

Local Global Approximations (working title)

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1 Introduction

We will continue to work on the Lasso because that is a good place to start.

2 The Lasso Distribution

If $x \sim \text{Lasso}(a, b, c)$ with then it has density given by

$$p(x, a, b, c) = Z^{-1} \exp \left(-\frac{1}{2}ax^2 + bx - c|x| \right)$$

where $x \in \mathbb{R}$, $a > 0$, $b \in \mathbb{R}$, $c > 0$, and Z is the normalizing constant. Then

$$\begin{aligned} Z(a, b, c) &= \int_{-\infty}^{\infty} \exp \left[-\frac{1}{2}ax^2 + bx - c|x| \right] dx \\ &= \int_0^{\infty} \exp \left[-\frac{1}{2}ax^2 + (b - c)x \right] dx + \int_{-\infty}^0 \exp \left[-\frac{1}{2}ax^2 + (b + c)x \right] dx \\ &= \int_0^{\infty} \exp \left[-\frac{1}{2}ax^2 + (b - c)x \right] dx + \int_0^{\infty} \exp \left[-\frac{1}{2}ay^2 - (b + c)y \right] dy \\ &= \int_0^{\infty} \exp \left[-\frac{(x - \mu_1)^2}{2\sigma^2} + \frac{\mu_1^2}{2\sigma^2} \right] dx + \int_0^{\infty} \exp \left[-\frac{(x - \mu_2)^2}{2\sigma^2} + \frac{\mu_2^2}{2\sigma^2} \right] dy \\ &= \sqrt{2\pi\sigma^2} \left[\exp \left\{ \frac{\mu_1^2}{2\sigma^2} \right\} \int_0^{\infty} \phi(x; \mu_1, \sigma^2) dx + \exp \left\{ \frac{\mu_2^2}{2\sigma^2} \right\} \int_0^{\infty} \phi(y; \mu_2, \sigma^2) dy \right] \\ &= \sqrt{2\pi\sigma^2} \left[\exp \left\{ \frac{\mu_1^2}{2\sigma^2} \right\} \{1 - \Phi(-\mu_1/\sigma)\} + \exp \left\{ \frac{\mu_2^2}{2\sigma^2} \right\} \{1 - \Phi(-\mu_2/\sigma)\} \right] \\ &= \sqrt{2\pi\sigma^2} \left[\exp \left(\frac{\mu_1^2}{2\sigma^2} \right) \Phi \left(\frac{\mu_1}{\sigma} \right) + \exp \left(\frac{\mu_2^2}{2\sigma^2} \right) \Phi \left(\frac{\mu_2}{\sigma} \right) \right] \\ &= \sigma \left[\frac{\Phi(\mu_1/\sigma)}{\phi(\mu_1/\sigma)} + \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \right] \end{aligned}$$

where $\mu_1 = (b - c)/a$, $\mu_2 = -(c + b)/a$ and $\sigma^2 = 1/a$. Care should be taken when evaluating Z (which is prone to overflow and divide by zero problems) and is a function of the [Mills ratio](#).

2.1 The Moment Generating function and Moments

The moment generating function requires almost identical calculations with b replaced with $b + t$.

$$M(t) = \frac{Z(a, b + t, c)}{Z(a, b, c)}$$

While this is true it doesn't look useful for calculating moments.

2.2 Moments

The moments of the Lasso distribution are:

