

Exercise 1.

$$1) \pi(\tau^2 | y^{(T)}, \sigma^{(T)}, \sigma_0, \beta_0, \beta_1) \propto f(y^{(T)} | \sigma^{(T)}, \sigma_0, \beta_0, \beta_1, \tau^2) \cdot \pi(\tau^2)$$

The likelihood: $f(y^{(T)} | \sigma^{(T)}, \sigma_0, \beta_0, \beta_1, \tau^2)$

$$= \prod_{t=1}^T f(y_t | \sigma^{(T)}, \sigma_0, \beta_0, \beta_1, \tau^2)$$

$$= \frac{1}{(2\pi)^{\frac{T}{2}} \cdot \exp(\frac{1}{2} \cdot \sum_{t=1}^T \sigma_t)} \exp\left(-\sum_{t=1}^T \frac{(y_t)^2}{2 \cdot \exp(\sigma_t)}\right)$$

A posteriori:

$$\pi(\tau^2 | y^{(T)}, \sigma^{(T)}, \sigma_0, \beta_0, \beta_1) \propto \frac{1}{(2\pi)^{\frac{T}{2}} \cdot \exp(\frac{1}{2} \cdot \sum_{t=1}^T \sigma_t)} \exp\left(-\sum_{t=1}^T \frac{(y_t)^2}{2 \cdot \exp(\sigma_t)}\right) \frac{(d_0)^{c_0}}{\Gamma(c_0)} \cdot \left(\frac{1}{\tau^2}\right)^{c_0+1} \cdot \exp\left(-\frac{d_0}{\tau^2}\right)$$

2)

$$\pi(\beta_0 | y^{(T)}, \sigma^{(T)}, \sigma_0, \tau^2, \beta_1) \propto f(y^{(T)} | \sigma^{(T)}, \sigma_0, \beta_0, \beta_1, \tau^2) \cdot \pi(\beta_0)$$

$$\propto \frac{1}{(2\pi)^{\frac{T}{2}} \cdot \exp(\frac{1}{2} \cdot \sum_{t=1}^T \sigma_t)} \exp\left(-\sum_{t=1}^T \frac{(y_t)^2}{2 \cdot \exp(\sigma_t)}\right) \cdot \frac{1}{\sqrt{2\pi} \cdot \gamma_0} \cdot \exp\left(-\frac{(\beta_0 - \alpha_0)^2}{2 \gamma_0^2}\right)$$

3)

$$\pi(\beta_1 | y^{(T)}, \sigma^{(T)}, \sigma_0, \tau^2, \beta_0) \propto f(y^{(T)} | \sigma^{(T)}, \sigma_0, \beta_0, \beta_1, \tau^2) \cdot \pi(\beta_1)$$

$$\propto \frac{1}{(2\pi)^{\frac{T}{2}} \cdot \exp(\frac{1}{2} \cdot \sum_{t=1}^T \sigma_t)} \exp\left(-\sum_{t=1}^T \frac{(y_t)^2}{2 \cdot \exp(\sigma_t)}\right) \cdot \frac{1}{\sqrt{2\pi} \cdot \gamma_1} \cdot \exp\left(-\frac{(\beta_1 - \alpha_1)^2}{2 \gamma_1^2}\right) \cdot \mathbb{1}_{(-1,1)}(\beta_1)$$

Exercise 2

1) The likelihood:

$$\begin{aligned}
 f(y^T | \gamma, \delta, \lambda) &= \prod_{t=1}^{\lambda} f(y_t | \gamma, \delta, \lambda) \cdot \prod_{t=\lambda+1}^T f(y_t | \gamma, \delta, \lambda) \\
 &= \prod_{t=1}^{\lambda} \frac{\gamma^{y_t} \cdot e^{-\gamma}}{(y_t)!} \cdot \prod_{t=\lambda+1}^T \frac{\delta^{y_t} \cdot e^{-\delta}}{(y_t)!} \\
 &= \frac{\gamma^{\sum_{t=1}^{\lambda} y_t} \cdot e^{-\lambda\gamma} \cdot \delta^{\sum_{t=\lambda+1}^T y_t} \cdot e^{-(T-\lambda)\delta}}{\prod_{t=1}^T (y_t)!}
 \end{aligned}$$

$$y_t = 0, 1, 2, \dots \quad \forall t \in [1, T]$$

2) Gibbs:

