

#### Lecture 4: Random Variable, Part II

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EE210: Probability and Introductory Random Processes KAIST EE

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## Outline



- Continuous Random Variable
- PDF (Probability Density Function)
- CDF (Cumulative Distribution Function)
- Exponential and Normal Distribution
- Joint PDF, Conditional PDF
- Bayes' rule for continous and even mixed cases

#### Roadmap



- Famous discrete random variables used in the community
  - Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- Functions of a single random variable, Functions of multiple random variables
- Conditioning for random variables, Independence for random variables
- Continuous random variables
  - Normal, Uniform, Exponential, etc.
- Bayes' rule for random variables

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### Continuous RV and Probability Density Function



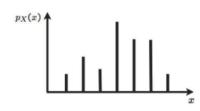
- Many cases when random variable have "continuous values", e.g., velocity of a car

#### Continuous Random Variable

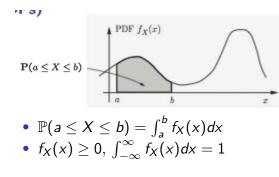
A rv X is continuous if  $\exists$  a function  $f_X$ , called probability density function (PDF), s.t.

$$\mathbb{P}(X \in B) = \int_B f_X(x) dx$$

- All of the concepts and methods (expectation, PMFs, and conditioning) for discrete rvs have continuous counterparts

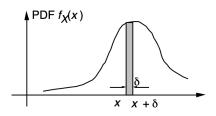


- $\mathbb{P}(a \le X \le b) = \sum_{x:a \le x \le b} p_X(x)$   $p_X(x) \ge 0$ ,  $\sum_x p_X(x) = 1$



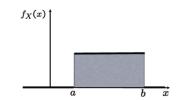
# PDF and Examples

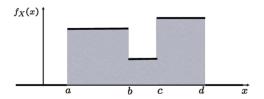




- $\mathbb{P}(a \leq X \leq a + \delta) \approx \boxed{f_X(a) \cdot \delta}$
- $\mathbb{P}(X = a) = 0$

#### Examples

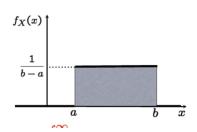




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# Expectation and Variance





- $\mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_a^b \frac{x}{b-a} dx = \frac{1}{b-a} \frac{b^2 a^2}{2} = \frac{b+a}{2}$
- $\mathbb{E}[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_a^b \frac{x^2}{b-a} dx = \frac{1}{b-a} \frac{b^3 a^3}{3} = \frac{a^2 + ab + b^2}{3}$
- $var[X] = \frac{a^2 + ab + b^2}{3} \frac{a^2 + 2ab + b^2}{4}$

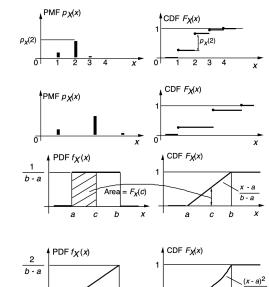
# Cumulative Distribution Function (CDF)



- Discrete: PMF, Continuous: PDF
- Can we describe all rvs with a single mathematical concept?

$$F_X(x) = \mathbb{P}(X \le x) =$$
 
$$\begin{cases} \sum_{k \le x} p_X(k), & \text{discrete} \\ \int_{-\infty}^x f_X(t) dt, & \text{continuous} \end{cases}$$

- always well defined, because we can always compute the probability for the event {X ≤ x}
- CCDF (Complementary CDF):  $\mathbb{P}(X > x)$



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# **CDF** Properties



- Non-decreasing
- $F_X(x)$  tends to 1, as  $x \to \infty$
- $F_X(x)$  tends to 0, as  $x \to -\infty$

Now, let's look at famous continuous random variables popularly used in our life.

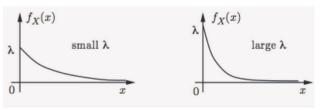
## Exponential RV with parameter $\lambda > 0$ : $exp(\lambda)$



• A rv X is called exponential with  $\lambda$ , if

$$f_X(x) = egin{cases} \lambda e^{-\lambda x}, & x \geq 0 \ 0, & x < 0 \end{cases} ext{ or } F_X(x) = 1 - e^{-\lambda x}$$

- Models a waiting time
- CCDF  $\mathbb{P}(X \ge x) = e^{-\lambda x}$  (waiting time decays exponentially)
- $\mathbb{E}[X] = 1/\lambda$ ,  $\mathbb{E}[X^2] = 2/\lambda^2$ ,  $\text{var}[X] = 1/\lambda^2$
- (Q) What is the discrete rv which models a waiting time?



