Local Global Approximations (working title)

John T. Ormerod

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1 Introduction

We will continue to work on the Lasso becuse that is a good place to start.

2 The Lasso Distribution

If $x \sim \text{Lasso}(a, b, c)$ with then it has density given by

$$p(x, a, b, c) = Z^{-1} \exp\left(-\tfrac{1}{2}ax^2 + bx - c|x|\right)$$

where $x \in \mathbb{R}$, $a > 0$, $b \in \mathbb{R}$, $c > 0$, and Z is the normalizing constant. Then

$$\begin{aligned}
Z(a, b, c) &= \int_{-\infty}^{\infty} \exp\left[-\tfrac{1}{2}ax^2 + bx - c|x|\right] dx \\
&= \int_0^{\infty} \exp\left[-\tfrac{1}{2}ax^2 + (b-c)x\right] dx + \int_{-\infty}^{0} \exp\left[-\tfrac{1}{2}ax^2 + (b+c)x\right] dx \\
&= \int_0^{\infty} \exp\left[-\tfrac{1}{2}ax^2 + (b-c)x\right] dx + \int_0^{\infty} \exp\left[-\tfrac{1}{2}ay^2 - (b+c)y\right] dy \\
&= \int_0^{\infty} \exp\left[-\frac{(x-\mu_1)^2}{2\sigma^2} + \frac{\mu_1^2}{2\sigma^2}\right] dx + \int_0^{\infty} \exp\left[-\frac{(x-\mu_2)^2}{2\sigma^2} + \frac{\mu_2^2}{2\sigma^2}\right] dy \\
&= \sqrt{2\pi\sigma^2}\left[\exp\left\{\frac{\mu_1^2}{2\sigma^2}\right\}\int_0^{\infty} \phi(x; \mu_1, \sigma^2) dx + \exp\left\{\frac{\mu_2^2}{2\sigma^2}\right\}\int_0^{\infty} \phi(y; \mu_2, \sigma^2) dy\right] \\
&= \sqrt{2\pi\sigma^2}\left[\exp\left\{\frac{\mu_1^2}{2\sigma^2}\right\}\{1 - \Phi(-\mu_1/\sigma)\} + \exp\left\{\frac{\mu_2^2}{2\sigma^2}\right\}\{1 - \Phi(-\mu_2/\sigma)\}\right] \\
&= \sqrt{2\pi\sigma^2}\left[\exp\left(\frac{\mu_1^2}{2\sigma^2}\right)\Phi\left(\frac{\mu_1}{\sigma}\right) + \exp\left(\frac{\mu_2^2}{2\sigma^2}\right)\Phi\left(\frac{\mu_2}{\sigma}\right)\right] \\
&= \sigma\left[\frac{\Phi(\mu_1/\sigma)}{\phi(\mu_1/\sigma)} + \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)}\right]
\end{aligned}$$

where $\mu_1 = (b-c)/a$, $\mu_2 = -(c+b)/a$ and $\sigma^2 = 1/a$. Care should be taken when evaluating Z (which is prone to overflow and divide by zero problems) and is a function of the Mills ratio.

2.1 The Moment Generating function and Moments

The moment generating function requires almost identical calculations with b replaced with $b+t$.

$$M(t) = \frac{Z(a, b+t, c)}{Z(a, b, c)}$$

While this is true it doesn't look useful for calculating moments.

2.2 Moments

The moments of the Lasso distribution are:

