MATH 285: Stochastic Processes

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Today: Hitting times. First step analysis. Stopping times

Homework 1 is due on Friday, January 14, 11:59 PM

Expected hitting times Let (Xn), be a Markov chain with transition probabilities p(iij) and state space S. Notation: $P[Y] = P[Y|X_{\circ}=i]$, $E_{i}[Y] = E[Y|X_{\circ}=i]$ Let ACS, TA := min {n20: Xn ∈ A} Q: How long (on average) does it take to reach A? Compute Eilta] = E[TalXo = i] =: g(i) By definition $E:[Y] = \sum_{k=1}^{\infty} k P[Y=k | X_o=i]$ (Ye {0,1,2,---}) First step analysis (conditioning on the first step)

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$$g(i) = E_i(T_A) = \sum_{j \in S} E[T_A \mid X_{i=j}, X_o = i] P[X_{i=j} \mid X_o = i]$$

Expected hitting times If i A, then g(i) = 0. Suppose i & A. $\mathbb{P}[T_{A}=k\mid X_{1}=j,X_{0}=i]=\mathbb{P}[X_{0}\notin A,X_{1}\notin A,\dots,X_{k-1}\notin A,X_{k}\in A\mid X_{1}=j,X_{0}=i]$ = P[X, & A, X, & A, ..., XK-2 & A, Xk-1 & A | X =]] = P[TA = K-1 | Xo=j] Compute the expectation $g(i) = \sum_{i \in S} E[T_A | X_{i=i} | X_{o=i}] P[X_{i=j} | X_{o=i}]$ $= Z Z k P[T_A = k | X_1 = j, X_0 = i] p(i,j) = Z Z k P[T_A = k-1 | X_0 = j]$ $\int_{\infty}^{e} \sum_{k=1}^{e} | x_k | X_1 = j, X_0 = i$ $\int_{\infty}^{e} \sum_{k=1}^{e} | x_k | X_1 = j, X_0 = i$ $\int_{\infty}^{e} | x_k | X_1 = j, X_0 = i$ $\int_{\infty}^{e} | x_k | X_1 = j, X_0 = i$ $\int_{\infty}^{e} | x_k | X_1 = j, X_0 = i$ = Z Z l P(TA = e | Xo = j] p(i,j) + Z Z P(TA = e | Xo = j] p(i,j) $= \sum_{j \in S} \mathbb{E}[T_A] P(i,j) + 1 = \sum_{j \in S} g(j) P(i,j) + 1$

Expected hitting times Conclusion: $(g(i) = 1 + \sum_{j \in S} p(i,j)g(j))$ if $i \notin A$ g(i) = 0 if $i \in A$ Example 3.2 On average how many times do we need to toss a coin to get two consecutive heads? Denote by Xn the number of consecutive heads after nth toss. $X_{n} \in \{0, 1, 2\}, \quad P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 2 & 0 & 0 \end{bmatrix}$ $A = \{2\} \qquad 2\begin{bmatrix} 0 & 0 & 1 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 2 & 0 & 0 \end{bmatrix} \qquad 9(0) = \frac{3}{2}g(1)$ $g(0) = 1 + \frac{1}{2}g(0) + \frac{1}{2}g(1)$ g(1) = 4 g(0) = 6on average 6 tosses to reach state 2. Starting from state o it takes

Stopping Times Def 3.3 Let (Xn)_{n20} be a discrete time stochastic process. A stopping time is a random variable T∈ {0,1,2,- }u{∞} such that for each n the event { T=n} depends only on Xo, X, ..., Xn Examples T = min { n > 0 : Xn = i } is a stopping time $\{T_i = n\} = \{X_0 \neq i, X_1 \neq i, \dots, X_{n-1} \neq i, X_n = i\}$ T2 = max { n ≥ 0 : Xn = i} is not a stopping time $\{T_2 = n\} = \{X_n = i_1 X_{n+1} \neq i_1 - -- \}$ Recall Markov property: If (Xn) is Markov (A.P), then conditional on Xm=l, the process (Xmin)neN is Markov (Se,P) independent of Xo, X,,--, Xm

