Lecture 3: Random Variable, Part I

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EE210: Probability and Introductory Random Processes KAIST EE

MONTH DAY, 2021

• Random Variable: Discrete

- PMF (Probability Mass Function)
- Representative Discrete Random Variables
- Expectation and Variance
- Functions of Random Variables
- Conditioning and Independence for Random Variables

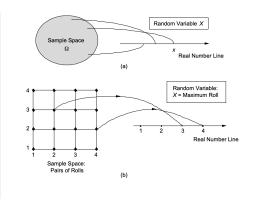
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Random Variable: Idea

- In reality, many outcomes are numerical, e.g., stock price.
- Even if not, very convenient if we map numerical values to random outcomes, e.g., '0' for male and '1' for female.





Random Variable: More Formally



- \circ Mathematically, a random variable X is a function which maps from Ω to \mathbb{R} .
- Notation. Random variable X, numerical value x.
- \circ Different random variables X, Y_{\circ} , etc can be defined on the same sample space.
- For a fixed value x, we can associate an event that a random variable X has the value x, i.e., $\{\omega \in \Omega \mid X(w) = x\}$
- Assume that values x are discrete¹ such as $1, 2, 3, \ldots$. For notational convenience,

$$p_X(x) \triangleq \mathbb{P}(X=x) \triangleq \mathbb{P}(\{\omega \in \Omega \mid X(w)=x\})$$

• For a discrete random variable X, we call $p_X(x)$ probability mass function (PMF).

¹Finite or countably infinite.

Roadmap



Roadmap



- Famous discrete random variables used in the community
- Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- Functions of a single random variable, Functions of multiple random variables
- Conditioning for random variables, Independence for random variables
- · Continuous random variables
- Normal, Uniform, Exponential, etc.
- Bayes' rule for random variables

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Bernoulli X with parameter $p \in [0, 1]$

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Only binary values

$$X = \begin{cases} 0, & \text{w.p.}^2 & 1 - p, \\ 1, & \text{w.p.} & p \end{cases}$$

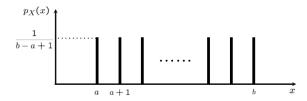
In other words, $p_X(0) = 1 - p$ and $p_X(1) = p$ from our PMF notation.

- Models a trial that results in binary results, e.g., success/failure, head/tail
- Very useful for an indicator rv of an event A. Define a rv 1_A as:

$$1_{\mathcal{A}} = egin{cases} 1, & ext{if A occurs,} \ 0, & ext{otherwise} \end{cases}$$



- integers a, b, where a < b
- Choose a number of $\Omega = \{a, a+1, \ldots, b\}$ uniformly at random.
- $p_X(i) = \frac{1}{b-a+1}, i \in \Omega.$



• Models complete ignorance (I don't know anything about X)

Uniform X with parameter a, b

²with probability

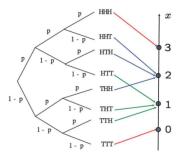
Binomial X with parameter n, p

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- Poisson X with parameter λ



- Models the number of successes in a given number of independent trials
- *n* independent trials, where one trial has the success probability *p*.

$$p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$$



• Binomial(n, p): Models the number of successes in a given number of independent trials with success probability p.

• Very large n and very small p, such that $np = \lambda$

$$p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!}, \quad k = 0, 1, \dots$$

• Is this a legitimate PMF?

$$\sum_{k=0}^{\infty} e^{-\lambda} \frac{\lambda^k}{k!} = e^{-\lambda} \left(1 + \lambda + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} \dots \right) = e^{-\lambda} e^{\lambda} = 1$$

• Prove this:

$$\lim_{n\to\infty} p_X(k) = \binom{n}{k} (1/n)^k (1-1/n)^{n-k} = e^{-\lambda} \frac{\lambda^k}{k!}$$

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Geometric X with parameter p

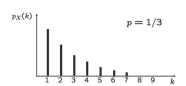
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Roadmap



- Experiment: infinitely many independent Bernoulli trials, where each trial has success probability p
- Random variable: number of trials until the first success.
- Models waiting times until something happens.

$$p_X(k) = (1-p)^{k-1}p$$



Поаціпар

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Expectation/Mean



Properties of Expectation



Average.

