

ML 7641 Homework 1

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1 Linear Algebra

1.1 Determinant and Inverse of Matrix

1. The determinant of M

$$\begin{aligned}|M| &= r \times \begin{vmatrix} r & 3 \\ 4 & 1 \end{vmatrix} - 0 \times \begin{vmatrix} 0.5 & 3 \\ -9 & 1 \end{vmatrix} + 1 \times \begin{vmatrix} 0.5 & r \\ -9 & 4 \end{vmatrix} \\ &= r(r - 12) + 1(2 - (-9r)) \\ &= r^2 - 12r + 2 + 9r \\ &= r^2 - 3r + 2\end{aligned}$$

2. The value(s) of r for which M^{-1} does not exist and the means in terms of rank and singularity for values of r.

According to the properties of the matrix, M^{-1} does not exist When $|M| = 0$

$$|M| = 0$$

$$r^2 - 3r + 2 = 0$$

$$r = 1, 2$$

When $r = 1, 2$, the determinant of M is 0, which means $\text{rank}(M) = 0$, and M is a singular matrix.

3. Calculate M^{-1} for $r = 3$ and show that $MM^{-1} = I$.

When $r = 3$,

$$\begin{aligned}M &= \begin{bmatrix} 3 & 0 & 1 \\ 0.5 & 3 & 3 \\ -9 & 4 & 1 \end{bmatrix} \\ \text{Adj}(M) &= \begin{pmatrix} +\begin{vmatrix} 3 & 3 \\ 4 & 1 \end{vmatrix} & -\begin{vmatrix} 0 & 1 \\ 4 & 1 \end{vmatrix} & +\begin{vmatrix} 0 & 1 \\ 3 & 3 \end{vmatrix} \\ -\begin{vmatrix} 0.5 & 3 \\ -9 & 1 \end{vmatrix} & +\begin{vmatrix} 3 & 1 \\ -9 & 1 \end{vmatrix} & -\begin{vmatrix} 3 & 1 \\ 0.5 & 3 \end{vmatrix} \\ +\begin{vmatrix} 0.5 & 3 \\ -9 & 4 \end{vmatrix} & -\begin{vmatrix} 3 & 0 \\ -9 & 4 \end{vmatrix} & +\begin{vmatrix} 3 & 0 \\ 0.5 & 3 \end{vmatrix} \end{pmatrix} = \begin{bmatrix} -9 & 4 & -3 \\ -27.5 & 12 & -8.5 \\ 29 & -12 & 9 \end{bmatrix} \\ |M| &= r^2 - 3r + 2 = 9 - 9 + 2 = 2 \\ M^{-1} &= \frac{\text{adj}(M)}{\det(M)} = \frac{\begin{bmatrix} -9 & 4 & -3 \\ -27.5 & 12 & -8.5 \\ 29 & -12 & 9 \end{bmatrix}}{2} = \begin{bmatrix} -4.5 & 2 & -1.5 \\ -13.75 & 6 & -4.25 \\ 14.5 & -6 & 4.5 \end{bmatrix}\end{aligned}$$

Thus,

$$MM^{-1} = \begin{bmatrix} 3 & 0 & 1 \\ 0.5 & 3 & 3 \\ -9 & 4 & 1 \end{bmatrix} \begin{bmatrix} -4.5 & 2 & -1.5 \\ -13.75 & 6 & -4.25 \\ 14.5 & -6 & 4.5 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

4. Find the determinant of M^{-1} for $r = 3$. What is the relationship between the determinant of M and the determinant of M^{-1} ?