$$\theta_0 = (\gamma_0, \delta_0, \lambda_0)$$

Sample from:

$$\bullet \gamma_i \sim \pi(\gamma | y^T, \delta_0, \lambda_0) \propto f(y^T | \lambda_0, \delta_0, \gamma) \cdot \pi(\gamma)$$

$$\propto \frac{\gamma^{\sum_{t=1}^{\lambda_0} y_t} \cdot e^{-\lambda_0 \gamma} \cdot \delta_0^{\sum_{t=\lambda_0+1}^T y_t} \cdot e^{-(T-\lambda_0)\delta_0}}{\prod_{t=1}^T (y_t)!} \cdot \gamma^{a_1} \cdot e^{-\frac{\gamma}{a_2}}$$

$$\propto \gamma^{\sum_{t=1}^{\lambda_0} y_t + a_1} \cdot e^{-\lambda_0 \gamma} \cdot e^{-\frac{\gamma}{a_2}}$$

$$= \gamma^{\sum_{t=1}^{\lambda_0} y_t + a_1} \cdot e^{-\gamma(\frac{1}{a_2} + \lambda_0)}$$

$$\Gamma\left(\sum_{t=1}^{\lambda_0} y_t + a_1 + 1, \frac{1}{a_2} + \lambda_0\right)$$

$$\begin{aligned}
\bullet \delta_1 &\sim \pi(\delta | y^T, \gamma, \lambda_0) \propto f(y^T | \gamma, \lambda_0, \delta) \cdot \pi(\delta) \\
&\propto \delta^{\sum_{t=\lambda_0+1}^T y_t} \cdot e^{-(T-\lambda_0) \cdot \delta} \cdot \delta^{d_1} \cdot e^{-\frac{\delta}{d_2}} \\
&= \delta^{(\sum_{t=\lambda_0+1}^T y_t + d_1)} \cdot e^{-\delta \cdot (\frac{1}{d_2} + (T-\lambda_0))} \\
&\quad \Gamma\left(\sum_{t=\lambda_0+1}^T y_t + d_1 + 1, \frac{1}{d_2} + T - \lambda_0\right)
\end{aligned}$$

$$\begin{aligned}
\bullet \lambda_1 &\sim \pi(\lambda | y^T, \gamma, \delta_1) \propto f(y^T | \gamma, \delta_1, \lambda) \cdot \pi(\lambda) \\
&\propto \gamma_1^{\sum_{t=1}^{\lambda} y_t} \cdot \delta_1^{\sum_{t=\lambda+1}^T y_t} \cdot e^{-(T-\lambda)\delta_1} \cdot \frac{1}{T} \\
&\propto \gamma_1^{\sum_{t=1}^{\lambda} y_t} \cdot \delta_1^{\sum_{t=\lambda+1}^T y_t} \cdot e^{-(T-\lambda)\delta_1}
\end{aligned}$$

$$\lambda_1 \in \{1, 2, \dots, T-1\}$$

In this way, we can obtain the first iterate.

We can repeat iterations as we need.

Statistiques Bayésiens DM1

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1 Exercice 2.3

1.1 Code

```
import numpy as np
import matplotlib.pyplot as plt
rng = np.random.default_rng()

with open('acc_usines.txt') as f:
    contents = f.readlines()

y = [] # store the number of accidents
for i in contents:
    y.append(int(i[0]))

T = len(y)
# hyperparameters
a1, a2, d1, d2 = 1, 1, 1, 1
epochs = 100

# initialization
gamma_ = rng.gamma(a1, scale=1/a2)
delta_ = rng.gamma(d1, scale=1/d2)
lambda_ = np.random.choice(range(1, T+1), p=[1/T]*T)

# store the sampling history
l_gamma, l_lambda, l_delta = [gamma_], [lambda_], [delta_]

# Gibbs sampler
for e in range(epochs):
    # define the new parameters for the law of parameter gamma, which follows the law gamma
    shape_gamma_ = sum(y[:lambda_]) + a1 + 1
    rate_gamma_ = 1 / a2 + lambda_
    gamma_ = rng.gamma(shape_gamma_, scale=1/rate_gamma_)
    l_gamma.append(gamma_)

    # define the new parameters for the law of parameter delta, which follows the law gamma
    shape_delta_ = sum(y[lambda_:]) + d1 + 1
    rate_delta_ = 1 / d2 + T - lambda_
    delta_ = rng.gamma(shape_delta_, scale=1/rate_delta_)
    l_delta.append(delta_)

    proba_lambda = [gamma_**(sum(y[:lamb]))*delta_**(sum(y[lamb:]))*np.exp(-(T-lamb)*delta_)
                    for lamb in range(1, T)]
    proba_lambda = proba_lambda / sum(proba_lambda)
    lambda_ = np.random.choice(range(1,T), p=proba_lambda)
    l_lambda.append(lambda_)
```

