## **Problem Set 4 Solution**

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## Question 1.

(a)

$$\begin{split} E\left[\Delta u_{it}(\Delta u_{it-1} + \Delta u_{it} + \Delta u_{it+1})\right] &= & E\left[(\zeta_{it} + m_{it})(\zeta_{it-1} + \zeta_{it} + \zeta_{it+1} - m_{it-1} + m_{it+1})\right] = \sigma_{\zeta}^{2} \\ E\left[\Delta u_{it}\Delta u_{it-1}\right] &= & E\left[(\zeta_{it} + m_{it} - m_{it-1})(\zeta_{it-1} + m_{it-1} - m_{it-2})\right] = -\sigma_{m}^{2} \end{split}$$

(b) Note that  $\alpha_{it}$  is a random variable.

$$\begin{split} E\left[\Delta u_{it}|L_{it} = 1, L_{it-1} = 1\right] &= \sigma_{\zeta\eta} E\left[\frac{\varphi(\alpha_{it})}{1 - \Phi(\alpha_{it})}\right] \\ E\left[\Delta u_{it}(\Delta u_{it-1} + \Delta u_{it} + \Delta u_{it+1})|L_{it-2} = 1, L_{it-1} = 1, L_{it} = 1, L_{it+1} = 1\right] &= \sigma_{\zeta}^2 + \sigma_{\zeta\eta}^2 E\left[\frac{\varphi(\alpha_{it})}{1 - \Phi(\alpha_{it})}\alpha_{it}\right] \\ E\left[\Delta u_{it}^2|L_{it} = 1, L_{it-1} = 1\right] &= \sigma_{\zeta}^2 + \sigma_{\zeta\eta}^2 E\left[\frac{\varphi(\alpha_{it})}{1 - \Phi(\alpha_{it})}\alpha_{it}\right] + 2\sigma_m^2 \end{split}$$

To see this,

$$\begin{split} E\left[\Delta u_{it}|L_{it}=1,L_{it-1}=1\right] &= & E\left[E\left[\zeta_{it}+m_{it}-m_{it-1}|\eta_{it}>\alpha_{it},\eta_{it-1}>\alpha_{it-1}\right]\right] \\ &= & E\left[E\left[\zeta_{it}|\eta_{it}>\alpha_{it}\right]\right] \\ &= & \sigma_{\zeta\eta}E\left[\frac{\varphi(\alpha_{it})}{1-\Phi(\alpha_{it})}\right] (\because truncated bivariate normal distribution formula) \end{split}$$

$$\begin{split} & E\left[\Delta u_{it}(\Delta u_{it-1} + \Delta u_{it} + \Delta u_{it+1})|L_{it-2} = 1, L_{it-1} = 1, L_{it} = 1, L_{it+1} = 1\right] \\ & = E\left[E\left[\zeta_{it}^2|\eta_{it} > \alpha_{it}\right]\right] \\ & = \sigma_\zeta^2 E\left[E\left[\zeta_{it}^2|\eta_{it} > \alpha_{it}\right]\right] \quad (\because \tilde{\zeta}_{it} \equiv \zeta_{it}/\sigma_\zeta) \\ & = \sigma_\zeta^2 E\left[E\left[E\left[\zeta_{it}^2|\eta_{it}\right]|\eta_{it} > \alpha_{it}\right]\right] \\ & = \sigma_\zeta^2 E\left[E\left[Var(\tilde{\zeta}_{it}|\eta_{it}) + E\left[\tilde{\zeta}_{it}|\eta_{it}\right]^2|\eta_{it} > \alpha_{it}\right]\right] \\ & = \sigma_\zeta^2 E\left[E\left[1 - \rho_{\zeta\eta}^2 + \rho_{\zeta\eta}^2 \eta_{it}^2|\eta_{it} > \alpha_{it}\right]\right] \\ & = \sigma_\zeta^2 \left[1 - \rho_{\zeta\eta}^2 + \rho_{\zeta\eta}^2 E\left[E\left[\eta_{it}^2|\eta_{it} > \alpha_{it}\right]\right]\right) \end{split}$$

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$$\begin{split} &= \sigma_{\zeta}^{2} (1 - \rho_{\zeta\eta}^{2} + \rho_{\zeta\eta}^{2} (E \left[ \alpha_{it} \frac{\varphi(\alpha_{it})}{1 - \Phi(\alpha_{it})} \right] + 1)) \\ &= \sigma_{\zeta}^{2} (1 + \rho_{\zeta\eta}^{2} E \left[ \alpha_{it} \frac{\varphi(\alpha_{it})}{1 - \Phi(\alpha_{it})} \right]) \\ &= E \left[ \Delta u_{it}^{2} | L_{it} = 1, L_{it-1} = 1 \right] \\ &= E \left[ E \left[ (\zeta_{it} + m_{it} - m_{it-1})^{2} | \eta_{it} > \alpha_{it}, \eta_{it-1} > \alpha_{it-1} \right] \right] \\ &= E \left[ E \left[ \zeta_{it}^{2} | \eta_{it} > \alpha_{it} \right] \right] + 2 \sigma_{m}^{2} \\ &= \sigma_{\zeta}^{2} (1 + \rho_{\zeta\eta}^{2} E \left[ \alpha_{it} \frac{\varphi(\alpha_{it})}{1 - \Phi(\alpha_{it})} \right]) + 2 \sigma_{m}^{2} \end{split}$$

## Question 2.

(a)

```
workeq <- glm(work ~ age + age2 + factor(educ) + noveliv, family = binomial(link = "probit"), data
  # (a)
  alphahat <- -cbind(rep(1,nobs),data$age,data$age^2,(data$educ=="College_Graduate"),</pre>
                        (data$educ=="Postgraduate"), (data$educ=="Some_College"),
                       data$noveliv)%*%workeq$coefficients
  lambdahat <- dnorm(alphahat)/(1-pnorm(alphahat))</pre>
(b)
                          \log w_{it} - \log w_{it-1} = \Delta Z'_{it} \beta + \zeta_{it} + m_{it} + m_{it-1}
                                                                                              (0.1)
reg1 <- lm(d1lnwage ~ d1age2 + lambdahat)
# compute residual and find uhat
duhat <- d1lnwage - reg1$coefficients[1] - reg1$coefficients[2]*d1age2</pre>
(c)
> paramest
[1] 0.03119798 0.03159257 0.01941181
> stderr
[1] \ 0.0010169239 \ 0.0125732414 \ 0.0006523193
```