When r is 3, according to the problem 3 the determinant of M is 2.

$$\begin{aligned} |M^{-1}| &= -4.5 \times \begin{vmatrix} 6 & -4.25 \\ -6 & 4.5 \end{vmatrix} - 2 \times \begin{vmatrix} -13.75 & -4.25 \\ 14.5 & 4.5 \end{vmatrix} - 1.5 \times \begin{vmatrix} -13.75 & 6 \\ 14.5 & -6 \end{vmatrix} \\ &= -4.5 \times 1.5 - 2 \times (-0.25) - 1.5 \times (-4.5) \\ &= 0.5 \end{aligned}$$

Thus, the determinant of M^{-1} is 0.5. The relationship between the determinant of M and the determinant of M^{-1} is that $\det(M) = \frac{1}{\det(M^{-1})}$

1.2 Characteristic Equation

Consider the eigenvalue problem:

$$Ax = \lambda x, x \neq 0$$

where x is a non-zero eigenvector and λ is eigenvalue of A . Prove that the determinant $|A - \lambda I| = 0$. Multiply the identity matrix I on the both sides of the equation, then we get:

$$IAx = I\lambda x$$

$$Ax = (\lambda I)x$$

Then,

$$(A - \lambda I)x = 0$$

If $B = A - \lambda I$ is a inverse matrix, then we could get:

$$x = B^{-1}0$$

Which means x is a zero vector.

However, x is a non-zero eigenvector, so $B = A - \lambda I$ is not a inverse matrix. It means that the determinant of $B = A - \lambda I$ is zero, $|A - \lambda I| = 0$.

1.3 Eigenvalues and Eigenvectors

1. The eigenvalues of A in terms of z.
Organize the equation:

$$Ax = \lambda x$$

which x is the eigenvector, λ is the eigenvalue

$$(A - \lambda I)x = 0$$

$$\det(A - \lambda I) = 0$$

$$\begin{aligned}\det(A - \lambda I) &= \begin{vmatrix} z - \lambda & 2 \\ 3 & z - \lambda \end{vmatrix} \\ &= (z - \lambda)^2 - 6 \\ \lambda &= z \pm \sqrt{6}\end{aligned}$$

2. Find the normalized eigenvectors of matrix A.

$$A - \lambda I = \begin{bmatrix} \pm\sqrt{6} & 2 \\ 3 & \pm\sqrt{6} \end{bmatrix}$$

$$(A - \lambda I)x = 0$$

$$\begin{bmatrix} \pm\sqrt{6} & 2 \\ 3 & \pm\sqrt{6} \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = 0$$

From this equation, we could get:

$$3x_1 \pm \sqrt{6}x_2 = 0$$

Due to the normalized eigenvectors, which means:

$$x_1^2 + x_2^2 = 1$$

Thus,

$$x = \begin{bmatrix} \frac{\sqrt{10}}{5} \\ \frac{\sqrt{15}}{5} \end{bmatrix}, \begin{bmatrix} -\frac{\sqrt{10}}{5} \\ \frac{\sqrt{15}}{5} \end{bmatrix}, \begin{bmatrix} \frac{\sqrt{10}}{5} \\ -\frac{\sqrt{15}}{5} \end{bmatrix}, \begin{bmatrix} -\frac{\sqrt{10}}{5} \\ -\frac{\sqrt{15}}{5} \end{bmatrix}$$

2 Expectation, Co-variance, Independence, and Bayes' Rule

1. Calculate the Expectation and Variance of X

$$\begin{aligned}E(X) &= 0 \times P(x = 0) + 1 \times P(x = 1) \\&= 0(1 - p) + 1 \times p \\&= p \\E(X^2) &= 0 \times P(x^2 = 0) + 1 \times P(x^2 = 1) \\&= 0(1 - p) + 1 \times p \\&= p \\Var(X) &= E(X^2) - E(X)^2 = p - p^2\end{aligned}$$

2. Calculate the Expectation and Variance of Z.

According to the question, Z is a variable following the Binomial distribution with probability p. The probability density function for Z is given by:

$$P(Z = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

where k is the number of x = 1

$$\begin{aligned}E(Z) &= E(Y_1 + Y_2 + \dots + Y_n) \\&= E(Y_1) + E(Y_2) + \dots + E(Y_n) \\&= nE(X) \\&= np\end{aligned}$$

Due to each element of Y is independent of the other elements, the Variance of Z is:

$$\begin{aligned}Var(Z) &= Var(Y_1 + Y_2 + \dots + Y_n) \\&= Var(Y_1) + Var(Y_2) + \dots + Var(Y_n) \\&= nVar(X) \\&= np(1 - p)\end{aligned}$$