$$\begin{aligned}
E(x^r) &= Z^{-1} \int_{-\infty}^{\infty} x^r \exp\left[-\tfrac{1}{2}ax^2 + bx - c|x|\right] dx \\
&= Z^{-1} \int_{0}^{\infty} x^r \exp\left[-\tfrac{1}{2}ax^2 + (b-c)x\right] dx + \int_{-\infty}^{0} x^r \exp\left[-\tfrac{1}{2}ax^2 + (b+c)x\right] dx \\
&= Z^{-1} \int_{0}^{\infty} x^r \exp\left[-\tfrac{1}{2}ax^2 + (b-c)x\right] dx + (-1)^r \int_{0}^{\infty} y^r \exp\left[-\tfrac{1}{2}ay^2 - (b+c)y\right] dy \\
&= Z^{-1}\sqrt{2\pi\sigma^2} \exp\left(\frac{\mu_1^2}{2\sigma^2}\right) \int_0^\infty x^r \phi(x; \mu_1, \sigma^2) dx \\
&\qquad +(-1)^r \sqrt{2\pi\sigma^2} \exp\left(\frac{\mu_2^2}{2\sigma^2}\right) \int_0^\infty y^r \phi(y; \mu_2, \sigma^2) dy \\
&= \frac{\sigma}{Z}\left[\frac{\Phi(\mu_1/\sigma)}{\phi(\mu_1/\sigma)} \frac{\int_0^\infty x^r \phi(x; \mu_1, \sigma^2)dx}{\Phi(\mu_1/\sigma)} + (-1)^r \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \frac{\int_0^\infty y^r \phi(y; \mu_2, \sigma^2)dy}{\Phi(\mu_2/\sigma)}\right] \\
&= \frac{\sigma}{Z}\left[\frac{\Phi(\mu_1/\sigma)}{\phi(\mu_1/\sigma)} \mathbb{E}(A^r) + (-1)^r \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \mathbb{E}(B^r)\right]
\end{aligned}$$

where $A \sim TN_+(\mu_1, \sigma^2)$, $B \sim TN_+(\mu_2, \sigma^2)$ and TN_+ is denotes the positively truncated normal distribution. Note that

$$\mathbb{E}(A) = \mu_1 + \frac{\sigma\phi(\mu_1/\sigma)}{\Phi(\mu_1/\sigma)} = \mu_1 + \sigma\zeta_1(\mu_1/\sigma)$$

and

$$\mathbb{V}(A) = \sigma^2\left[1 + \zeta_2(\mu_1/\sigma)\right]$$

where $\zeta_k(x) = d^k \log\Phi(x)/dx^k$, $\zeta_1(t) = \phi(t)/\Phi(t)$, $\zeta_2(t) = -t\,\zeta_1(t) - \zeta_1(t)^2$. Here $\zeta_1(x)$ is the inverse Mills ratio which too needs to be treated with care. Hence,

$$\mathbb{E}(A^2) = \mathbb{V}(A) + \mathbb{E}(A)^2 = \sigma^2\left[1 + \zeta_2(\mu_1/\sigma)\right] + \left[\mu_1 + \sigma\zeta_1(\mu_1/\sigma)\right]^2$$

We now have sufficient information to calculate the moments of the Lasso distribution. We also have sufficient information to implement a VB approximation.

2.3 CDF

Similarly if $z \leq 0$ the CDF is given by

$$
\begin{aligned}
P(Z < z) &= Z^{-1} \int_{-\infty}^{z} \exp\left[-\tfrac{1}{2}ax^2 + (b+c)x\right] dx \\
&= Z^{-1}\sqrt{2\pi\sigma^2} \exp\left(\frac{\mu_2^2}{2\sigma^2}\right) \int_{-\infty}^{z} \phi(x; -\mu_2, \sigma^2) \\
&= Z^{-1}\sqrt{2\pi\sigma^2} \exp\left(\frac{\mu_2^2}{2\sigma^2}\right) \Phi\left(\frac{z+\mu_2}{\sigma}\right) \\
&= \frac{\sigma}{Z} \frac{\Phi\left(\dfrac{z+\mu_2}{\sigma}\right)}{\phi(\mu_2/\sigma)}
\end{aligned}
$$

and if $z > 0$ we have

$$
\begin{aligned}
P(Z < z) &= Z^{-1} \int_{-\infty}^{z} \exp\left(-\tfrac{1}{2}ax^2 + bx - c|x|\right) dx \\
&= Z^{-1}\sqrt{2\pi\sigma^2} \left[\exp\left(\frac{\mu_1^2}{2\sigma^2}\right) \int_{0}^{z} \phi(x; \mu_1, \sigma^2) dx + \exp\left(\frac{\mu_2^2}{2\sigma^2}\right) \Phi\left(\frac{\mu_2}{\sigma}\right)\right] \\
&= \frac{\sigma}{Z}\left[\frac{\Phi\left(\frac{z-\mu_1}{\sigma}\right) - \Phi\left(\frac{-\mu_1}{\sigma}\right)}{\phi(\mu_1/\sigma)} + \frac{\Phi\left(\mu_2/\sigma\right)}{\phi(\mu_2/\sigma)}\right]
\end{aligned}
$$

