MATH 180A (Lecture A00)

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Today: Expectation of a function of a random variable. Variance Next: ASV 3.5

We'ek 6:

Homework 4 due Friday, February 17

Expectation of continuous random variables

Example Consider function
$$f(t) = \begin{cases} 0, & t \leq 1 \\ \frac{1}{t^2}, & t > 1 \end{cases}$$

Is
$$f(t)$$
 a probability density?

$$f(t) \ge 0 \qquad \int f(t) dt = \int_{t^2}^{t} dt = -\frac{1}{t} \Big|_{t}^{t} = 1$$

What is E(X)?

$$E(X) = \int_{-\infty}^{+\infty} t f_X(t) dt = \int_{1}^{+\infty} t \cdot \frac{1}{t^2} dt = \int_{1}^{+\infty} \frac{1}{t} dt = \log t \Big|_{1}^{+\infty} = +\infty$$

Expectations of functions of random variables

Proposition For a continuous random variable X with

density
$$f_x$$

$$E(g(x)) = \int_{-\infty}^{+\infty} g(t) f_x(t) dt$$

Example Let
$$U$$
 be a uniform random variable on $[a,b]$
Then $E[U^2] = \int t^2 \cdot f_U(t) dt$ $f_U(t) = \begin{cases} \frac{1}{b-a}, t \in [a,b] \\ o, otherwise \end{cases}$

Important class of functions:
$$g(x) = x^n$$

discrete: $E(X^n) = \sum_{t=0}^{n} t^n P(X=t)$, continuous: $E(X^n) = \int_{-\infty}^{\infty} t^n f_X(t) dt$

Expectations of functions of random variables Example (Car accident/insurance example) An accident causes Y dollars of damage to your car where insurance deductible is 500 dollars. Y~ Unif ([100,1500]). What is the expected amount you pay? X = amount you pay = min {Y, 500} = g(Y), g(f) = min(t,500) (neither discrete nor continuous) E(X) = E(g(Y)) = \inn\{t, 500\}. fy(t) dt = \inn\{t, 500\}. \frac{1}{1400} dt $= \int \frac{1}{1400} dt + \int \frac{500}{1400} dt = \frac{1}{1400} \left(\frac{500^2}{2} - \frac{100^2}{2} \right) \quad \min\{t, 500\} = \begin{cases} t, & t \in [100, 500] \end{cases}$ $+\frac{500}{1400} \cdot (1500-500) = 442.86$ 1500

Definition The variance of a random variable X is

Var
$$(X) = E((X - E(X))^2)$$

• first compute $\mu = E(X)$, then compute $E(g(X))$ with $g(t) = (t - \mu)^2$

- if X is discrete, $Var(X) = \sum_{t=\pm 1}^{\infty} (t-\mu)^2 P(X=\pm)$ • if X is continuous, $Var(X) = \int_{\infty}^{\infty} (t-\mu)^2 f_X(t) dt$
- ~ (t) / (3 continuous, voet (x)) (t) / (r) dt

The square root of the variance is called standard deviation $G(x) = \sqrt{Var(x)}$

Var (X) ≥ 0 (always)

Variance Example

$$Var(X) = E((X-p)^2) = Z(t-p)^2 P(X=t)$$

$$= (I-p)^2 \cdot p + (0-p)^2 \cdot (I-p) = p(I-p)^2$$

$$= (I-P)^{2} \cdot P + (0-P)^{2} \cdot (I-P) = P(I-P)^{2} + P^{2}(I-P)$$

$$= P(I-P)(I-P+P) = P(I-P)$$

$$\in (X) = IP(I-P)$$

Example
$$U \sim Unif[a,b]$$
, $E(U) = \frac{a+b}{2}$

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$$U \sim U \cap f [a,b]$$
, $E(U) = \frac{a+b}{2}$

$$Var(U) = \int_{-\infty}^{\infty} (t - \frac{a+b}{2})^2 \int_{u}^{2} (t) dt = \int_{a}^{\infty} (t - \frac{a+b}{2})^2 \frac{1}{b-a} dt = \frac{(b-a)^2}{12}$$

$$G(X) = \frac{1b-a}{12}$$

Alternative formula for variance

Proposition. Let X be a random variable. Then

 $Var(X) = E(X^2) - (E(X))^2$ Proof (For continuous random variables) Let $\mu := E(X)$.

Then $Var(X) = E((X-\mu)^2) = \int_{\mathbb{R}} (t-\mu)^2 f_X(t) dt$

$$= \int_{-\infty}^{\infty} (t^2 - 2t\mu + \mu^2) f_X(t) dt$$

$$= \int_{-\infty}^{\infty} t^2 f_X(t) dt - 2\mu \int_{-\infty}^{\infty} t f_X(t) dt + \mu^2 \int_{-\infty}^{\infty} f_X(t) dt$$

 $= E(X^{2}) - 2 \cdot \mu \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - 2 \cdot \mu \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu + \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2}$ $= E(X^{2}) - \mu^{2} \cdot \mu^{2} \cdot 1 = E(X^{2}) - \mu^{2} \cdot 1 = E(X^{2})$

Variance Variance is a measure of how "spread out from the mean" the distribution is. Proposition Let X be a random variable with finite expectation E(X)= u. Then Proof (=) Exercise (⇒) (Assume X is discrete). 0 = Var (X) = => For all t, , so if For all t, either 00 therefore,