3. Are Z and X independent?

The conclusion is that Z and X is dependent, the prove is below:

$$\begin{aligned}P(Z = k, X = 1) &= P(Z = k | X = 1)P(X = 1) \\&= \binom{n-1}{k-1} p^{k-1} (1-p)^{n-1-k+1} \\&= \frac{(n-1)!}{(k-1)!(n-k)!} p^{k-1} (1-p)^{n-k} \\P(Z = k)P(X = 1) &= \binom{n}{k} p^k (1-p)^{n-k} \times p \\&= \frac{n!}{k!(n-k)!} p^{k+1} (1-p)^{n-k}\end{aligned}$$

Thus,

$$P(Z = k)P(X = 1) = \frac{np^2}{k} P(Z = k, X = 1)$$

For each k, it is impossible for $np^2 = k$, which means Z and X are dependent.

4. Find $P(X = 1|Z = k)$ using Bayes' Rule. Simplify your answer.

$$\begin{aligned}
 P(X = 1|Z = k) &= \frac{P(X = 1, Z = k)}{P(Z = k)} \\
 &= \frac{\frac{(n-1)!}{(k-1)!(n-k)!} p^{k-1} (1-p)^{n-k}}{\frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}} \\
 &= \frac{k}{np}
 \end{aligned}$$

5. Calculate $Cov(X, Z)$. Are X and Z uncorrelated or not? If they are correlated, interpret the sign of $Cov(X, Z)$.

For $Z = Y_1 + Y_2 + \cdots + Y_n = X + Y_2 + \cdots + Y_n$, it is obvious that X is correlated with Z .

$$\begin{aligned}
 E(XZ) &= E[X(Y_1 + Y_2 + \cdots + Y_n)] \\
 &= E(Y_1^2 + Y_1 Y_2 + \cdots + Y_1 Y_n) \\
 &= E(Y_1^2) + E(Y_1)E(Y_2) + \cdots + E(Y_1)E(Y_n) \\
 &= p + (n-1)p^2 \\
 Cov(X, Z) &= E(XZ) - E(X)E(Z) \\
 &= p + (n-1)p^2 - p \times np \\
 &= p - p^2
 \end{aligned}$$

3 Optimization

1. Specify the Lagrange function.

According to the question, the Lagrange function is:

$$\begin{aligned} L(x, y, a, b) &= f(x, y) - a[g_1(x, y) - 4] - b[g_2(x, y) - 4] \\ &= 2x^2 + 3xy - a\left(\frac{1}{2}x^2 - y - 4\right) - b(y - 4) \end{aligned}$$

where $a, b \geq 0$

2. List the KTT conditions.

$$\frac{\partial L}{\partial x} = 4x - ax + 3y = 0$$

$$\frac{\partial L}{\partial y} = a + 3x - b = 0$$

$$a \geq 0$$

$$b \geq 0$$

$$a\left(\frac{1}{2}x^2 - y - 4\right) = 0$$

$$b(y - 4) = 0$$

3. 4 possibilities formed could be

(a) Both constraint are active, the optimal solution that lies at the intersection point of two constraints, so

$$a\left(\frac{1}{2}x^2 - y - 4\right) = 0 \quad a > 0$$

Therefore,

$$\frac{1}{2}x^2 - y - 4 = 0$$

And

$$b(y - 4) = 0 \quad b > 0$$

Therefore,

$$y = 4$$

and to solve the equal we could gain that $x = 4, -4$ and $y = 4$ Take the results to the KTT condition, we could obtain that

$$\begin{cases} x = 4 & a = 7, b = 19 \\ x = -4 & a = 1, b = -11 \text{ [invalid]} \end{cases}$$