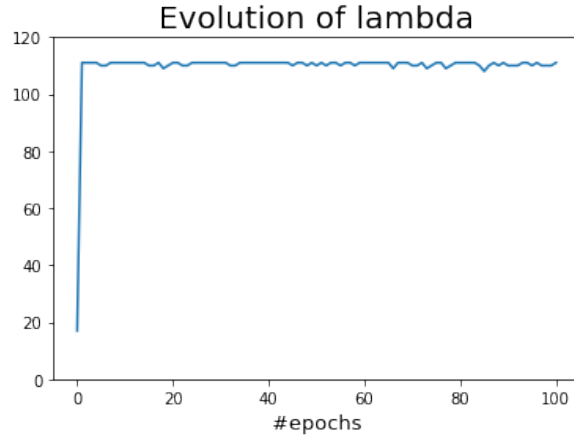


Figure 1: Lambda

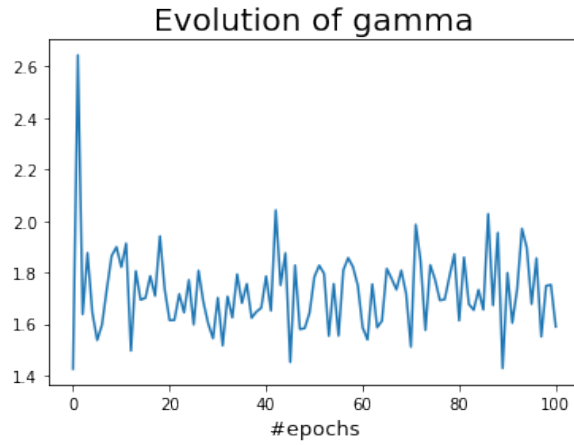


Figure 2: Gamma

```
plt.plot(l_lambda)
plt.xlabel("#epochs", size=13)
plt.ylim((0,120))
plt.title("Evolution of lambda", size=20)

plt.plot(l_gamma)
plt.xlabel("#epochs", size=13)
plt.title("Evolution of gamma", size=20)

plt.plot(l_delta)
plt.xlabel("#epochs", size=13)
plt.title("Evolution of delta", size=20)
```

1.2 Results

There are 100 epochs in Gibbs sampler. Fig. 1, Fig. 2 and Fig. 3 give the evolution of parameters (λ, γ, δ) in the model:

After 20 epochs, λ arrives at a stable state. We can then calculate its mean and variance using command `np.mean(l_lambda[20:])` and `np.var(l_lambda[20:])`. We respectively get **110.6** and **0.39** as mean and variance of λ . Therefore, the change point is **around 110 and 111**.

The number of accidents is given by the parameter in law of Poisson. That is to say, before change point, the anticipated number of accident equals to γ , which is around **1.7**; After change point, the

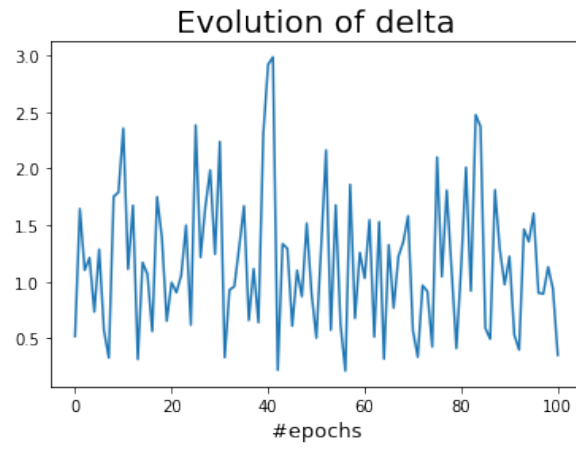


Figure 3: Delta

anticipated number of accident drop to around **1.2** (we should notice that the δ is in fact not as stable as γ and λ).