

Roadmap



Lecture 5: Random Variable, Part III

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EE210: Probability and Introductory Random Processes KAIST EE

MONTH DAY, 2021

o Famous discrete random variables used in the community

- Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- o Functions of a single random variable, Functions of multiple random variables
- o Conditioning for random variables, Independence for random variables
- o Continuous random variables
- Normal, Uniform, Exponential, etc.
- o Bayes' rule for random variables
- (Derived) Distribution of Y = g(X) or Z = g(X, Y)
- o Quantifying the degree of dependence between two rvs.
- Conditional expectation/variance
- o (Random) Sum of random variables

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Derived Distribution: Y = g(X)







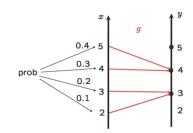
- Given the PDF of X, What is the PDF of Y = g(X)?
- Wait! Didn't we cover this topic? No. We covered just $\mathbb{E}[g(X)]$.
- Examples: Y = X, Y = X + 1, $Y = X^2$, etc.
- What are easy or difficult cases?
- Easy cases
 - Discrete
 - Linear: Y = aX + b

• Take all values of x such that g(x) = y, i.e.,

$$p_Y(y) = \mathbb{P}(g(X) = y)$$
$$= \sum_{x:g(x)=y} p_X(x)$$

$$p_Y(3) = p_X(2) + p_X(3) = 0.1 + 0.2 = 0.3$$

 $p_Y(4) = p_X(4) + p_X(5) = 0.3 + 0.4 = 0.7$



Linear: Y = aX + b, $a \neq 0$



If
$$a > 0$$
, $F_Y(y) = \mathbb{P}(aX + b \le y) = \mathbb{P}(X \le \frac{y - b}{a}) = F_X(\frac{y - b}{a})$
 $\to f_Y(y) = \frac{1}{a} f_X\left(\frac{y - b}{a}\right)$

If
$$a < 0$$
, $F_Y(y) = \mathbb{P}(aX + b \le y) = \mathbb{P}(X > \frac{y - b}{a}) = 1 - F_X(\frac{y - b}{a})$

$$\to f_Y(y) = -\frac{1}{a}f_X\left(\frac{y - b}{a}\right)$$

Therefore, $f_Y(y) = \frac{1}{|a|} f_X(\frac{y-b}{a})$

Special case. X is normal. Then, Y is also normal, i.e., $Y \sim N(a\mu + b, a^2\sigma^2)$

Generally, Y = g(X)

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Step 1. Find the CDF of Y:

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(g(X) \leq y)$$

Step 2. Differentiate: $f_Y(y) = \frac{dF_Y}{dy}(y)$

** When Y = g(X) is monotonic, a general formula can be drawn (see the textbook at pp 207)

Ex1. $X \sim uniform[0,1]$. $Y = \sqrt{X}$.

$$F_Y(y) = \mathbb{P}(\sqrt{X} \le y) = \mathbb{P}(X \le y^2) = y^2$$

$$f_Y(y) = 2y, \quad 0 \le y \le 1$$

Ex2. $X \sim uniform[0, 2]$. $Y = X^3$.

$$F_Y(y) = \mathbb{P}(X^3 \le y) = \mathbb{P}(X \le \sqrt[3]{y}) = \frac{1}{2}y^{1/3}$$

$$f_Y(y) = \frac{1}{6}y^{-2/3}, \quad 0 \le y \le 8$$

Ex3. X with $f_X(x)$. $Y = X^2$.

$$F_Y(y) = \mathbb{P}(X^2 \le y) = \mathbb{P}(-\sqrt{y} \le X \le \sqrt{y})$$
$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

$$f_Y(y) = \frac{1}{2\sqrt{y}}f_X(\sqrt{y}) +$$

$$\frac{1}{2\sqrt{y}}f_X(-\sqrt{y}), \quad y \ge 0$$

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Functions of multiple rvs: Y = g(X, Y) (1)



Basically, follow two step approach: (i) CDF and (ii) differentiate.

Ex1. $X, Y \sim uniform[0, 1]$, and $X \perp \!\!\!\perp Y$. Z = max(X, Y).

*
$$\mathbb{P}(X \leq z) = \mathbb{P}(Y \leq z) = z, \ z \in [0,1].$$

$$\begin{aligned} F_Z(z) &= \mathbb{P}(\max(X,Y) \leq z) = \mathbb{P}(X \leq z, Y \leq z) \\ &= \mathbb{P}(X \leq z) \mathbb{P}(Y \leq z) = z^2 \\ f_Z(z) &= \begin{cases} 2z, & \text{if } 0 \leq z \leq 1 \\ 0, & \text{otherwise} \end{cases} \end{aligned}$$

Functions of multiple rvs: Y = g(X, Y) (2)



Basically, follows two step approach: (i) CDF and (ii) differentiate.