Definition

$$\mathbb{E}[X] = \sum_{x} x p_X(x)$$

- $p_X(x)$: relative frequency of value x (trials with x/total trials)
- Example 1: Bernoulli r.v. with p

$$\mathbb{E}[X] = 1 \times p + 0 \times (1 - p) = p_X(1)$$

Not very surprising. Easy to prove using the definition.

- If $X \geq 0$, $\mathbb{E}[X] \geq 0$.
- If $a \le X \le b$, $a \le \mathbb{E}[X] \le b$.
- For a constant c, $\mathbb{E}[c] = c$.

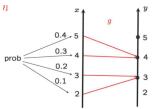
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Expectation of a function of a RV

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- For a rv X, Y = g(X) is also a r.v.
- $\mathbb{E}[Y] = \mathbb{E}[g(X)] = \sum_{x} g(x) p_X(x)$
- Compute $\mathbb{E}[Y]$ for the following:



$$4 \times (0.4 + 0.3) + 3 \times (0.1 + 0.2)$$

= 2.8 + 0.9 = 3.7

Linearity of Expectation

$$\mathbb{E}[aX+b]=a\mathbb{E}[X]+b$$

<u>Varia</u>nce



- Measures how much the spread of a PMF is.
- What about $\mathbb{E}[X \mu]$, where $\mu = \mathbb{E}[X]$? Then, what about $\mathbb{E}[(X \mu)^2]$?

Variance, Standard Deviation

$$\mathsf{var}[X] = \mathbb{E}[(X - \mu)^2]$$

$$\sigma_X = \sqrt{\operatorname{var}[X]}$$

Variance: Useful Property

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Roadmap



- $\operatorname{var}[X] = \mathbb{E}[X^2] (\mathbb{E}[X])^2$ $\operatorname{var}[X] = \mathbb{E}[X^2 - 2\mu X + \mu^2]$ $= \mathbb{E}[X^2] - 2\mu \mathbb{E}[X] + \mu^2 = \mathbb{E}[X^2] - \mu^2$
- Y = X + b, var[Y] = var[X] $var[Y] = \mathbb{E}[(X + b)^2] - (\mathbb{E}[X + b])^2$
- Y = aX, $var[Y] = a^2var[X]$ $var[Y] = \mathbb{E}[a^2X^2] - (a\mathbb{E}[X])^2$

Example: Variance of a Bernoulli rv (p)

$$\mathbb{E}[X] = 1 \times p + 0 \times (1 - p) = p$$
 $\mathbb{E}[X^2] = 1 \times p + 0 \times (1 - p) = p$
 $\text{var}[X] = \mathbb{E}[X^2] - \mu^2 = p - p^2$
 $= p(1 - p)$

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Joint PMF

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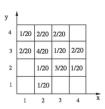
• Joint PMF. For two random variables X, Y, consider two events $\{X = x\}$ and $\{Y = y\}$, and

$$p_{X,Y}(x,y) \triangleq \mathbb{P}(X=x) \cap \{Y=y\}$$

- $\sum_{x} \sum_{y} p_{X,Y}(x,y) = 1$
- Marginal PMF.

$$p_X(x) = \sum_{y} p_{X,Y}(x,y),$$
$$p_Y(y) = \sum_{y} p_{X,Y}(x,y)$$

Example.



$$p_{X,Y}(1,3)=2/20$$

$$p_X(4) = 2/20 + 1/20 = 3/20$$

$$\mathbb{P}(X = Y) = 1/20 + 4/20 + 3/20 = 8/20$$

Functions of Multiple R.V.s



• Consider a rv Z = g(X, Y). (Ex) X + Y, $X^2 + Y^2$. Then, PMF of Z is:

$$p_Z(z) = \mathbb{P}(g(X, Y) = z) = \sum_{(x,y):g(x,y)=z} p_{X,Y}(x,y)$$

Similarly.

