

MATH180C: Introduction to Stochastic Processes II

www.math.ucsd.edu/~ynemish/teaching/180c

Today: Kolmogorov's equations

> Q&A: October 21

Next: PK 6.4, 6.6, Durrett 4.3

This week:

- Quiz 2 on Wednesday, October 21 (lectures 4-6)
- Homework 2 (due Friday, October 23, 11:59 PM)



Kolmogorov equations

Jump and hold description is very intuitive, gives a very clear picture of the process, but does not answer to some very basic questions, e.g., computing $P_{ij}(t) := P(X_t = j | X_0 = i)$.

For computing the transition probabilities the differential equation approach is more appropriate.

In order to derive the system of differential equations for $P_{ij}(t)$ from the infinitesimal description, we start from the familiar relation:

Chapman-Kolmogorov equation (semigroup property)

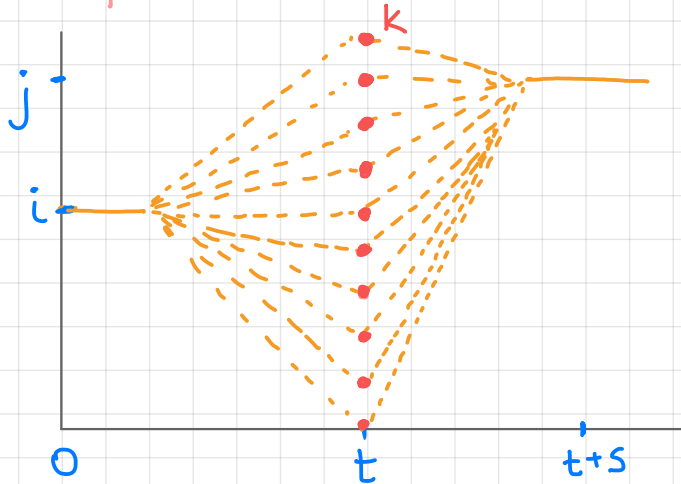
Chapman-Kolmogorov equation

$$P_{ij}(t+s) = P(X_{t+s} = j \mid X_0 = i) \quad \text{condition on the value of } X_t$$

=

Markov =

stationary
trans. prob.



Or in matrix form

Kolmogorov forward equations

Apply Chapman-Kolmogorov equations to compute

$$P_{ij}(t+h):$$

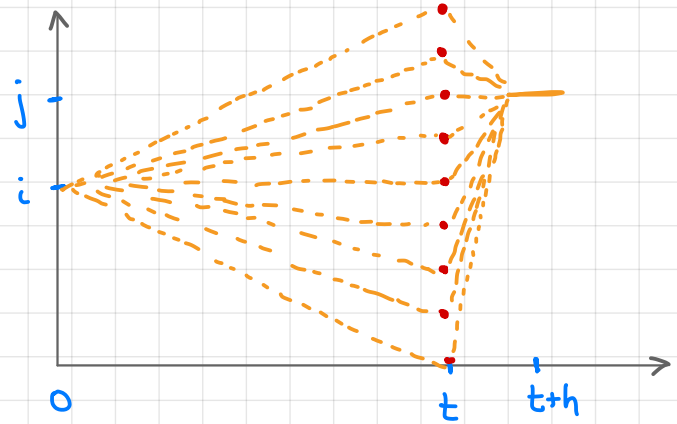
$$P_{ij}(t+h) =$$

Use infinitesimal description:

$$P_{kj}(h) = \begin{cases} q_{kj}h + o(h), & k \neq j \\ 1 + q_{jj}h + o(h), & k = j \end{cases}$$

$$(*) =$$

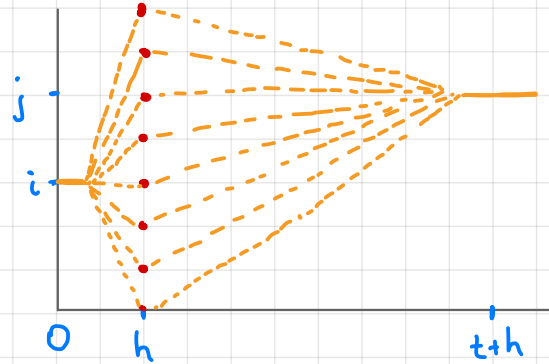
$$=$$



$$\frac{d}{dt}P(t) = P(t)Q$$

Kolmogorov backward equations

$$\begin{aligned} P_{ij}(t+h) &= \sum_{k=0}^N P_{ik}(h) P_{kj}(t) \\ &= (1 + q_{ii}h + o(h)) P_{ij}(t) \\ &\quad + \sum_{\substack{k=0 \\ k \neq i}}^N (q_{ik}h + o(h)) P_{kj} \end{aligned}$$



$$= P_{ij}(t) + \sum_{k=0}^N q_{ik} P_{kj}(t) h + o(h)$$

↳



Kolmogorov equations. Remarks

1. e^{tQ} satisfies both (forward and backward) equations. Indeed, omitting technical details, differentiate term-by-term

$$\frac{d}{dt} e^{tQ} = \frac{d}{dt} \left(\sum_{k=0}^{\infty} \frac{Q^k t^k}{k!} \right) =$$

$$\text{Now } \sum_{k=1}^{\infty} \frac{Q^k}{(k-1)!} t^{k-1} \stackrel{\ell=k-1}{=} \sum_{\ell=0}^{\infty} \frac{Q^{\ell+1}}{\ell!} t^{\ell} =$$