Strong Markov property Proposition 3.6 Let (Xn) be a time-homogeneous Markov chain with state space 5 and transition probabilities p (i,j). Let T be a stopping time, les and P[XT=l]>0. Then, conditional on XT=e, (XT+n)nzo is a timehomogeneous Markou chain with transition probabilities plij) independent of Xo, ---, XT. In other words, if A is an event that depends only on Xo, XI, --, XT and P[An[XT=1]]>0 then for all n≥o and all io, i,..., in ∈ S $P[X_{T+1} = i, X_{T+2} = i_2, ..., X_{T+n} = i, IA \cap [X_T = \ell] = p(\ell, i, i_2) - -- p(i_n, i_n)$ Proof Use the partition {{T=m}}m=s (see the notes)

Classification of states: recurrence and transience Let (Xn) be a Markov chain with state space S. Def 4.1 A state i & S is called recurrent if Pi[Xn=i for infinitely many n]=1 A state ies is called transient if P: (Xn=i for infinitely many n] = 0 Remark Let Tilk = time Xn (starting from i) visits state i kth time Ti, 1=0, Ti, k+1 = min {n> Ti, k: Xn = i} Then, for k = 2, Ti,k are stopping times. Indeed, $\left\{ \begin{array}{c} \text{Ti.}_{2} = m \end{array} \right\} = \left\{ \begin{array}{c} X_{1} \neq i, \quad X_{2} \neq i, \quad -1, \quad X_{m-1} \neq i, \quad X_{m-1} \neq$ $\left\{ \begin{array}{ll} T_{i,k-1} = \ell & T_{i,k-1} = \ell & T_{i,k-1} = \ell & X_{\ell+1} \neq i, \dots, X_{m-1} \neq i, X_{m} = i \right\} \\ \ell = k-2 & \ell$

Classification of states: recurrence and transience Denote Ti = Ti, = min 1 n>0: Xn = i g and ri = Pi [Ti <∞] Theorem 4.2 Ti,2 Ti3 Ti14 M Let i & S. Then (1) i is recurrent (=> ri=1 (=> ∑ pn (iii) = ∞ (2) i is transient ⇔ ri ∠1 ⇔ ∑ pn (iii) <∞ Proof Step1: By the Strong Markov property $P[T_{i,k+1} < \infty | T_{i,k} < \infty] = P[T_{i,2} < \infty] = \Gamma_i$ P[Ti,k+1 < 00] = P[Ti,k+1 < 00 | Ti,k < 00] P[Ti,k < 00] = -- = rik Step 2: Denote Ni := \(\frac{1}{2} \) \(\tau_{n=i} \) \(\tau \) times (Xn) visits state i $\forall k \geq 1, \{N_i \geq k\} = \{T_{i,k} < \infty\}, so P_{i}[N_i \geq k] = P_{i}[T_{i,k} < \infty] = \Gamma_{i}^{k-1}$

Classification of states: recurrence and transience

Thus
$$\mathbb{E}[Ni] = \sum_{k=1}^{\infty} \mathbb{P}[Ni \ge k] = \sum_{k=1}^{\infty} \mathbb{P}[Ni \ge k] = \sum_{k=1}^{\infty} \mathbb{P}[X] = \sum_{k=1}^{\infty} \mathbb{P}$$

$$\mathbb{E}_{i}[N_{i}] = \mathbb{E}_{i}\left[\sum_{n=0}^{\infty} \mathbb{1}_{X_{n}=i}\right] = \sum_{n=0}^{\infty} \mathbb{P}_{i}[X_{n}=i] = \sum_{n=0}^{\infty} \mathbb{P}_{n}(i,i)$$

Since
$$C_i \in [0,1]$$
, $\sum_{\ell=0}^{\infty} C_i^{\ell} = \infty \Leftrightarrow C_{i=1}$, $\sum_{\ell=0}^{\infty} C_i^{\ell} < \infty \Leftrightarrow C_i < 1$

Step 3:
$$\Gamma_i = 1 \iff \forall k \ P_i[\ N_i \ge k] = 1$$
, i.e., i is recurrent
Step 4: $\Gamma_i < 1 \iff P_i[\ N_i \ge k] = \Gamma_i^k \to 0$, $k \to \infty$,

so
$$P_i[N_i = \infty] = 0$$
, i.e., i is transient $\sum_{n=0}^{\infty} P_n(iii) = \sum_{\ell=0}^{\infty} \Gamma_i^{\ell} = \frac{1}{1-\Gamma_i}$