$$\begin{aligned} E(x^r) &= Z^{-1} \int_{-\infty}^{\infty} x^r \exp \left[-\frac{1}{2}ax^2 + bx - c|x| \right] dx \\ &= Z^{-1} \int_0^{\infty} x^r \exp \left[-\frac{1}{2}ax^2 + (b - c)x \right] dx + \int_{-\infty}^0 x^r \exp \left[-\frac{1}{2}ax^2 + (b + c)x \right] dx \\ &= Z^{-1} \int_0^{\infty} x^r \exp \left[-\frac{1}{2}ax^2 + (b - c)x \right] dx + (-1)^r \int_0^{\infty} y^r \exp \left[-\frac{1}{2}ay^2 - (b + c)y \right] dy \\ &= Z^{-1} \sqrt{2\pi\sigma^2} \exp \left(\frac{\mu_1^2}{2\sigma^2} \right) \int_0^{\infty} x^r \phi(x; \mu_1, \sigma^2) dx \\ &\quad + (-1)^r \sqrt{2\pi\sigma^2} \exp \left(\frac{\mu_2^2}{2\sigma^2} \right) \int_0^{\infty} y^r \phi(y; \mu_2, \sigma^2) dy \\ &= \frac{\sigma}{Z} \left[\frac{\Phi(\mu_1/\sigma)}{\phi(\mu_1/\sigma)} \frac{\int_0^{\infty} x^r \phi(x; \mu_1, \sigma^2) dx}{\Phi(\mu_1/\sigma)} + (-1)^r \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \frac{\int_0^{\infty} y^r \phi(y; \mu_2, \sigma^2) dy}{\Phi(\mu_2/\sigma)} \right] \\ &= \frac{\sigma}{Z} \left[\frac{\Phi(\mu_1/\sigma)}{\phi(\mu_1/\sigma)} \mathbb{E}(A^r) + (-1)^r \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \mathbb{E}(B^r) \right] \end{aligned}$$

where $A \sim TN_+(\mu_1, \sigma^2)$, $B \sim TN_+(\mu_2, \sigma^2)$ and TN_+ is denotes the positively truncated normal distribution. Note that

$$\mathbb{E}(A) = \mu_1 + \frac{\sigma\phi(\mu_1/\sigma)}{\Phi(\mu_1/\sigma)} = \mu_1 + \sigma\zeta_1(\mu_1/\sigma)$$

and

$$\mathbb{V}(A) = \sigma^2 [1 + \zeta_2(\mu_1/\sigma)]$$

where $\zeta_k(x) = d^k \log \Phi(x)/dx^k$, $\zeta_1(t) = \phi(t)/\Phi(t)$, $\zeta_2(t) = -t\zeta_1(t) - \zeta_1(t)^2$. Here $\zeta_1(x)$ is the inverse Mills ratio which too needs to be treated with care. Hence,

$$\mathbb{E}(A^2) = \mathbb{V}(A) + \mathbb{E}(A)^2 = \sigma^2 [1 + \zeta_2(\mu_1/\sigma)] + [\mu_1 + \sigma\zeta_1(\mu_1/\sigma)]^2$$

We now have sufficient information to calculate the moments of the Lasso distribution. We also have sufficient information to implement a VB approximation.

2.3 CDF

Similarly if $z \leq 0$ the CDF is given by

$$\begin{aligned}
 P(Z < z) &= Z^{-1} \int_{-\infty}^z \exp \left[-\frac{1}{2}ax^2 + (b+c)x \right] dx \\
 &= Z^{-1} \sqrt{2\pi\sigma^2} \exp \left(\frac{\mu_2^2}{2\sigma^2} \right) \int_{-\infty}^z \phi(x; \mu_2, \sigma^2) \\
 &= Z^{-1} \sqrt{2\pi\sigma^2} \exp \left(\frac{\mu_2^2}{2\sigma^2} \right) \Phi \left(\frac{z + \mu_2}{\sigma} \right) \\
 &= \frac{\sigma}{Z} \frac{\Phi \left(\frac{z + \mu_2}{\sigma} \right)}{\phi(\mu_2/\sigma)}
 \end{aligned}$$

and if $z > 0$ we have

$$\begin{aligned}
 P(Z < z) &= Z^{-1} \int_{-\infty}^z \exp \left(-\frac{1}{2}ax^2 + bx - c|x| \right) dx \\
 &= Z^{-1} \sqrt{2\pi\sigma^2} \left[\exp \left(\frac{\mu_1^2}{2\sigma^2} \right) \int_0^z \phi(x; \mu_1, \sigma^2) dx + \exp \left(\frac{\mu_2^2}{2\sigma^2} \right) \Phi \left(\frac{\mu_2}{\sigma} \right) \right] \\
 &= \frac{\sigma}{Z} \left[\frac{\Phi \left(\frac{z - \mu_1}{\sigma} \right) - \Phi \left(\frac{-\mu_1}{\sigma} \right)}{\phi(\mu_1/\sigma)} + \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \right]
 \end{aligned}$$