(b) Both constraint are inactive, any of the constraint lines do not pass through the optimal point. We could get $a = 0$ and $b = 0$, so

$$\begin{cases} \frac{1}{2}x^2 - y - 4 \leq 0 \\ y - 4 \leq 0 \end{cases}$$

Take the a, b to the KTT condition, we could obtain that

$$\begin{cases} x = 0 \\ y = 0 \end{cases}$$

(c) The first constraint is active and the second constraint is inactive, therefore, we could get

$$\begin{cases} \frac{1}{2}x^2 - y - 4 = 0 \quad a > 0 \\ y - 4 \leq 0 \quad b = 0 \end{cases}$$

Take the $y = \frac{1}{2}x^2 - 4, b = 0$ to the KTT condition, we could obtain that

$$\begin{cases} 4x + 3(\frac{1}{2}x^2 - 4) - ax = 0 \\ 3x + a = 0 \end{cases}$$

Therefore,

$$x = -\frac{4 + 2\sqrt{58}}{9} = -2.14, \quad a = \frac{4 + 2\sqrt{58}}{3} = 6.41, \quad y = \frac{4(31 + 2\sqrt{58})}{81} - 4 = -1.72$$

(c) The first constraint is inactive and the second constraint is active, therefore, we could get

$$\begin{cases} \frac{1}{2}x^2 - y - 4 \leq 0 \quad a = 0 \\ y - 4 = 0 \quad b > 0 \end{cases}$$

Take the $y = 4, a = 0$ to the KTT condition, we could obtain that

$$\begin{cases} x = -3 \\ b = -9[invalid] \end{cases}$$

4. All candidate point is

$$\begin{cases} x = 4, y = 4 & a = 7, b = 19 \\ x = 0, y = 0 & a = 0, b = 0 \\ x = -2.14, y = -1.72 & a = 6.41, b = 0 \end{cases}$$

5. Take the all candidate point for $f(x, y)$

$$\begin{cases} f(4, 4) = 80 & a=7, b=19 \\ f(0, 0) = 0 & a=0, b=0 \\ f(-2.14, -1.72) = 20.14 & a = 6.41, b = 0 \end{cases}$$

for those candidate point, the maximum value is 80 when $x = 4$, and $y = 4$ To find the at this point, the $L(x, y)$ is concave or convex we need to take the derivative for $a = 7, b = 19$

$$\frac{\partial L}{\partial x} = 3y - 3x$$

Take the second derivative, we could gain

$$\frac{\partial^2 L}{\partial x^2} = -3$$

For y,

$$\frac{\partial L}{\partial y} = 3x - 12$$

Take the second derivative, we could gain

$$\frac{\partial^2 L}{\partial y^2} = 0$$

Therefore, at this point, the $L(x, y)$ is concave.

4 Maximum Likelihood

4.1 Discrete Example

1. When you roll the die 6 times and each side appears exactly once, the likelihood of the result will use number of success divide the total number:

$$\theta = \frac{1}{6}$$

2. The maximum likelihood estimation(MLE): Based on Bernoulli Distribution, we took into two conditions, one is for when dice landing on 1 and another is for dice landing on other numbers.

$$P(X = x) = \begin{cases} \theta & x=1 \\ 1 - \theta & x=\text{other} \end{cases}$$

For Bernoulli: objective function:

$$f(x_i; \theta) = \theta(1 - \theta)^5, x_i \in \{1, \text{other}\}$$

$$L(\theta) = p(X = 1, X = 2, X = 3, X = 4, X = 5, X = 6)$$

Based on the i.i.d assumption

$$L(\theta) = \prod_{i=1}^6 \theta(1 - \theta)^5 = \underset{\theta}{\operatorname{argmax}} \theta(1 - \theta)^5$$