$$\mathbb{E}[Z] = \mathbb{E}[g(X,Y)] = \sum_{x} \sum_{y} g(x,y) p_{X,Y}(x,y)$$

Linearity of Expectation for multiple rvs

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- Remember: $\mathbb{E}[aX + b] = a\mathbb{E}[X] + b$
- Similarly,

$$\mathbb{E}[X+Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

(easy to prove, using the definition.)

- $\mathbb{E}[X_1 \ldots + X_n] = \mathbb{E}[X_1] + \ldots + \mathbb{E}[X_n]$
- $\mathbb{E}[2X+3Y-Z] = 2\mathbb{E}[X]+3\mathbb{E}[Y]-\mathbb{E}[Z]$

- Example. Mean of a binomial rv Y with (n, p)
- Y: number of successes in n Bernoulli trials with p
- $Y = X_1 + ... X_n$, where X_i is a Bernoulli rv.
- $\mathbb{E}[Y] = n\mathbb{E}[X_i] = n\mathbb{P}(X_i = 1) = np$

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Roadmap

- Normal, Uniform, Exponential, etc.
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Conditional PMF: Conditioning on an event



Remember two probability laws: $\mathbb{P}(\cdot)$ and $\mathbb{P}(\cdot|A)$, for an event A.

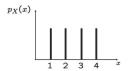
- $p_X(x) = \mathbb{P}(X = x)$
- $\mathbb{E}[X] = \sum_{x} x p_X(x)$
- $\mathbb{E}[g(X)] = \sum_{x} g(x) p_X(x)$
- $\operatorname{var}[X] = \mathbb{E}[X^2] (\mathbb{E}[X])^2$

- $p_{X|A}(x) \triangleq \mathbb{P}(X = x|A)$
- $\mathbb{E}[X|A] \triangleq \sum_{x} x p_{X|A}(x)$
- $\mathbb{E}[g(X)|A] \triangleq \sum_{x} g(x) p_{X|A}(x)$
- $\operatorname{var}[X|A] \triangleq \mathbb{E}[X^2|A] (\mathbb{E}[X|A])^2$
- Note. $p_{X|A}(x)$, $\mathbb{E}[X|A]$, $\mathbb{E}[g(X)|A]$, and var[X|A] are all just notations!

Example: Conditional PMF



$$A = \{X \ge 2\}$$



$$\mathbb{E}[X] = \frac{1}{4} (1 + 2 + 3 + 4) = 2.5$$

$$var[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$
$$= \frac{1}{4}(1 + 2^2 + 3^2 + 4^2) - 2.5^2$$



$$\mathbb{E}[X|A] = \frac{1}{3}(2+3+4) = 3$$

$$var[X|A] = \mathbb{E}[X^2|A] - (\mathbb{E}[X|A])^2$$
$$= \frac{1}{3}(2^2 + 3^2 + 4^2) - 3^2 = 2/3$$

Conditional PMF: Conditioning on a RV

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Conditional PMF

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What do we mean by "conditioning on a rv"? Consider $A = \{Y = y\}$ for a rv Y.