2.4 Inverse CDF

For the inverse CDF we again have two cases. Let $u = P(Z < z)$. When

$$
u \leq \frac{\sigma}{Z} \frac{\Phi\left(\mu_2/\sigma\right)}{\phi(\mu_2/\sigma)}
$$

we solve

$$
u = \frac{\sigma \Phi\left(\dfrac{z+\mu_2}{\sigma}\right)}{Z\phi(\mu_2/\sigma)}
$$

for z to obtain

$$
z = \mu_2 + \sigma\, \Phi^{-1}\left[(Z/\sigma)\phi_\sigma(\mu_2/\sigma)u\right]
$$

When

$$
u > \frac{\sigma}{Z} \frac{\Phi\left(\mu_2/\sigma\right)}{\phi(\mu_2/\sigma)}
$$

we need to solve

$$
u = \frac{\sigma}{Z}\left[\frac{\Phi\left(\frac{z-\mu_1}{\sigma}\right) - \Phi\left(\frac{-\mu_1}{\sigma}\right)}{\phi(\mu_1/\sigma)} + \frac{\Phi\left(\mu_2/\sigma\right)}{\phi(\mu_2/\sigma)}\right]
$$

for z to obtain

$$
z = \mu_1 + \sigma\, \Phi^{-1}\left[\phi(\mu_1/\sigma)\left\{\frac{Zu}{\sigma} - \frac{\Phi\left(\mu_2/\sigma\right)}{\phi(\mu_2/\sigma)}\right\} + \Phi\left(\frac{-\mu_1}{\sigma}\right)\right]
$$

which also involves the Mills ratio.

3 Tasks

Suppose that conformably we have

$$\boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{bmatrix} \qquad \text{and} \qquad \boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{bmatrix}$$

and $\boldsymbol{\theta} = (\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)$ and $q(\boldsymbol{\theta}) \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ approximates $p(\boldsymbol{\theta} \mid \mathcal{D})$.

1. Verify the algebra in Section 2 above for the Lasso distribution.
2. Show that
$$p(\boldsymbol{\theta}_1 \mid \mathcal{D}) = \int p(\boldsymbol{\theta}_1 \mid \mathcal{D}, \boldsymbol{\theta}_2) p(\boldsymbol{\theta}_2 \mid \mathcal{D}) d\boldsymbol{\theta}_2$$
3. Find $q(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)$.

   Solution: The conditional distribution of $\boldsymbol{\theta}_1 | \boldsymbol{\theta}_2$ is
$$\phi\left(\boldsymbol{\theta}_2; \boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1\right), \boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{21}\right)$$
4. Suppose that we approximate $p(\boldsymbol{\theta}_2 \mid \mathcal{D})$ by $q(\boldsymbol{\theta}_2) = N(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$. Suppose we use this to approximate $p(\boldsymbol{\theta}_1 \mid \mathcal{D})$ by
$$q^*(\boldsymbol{\theta}_1) = \int p(\boldsymbol{\theta}_1 \mid \mathcal{D}, \boldsymbol{\theta}_2) q(\boldsymbol{\theta}_2) d\boldsymbol{\theta}_2$$

   Suppose that the mean of $q^*(\boldsymbol{\theta}_1)$ is $\boldsymbol{\mu}_1^*$, covariance is $\boldsymbol{\Sigma}_{11}^*$, and $q^*(\boldsymbol{\theta}_1) \approx N(\boldsymbol{\mu}_1^*, \boldsymbol{\Sigma}_{11}^*)$. We update $q(\boldsymbol{\theta})$ via
$$q^*(\boldsymbol{\theta}) = q(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)\phi(\boldsymbol{\theta}_1; \boldsymbol{\mu}_1^*, \boldsymbol{\Sigma}_{11}^*)$$

   Find $q^*(\boldsymbol{\theta})$.

