Bivariate Lasso Distribution Derivation

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1 Setting

If $x \sim \text{MultiLasso}(A, b, c)$ with then it has density given by

$$p(x, a, b, c) = Z^{-1} exp(-\frac{1}{2}x^{T}Ax + b^{T}x - c||x||_{1})$$

where $A \in S_d^+$: positive definite matrix with dimension d, $b \in \mathbb{R}^2$, c > 0

2 Finding Normalizing Constant

$$\begin{split} Z(a,b,c) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + \mathbf{b}^{T}x - c\mathbf{1}^{T}|x|\right] d\mathbf{x} \\ &= \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{-\infty}^{0} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{-\infty}^{0} \int_{-\infty}^{0} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{-\infty}^{0} \int_{-\infty}^{0} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{-\infty}^{0} \int_{-\infty}^{0} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax - (\mathbf{b}^{T} + c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax - (\mathbf{b}^{T} + c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} + \int_{0}^{\infty} \int_{0}^{\infty} \exp\left[-\frac{1}{2}(x^{T}Ax - 2(\mathbf{b$$

where $\mu_1 = A^{-1}(b-c\mathbf{1})^T$, $\mu_2 = A^{*-1}(b_1-c, -b_2-c)^T$, $\mu_3 = A^{*-1}(-b_1-c, b_2-c)^T$ $\mu_4 = A^{-1}(-b-c\mathbf{1}^T)^T$ and $\Sigma_1 = A^{-1}$, $\Sigma_2 = A^{*-1}$ $A^* = A \otimes \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix}$.

3 Find Expectation

Follow similar step as before

$$\begin{split} E[X] &= Z^{-1} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \otimes \exp\left[-\frac{1}{2}x^{T}Ax + \mathbf{b}^{T}x - c\mathbf{1}^{T}||x||_{1}\right] d\mathbf{x} \\ &= Z^{-1} \int_{0}^{\infty} \int_{0}^{\infty} x \otimes \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} - c\mathbf{1}^{T})x\right] d\mathbf{x} \\ &+ \int_{0}^{\infty} \int_{0}^{\infty} [1, -1]^{T} \otimes x \otimes \exp\left[-\frac{1}{2}x^{T}A \otimes \begin{bmatrix} 1 & -1 \\ -1 & 1 \\ 1 & 1 \end{bmatrix} x + (b_{1} - c, -b_{2} - c)^{T}x \right] d\mathbf{x} \\ &+ \int_{0}^{\infty} \int_{0}^{\infty} [-1, 1]^{T} \otimes x \otimes \exp\left[-\frac{1}{2}x^{T}A \otimes \begin{bmatrix} 1 & -1 \\ -1 & 1 \\ 1 & 1 \end{bmatrix} x + (-b_{1} - c, b_{2} - c)^{T}x \right] d\mathbf{x} \\ &- \int_{0}^{\infty} \int_{0}^{\infty} x \otimes \exp\left[-\frac{1}{2}x^{T}Ax + (\mathbf{b}^{T} + c\mathbf{1}^{T})x\right] d\mathbf{x} \\ &= Z^{-1}[|\Sigma_{1}|(\frac{\int_{0}^{\infty} \int_{0}^{\infty} x \otimes \phi_{2}(x;\mu_{1},\Sigma) d\mathbf{x}}{\phi_{2}(A\mu_{1},\Sigma_{1}^{-1})} - \frac{\int_{0}^{\infty} \int_{0}^{\infty} x \otimes \phi_{2}(x;\mu_{4},\Sigma) d\mathbf{x}}{\phi_{2}(A\mu_{4},\Sigma_{1}^{-1})}) + |\Sigma_{2}|[1,-1]^{T}(\frac{\int_{0}^{\infty} \int_{0}^{\infty} x \otimes \phi_{2}(x;\mu_{2},\Sigma) d\mathbf{x}}{\phi_{2}(A^{*}\mu_{2},\Sigma_{2}^{-1})}) \\ &+ [-1,1]^{T} \frac{\int_{0}^{\infty} \int_{0}^{\infty} x \otimes \phi_{2}(x;\mu_{1},\Sigma) d\mathbf{x}}{\phi_{2}(A^{\mu}_{1},\Sigma_{1}^{-1})} - \frac{E[D] \int_{0}^{\infty} \int_{0}^{\infty} x \otimes \phi_{2}(x;\mu_{4},\Sigma_{1}) d\mathbf{x}}{\phi_{2}(A\mu_{1},\Sigma_{1}^{-1})}) \\ &+ |\Sigma_{2}|([1,-1]^{T} \frac{E[B] \int_{0}^{\infty} \int_{0}^{\infty} x \otimes \phi_{2}(x;\mu_{2},\Sigma_{2}) d\mathbf{x}}{\phi_{2}(A^{*}\mu_{2},\Sigma_{2}^{-1})} + [-1,1]^{T} \frac{E[C] \int_{0}^{\infty} \int_{0}^{\infty} \phi_{2}(x;\mu_{3},\Sigma_{2}) d\mathbf{x}}{\phi_{2}(A^{*}\mu_{2},\Sigma_{2}^{-1})})] \end{split}$$