- $p_{X|A}(x) \triangleq \mathbb{P}(X = x|A)$
- $\mathbb{E}[X|A] \triangleq \sum_{x} x p_{X|A}(x)$
- $\mathbb{E}[g(X)|A] \triangleq \sum_{x} g(x) p_{X|A}(x)$
- $\operatorname{var}[X|A] \triangleq \mathbb{E}[X^2|A] (\mathbb{E}[X|A])^2$

- $p_{X|Y}(x|y) \triangleq \mathbb{P}(X=x|Y=y)$
- $\mathbb{E}[X|Y=y] \triangleq \sum_{x} x p_{X|Y}(x|y)$
- $\mathbb{E}[g(X)|Y=y] \triangleq \sum_{x} g(x)p_{X|Y}(x|y)$
- $\operatorname{var}[X|Y = y] \triangleq \mathbb{E}[X^2|Y = y] (\mathbb{E}[X|Y = y])^2$

Conditional PMF

$$p_{X|Y}(x|y) \triangleq \mathbb{P}(X=x|Y=y) = \frac{p_{X,Y}(x,y)}{p_{Y}(y)}$$

for y such that $p_Y(y) > 0$.

- $\sum_{x} p_{X|Y}(x|y) = 1$
- Multiplication rule.

$$p_{X,Y}(x,y) = p_Y(y)p_{X|Y}(x|y)$$
$$= p_X(x)p_{Y|X}(y|x)$$

• $p_{X,Y,Z}(x,y,z) = p_X(x)p_{Y|X}(y|x)p_{Z|X,Y}(z|x,y)$

$$p_{X|Y}(2|2) = \frac{1}{1+3+1}$$

$$p_{X|Y}(3|2) = \frac{3}{1+3+1}$$

$$\mathbb{E}[X|Y=3] = 1(2/9) + 2(4/9) + 3(1/9) + 4(2/9)$$

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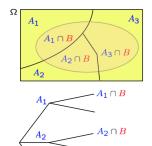
Remind: Total Probability Theorem (from Lecture 2)



- Partition of Ω into A_1, A_2, A_3
- o Known: $\mathbb{P}(A_i)$ and $\mathbb{P}(B|A_i)$
- What is $\mathbb{P}(B)$? (probability of result)

Total Probability Theorem

$$\mathbb{P}(B) = \sum_{i} \mathbb{P}(A_i) \mathbb{P}(B|A_i)$$



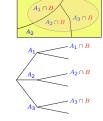
Total Probability Theorem: $B = \{X = x\}$

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• Partition of Ω into A_1, A_2, A_3

Total Probability Theorem

$$p_X(x) = \sum_i \mathbb{P}(A_i)\mathbb{P}(X = x|A_i) = \sum_i \mathbb{P}(A_i)p_{X|A_i}(x)$$



Total Expectation Theorem for $\{A_i\}$

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Total Expectation Theorem for $\{Y = y\}$

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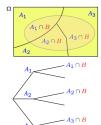
• Partition of Ω into A_1, A_2, A_3

Total Probability Theorem

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Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{i} \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$



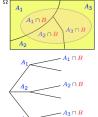
 \circ Partition of Ω into A_1,A_2,A_3

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{i} \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{y} \mathbb{P}(Y = y) \mathbb{E}[X|Y = y] = \sum_{y} p_{Y}(y) \mathbb{E}[X|Y = y]$$



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Example 1: Total Expectation Theorem



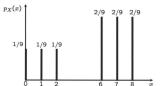


- $A_1 = \{X \in \{0, 1, 2\}\}, A_2 = \{X \in \{6, 7, 8\}\}$
- Using TET,

$$\mathbb{E}[X] = \sum_{i=1,2} \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$
$$= 1/3 \cdot 1 + 2/3 \cdot 7 = 7$$

• Without using TET,

$$\mathbb{E}[X] = \frac{1}{9}(0+1+2) + \frac{2}{9}(6+7+8)$$



Background: Memoryless Property of Geometric rv (1)



- · Some random variable often does not have memory.
- Definition. A random variable X is called memoryless if, for any $n, m \ge 0$,

$$\mathbb{P}(X > n + m | X > m) = \mathbb{P}(X > n)$$

- Meaning. Conditioned on X > m, X m's distribution is the same as the original X.
- Remind. Geometric rv X with parameter p

$$\mathbb{P}(X=k)=(1-p)^{k-1}p$$

$$\mathbb{P}(X > k) = 1 - \sum_{k'=1}^{k} (1 - p)^{k'-1} p = (1 - p)^{k}$$

Background: Memoryless Property of Geometric rv (2)



