

# MATH180C: Introduction to Stochastic Processes II

[www.math.ucsd.edu/~ynemish/teaching/180c](http://www.math.ucsd.edu/~ynemish/teaching/180c)

Today: Renewal processes.

Examples

> Q&A: November 6

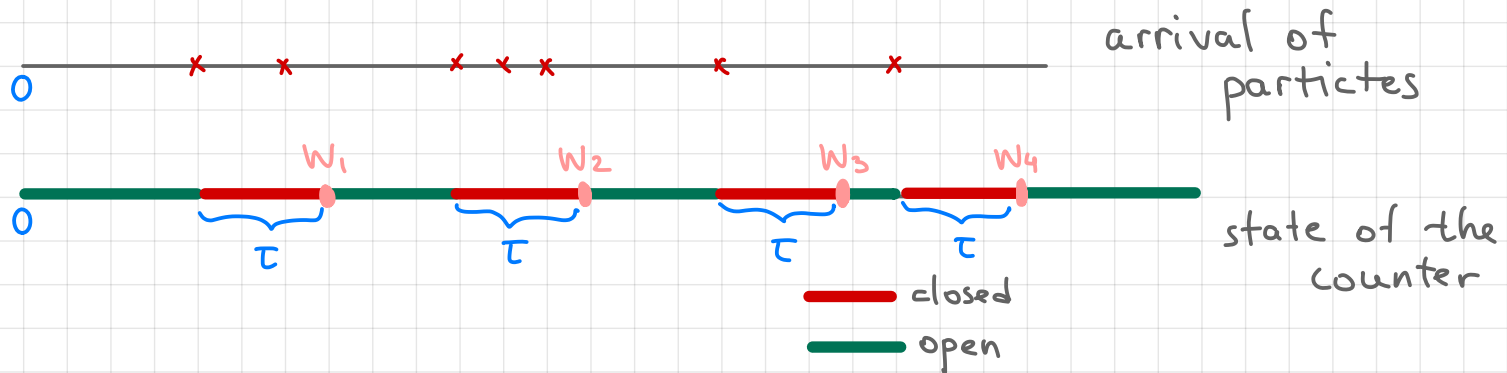
Next: PK 7.4-7.5, Durrett 3.1

This week:

- Homework 4 (due Friday, November 6, 11:59 PM)

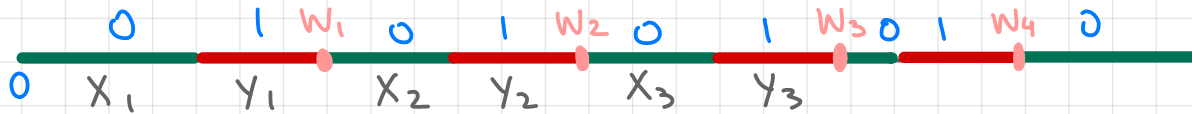
## Other renewal processes

- traffic flow : distances between successive cars are assumed to be i.i.d. random variables
- counter process: particles/signals arrive on a device and lock it for time  $\tau$  ; particles arrive according to a PP; times at which the counter unlocks form a renewal process



## Other renewal processes

- more generally, if a component has two states (0/1, operating/non-operating etc), switches between them, times spent in 0 are  $X_i$ , times spent in 1 are  $Y_i$ ,  $(X_i)_{i=1}^{\infty}$  i.i.d.,  $(Y_i)_{i=1}^{\infty}$  i.i.d., then the times of switching from 0 to 1 form a renewal process with interrenewal times  $X_i + Y_i$

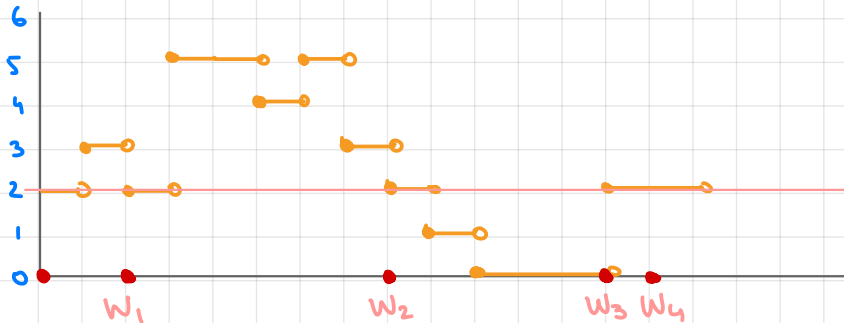


## Other renewal processes

- Markov chains: if  $(Y_n)_{n \geq 0}$ ,  $Y_n \in \{0, 1, \dots\}$  is a recurrent MC starting from  $Y_0 = k$ , then the times of returns to state  $k$  form a renewal process. More precisely