Solution [The easy way]: We know that $q(\boldsymbol{\theta}_2|\boldsymbol{\theta}_1)$ and $q(\boldsymbol{\theta}_1)$ are Gaussian. Hence, their joint distribution will also be Gaussian. The marginal mean and variance of the joint distribution for $\boldsymbol{\theta}_1$ will be $\boldsymbol{\mu}_1^*$ and $\boldsymbol{\Sigma}_{11}^*$. Suppose the mean and covariance of the joint distribution are $\widetilde{\boldsymbol{\mu}}$ and $\widetilde{\boldsymbol{\Sigma}}$ respectively. Then $\widetilde{\boldsymbol{\mu}}_1 = \boldsymbol{\mu}_1^*$, and $\widetilde{\boldsymbol{\Sigma}}_{11} = \boldsymbol{\Sigma}_{11}^*$. All we need to do is determine $\widetilde{\boldsymbol{\mu}}_2$, $\widetilde{\boldsymbol{\Sigma}}_{22}$ and $\widetilde{\boldsymbol{\Sigma}}_{12}$.

We have

$$\begin{aligned}
\mathbb{E}(\boldsymbol{\theta}_2) &= \mathbb{E}[\mathbb{E}(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)] \\
&= \mathbf{E}[\boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1\right)] \\
&= \boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\mu}_1^* - \boldsymbol{\mu}_1\right).
\end{aligned}$$

Similarly,

$$\begin{aligned}
\mathrm{Cov}(\boldsymbol{\theta}_2) &= \mathbb{E}[\mathrm{Cov}(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)] + \mathrm{Cov}[\mathbb{E}(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)] \\
&= \mathbb{E}(\boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{22}) + \mathrm{Cov}[\boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1\right)] \\
&= \boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{22} + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\mathrm{Cov}(\boldsymbol{\theta}_1)\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12} \\
&= \boldsymbol{\Sigma}_{22} + \boldsymbol{\Sigma}_{21}(\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{11}^*\boldsymbol{\Sigma}_{11}^{-1} - \boldsymbol{\Sigma}_{11}^{-1})\boldsymbol{\Sigma}_{12}.
\end{aligned}$$

Lastly,

$$\begin{aligned}
\mathrm{Cov}(\boldsymbol{\theta}_1, \boldsymbol{\theta}_2) &= \mathbb{E}[(\boldsymbol{\theta} - \mathbb{E}(\boldsymbol{\theta}_1))(\boldsymbol{\theta}_2 - \mathbb{E}(\boldsymbol{\theta}_2))^T] \\
&= \mathbb{E}\left[\mathbb{E}\left\{(\boldsymbol{\theta}_1 - \mathbb{E}(\boldsymbol{\theta}_1))(\boldsymbol{\theta}_2 - \mathbb{E}(\boldsymbol{\theta}_2))^T \mid \boldsymbol{\theta}_1\right\}\right] \\
&= \mathbb{E}\left[(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1^*)\left(\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1^*)\right)^T\right] \\
&= \mathbb{E}\left[(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1^*)\left(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1^*\right)^T\right]\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12} \\
&= \boldsymbol{\Sigma}_{11}^*\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}.
\end{aligned}$$

Hence, $q^*(\boldsymbol{\theta}) = N(\widetilde{\boldsymbol{\mu}}, \widetilde{\boldsymbol{\Sigma}})$ where

$$\widetilde{\boldsymbol{\mu}} = \begin{bmatrix} \boldsymbol{\mu}_1^* \\ \boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\mu}_1^* - \boldsymbol{\mu}_1\right) \end{bmatrix}$$

and

$$\widetilde{\boldsymbol{\Sigma}} = \begin{bmatrix} \boldsymbol{\Sigma}_{11}^* & \boldsymbol{\Sigma}_{11}^*\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{11}^* & \boldsymbol{\Sigma}_{22} + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}(\boldsymbol{\Sigma}_{11}^* - \boldsymbol{\Sigma}_{11})\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12} \end{bmatrix}.$$