Theorem. Any geometric random variable is memoryless.

$$\mathbb{P}(X > n + m | X > m) = \frac{\mathbb{P}(X > n + m \text{ and } X > m)}{\mathbb{P}(X > m)}$$

$$= \frac{\mathbb{P}(X > n + m)}{\mathbb{P}(X > m)}$$

$$= \frac{(1 - p)^{n+m}}{(1 - p)^m} = (1 - p)^n = \mathbb{P}(X > n)$$

• Meaning. Conditioned on X > m, X - m is geometric with the same parameter.

Example 2: Mean and Variance of Geometric rv



- Write softwares over and over, and each time w.p. p of working correctly (independent from prev. programs)
- X: number of tries until the program works correctly.
- Q) mean and variance of X
- *X* is geometric
- Direct computation is boring.

$$\mathbb{E}[X] = \sum_{k=1}^{\infty} k(1-p)^{k-1}p$$

• Total expectation theorem and memorylessness nhelps a lot.

• $A_1 = \{X = 1\}$ (first try is success) $A_2 = \{X > 1\}$ (first try is failure).

$$\begin{split} \mathbb{E}[X] &= 1 + \mathbb{E}[X - 1] \\ &= 1 + \mathbb{P}(A_1)\mathbb{E}[X - 1|X = 1] \\ &+ \mathbb{P}(A_2)\mathbb{E}[X - 1|X > 1] \\ &= 1 + (1 - \rho)\mathbb{E}[X] \end{split}$$

$$\mathbb{E}[X] = 1 + (1 - p)\frac{1}{p} = 1/p.$$

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Independence, Conditional Independence



Two events

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$
$$\mathbb{P}(A \cap B | C) = \mathbb{P}(A | C) \cdot \mathbb{P}(B | C)$$

• A rv and an event

$$\mathbb{P}(\{X = x\} \cap B) = \mathbb{P}(X = x) \cdot \mathbb{P}(B), \text{ for all } x$$

$$\mathbb{P}(\{X = x\} \cap B | C) = \mathbb{P}(X = x | C) \cdot \mathbb{P}(B | C), \text{ for all } x$$

Two rvs

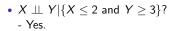
$$\mathbb{P}(\{X = x\} \cap \{Y = y\}) = \mathbb{P}(X = x) \cdot \mathbb{P}(Y = y), \text{ for all } x, y$$
$$p_{X,Y}(x,y) = p_X(x) \cdot p_Y(y)$$

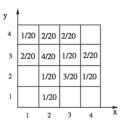
$$\mathbb{P}(\{X = x\} \cap \{Y = y\} | C) = \mathbb{P}(X = x | C) \cdot \mathbb{P}(Y = y | C), \text{ for all } x, y$$
$$p_{X,Y|C}(x, y) = p_{X|C}(x) \cdot p_{Y|C}(y)$$

Example



• *X* ⊥⊥ *Y*? $p_{X,Y}(1,1) = 0, \quad p_X(1) = 3/20$ $p_Y 1 = 1/20.$





Y =	= 4 (1/3)	1/9	2/9
Y =	= 3 (2/3)	2/9	4/9
		X = 1 (1/3)	X = 2(2/3)

Expectation and Variance

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- Always true. $\mathbb{E}[aX + b], \ \mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$
- Generally, $\mathbb{E}[g(X,Y)] \neq g(\mathbb{E}[X],\mathbb{E}[Y])$
- However, if $X \perp \!\!\! \perp Y$, $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$ $\mathbb{E}[g(X)h(Y)] = \mathbb{E}[g(X)]\mathbb{E}[g(Y)]$
- Proof.