define  $W_1 = \min \{n > 0 : Y_n = k\}$

$$W_p = \min \{n > W_{p-1} : Y_n = k\}$$



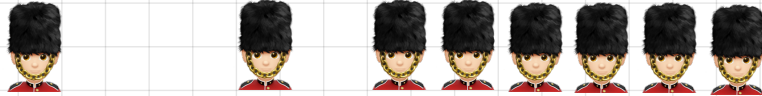
Example with  $k=2$

Similarly for continuous time MCs.

Strong Markov property!

## Other renewal processes

- Queues. Consider a single-server queueing process



customers arriving



server busy/idle  
service time

- (i) if customer arrival times form a renewal process  
then the times of the starts of successive idle periods  
generate a second renewal time
- (ii) if customers arrive according to a Poisson process,  
then the times when the server passes from  
busy to free form a renewal process

# Asymptotic behavior

# Asymptotic behavior of renewal processes

Let  $N(t)$  be a renewal process with interrenewal times  $X_i$ ,  $X_i \in (0, \infty)$ .

Thm.

$$P\left(\lim_{t \rightarrow \infty} N(t) = \infty\right) = 1$$

Proof.  $N(t)$  is nondecreasing, therefore  $\exists \lim_{t \rightarrow \infty} N(t) =: N_\infty$

$N_\infty$  is the total number of events ever happened.

$N_\infty \leq k$  if and only if  $W_{k+1} = \infty$

if and only if  $X_i = \infty$  for some  $i \leq k+1$

$$P(N_\infty < \infty) = P(X_i = \infty \text{ for some } i) \leq \sum_{i=1}^{\infty} P(X_i = \infty) = 0 \quad \square$$

Thm (Pointwise renewal thm).

$$P\left(\lim_{t \rightarrow \infty} \frac{N(t)}{t} = \frac{1}{\mu}\right) = 1$$

$$(\mu = E(X_1))$$

## Elementary Renewal Theorem

Thm. If  $M(t) = E(N(t))$  and  $E(X_1) = \mu$ , then

$$\lim_{t \rightarrow \infty} \frac{M(t)}{t} = \frac{1}{\mu}$$

Proof. (Only for bounded  $X_i$ :  $\exists K$  s.t.  $P(X_i \leq K) = 1$ )

First note that  $W_{N(t)+1} = t + \gamma_t$

In lecture 11 we showed that  $E(W_{N(t)+1}) = \mu(M(t) + 1)$

so

$$M(t) = \frac{t + E(\gamma_t)}{\mu} - 1$$

$$\frac{M(t)}{t} = \frac{1}{\mu} + \frac{1}{t} \left( \frac{E(\gamma_t)}{\mu} - 1 \right) \rightarrow \frac{1}{\mu} \text{ as } t \rightarrow \infty$$

If  $X_i \leq K$ , then  $\gamma_t \leq K \Rightarrow E(\gamma_t) \leq K$

Exercise:  $(X_n)_{n \geq 0}$ : 1)  $P(\lim_{n \rightarrow \infty} X_n = 0) = 1$  2)  $\lim_{n \rightarrow \infty} E(X_n) \geq c > 0$ .



## Asymptotic distribution of $N(t)$

Thm. Let  $N(t)$  be a renewal process with  $E(X_1) = \mu$  and  $\text{Var}(X_1) = \sigma^2$ , then

$$1) \quad \lim_{t \rightarrow \infty} \frac{\text{Var}(N(t))}{t} = \frac{\sigma^2}{\mu^3}$$

$$2) \quad \lim_{t \rightarrow \infty} P\left(\frac{N(t) - E(N(t))}{\sqrt{\text{Var}(N(t))}} \leq x\right)$$

$$= \lim_{t \rightarrow \infty} P\left(\frac{N(t) - \frac{t}{\mu}}{\sqrt{\frac{\sigma^2}{\mu^3} t}} \leq x\right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy$$

$$N(t) \approx \frac{t}{\mu} + \sqrt{\frac{\sigma^2}{\mu^3} t} Z, \text{ where } Z \sim N(0,1) \text{ for large } t$$

No proof.