Solution [The hard way]: Consider,

$$
\begin{aligned}
&q(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)\phi(\boldsymbol{\theta}_1; \boldsymbol{\mu}_1^*, \boldsymbol{\Sigma}_{11}^*) \\
&\quad \propto \exp\left[-\tfrac{1}{2}\left(\boldsymbol{\theta}_2 - \boldsymbol{\mu}_2 - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1\right)\right)^T \mathbf{Q}\left(\boldsymbol{\theta}_2 - \boldsymbol{\mu}_2 - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\left(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1\right)\right)\right] \\
&\qquad \times \exp\left[-\tfrac{1}{2}(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1^*)^T(\boldsymbol{\Sigma}_{11}^*)^{-1}(\boldsymbol{\theta}_1 - \boldsymbol{\mu}_1^*)\right] \\
&\quad = \exp\left[-\tfrac{1}{2}\boldsymbol{\theta}_2^T\mathbf{Q}\boldsymbol{\theta}_2 + \left(\boldsymbol{\mu}_2 - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right)^T\mathbf{Q}\boldsymbol{\theta}_2\right] \\
&\qquad \times \exp\left[-\tfrac{1}{2}\boldsymbol{\theta}_1^T\left((\boldsymbol{\Sigma}_{11}^*)^{-1} + \boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}\mathbf{Q}\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\right)\boldsymbol{\theta}_1\right] \\
&\qquad \times \exp\left[\left((\boldsymbol{\Sigma}_{11}^*)^{-1}\boldsymbol{\mu}_1^* + \boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}\mathbf{Q}\left(\boldsymbol{\mu}_2 - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right)\right)^T\boldsymbol{\theta}_1\right] \\
&\qquad \times \exp\left[\boldsymbol{\theta}_2^T\mathbf{Q}\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\theta}_1\right]
\end{aligned}
$$

where $\mathbf{Q} = (\boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{22})^{-1}$
Now suppose that $q^*(\boldsymbol{\theta}) = N(\widetilde{\boldsymbol{\mu}}, \widetilde{\boldsymbol{\Sigma}})$ then

$$
\begin{aligned}
q(\boldsymbol{\theta}) &\propto \exp\left[-\frac{1}{2}\left(\begin{pmatrix}\boldsymbol{\theta}_1\\ \boldsymbol{\theta}_2\end{pmatrix} - \begin{pmatrix}\widetilde{\boldsymbol{\mu}}_1\\ \boldsymbol{\mu}_2\end{pmatrix}\right)^T\begin{pmatrix}\widetilde{\mathbf{Q}}_{11} & \widetilde{\mathbf{Q}}_{12}\\ \widetilde{\mathbf{Q}}_{21} & \widetilde{\mathbf{Q}}_{22}\end{pmatrix}\left(\begin{pmatrix}\boldsymbol{\theta}_1\\ \boldsymbol{\theta}_2\end{pmatrix} - \begin{pmatrix}\widetilde{\boldsymbol{\mu}}_1\\ \widetilde{\boldsymbol{\mu}}_2\end{pmatrix}\right)\right] \\
&\propto \exp\left[-\tfrac{1}{2}(\boldsymbol{\theta}_1 - \widetilde{\boldsymbol{\mu}}_1)^T\widetilde{\mathbf{Q}}_{11}(\boldsymbol{\theta}_1 - \widetilde{\boldsymbol{\mu}}_1)\right] \\
&\qquad \times \exp\left[-\tfrac{1}{2}(\boldsymbol{\theta}_2 - \widetilde{\boldsymbol{\mu}}_2)^T\widetilde{\mathbf{Q}}_{22}(\boldsymbol{\theta}_2 - \widetilde{\boldsymbol{\mu}}_2)\right] \\
&\qquad \times \exp\left[-(\boldsymbol{\theta}_1 - \widetilde{\boldsymbol{\mu}}_1)^T\widetilde{\mathbf{Q}}_{12}(\boldsymbol{\theta}_2 - \widetilde{\boldsymbol{\mu}}_2)\right] \\
&\propto \exp\left[-\tfrac{1}{2}\boldsymbol{\theta}_1^T\widetilde{\mathbf{Q}}_{11}\boldsymbol{\theta}_1 + \boldsymbol{\theta}_1^T(\widetilde{\mathbf{Q}}_{11}\widetilde{\boldsymbol{\mu}}_1 + \widetilde{\mathbf{Q}}_{12}\widetilde{\boldsymbol{\mu}}_2)\right] \\
&\qquad \times \exp\left[-\tfrac{1}{2}\boldsymbol{\theta}_2^T\widetilde{\mathbf{Q}}_{22}\boldsymbol{\theta}_2 + \boldsymbol{\theta}_2^T(\widetilde{\mathbf{Q}}_{22}\widetilde{\boldsymbol{\mu}}_2 + \widetilde{\mathbf{Q}}_{21}\widetilde{\boldsymbol{\mu}}_1) - \boldsymbol{\theta}_2^T\widetilde{\mathbf{Q}}_{21}\boldsymbol{\theta}_1\right]
\end{aligned}
$$