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## Modeling Waiting Time? A Discrete Twin (1)



- A discrete twin for modeling waiting times is geometric rvs.
- Models a system evolution over time: Continuous time vs. Discrete time. In many cases, continuous case is the some type of limit of its corresponding discrete case.
- Can you make mathematical description, where geometric and exponential rvs meet each other in the limit?
- Key idea.
  - Continuous system: Discrete system with
     infinitely many slots whose duration is infinitely small.
- limiting system:  $X_{exp}(\lambda)$  with CDF  $F_{exp}(\cdot)$
- *n*-th system:  $X_{geo}^n(p_n)$  with CDF  $F_{geo}^n(\cdot)$

## Modeling Waiting Time? A Discrete Twin (2)



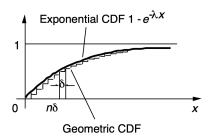
For a given x > 0,

- Define  $\delta = \frac{x}{n}$  (a slot length in the *n*-th system)
- Remember

$$F_{\text{exp}}(x) = 1 - e^{-\lambda x}$$
  
 $F_{\text{geo}}^{n}(n) = 1 - (1 - p_n)^{n}$ 

- Choose  $p_n = 1 e^{-\lambda \delta} = 1 e^{-\lambda \frac{x}{n}}$ .
- As  $n \to \infty$ , the slot length  $\delta \to 0$  thus  $p_n \to 0$
- The CDF values of exponential and *n*-th geometric rvs become equal whenever  $x = \delta, 2\delta, 3\delta, \ldots$ , i.e.,

$$F_{exp}(n\delta) = F_{geo}^n(n), \quad n = 1, 2, \dots$$



- As n grows, the number of slots grows, but the success probability over one slot decreases, so that everything is balanced up.
- As n grows,  $F_{geo}^n(n)$  approaches  $F_{exp}(n\delta)$ .

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# Normal (also called Gaussian) Random Variable



#### Why important?

- Central limit theorem (중심극한정리)
  - One of the most remarkable findings in the probability theory
- Convenient analytical properties
- · Modeling aggregate noise with many small, independent noise terms

## Normal: PDF, Expectation, Variance



Standard Normal N(0, 1)

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

- $\mathbb{E}[X] = 0$
- var[X] = 1

Need to check:

- a legitimate PDF or not
- expectation/variance

• General Normal  $N(\mu, \sigma^2)$ 

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-(x-\mu)^2/2}$$
•  $\mathbb{E}[X] = \mu$ 
•  $\operatorname{var}[X] = \sigma^2$ 

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# Normal: Useful Property



Linear transformation preserves normality

#### Linear transformation of Normal

If  $X \sim \textit{Norm}(\mu, \sigma^2)$ , then for  $a \neq 0$  and b  $Y = aX + b \sim \textit{Norm}(a\mu + b, a^2\sigma^2)$ .

- Thus, every normal rv can be standardized : If  $X \sim \textit{Norm}(\mu, \sigma^2)$ , then  $Y = \frac{\mathsf{X} - \mu}{\sigma} \sim \textit{Norm}(0, 1)$
- Thus, we can make the table which records the following CDF values:

$$\Phi(y) = \mathbb{P}(Y \le y) = \mathbb{P}(Y < y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{y} e^{-t^2/2} dt$$

#### Example



- Annual snowfall X is modeled as Norm(60, 20<sup>2</sup>). What is the probability that this year's snowfall is at least 80 inches?
- $Y = \frac{X-60}{20}$ .

$$\mathbb{P}(X \ge 80) = \mathbb{P}(Y \ge \frac{80 - 60}{20})$$
$$= \mathbb{P}(Y \ge 1) = 1 - \Phi(1)$$
$$= 1 - 0.8413 = 0.1587$$