2. Redundancy is related to the stationarity of transition probabilities. If transition probabilities

$P_{ij}(s, t) = P(X_t = j \mid X_s = i)$ are not stationary, then

$$\frac{\partial}{\partial t} P_{ij}(s, t) \rightarrow \text{forward equation}, \quad \frac{\partial}{\partial s} P_{ij}(s, t) \rightarrow \text{backward equation}$$

Example

Two-state MC

$$Q = \begin{pmatrix} -\alpha & \alpha \\ \beta & -\beta \end{pmatrix}$$

$$Q^2 = \begin{pmatrix} -\alpha & \alpha \\ \beta & -\beta \end{pmatrix} \begin{pmatrix} -\alpha & \alpha \\ \beta & -\beta \end{pmatrix} = \begin{pmatrix} \alpha(\alpha+\beta) & -\alpha(\alpha+\beta) \\ -\beta(\alpha+\beta) & \beta(\alpha+\beta) \end{pmatrix} =$$

\hookrightarrow

$$e^{tQ} = \sum_{k=0}^{\infty} \frac{Q^k t^k}{k!} =$$

=

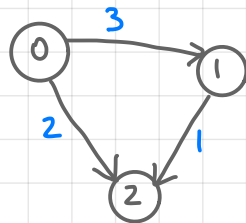
=

$$= I + \frac{1}{\alpha+\beta} Q - \frac{1}{\alpha+\beta} e^{-(\alpha+\beta)t} Q$$

Example

Let $(X_t)_{t \geq 0}$ be a MC with generator Q

$$Q = \begin{pmatrix} -5 & 3 & 2 \\ 0 & -1 & 1 \\ 0 & 0 & 0 \end{pmatrix}$$



Compute $P_{01}(t)$

For any k , $Q^k = \begin{pmatrix} * & * & * \\ 0 & * & * \\ 0 & 0 & * \end{pmatrix}$, \Rightarrow

$$P'(t) = \begin{pmatrix} P_{00} & P_{01} & P_{02} \\ 0 & P_{11} & P_{12} \\ 0 & 0 & P_{22} \end{pmatrix} \begin{pmatrix} -5 & 3 & 2 \\ 0 & -1 & 1 \\ 0 & 0 & 0 \end{pmatrix}$$

$$P'_{00}(t) =$$

$$P'_{11}(t) = -P_{11}(t), P_{11}(0) = 1 \Rightarrow P_{11}(t) = e^{-t}$$

$$P'_{22}(t) = 0, P_{22}(0) = 1 \Rightarrow P_{22}(t) = 1$$

$$P'_{01}(t) =$$

$$P_{01}(t) =$$

$$P_{01}(t) =$$

Forward and backward equations for B&D processes

Forward equation:

$$P_{ij}(t+h) = \sum_{k=0}^{\infty} P_{ik}(t) P_{kj}(h)$$

=

If $\Theta_{ij} = o(h)$ (requires additional technical assumptions)

$$\begin{cases} P'_{ij}(t) = \lambda_{j-1} P_{i,j-1}(t) - (\lambda_j + \mu_j) P_{ij}(t) + \mu_{j+1} P_{i,j+1}(t) \\ P'_{i0}(t) = -\lambda_0 P_{i0}(t) + \mu_1 P_{i1}(t) \end{cases}, \quad \text{with } P_{ij}(0) = \delta_{ij}$$

Forward and backward equations for B&D processes

Similarly, we derive the backward equations

$$\begin{cases} P'_{ij}(t) = \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t) + \lambda_i P_{i+1,j}(t) \\ P_{0j}(t) = -\lambda_0 P_{0j}(t) - \lambda_0 P_{1j}(t) \end{cases}, \quad \text{with} \quad P_{ij}(0) = \delta_{ij}$$

Example Linear growth with immigration.

Recall $\lambda_k = \lambda \cdot k + a$ \leftarrow immigration
 \uparrow linear birth rate

$$\mu_k = \mu \cdot k$$

↑ linear death rate

Example: Linear growth with immigration.

Use forward equations to compute $E(X_t | X_0 = i)$

$$\begin{cases} P'_{ij}(t) = \lambda_{j-1} P_{i,j-1}(t) - (\lambda_j + \mu_j) P_{ij}(t) + \mu_{j+1} P_{i,j+1}(t) \\ P'_{i0}(t) = -\lambda_0 P_{i0}(t) + \mu_1 P_{i1}(t) \end{cases}$$

$$E(X_t | X_0 = i) =$$

$$P'_{ij}(t) = (\lambda(j-1) + \alpha) P_{i,j-1}(t) - ((\lambda + \mu)j + \alpha) P_{ij}(t) + \mu(j+1) P_{i,j+1}(t)$$

Example: Linear growth with immigration.

$$M'(t) =$$

=

=

$$\begin{cases} M'(t) = \\ M(0) = \end{cases}$$

$$M(t) = i + at \quad \text{if } \lambda = \mu$$

$$M(t) = \frac{a}{\lambda - \mu} (e^{(\lambda - \mu)t} - 1) + i e^{(\lambda - \mu)t} \quad \text{if } \lambda \neq \mu$$