$$\mathbb{E}[g(X)h(Y)] = \sum_{x} \sum_{y} g(x)h(y)p_{X,Y}(x,y)$$
$$= \sum_{x} xp_{X}(x) \sum_{y} yp_{Y}(y)$$

- Always true. $var[aX] = a^2 var[X], var[X + a] = var[X]$
- Generally, $var[X + Y] \neq var[X] + var[Y]$
- However, if $X \perp \!\!\! \perp Y$. var[X + Y] = var[X] + var[Y]
- Practice.
 - $\circ X = Y \Longrightarrow var[X + Y] = 4var[X]$
- $X = -Y \Longrightarrow var[X + Y] = 0$ $X \perp Y \Longrightarrow var[X 3Y] = var[X] + 9var[Y]$

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$\mathsf{var}[X+Y] eq \mathsf{var}[X] + \mathsf{var}[Y]$



• Why not generally true?

$$\begin{aligned} \operatorname{var}[X+Y] &= \mathbb{E}[(X+Y)^2] - (\mathbb{E}[X+Y])^2 \\ &= \mathbb{E}[X^2+Y^2+2XY] - \left((\mathbb{E}[X])^2 + (\mathbb{E}[Y])^2 + 2\mathbb{E}[X]\mathbb{E}[Y]\right) \\ &= \operatorname{var}[X] + \operatorname{var}[Y] + 2\left(\mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]\right) \end{aligned}$$

- $\circ \mid X \perp \!\!\!\perp Y \mid$ is a sufficient condition for $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$
- Also, a necessary condition? we will see later, when we study covariance.

Example: The hat problem (1)



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- n people throw their hats in a box and then pick one at random
- X: number of people with their own hat
- $\mathbb{E}[X]$? var[X]?
- All permutations are equally likely as 1/n!. Thus, this equals to picking one hat at a
- Key step 1. Define a rv $X_i = 1$ if i selects own hat and 0 otherwise.

$$X = \sum_{i=1}^{n} X_i.$$

• $\{X_i\}, i = 1, 2, \dots, n$: identically distributed (symmetry)

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Example: The hat problem (2)



- $\mathbb{E}[X] = n\mathbb{E}[X_1] = n\mathbb{P}(X_1 = 1) = n \times \frac{1}{n} = 1.$
- Key step 2. Are X_i s are independent? If yes, easy to get var(X).
- Assume n=2. Then, $X_1=1 \rightarrow X_2=1$, and $X_1=0 \rightarrow X_2=0$. Thus, dependent.

$$\begin{aligned} \text{var}(X) &= \mathbb{E}[X^2] - (\mathbb{E}[X])^2 \\ &= \mathbb{E}\Big[\sum_i X_i^2 + \sum_{i,j:i \neq j} X_i X_j\Big] - (\mathbb{E}[X])^2 \\ \mathbb{E}[X_i^2] &= 1 \times \frac{1}{n} + 0 \times \frac{n-1}{n} = \frac{1}{n} \\ \mathbb{E}[X_i X_j] &= \mathbb{E}[X_1 X_2] = 1 \times \mathbb{P}(X_1 X_2 = 1) = \mathbb{P}(X_1 = 1)\mathbb{P}(X_2 = 1 | X_1 = 1), \quad (i \neq j) \end{aligned}$$

- $\mathbb{E}[X^2] = n\mathbb{E}[X_1^2] + n(n-1)\mathbb{E}[X_1X_2] = n\frac{1}{n} + n(n-1)\frac{1}{n(n-1)} = 2$
- var(X) = 2 1 = 1

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Review Questions



- 1) What is Random Variable? Why is it useful?
- 2) What is PMF (Probability Mass Function)?
- 3) Explain Bernoulli, Binomial, Poisson, Geometric rvs, when they are used and what their PMFs are.
- 4) What are joint and marginal PMFS?
- 5) Describe and explain the total probability/expectation theorem for random variables?
- 6) When is it useful to use total probability/expectation theorem?
- 7) What is conditional independence?

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Questions?