|     | .00   | .01   | .02   | .03   | .04   | .05   | .06   | .07   | .08   | .09   |
|-----|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5675 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7580 | .7611 | .7642 | .7673 | .7704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .8830 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9013 |
| 1.3 | .9032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9279 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9543 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9708 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |
| 2.0 | .9772 | .9778 | .9783 | .9788 | .9793 | .9798 | .9803 | .9808 | .9812 | .9817 |
| 2.1 | .9821 | .9826 | .9830 | .9834 | .9838 | .9842 | .9846 | .9850 | .9854 | .9857 |
| 2.2 | .9861 | .9864 | .9868 | .9871 | .9875 | .9878 | .9881 | .9884 | .9887 | .9890 |
| 2.3 | .9893 | .9896 | .9898 | .9901 | .9904 | .9906 | .9909 | .9911 | .9913 | .9916 |
| 2.4 | .9918 | .9920 | .9922 | .9925 | .9927 | .9929 | .9931 | .9932 | .9934 | .9936 |
| 2.5 | .9938 | .9940 | .9941 | .9943 | .9945 | .9946 | .9948 | .9949 | .9951 | .9952 |
| 2.6 | .9953 | .9955 | .9956 | .9957 | .9959 | .9960 | .9961 | .9962 | .9963 | .9964 |
| 2.7 | .9965 | .9966 | .9967 | .9968 | .9969 | .9970 | .9971 | .9972 | .9973 | .9974 |
| 2.8 | .9974 | .9975 | .9976 | .9977 | .9977 | .9978 | .9979 | .9979 | .9980 | .9981 |
| 2.9 | .9981 | .9982 | .9982 | .9983 | .9984 | .9984 | .9985 | .9985 | .9986 | .9986 |

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## Roadmap



- Famous discrete random variables used in the community
  - Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- o Functions of a single random variable, Functions of multiple random variables
- o Conditioning for random variables, Independence for random variables
- Continuous random variables
  - Normal, Uniform, Exponential, etc.
- Bayes' rule for random variables
- \*\* Continuous counterparts are intuitively understandable. So, we will be quick at reviewing them.

#### Continuous: Joint PDF and CDF (1)



#### Jointly Continuous

Two continuous rvs are jointly continuous if a non-negative function  $f_{X,Y}(x,y)$  (called joint PDF) satisfies: for every subset B of the two dimensional plane,

$$\mathbb{P}((X,Y)\in B)=\iint_{(x,y)\in B}f_{X,Y}(x,y)dxdy$$

1. The joint PDF is used to calculate probabilities

$$\mathbb{P}((X,Y)\in B)=\iint_{(x,y)\in B}f_{X,Y}(x,y)dxdy$$

Our particular interest:  $B = \{(x, y) \mid a \le x \le b, c \le y \le d\}$ 

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## Continuous: Joint PDF and CDF (2)



2. The marginal PDFs of X and Y are from the joint PDF as:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy, \quad f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

3. The joint CDF is defined by  $F_{X,Y}(x,y) = \mathbb{P}(X \le x, Y \le y)$ , and determines the joint PDF as:

$$f_{X,Y}(x,y) = \frac{\partial^2 F_{x,y}}{\partial x \partial y}(x,y)$$

4. A function g(X, Y) of X and Y defines a new random variable, and

$$\mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dxdy$$

## Continuous: Conditional PDF given an event



\* Conditional PDF, given an event

• 
$$f_X(x) \cdot \delta \approx \mathbb{P}(x \le X \le x + \delta)$$
  
 $f_{X|A}(x) \cdot \delta \approx \mathbb{P}(x \le X \le x + \delta|A)$ 

• 
$$\mathbb{P}(X \in B) = \int_B f_X(x) dx$$
  
 $\mathbb{P}(X \in B|A) = \int_B f_{X|A}(x) dx$ 

Note: A is an event, but B is a subset that includes the possible values which can be taken by the rv X.

• 
$$\int f_{X|A}(x) = 1$$

\* Conditional PDF, given  $X \in B$ 

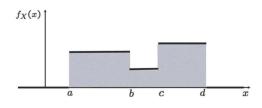
$$\mathbb{P}(x \le X \le x + \delta | X \in B) \approx f_{X|X \in B}(x) \cdot \delta$$

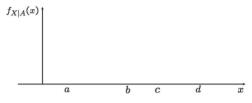
$$f_{X|X\in B}(x) = \begin{cases} 0, & \text{if } x \notin B \\ \frac{f_X(x)}{\mathbb{P}(B)}, & \text{if } x \in B \end{cases}$$

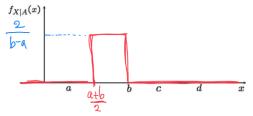
(Q) In the discrete, we consider the event  $\{X = x\}$ , not  $\{X \in B\}$ . Why?