On here, we covert multiplication into summation by taking the log

$$\log L(\theta) = l(\theta) = \log \theta + 5 \log(1 - \theta)$$

Take the derivative of the θ , we can gain,

$$\frac{1}{\theta} - \frac{5}{1 - \theta} = 0$$

Therefore,

$$\theta = \frac{1}{6}$$

3. When you roll the die 6 times and side 1 appears 4 times, the MLE for θ is:

$$\theta = \frac{2}{3}$$

4.2 Normal distribution

1. We assume that the noise is Gaussian with mean 0 and variance σ^2 . The z is in normal distribution, for y' it is also a noisy value which is equal to $y + z$. Therefore, if we re-write the model, it will be:

$$y' \sim \mathcal{N}(mx + c, \sigma^2)$$

Therefore, we can write the likelihood of this sample at (x_1, y'_1) as,

$$L(y'_1|x_1, m, c, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y'_1 - (mx_1 + c))^2}{2\sigma^2}}$$

2. When there are given n values and make $c=0$, the likelihood will be like,

$$L(y'_i|x_i, m, \sigma^2) = \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^n e^{-\frac{\sum_{i=1}^n (y'_i - mx_i)^2}{2\sigma^2}}$$

3. The maximum likelihood estimator for m is, first covert the multiplication into summation by take log

$$\log L(y'_i|x_i, m, \sigma^2) = -n \log\left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) - \frac{\sum_{i=1}^n (y'_i - mx_i)^2}{2\sigma^2}$$

Then, make derivative for m ,

$$\begin{aligned} \frac{\partial \log(L)}{\partial m} &= 0 \\ -\frac{2 \sum_{i=1}^n x_i (y'_i - mx_i)}{\sigma^2} &= -\frac{2 \sum_{i=1}^n x_i y'_i + \sum_{i=1}^n x_i^2}{\sigma^2} \\ m &= \frac{\sum_{i=1}^n x_i y'_i}{\sum_{i=1}^n x_i^2} \end{aligned}$$

4.3 Exponential Distribution

1. Find the P.D.F for above C.D.C, we need to take the derivative,

$$f(x) = \frac{d}{dx_i} P(X_i \leq x | \lambda_i) = \frac{d}{dx_i} (1 - e^{-\lambda_i x}) = -(-\lambda_i \frac{d}{dx_i} x) e^{-\lambda_i x} = \lambda_i e^{-\lambda_i x}$$

$$P(X_i \leq x | \lambda_i) = \begin{cases} 0 & x < 0 \\ \lambda_i e^{-\lambda_i x} & 0 \leq x \end{cases}$$

2. The expected value of distribution for a given i is

$$\int_0^{\infty} p(x) dx = \frac{1}{\lambda_i}$$

3. Find Maximum Likelihood Estimator for λ ,

$$\begin{aligned} L(\lambda | x_1, x_2, \dots, x_n) &= \prod_{i=1}^{\frac{n}{2}} \lambda \exp^{-\lambda x_i} \prod_{i=\frac{n}{2}+1}^n 2\lambda \exp^{-2\lambda x_i} \\ &= \lambda^{\frac{n}{2}} \exp^{-\lambda \sum_{i=1}^{\frac{n}{2}} x_i} \times 2\lambda^n \exp^{-2\lambda \sum_{i=\frac{n}{2}+1}^n x_i} \end{aligned}$$

On here, we convert multiplication into summation by taking the log,

$$\ln L = \frac{n \ln(\lambda)}{2} - \lambda \sum_{i=1}^{\frac{n}{2}} x_i - (\ln(2) + n \ln(\lambda)) - 2\lambda \sum_{i=\frac{n}{2}+1}^n x_i$$

Take the derivative,

$$\frac{\partial \log(L)}{\partial \lambda} = 0$$

$$\frac{n}{2\lambda} - \sum_{i=1}^{\frac{n}{2}} x_i + \frac{n}{\lambda} - 2 \sum_{i=\frac{n}{2}+1}^n x_i = 0$$

Based on the n is even and all values are positive, we could gain,

$$\frac{3n}{2\lambda} = \sum_{i=1}^{\frac{n}{2}} x_i + \sum_{i=\frac{n}{2}+1}^n x_i$$

$$\lambda = \frac{3n}{2(\sum_{i=1}^{\frac{n}{2}} x_i + \sum_{i=\frac{n}{2}+1}^n x_i)}$$

4. the MLE for λ for $n = 1$ variable,

$$\lambda_1 = \frac{1}{x}$$

The reciprocal of the expected value is this value.