where $\widetilde{\mathbf{Q}} = \widetilde{\boldsymbol{\Sigma}}^{-1}$. Matching term by term with the first expression above we find that

$$
\begin{aligned}
\widetilde{\mathbf{Q}}_{21} &= -\mathbf{Q}\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1} \\
\widetilde{\mathbf{Q}}_{11} &= (\boldsymbol{\Sigma}_{11}^*)^{-1} + \boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}\mathbf{Q}\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1} \\
\widetilde{\mathbf{Q}}_{22} &= \mathbf{Q} \\
\widetilde{\mathbf{Q}}_{11}\widetilde{\boldsymbol{\mu}}_1 + \widetilde{\mathbf{Q}}_{12}\widetilde{\boldsymbol{\mu}}_2 &= (\boldsymbol{\Sigma}_{11}^*)^{-1}\boldsymbol{\mu}_1^* + \boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}\mathbf{Q}\left(\boldsymbol{\mu}_2 - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right) \\
\widetilde{\mathbf{Q}}_{22}\widetilde{\boldsymbol{\mu}}_2 + \widetilde{\mathbf{Q}}_{21}\widetilde{\boldsymbol{\mu}}_1 &= \mathbf{Q}\left(\boldsymbol{\mu}_2 - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right)
\end{aligned}
$$

We need to solve this system of equations to find $\widetilde{\mathbf{Q}}_{11}$, $\widetilde{\mathbf{Q}}_{12}$, $\widetilde{\mathbf{Q}}_{22}$, $\widetilde{\boldsymbol{\mu}}_1$ and $\widetilde{\boldsymbol{\mu}}_2$. We then need to use $\widetilde{\mathbf{Q}}_{11}$, $\widetilde{\mathbf{Q}}_{12}$, and $\widetilde{\mathbf{Q}}_{22}$ to find $\widetilde{\boldsymbol{\Sigma}}_{11}$, $\widetilde{\boldsymbol{\Sigma}}_{12}$, $\widetilde{\boldsymbol{\Sigma}}_{22}$.

The block inverse formula states that the inverse of a real matrix can be written as

$$\begin{aligned}\begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix}^{-1} &= \begin{bmatrix} \mathbf{I} & \mathbf{0} \\ -\mathbf{D}^{-1}\mathbf{C} & \mathbf{I} \end{bmatrix} \begin{bmatrix} \widetilde{\mathbf{A}} & \mathbf{0} \\ \mathbf{0} & \mathbf{D}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{I} & -\mathbf{B}\mathbf{D}^{-1} \\ \mathbf{0} & \mathbf{I} \end{bmatrix} &(1)\\ &= \begin{bmatrix} \widetilde{\mathbf{A}} & -\widetilde{\mathbf{A}}\mathbf{B}\mathbf{D}^{-1} \\ -\mathbf{D}^{-1}\mathbf{C}\widetilde{\mathbf{A}} & \mathbf{D}^{-1} + \mathbf{D}^{-1}\mathbf{C}\widetilde{\mathbf{A}}\mathbf{B}\mathbf{D}^{-1}, \end{bmatrix} &(2)\end{aligned}$$

where $\widetilde{\mathbf{A}} = (\mathbf{A} - \mathbf{B}\mathbf{D}^{-1}\mathbf{C})^{-1}$, provided all inverses in (1) and (2) exist.
Using the block inverse formula for the top left element of $\mathbf{\Sigma}$ we have