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# Continuous: Conditional Expectation $A = \{a+b < X < b\}$







# KAIST EE

- $\mathbb{E}[X] = \int x f_X(x) dx$  $\mathbb{E}[X|A] = \int x f_{X|A}(x) dx$
- $\mathbb{E}[g(X)] = \int g(x) f_X(x) dx$  $\mathbb{E}[g(X)|A] = \int g(x) f_{X|A}(x) dx$

$$\mathbb{E}[X|A] = \int_{(a+b)/2}^{b} x \frac{2}{b-a} dx = \frac{a}{4} + \frac{3b}{4}$$

$$\mathbb{E}[X^{2}|A] = \int_{(a+b)/2}^{b} x^{2} \frac{2}{b-a} dx =$$

#### Exponential RV: Memoryless



- Exponential rv is a continous counterpart of geometric rv.
- Thus, expected to be memeoryless.

Definition. A random variable X is called memoryless if, for any  $n, m \ge 0$ ,

$$\mathbb{P}(X > n + m | X > m) = \mathbb{P}(X > n)$$

• Proof. Note that  $\mathbb{P}(X > x) = e^{-\lambda x}$ . Then,

$$\mathbb{P}(X>n+m|X>m)=\frac{\mathbb{P}(X>n+m)}{\mathbb{P}(X>m)}=\frac{e^{-\lambda(n+m)}}{e^{-\lambda m}}=e^{-\lambda n}=\mathbb{P}(X>n)$$

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## Total Probability/Expectation Theorem



Partition of  $\Omega$  into  $A_1, A_2, A_3, \ldots$ 

\* Discrete case

#### Total Probability Theorem

$$p_X(x) = \sum_i \mathbb{P}(A_i)\mathbb{P}(X = x|A_i)$$
  
=  $\sum_i \mathbb{P}(A_i)p_{X|A_i}(x)$ 

#### Total Expectation Theorem

$$\mathbb{E}[X] = \sum_i \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$

\* Continuous case

#### Total Probability Theorem

$$f_X(x) = \sum_i \mathbb{P}(A_i) f_{X|A_i}(x)$$

#### Total Expectation Theorem

$$\mathbb{E}[X] = \sum_i \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$

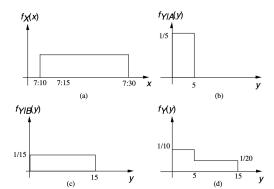
#### Example

# **KAIST EE**

- Your train's arrival every quarter hour (0, 15min, 30min, 45min).
- Your arrival  $\sim$  uniform(7:10, 7:30) am.
- What is the PDF of waiting time for the first train?
- X : your arrival time, Y : waiting time.
- The value of X makes a different waiting time. So, consider two events:

$$A = \{7:10 \le X \le 7:15\}$$

$$B = \{7:15 \le X \le 7:30\}$$



$$f_Y(y) = \mathbb{P}(A)f_{Y|A}(y) + \mathbb{P}(B)f_{Y|B}(y)$$

$$f_Y(y) = \frac{1}{4}\frac{1}{5} + \frac{3}{4}\frac{1}{15} = \frac{1}{10}, \quad \text{for } 0 \le y \le 5$$

$$f_Y(y) = \frac{1}{4}0 + \frac{3}{4}\frac{1}{15} = \frac{1}{20}, \quad \text{for } 5 < y \le 15$$

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## Continuous: Conditional PDF given a RV



- $p_{X|Y}(x|y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}$
- Similarly, for  $f_Y(y) > 0$ ,

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)}$$

- Remember: For a fixed event A,  $\mathbb{P}(\cdot|A)$  is a legitimate probability law.
- Similarly, For a fixed y,  $f_{X|Y}(x|y)$  is a legitimate PDF, since

$$\int_{-\infty}^{\infty} f_{X|Y}(x|y) \frac{dx}{dx} = \frac{\int_{-\infty}^{\infty} f_{X,Y}(x,y) dx}{f_{Y}(y)} = 1$$

Multiplication rule.

$$f_{X,Y}(x,y) = f_Y(y) \cdot f_{X|Y}(x|y)$$
  
=  $f_X(x) f_{Y|X}(y|x)$ 

Total prob./exp. theorem.

$$f_X(x) = \int_{-\infty}^{\infty} f_Y(y) f_{X|Y}(x|y) dy$$

$$\mathbb{E}[X|Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} f_Y(y) \mathbb{E}[X|Y = y] dy$$