where $\mu_1 = A^{-1}(b-c\mathbf{1})^T$, $\mu_2 = A^{*-1}(b_1-c, -b_2-c)^T$, $\mu_3 = A^{*-1}(-b_1-c, b_2-c)^T$, $\mu_4 = A^{-1}(-b-c\mathbf{1}^T)^T$ and $\Sigma_1 = A^{-1}$, $\Sigma_2 = A^{*-1}$, $A^* = A \otimes \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix}$. $\mathbf{A} \sim MTN_+(\mu_1, \Sigma_1)$, $B \sim MTN_+(\mu_2, \Sigma)$, $C \sim MTN_+(\mu_3, \Sigma_2)$, $D \sim MTN_+(\mu_4, \Sigma_1)$ is denotes the multivariate positively truncated normal distribution.

4 Find Covariance Matrix

Follow similar steps as before

$$\begin{split} Cov(X) &= E[XX^T] - E[X]E[X]^T \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xx^T \otimes \exp\left[-\frac{1}{2}x^TAx + \mathbf{b}^Tx - c\mathbf{1}^T||x||_1\right] d\mathbf{x} \\ &= Z^{-1}2\pi|\Sigma|^{\frac{1}{2}} [\exp\left[\frac{(A\mu_1)^T\Sigma(A\mu_1)]}{2}\right] \int_{0}^{\infty} \int_{0}^{\infty} xx^T \otimes \phi_2(x;\mu_1,\Sigma) d\mathbf{x} \\ &+ \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix} \otimes \exp\left[\frac{(A\mu_2)^T\Sigma(A\mu_2)]}{2}\right] \int_{0}^{\infty} \int_{0}^{\infty} xx^T \otimes \phi_2(x;\mu_2,\Sigma) d\mathbf{x} \\ &+ \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix} \otimes \exp\left[\frac{(A\mu_3)^T\Sigma(A\mu_3)]}{2}\right] \int_{0}^{\infty} \int_{0}^{\infty} xx^T \otimes \phi_2(x;\mu_3,\Sigma) d\mathbf{x} \\ &- \exp\left[\frac{(A\mu_4)^T\Sigma(A\mu_4)]}{2}\right] \int_{0}^{\infty} \int_{0}^{\infty} xx^T \otimes \phi_2(x;\mu_4,\Sigma) d\mathbf{x} \\ &= Z^{-1}|\Sigma| [\frac{\int_{0}^{\infty} \int_{0}^{\infty} xx^T\otimes\phi_2(x;\mu_1,\Sigma) d\mathbf{x}}{\phi_2(A\mu_1,\Sigma))} + \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix} \otimes \frac{\int_{0}^{\infty} \int_{0}^{\infty} xx^T\otimes\phi_2(x;\mu_2,\Sigma) d\mathbf{x}}{\phi_2(A\mu_4,\Sigma))} \\ &+ \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix} \otimes \frac{\int_{0}^{\infty} \int_{0}^{\infty} xx^T\otimes\phi_2(x;\mu_3,\Sigma) d\mathbf{x}}{\phi_2(A\mu_3,\Sigma))} - \frac{\int_{0}^{\infty} \int_{0}^{\infty} xx^T\otimes\phi_2(x;\mu_4,\Sigma) d\mathbf{x}}{\phi_2(A\mu_4,\Sigma)}] \\ &= Z^{-1}[|\Sigma_1|(\frac{E[A]\Phi_2(\frac{\mu_1}{\Sigma_1},\rho_1)}{\phi_2(A\mu_1,\Sigma_1^{-1})}) - \frac{E[D]\Phi_2(\frac{\mu_4}{\Sigma_2},\rho_2)}{\phi_2(A\mu_4,\Sigma_1^{-1})}) \\ &+ |\Sigma_2|(\begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix} \otimes \frac{E[C]\Phi_2(\frac{\mu_3}{\Sigma_2},\rho_2)}{\phi_2(A^*\mu_2,\Sigma_2^{-1})}) + \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix} \otimes \frac{E[C]\Phi_2(\frac{\mu_3}{\Sigma_2},\rho_2)}{\phi_2(A^*\mu_3,\Sigma_2^{-1})})] \end{split}$$

where $\mu_1 = A^{-1}(b-c\mathbf{1})^T$, $\mu_2 = A^{*-1}(b_1-c, -b_2-c)^T$, $\mu_3 = A^{*-1}(-b_1-c, b_2-c)^T\mu_4 = A^{-1}(-b-c\mathbf{1}^T)^T$ and $\Sigma_1 = A^{-1}$, $\Sigma_2 = A^{*-1}$ $A^* = A \otimes \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix}$. $\mathbf{A} \sim MTN_+(\mu_1, \Sigma_1)$, $B \sim MTN_+(\mu_2, \Sigma)$, $C \sim MTN_+(\mu_3, \Sigma_2)$, $D \sim MTN_+(\mu_4, \Sigma_1)$ is denotes the multivariate positively truncated normal distribution.

$$E[AA^T] = Cov(A) - E[A]E[A]^T$$