5 Information Theory

5.1 Mutual Information and Entropy

(a) The joint entropy

$$H(X, Y) = - \sum_{x \in X} \sum_{y \in Y} P(x, y) \log_2 [P(x, y)] = H(X) + H(Y|X) = 3.375$$

(b) The marginal entropies is,

$$H(X) = \sum P(x) \log_2 \frac{1}{P(x)} = 1.75$$

$$H(Y) = \sum P(y) \log_2 \frac{1}{P(y)} = 2$$

(c) The conditional entropy formal is

$$H(X|Y) = H(X, Y) - H(Y) = 3.375 - 2 = 1.375$$

Mutual Information between X and Y is

$$I(X; Y) = H(X) - H(X|Y) = 1.75 - 1.375 = 0.375$$

5.2 Entropy Proofs

(a)

$$H(X | Y) = \sum_{i=1}^n \sum_{j=1}^m P(x_i, y_j) \log_2 \left(\frac{1}{P(x_i | y_j)} \right)$$

$$I(X; Y) = \sum_{i=1}^n \sum_{j=1}^m P(x_i, y_j) \log_2 \frac{P(x_i | y_j)}{P(x_i)}$$

$$I(X; Y) = \sum_{i=1}^n \sum_{j=1}^m p(x_i, y_j) \log_2 \left(\frac{1}{p(x_i)} \right) - \sum_{i=1}^n \sum_{j=1}^m P(x_i, y_j) \log_2 \left(\frac{1}{P(x_i | y_j)} \right)$$

On here the $H(X|Y) = \sum_{i=1}^n \sum_{j=1}^m P(x_i, y_j) \log_2 \left(\frac{1}{P(x_i | y_j)} \right)$

Therefore

$$I(X; Y) = \sum_{i=1}^n \sum_{j=1}^m P(x_i, y_j) \log_2 \left(\frac{1}{P(x_i)} \right) - H(X | Y)$$

On here

$$\sum_{j=1}^m p(x_i, y_j) = P(x_i)$$

$$I(X; Y) = \sum_{i=1}^n P(x_i) \log_2 \left(\frac{1}{p(x_i)} \right) - H(X | Y)$$

On here $\sum_{i=1}^n P(x_i) \log_2 \left(\frac{1}{p(x_i)} \right) = H(X)$

Therefore, $I(X, Y) = H(X) - H(X|Y)$. And according to the symmetric properties, $H(X) - H(X|Y) = H(Y) - H(Y|X)$, so $I(X, Y) = H(X) - H(X|Y) = H(Y) - H(Y|X)$

(b) Use the formula for joint entropy,

$$H(X, Y) = H(X) + H(Y|X)$$

Use the formula on (a),

$$I(X, Y) = H(X) - H(X|Y)$$

so

$$H(X) = I(X, Y) + H(X|Y)$$

Therefore,

$$H(X, Y) = H(X) + H(Y|X) = I(X, Y) + H(X|Y) + H(Y|X)$$

5.3 Bonus for undergrads

- (a) According to the conditional entropy formula, Given that $H(Y|Z) = 0$ could be write as

$$H(Y|Z) = \sum_i^n P(y_i, Z) \log_2 \frac{p(Z)}{p(y_i, Z)} = 0$$

Therefore, $\log_2 \frac{p(Z)}{p(y_i, Z)} = 0$, so $\frac{p(Z)}{p(y_i, Z)} = 1$ and $p(Z) = p(y, Z)$

Since for any value of Z such $P(Z) > 0$, therefore, for any value of y, there exist such $P(y, Z) > 0$
For any value of Z and Y,because

$$P_{Y|Z}(y_i | z_j) = \frac{P_{YZ}(y_i, z_j)}{P_z(z_j)}$$

, so $P(Y|Z)$ cannot equal to 0. Furthermore, due to

$$p(Z) = p(y, Z) = p(y_i, Z)p(y_i)$$

, there only exists one value of Y such that $P(Y|Z)$ cannot equal to 0.