binom{t}{2}} \\
< &\,1.
\end{aligned}$$

Therefore,

$$\begin{aligned}
&\mathbb{P}\left[\forall S \subset V\left(K_m\right) \text{of size } t, E\left(K_m[S]\right) \text{ is not monochromatic}\right] \\
= &\,1 - \mathbb{P}\left[\exists S \subset V\left(K_m\right), |S| = t, E\left(K_m[S]\right) \text{ is monochromatic}\right] \\
> &\,1 - 1 = 0.
\end{aligned}$$

This proves that there exists a 2-edge-coloring of K_m such that for all $S \subset V(K_m)$ of size t, $E(K_m[S])$ is not monochromatic. The proof is complete. $\square$