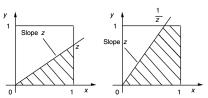
Ex2. $X, Y \sim uniform[0, 1]$, and $X \perp \!\!\! \perp Y$. Z = Y/X.

$$F_Z(z) = \mathbb{P}(Y/X \le z)$$

$$= \begin{cases} z/2, & 0 \le z \le 1 \\ 1 - 1/2z, & z > 1 \\ 0, & \text{otherwise} \end{cases}$$

$$f_Z(z) = \begin{cases} 1/2, & 0 \le z \le \\ 1/(2z^2), & z > 1 \\ 0, & \text{otherwise} \end{cases}$$

- Depending on the value of \boldsymbol{z} , two cases need to be considered separately.



Functions of multiple rvs: Z = X + Y, $X \perp \!\!\! \perp Y$

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- A very basic case with many applications
- Assume that $X, Y \in \mathbb{Z}$

$$p_{Z}(z) = \mathbb{P}(X + Y = z)$$

$$= \sum_{\{(x,y): x+y=z\}} \mathbb{P}(X = x, Y = y)$$

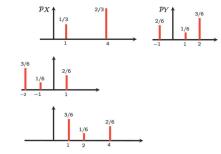
$$= \sum_{x} \mathbb{P}(X = x, Y = z - x)$$

$$= \sum_{x} \mathbb{P}(X = x)\mathbb{P}(Y = z - x)$$

$$= \sum_{x} p_{X}(x)p_{Y}(z - x)$$

- $p_Z(z)$ is called **convolution** of the PMFs of X and Y.

- Interpretation (for a given z)
- (i) Flip (horizontally) $p_Y(y)$ ($p_Y(-x)$)
- (ii) Put it underneath $p_X(x)$ $(p_Y(-x+z))$



Y = X + Y, $X \perp \!\!\!\perp Y$: Continuous

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• Same logic as the discrete case

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx$$

- Very special, but useful case
- X and Y are normal.

Sum of two independent normal rvs

$$X \sim N(\mu_x, \sigma_x^2)$$
 and $Y \sim N(\mu_x, \sigma_x^2)$
Then, $X + Y \sim N(\mu_x + \mu_y, \sigma_x^2 + \sigma_y^2)$

- Why normal rvs are used to model the sum of random noises.
- (Extension) The sum of finitely many independent normals is also normal.

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Making a Metric of Dependence Degree



- Goal: Given two rvs X and Y, assign some number that quantifies the degree of their dependence
- Regs.
 - a) Increases (resp. decreases) as they become more (resp. less) dependent.
 - b) 0 when they are independent.
- c) Shows the direction of dependence by + and -
- d) Always bounded by some numbers, e.g., [-1,1]
- Good engineers: Good at making good metrics
 - Metric of how our society is economically polarized
 - A lot of metrics in our professional sports leagues (baseball, basketball, etc)
 - Cybermetrics in MLB (Major League Baseball): http://m.mlb.com/glossary/advanced-stats

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OK. Let's Design!

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What If $\mu_X \neq 0, \mu_Y \neq 0$?

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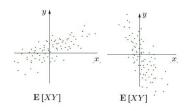
• Simple case: $\mathbb{E}[X] = \mu_X = 0$ and $\mathbb{E}[Y] = \mu_Y = 0$

• Dependent: Positive (If $X \uparrow$, $Y \uparrow$) or Negative (If $X \uparrow$, $Y \downarrow$)

• What about $\mathbb{E}[XY]$? Seems good.

•
$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y] = 0$$
 when $X \perp \!\!\! \perp Y$

• More data points (thus increases) when xy > 0 (both positive or negative)



(Q) What about $\mathbb{E}[X + Y]$?

Covariance

$$cov(X, Y) = \mathbb{E} [(X - \mathbb{E}[X]) \cdot (Y - \mathbb{E}[Y])]$$

• After some algebra, $cov(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$

• Solution: Centering. $X \to X - \mu_X$ and $Y \to Y - \mu_Y$

•
$$X \perp \!\!\!\perp Y \Longrightarrow cov(X,Y) = 0$$

•
$$cov(X, Y) = 0 \Longrightarrow X \perp\!\!\!\perp Y$$
? NO.

• When cov(X, Y) = 0, we say that X and Y are uncorrelated.

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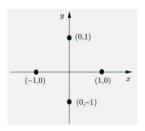
Example: cov(X, Y) = 0, but not independent

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• $p_{X,Y}(1,0) = p_{X,Y}(0,1) = p_{X,Y}(-1,0) = p_{X,Y}(0,-1) = 1/4.$

•
$$\mathbb{E}[X] = \mathbb{E}[Y] = 0$$
, and $\mathbb{E}[XY] = 0$. So, $cov(X, Y) = 0$

• Are they independent? No, because if X = 1, then we should have Y = 0.