2.4 Inverse CDF

For the inverse CDF we again have two cases. Let $u = P(Z < z)$. When

$$u \leq \frac{\sigma}{Z} \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)}$$

we solve

$$u = \frac{\sigma \Phi \left(\frac{z + \mu_2}{\sigma} \right)}{Z \phi(\mu_2/\sigma)}$$

for z to obtain

$$z = \mu_2 + \sigma \Phi^{-1} \left[\left(\frac{Z}{\sigma} \right) \phi(\mu_2/\sigma) u \right]$$

When

$$u > \frac{\sigma}{Z} \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)}$$

we need to solve

$$u = \frac{\sigma}{Z} \left[\frac{\Phi \left(\frac{z - \mu_1}{\sigma} \right) - \Phi \left(\frac{-\mu_1}{\sigma} \right)}{\phi(\mu_1/\sigma)} + \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \right]$$

for z to obtain

$$z = \mu_1 + \sigma \Phi^{-1} \left[\phi(\mu_1/\sigma) \left\{ \frac{Zu}{\sigma} - \frac{\Phi(\mu_2/\sigma)}{\phi(\mu_2/\sigma)} \right\} + \Phi \left(\frac{-\mu_1}{\sigma} \right) \right]$$

which also involves the Mills ratio.

3 Tasks

Suppose that conformably we have

$$\boldsymbol{\mu} = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \quad \text{and} \quad \boldsymbol{\Sigma} = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}$$

and $\boldsymbol{\theta} = (\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)$ and $q(\boldsymbol{\theta}) \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ approximates $p(\boldsymbol{\theta} \mid \mathcal{D})$.

1. Verify the algebra in Section 2 above for the Lasso distribution.
2. Show that

$$p(\boldsymbol{\theta}_1 \mid \mathcal{D}) = \int p(\boldsymbol{\theta}_1 \mid \mathcal{D}, \boldsymbol{\theta}_2) p(\boldsymbol{\theta}_2 \mid \mathcal{D}) d\boldsymbol{\theta}_2$$

3. Find $q(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1)$.
4. Suppose that we approximate $p(\boldsymbol{\theta}_2 \mid \mathcal{D})$ by $q(\boldsymbol{\theta}_2) = N(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$. Suppose we use this to approximate $p(\boldsymbol{\theta}_1 \mid \mathcal{D})$ by

$$q^*(\boldsymbol{\theta}_1) = \int p(\boldsymbol{\theta}_1 \mid \mathcal{D}, \boldsymbol{\theta}_2) q(\boldsymbol{\theta}_2) d\boldsymbol{\theta}_2$$

Suppose that the mean of $q^*(\boldsymbol{\theta}_1)$ is $\boldsymbol{\mu}_1^*$, covariance is $\boldsymbol{\Sigma}_{11}^*$, and $q^*(\boldsymbol{\theta}_1) \approx N(\boldsymbol{\mu}_1^*, \boldsymbol{\Sigma}_{11}^*)$. We update $q(\boldsymbol{\theta})$ via

$$q^*(\boldsymbol{\theta}) = q(\boldsymbol{\theta}_2 \mid \boldsymbol{\theta}_1) \phi(\boldsymbol{\theta}_1; \boldsymbol{\mu}_1^*, \boldsymbol{\Sigma}_{11}^*)$$

Find $q^*(\boldsymbol{\theta})$.

5. Consider the multivariate Lasso distribution

$$p(\mathbf{x}) \propto \exp\left(-\frac{1}{2} \mathbf{x}^T \mathbf{A} \mathbf{x} + \mathbf{b}^T \mathbf{x} - c \|\mathbf{x}\|_1\right)$$

where $\mathbf{A} \in \mathcal{S}_d^+$ is a positive definite matrix of dimension d , $\mathbf{b} \in \mathbb{R}^d$ and $c > 0$ with $\|\boldsymbol{\beta}\|_1 = \sum_{j=1}^d |x_j|$. For the case $d = 2$

- Find the normalizing constant.
- Find an expression for the expectation and the covariance.