#### MATH 285: Stochastic Processes

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# Today: Periodic, aperiodic, reducible, irreducible Markov chains with finite state space

Homework 2 is due on Friday, January 21 11:59 PM

Prop. 7.1 Let (Xn) be a MC with finite state space S. Suppose that there exists no EN s.t [P]; >0 for all i,jes Then for each i, Ti(j) is equal to the asymptotic expected fraction of time the chain spends in state j, i.e.,  $\lim_{n\to\infty} \mathbb{E} \left[ \frac{1}{n+1} \sum_{k=0}^{\infty} \mathbb{1}_{\{X_k=j\}} \right] = \pi(j)$ Proot.  $\mathbb{E}\left[\frac{1}{n+1}\sum_{k=0}^{n}\mathbb{1}_{\left\{X_{k}=j\right\}}\right] = \frac{1}{n+1}\sum_{k=0}^{n}\mathbb{P}\left[X_{k}=j\right] = \frac{1}{n+1}\sum_{k=0}^{n}\mathbb{E}\left[X_{k}=j\right] \mathbb{P}\left[X_{o}=i\right]$  $=\frac{1}{n+1}\sum_{k=0}^{\infty}\left[\pi_{o}P^{k}\right].$ By Cor. 6.6.  $[\Pi_0 P^K]_j \to \Pi(j)$ ,  $k \to \infty$ , for all jes and  $\Pi_0$ .

Therefore,  $\frac{1}{n+1} \sum_{k=0}^{n} [\Pi_0 P^K]_j \to \Pi(j)$  [if  $a_n \to a$ ,  $n \to \infty$ , then  $\frac{1}{n} \sum_{k=1}^{n} a_k \to a$ ]

Stationary distribution and long-run behavior

## Stationary distribution and expected return times

Recall that Ti,k denotes the time of the k-th visit to state i.

 $y_{k} = \frac{1}{2} \sum_{k=1}^{\infty} y_{k} = \frac{1}{2} \sum_{k=1}^{\infty}$ 

$$y_k \sim$$
. Notice that  $\sum_{k=1}^m y_k = \sum_{k=1}^m T_{i,k+1} - T_{i,k} = \sum_{k=1}^m T_{i,m+1} = \sum_{k=1}^m T_{i,m+1} \approx \sum_{k=1}^m T_{i,m$ 

Take m large, and let  $n = m E(T_i)$ . Then

so  $\sum_{k=0}^{n} \mathbb{1}_{\{X_k=i\}}$ . Then  $\frac{m+1}{n} \approx$ 

Periodic and aperiodic chains Let (Xn) be a MC with state space 5 and transition probability p(iij). Def For ies, denote Ji := We call d(i):= De (0,1)  $J_1 =$ d(1)= d(1) = 9(1)= Def If d(i)=1 for all i ∈ S, then (Xn) is called

Periodic and aperiodic chains Lemma 7.2 If P is the transition matrix for an irredusible Markov chain, then for all states i.j. Proof Fix ies. (1) If mine Ji, then (2) Let d=d(i). Then (definition of d(i)) Take j≠i. (3) Pirreducible => 3 m,n s.t. pm (i,j)>0, pn (j,i)>0.  $\Rightarrow P_{m+n}(i,i) > 0 \Rightarrow \Rightarrow \Rightarrow$ (4) If le J; then pe(j,j)>0 and thus => dis a common divisor of J; => (5) Swap i and j: ∃ q2 € N s.t. d(i) = q2 d(j) ⇒ d(i) = d(j)

### RW on bipartite graphs

Example 7.3 Let G=(V,E) be finite connected graph.

· SRW on G is irreducible (all vertices have the same period) - we call the common period the period of MC

Period IS 2 ITT
$$V = V_1 \coprod V_2, E C(V_1 \times V_2 \cup V_2 \times V_1)$$

$$V = H$$
,  $V_1 = \text{even numbers}$   
 $V_2 = \text{odd numbers}$ 

Irreducible aperiodic Markov chains Theorem 7.4 Let P be a transition matrix for a finite-state, irreducible, aperiodic Markov chain. Then there exists a unique stationary distribution II, II = IT P, and for any initial probability distribution )  $\lim_{n\to\infty} P^n = 11$ Proof (1) By PF theorem, enough to show that there exists no>o s.t. \ cij \ . Fix cije \ S (2) d(i) = 1 (aperiodic) => 3 Mi s.t. Ji contains all nz Mi (3) irreducible => 3 mij s.t. Pmij (i,j)>0 (2)+(3): Take no = max (Mi + mij) =>

Reducible Markov chains Not irreducible MC = reducible MC Def 7.5 Let (Xn) be a MC with state space 5. We say that states i and i , denoted if there exists nime Nulo) s.t. and DE(0,1) Lemma 7.6 Relation  $\leftrightarrow$  on S is an equivalence relation. (reflexivity, i +i) po(i,i)=1, so i +i (symmetry, i + j + j + i) Follows from Def 7.5 (transitivity, i +) j + k => i + k) i + j: pn(i,j)>0, pm (j,i)>0 j + k: pn(j,k)>0, pm(k,j)>0. Then

### Communication classes Equivalence relation \( \rightarrow \) splits the state space into communication classes (sets of states that communicate with each other). 2 PE(0,1) MC is irreducible iff it consists of one communication class Class properties: [proof as in Prop 4.8, Prop. 7.2] - transience or recurrence : either all states in one class are transient (class) or all are recurrent (class) - periodicity: all states in one class have the same period

## Communication classes Suppose i and j belong to different classes.

- If p(i,j) > 0, then for all  $n \in \mathbb{N}$  (otherwise  $i \leftrightarrow j$ .