Some Properties



$$cov(X,X)=0$$

$$cov(aX + b, Y) = \mathbb{E}[(aX + b)Y] - \mathbb{E}[aX + b]\mathbb{E}[Y] = a \cdot cov(X, Y)$$

$$cov(X,Y+Z) = \mathbb{E}[X(Y+Z)] - \mathbb{E}[X]\mathbb{E}[Y+Z] = cov(X,Y) + cov(X,Z)$$

$$var[X + Y] = \mathbb{E}[(X + Y)^2] - (\mathbb{E}[X + Y])^2 = var[X] + var[Y] - 2cov(X, Y)$$

Example: The hat problem in Lecture 3. Remember?



- n people throw their hats in a box and then pick one at random
- X: number of people with their own hat
- (Q) var[X]
- Key step 1. Define a rv X_i = 1 if i selects own hat and 0 otherwise. Then,
 X = \sum_{i=1}^n X_i.
- Key step 2. Are X_i s are independent?
- $X_i \sim Bernoulli(1/n)$. Thus, $\mathbb{E}[X_i] = 1/n$ and $var[X_i] = \frac{1}{n}(1 \frac{1}{n})$

• For $i \neq j$,

$$\begin{aligned} \text{cov}(X_i, X_j) &= \mathbb{E}[X_i X_j] - \mathbb{E}[X_i] \mathbb{E}[X_j] \\ &= \mathbb{P}(X_i = 1 \text{ and } X_j = 1) - \frac{1}{n^2} \\ &= \mathbb{P}(X_i = 1) \mathbb{P}(X_j = 1 | X_i = 1) - \frac{1}{n^2} \\ &= \frac{1}{n} \frac{1}{n-1} - \frac{1}{n^2} = \frac{1}{n^2(n-1)} \end{aligned}$$

$$var[X] = var\left[\sum X_i\right]$$

$$= \sum var[X_i] + \sum_{i \neq j} cov(X_i, X_j)$$

$$= n\frac{1}{n}(1 - \frac{1}{n}) + n(n - 1)\frac{1}{n^2(n - 1)} = 1$$

Bounding the metric: Correlation Coefficient



- Reqs. a), b), and c) satisfied.
- d) Always bounded by some numbers, e.g., [-1,1]
- Dimensionless metric. How? Normalization, but by what?

Correlation Coefficient

$$\rho(X,Y) = \mathbb{E}\left[\frac{(X - \mu_X)}{\sigma_X} \cdot \frac{Y - \mu_Y}{\sigma_Y}\right] = \frac{\mathsf{cov}(X,Y)}{\sqrt{\mathsf{var}[X]\mathsf{var}[Y]}}$$

- $-1 < \rho < 1$
- $|\rho| = 1 \Longrightarrow X \mu_X = c(Y \mu_Y)$ (linear relation, VERY related)

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A Special Random Variable



• Consider a rv Y, such that

$$Y = \begin{cases} 0, & \text{w.p. } 1/4 \\ 1, & \text{w.p. } 1/4 \\ 2, & \text{w.p. } 1/2 \end{cases}$$

• If $h(y) = y^2$, then a new rv h(Y) is:

$$h(Y) = \begin{cases} 0, & \text{w.p. } 1/4\\ 1, & \text{w.p. } 1/4\\ 4, & \text{w.p. } 1/2 \end{cases}$$

Consider other rv X, such that

$$g(y) = \mathbb{E}[X|Y = y] = \begin{cases} 3, & \text{if } y = 0 \\ 8, & \text{if } y = 1 \\ 9, & \text{if } y = 2 \end{cases}$$

• Then, a rv g(Y) is:

$$g(Y) = \begin{cases} 3, & \text{w.p. } 1/4 \\ 8, & \text{w.p. } 1/4 \\ 9, & \text{w.p. } 1/2 \end{cases}$$

- The rv g(Y) looks special, so let's notate it with some fancy one.
- What about? $X_{exp}(Y)$, $\mathbb{E}[X_Y]$, $\mathbb{E}_X[Y]$?

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Conditional Expectation $\mathbb{E}[X|Y]$



Expectation of $\mathbb{E}[X|Y]$



Conditional Expectation

A random variable $g(Y) = \mathbb{E}[X|Y]$, called conditional expectation of X given Y, takes the value $g(y) = \mathbb{E}[X|Y=y]$, if Y happens to take the value y.