$$\begin{aligned}\widetilde{\mathbf{\Sigma}}_{11} &= \left(\widetilde{\mathbf{Q}}_{11} - \widetilde{\mathbf{Q}}_{12}\widetilde{\mathbf{Q}}_{22}^{-1}\widetilde{\mathbf{Q}}_{21}\right)^{-1} \\ &= \left((\mathbf{\Sigma}_{11}^*)^{-1} + \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1} - \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\mathbf{Q}^{-1}\mathbf{Q}\mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\right)^{-1} \\ &= \mathbf{\Sigma}_{11}^*\end{aligned}$$

as expected. Similarly, using the bottom right block inverse formula leads to

$$\widetilde{\mathbf{\Sigma}}_{22} = \mathbf{\Sigma}_{22} + \mathbf{\Sigma}_{21}(\mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{11}^*\mathbf{\Sigma}_{11}^{-1} - \mathbf{\Sigma}_{11}^{-1})\mathbf{\Sigma}_{12}$$

so that if $\mathbf{\Sigma}_{11}^* = \mathbf{\Sigma}_{11}$ then there will be no changes to $\mathbf{\Sigma}_{22}$. Finally, using the top left block inverse formula leads to

$$\widetilde{\mathbf{\Sigma}}_{12} = \mathbf{\Sigma}_{11}^*\mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}$$

Using the last equation in the system of equations to solve

$$\begin{aligned}\widetilde{\boldsymbol{\mu}}_2 &= \widetilde{\mathbf{Q}}_{22}^{-1}\mathbf{Q}\left(\boldsymbol{\mu}_2 - \mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right) - \widetilde{\mathbf{Q}}_{22}^{-1}\widetilde{\mathbf{Q}}_{21}\widetilde{\boldsymbol{\mu}}_1 \\ &= \boldsymbol{\mu}_2 + \mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}(\widetilde{\boldsymbol{\mu}}_1 - \boldsymbol{\mu}_1)\end{aligned}$$

Using the 4th equation in the system of equations we have

$$\begin{aligned}\widetilde{\mathbf{Q}}_{11}\widetilde{\boldsymbol{\mu}}_1 &= (\mathbf{\Sigma}_{11}^*)^{-1}\boldsymbol{\mu}_1^* + \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\left(\boldsymbol{\mu}_2 - \mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right) - \widetilde{\mathbf{Q}}_{12}\widetilde{\boldsymbol{\mu}}_2 \\ &= (\mathbf{\Sigma}_{11}^*)^{-1}\boldsymbol{\mu}_1^* + \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\left(\boldsymbol{\mu}_2 - \mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\boldsymbol{\mu}_1\right) - \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\left(\boldsymbol{\mu}_2 + \mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}(\widetilde{\boldsymbol{\mu}}_1 - \boldsymbol{\mu}_1)\right) \\ &= (\mathbf{\Sigma}_{11}^*)^{-1}\boldsymbol{\mu}_1^* - \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\widetilde{\boldsymbol{\mu}}_1\end{aligned}$$

Using the 2nd of the system of equations leads to

$$\left[(\mathbf{\Sigma}_{11}^*)^{-1} + \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\right]\widetilde{\boldsymbol{\mu}}_1 = (\mathbf{\Sigma}_{11}^*)^{-1}\boldsymbol{\mu}_1^* - \mathbf{\Sigma}_{11}^{-1}\mathbf{\Sigma}_{12}\mathbf{Q}\mathbf{\Sigma}_{21}\mathbf{\Sigma}_{11}^{-1}\widetilde{\boldsymbol{\mu}}_1$$

After some cancellation we have $\widetilde{\boldsymbol{\mu}}_1 = \boldsymbol{\mu}_1^*$.

5. Consider the multivariate Lasso distribution

$$p\left(\mathbf{x}\right) \propto \exp\left(-\tfrac{1}{2}\mathbf{x}^T\mathbf{A}\mathbf{x} + \mathbf{b}^T\mathbf{x} - c\|\mathbf{x}\|_1\right)$$

where $\mathbf{A} \in \mathcal{S}_d^+$ is a positive definite matrix of dimension d, $\mathbf{b} \in \mathbb{R}^2$ and $c > 0$ with $\|\mathbf{x}\|_1 = \sum_{j=1}^d |x_j|$. For the case $d = 2$

- Find the normalizing constant.
- Find an expression for the expectation and the covariance.