• Independence.

$$f_{X,Y}(x,y) = f_X(x)f_Y(y)$$
, for all x and y

# Example: Stick-breaking (Ch 3. Prob 21)



- Break a stick of length / twice
  - first break at  $X \sim uniform[0.1]$
  - second break at  $Y \sim \textit{uniform}[0, X]$
- (Q) What is  $\mathbb{E}[Y]$ ?
- Since Y depends on X, the total expectation theorem seems useful.

$$\mathbb{E}[Y] = \int_{-\infty}^{\infty} f_X(x) \mathbb{E}[Y|X = x] dx$$

• Using the TET,

$$\mathbb{E}[Y] = \int_0^l \frac{1}{l} \mathbb{E}[Y|X = x] dx$$
$$= \int_0^l \frac{1}{l} \frac{x}{2} dx = \frac{l}{4}$$

•  $f_X(x)$  and  $f_{Y|X}(y|x)$  seems easy to compute. Thus,

$$f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x) = \frac{1}{I} \cdot \frac{1}{x}$$

You can do many other things with the joint PDF.

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#### Bayes Rule for Continuous



- X: state/cause/original value  $\rightarrow Y$ : result/resulting action/noisy measurement
- Model:  $\mathbb{P}(X)$  (prior) and  $\mathbb{P}(Y|X)$  (cause  $\to$  result)
- Inference:  $\mathbb{P}(X|Y)$ ?

$$p_{X,Y}(x,y) = p_X(x)p_{Y|X}(y|x)$$

$$= p_Y(y)p_{X|Y}(x|y)$$

$$p_{X|Y}(x|y) = \frac{p_X(x)p_{Y|X}(y|x)}{p_Y(y)}$$

$$p_Y(y) = \sum_{x'} p_X(x')p_{Y|X}(y|x')$$

$$f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x)$$

$$= f_Y(y)f_{X|Y}(x|y)$$

$$f_{X|Y}(x|y) = \frac{f_X(x)f_{Y|X}(y|x)}{f_Y(y)}$$

$$f_Y(y) = \int f_X(x')f_{Y|X}(y|x')dx'$$

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## Bayes Rule for Mixed Case



K: discrete, Y: continuous

Inference of K given Y

$$p_{K|Y}(k|y) = \frac{p_K(k)f_{Y|K}(y|k)}{f_Y(y)}$$
$$f_Y(y) = \sum_{k'} p_K(k')f_{Y|K}(y|k')$$

Inference of Y given K

$$f_{Y|K}(y|k) = \frac{f_{Y}(y)p_{K|Y}(k|y)}{p_{K}(k)}$$
$$p_{K}(k) = \int f_{Y}(y')p_{K|Y}(k|y')dy'$$

#### Example: Signal Detection (1)



Inference of discrete K given continuous Y:

$$p_{K|Y}(k|y) = \frac{p_K(k)f_{Y|K}(y|k)}{f_Y(y)}, \quad f_Y(y) = \sum_{k'} p_K(k')f_{Y|K}(y|k')$$

- K: -1, +1, original signal, equally likely.  $p_K(1) = 1/2, p_K(-1) = 1/2$ .
- Y: measured signal with Gaussian noise, Y = K + W,  $W \sim N(0,1)$
- Your received signal = 0.7. What's your guess about the original signal? +1
- Your received signal = -0.2. What's your guess about the original signal? -1

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## Example: Signal Detection (2)



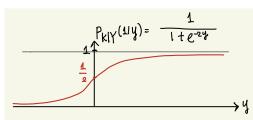
• 
$$Y|K = 1 \sim N(1,1)$$
 and  $Y|K = -1 \sim N(-1,1)$ .

$$f_{Y|K}(y|k) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y-k)^2}, \quad k = 1, -1$$

$$f_{Y}(y) = \frac{1}{2} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y+1)^2} + \frac{1}{2} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y-1)^2}$$

• Probability that K = 1, given Y = y? After some algebra,

$$p_{K|Y}(1|y) = \frac{1}{1 + e^{-2y}}$$





# Questions?

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## Review Questions



- 1) What is PDF and CDF?
- 2) Why do we need CDF?
- 3) What are joint/marginal/conditional PDFs?
- 4) Explain memorylessness of exponential random variables.
- 5) Explain the version of Bayes' rule for continuous and mixed random variables.