  If p(i,j) > 0 and p(j,i) = 0 for all  $n \in \mathbb{N}$ , then
- P<sub>i</sub>[Xn = i for infinitely many n] ≤ , and thus i is transient
  - Therefore, if i and j belong to different classes and i is recurrent, then (once in a

recurrent class, MC stays there forever)

If we split the state space into communication classes,

with Re denoting recurrent classes, then the transition matrix has the following form

General form of transition matrix with finite S Pe submatrix for the recurrent class Re Pe is a stochastic matrix, we can consider it as a Markov chain on Re [SIQ] transition probabilities starting from transient • If Pe is aperiodic, then Pe → (π(), n→∞ · What about transient states? · What if Pe is not aperiodic?

#### Transient states

Suppose there exists one transient class T Pi

• If Q is substochastic, then for all eigenvalues λ of Q 1λ/<1

$$\Rightarrow$$
 Q $^{\circ} \rightarrow 0$ ,  $n \rightarrow \infty$ , i.e. for i, jeT  $P_{i}[X_{n}=j] \rightarrow 0$ ,  $n \rightarrow \infty$ 

$$\int_{0}^{\infty} + Q + Q + \dots = \int_{0}^{\infty} + V D V + V D V$$
For  $i, j \in T$ ,  $E_{i} \left[ \sum_{\kappa=0}^{\infty} 1_{\kappa=0} \right] = 0$ 

]+Q+Q2+--= ]+VDV+VDV+--=V(I+D+D+--)V converges

## Transient states Conclusion: if TCS

O | 
$$\frac{1}{2}$$
 |  $\frac{1}{2}$  |

$$\lim_{n\to\infty} \mathbb{P}_{\bar{i}} \left[ X_n = j \right] =$$

$$\mathbb{E} \left[ \sum_{k=0}^{\infty} \mathbb{1}_{\left\{ X_n = j \right\}} \right] =$$

expected number of visits to j starting from i

is a transient class, then \tije T

Expected number of steps before absorption starting from O  $is = \frac{3}{2} + [t] = 3$ 

Transient states

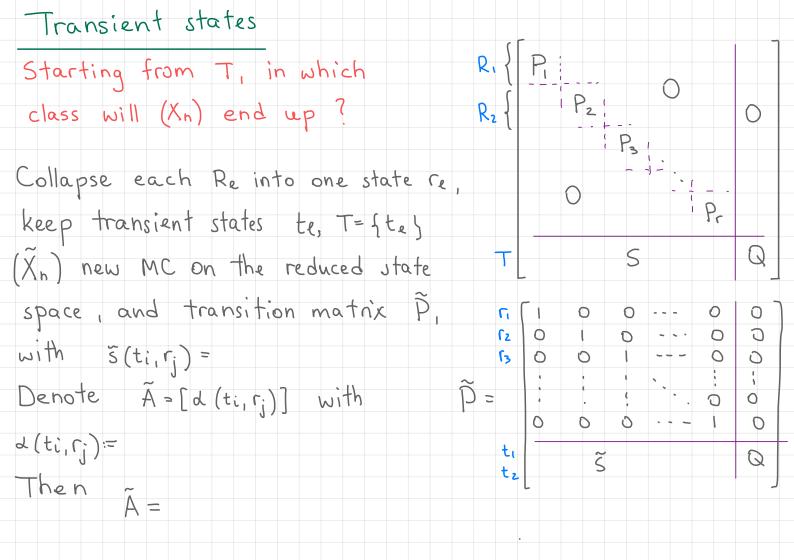
Recall, First step analysis for the mean hitting time

 $g_{i} = E_{i}[T_{A}] = \begin{cases} 0, & i \in A \\ 1 + \sum P(i,j)g_{j}, & i \notin A \end{cases}$   $T_{A} = \sum_{n=0}^{\infty} \Lambda_{\{X_{n} \notin A\}}$ 

Instead of adding I for each step, add I only when Xn visits j:

Denote SIA =: T, and for i,jeT gii = Then FSA gij = if ieA

G = [gij] then



Transient states

Example 8.2  $0 \ 1 \ 0 \ 0 \ 0$   $1 \ 2 \ 3$   $1 \ 2 \ 3$   $1 \ 2 \ 3$   $2 \ 0 \ 0 \ 2$   $2 \ 0 \ 0 \ 2$   $2 \ 0 \ 0 \ 2$   $2 \ 0 \ 0 \ 2$   $3 \ 0 \ 2$   $4 \ 0 \ 0 \ 2$   $4 \ 0 \ 0 \ 0$ 

What is the probability that starting from a transient state i we end up in a recurrent state i?

Use  $\tilde{A} = \begin{bmatrix} \frac{3}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} \begin{bmatrix} \frac{1}{2} & 0 \\ 0 & 0 \\ \frac{1}{2} \end{bmatrix} = \begin{bmatrix} \frac{3}{4} & \frac{1}{4} \\ \frac{1}{2} & \frac{1}{2} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix}$ 

Expected transit times from i to j (think about j as absorbing) --