- A function of Y
- A random variable
- Thus, having a distribution, expectation, variance, all the things that a random variable has
- Often confusing because of the notation

Expectation of Conditional Expectation

$$\mathbb{E}\big[\mathbb{E}[X|Y]\big] = \mathbb{E}[X],$$
 Law of iterated expectations

Proof.

$$\mathbb{E}\Big[\mathbb{E}[X|Y]\Big] = \sum_{y} \mathbb{E}[X|Y = y]\rho_{Y}(y)$$
$$= \mathbb{E}[X]$$

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Examples and Meaning



- Stick of length /
- Uniformly break at point Y, and break what is left uniformly at point X.
- $\mathbb{E}[X|Y = y] = y/2$
- $\mathbb{E}[X|Y] = Y/2$
- $\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}[Y/2] = \frac{1}{2}\frac{I}{2} = I/4$

- Forecasts on sales: calculating expected value, given any available information
- X : February sales
- Forecast in the beg. of the year: $\mathbb{E}[X]$
- End of Jan. new information Y = y
 (Jan. sales)
 Revised forecast: E[X|Y = y]
 Revised forecast ≠ E[X]
- Law of iterated expectations $\mathbb{E}[\text{revised forecast}] = \text{original one}$

Conditional Variance var[X|Y]



$$\operatorname{\mathsf{var}}[X] = \mathbb{E}[(X - \mathbb{E}[X])^2]$$

$$\operatorname{var}[X|Y = y] = \mathbb{E}[(X - \mathbb{E}[X|Y = y])^{2}|Y = y]$$

Conditional Variance

A random variable g(Y) = var[X|Y] and called conditional variance of X given Y takes the value g(y) = var[X|Y = y], if Y happens to take the value y.

- A function of Y
- A random variable
- Thus, having a distribution, expectation, variance, all the things that a random variable has

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Expectation and Variance of $\mathbb{E}[X|Y]$ and var[X|Y]



	$\mathbb{E}[X Y]$	var[X Y]
Expectation	$\mathbb{E}\Big[\mathbb{E}(X Y)\Big]$	$\mathbb{E}\Big[var(X Y)\Big]$
Variance	$var \Big[\mathbb{E}(X Y) \Big]$	var[var(X Y)]

Law of Total Variance

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Law of total variance

$$\mathsf{var}[X] = \mathbb{E}\Big[\mathsf{var}(X|Y)\Big] + \mathsf{var}[\mathbb{E}(X|Y)]$$

Proof.

$$\mathsf{var}(X|Y) = \mathbb{E}[X^2|Y] - (\mathbb{E}[X|Y])^2$$

$$\mathbb{E}\left[\operatorname{var}(X|Y)\right] = \mathbb{E}[X^2] - \mathbb{E}\left[\left(\mathbb{E}[X|Y]\right)^2\right] \tag{1}$$

$$\operatorname{var}\left[\mathbb{E}(X|Y)\right] = \mathbb{E}\left[\left(\mathbb{E}[X|Y]\right)^{2}\right] - \left(\mathbb{E}\left[\mathbb{E}(X|Y)\right]\right)^{2} = \mathbb{E}\left[\left(\mathbb{E}[X|Y]\right)^{2}\right] - \left(\mathbb{E}[X]\right)^{2}$$
(2)

$$(1) + (2) = \mathbb{E}[X^2] + (\mathbb{E}[X])^2 = \text{var}[X]$$

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Sum of a random number of rvs





- X_i : money spent in store i, independent of other X_i and N, X_i s are identically distributed with $\mathbb{E}[X_i] = \mu$
- $Y = X_1 + X_2 + \dots X_N$. What are $\mathbb{E}[Y]$ and var[Y]?

 $\mathbb{E}[\mathsf{var}(Y|N)] = \mathbb{E}[N\mathsf{var}[X_i]] = \mathbb{E}[N]\mathsf{var}[X_i]$

• $\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|N]] = \mathbb{E}[N\mathbb{E}[X_i]] = \mathbb{E}[N]\mathbb{E}[X_i] = \mu\mathbb{E}[N]$

•
$$\operatorname{var}[Y] = \mathbb{E}\left[\operatorname{var}(Y|N)\right] + \operatorname{var}[\mathbb{E}(Y|N)] = \mathbb{E}[N]\operatorname{var}[X_i] - \mu^2\operatorname{var}[N]$$

$$\operatorname{var}(\mathbb{E}[Y|N]) = \operatorname{var}(N\mu) = \mu^2\operatorname{var}[N]$$

$$\operatorname{var}[Y|N] = N\operatorname{var}[X_i]$$

Questions?

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Review Questions



- 1) What are the key steps to get the derived distributions of Y = g(X) or Z = g(X, Y)?
- 2) How can we compute the distribution of Z + X + Y when X and Y are independent?
- 3) What are covariance and correlation coefficient? Why do we need them?
- 4) Please explain the concepts of conditional expectation and conditional variance.