

Brief Article

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October 8, 2015

1 Problem 1

1.1

For OLS, we can derive the following properties:

$$Cov[X, Y] = \hat{\beta}_1 / \hat{\beta}_1 Cov[X, Y]$$

$$= Cov[\hat{\beta}_1 X, Y] / \hat{\beta}_1$$

$$= Cov[\hat{\beta}_1 X + \hat{\beta}_0, Y] / \hat{\beta}_1$$

$$= Cov[\hat{Y}, Y] / \hat{\beta}_1$$

$$\sigma[X] = \hat{\beta}_1 / \hat{\beta}_1 \sigma[X]$$

$$= \sigma[\hat{\beta}_1 X] / \hat{\beta}_1$$

$$= \sigma[\hat{\beta}_1 X + \hat{\beta}_0] / \hat{\beta}_1$$

$$= \sigma[\hat{Y}] / \hat{\beta}_1$$

$$\Rightarrow \rho_{xy} = Cov[X, Y] / \sigma[X] \sigma[Y] = Cov[\hat{Y}, Y] / \sigma[\hat{Y}] \sigma[Y]$$

$$= \sum (\hat{y}_i - \bar{y})(y_i - \bar{y}) / \sqrt{\sum (\hat{y}_i - \bar{y})^2} \sqrt{\sum (y_i - \bar{y})^2}$$

$$= \sum (\hat{y}_i - \bar{y})(y_i - \hat{y}_i + \hat{y}_i - \bar{y}) / \sqrt{\sum (\hat{y}_i - \bar{y})^2} \sqrt{\sum (y_i - \bar{y})^2}$$

$$= \sum [(\hat{y}_i - \bar{y})(y_i - \hat{y}_i) + (\hat{y}_i - \bar{y})(\hat{y}_i - \bar{y})] / \sqrt{\sum (\hat{y}_i - \bar{y})^2} \sqrt{\sum (y_i - \bar{y})^2}$$

$$= \sum [(\hat{y}_i - \bar{y})\hat{u}_i + (\hat{y}_i - \bar{y})^2] / \sqrt{\sum (\hat{y}_i - \bar{y})^2} \sqrt{\sum (y_i - \bar{y})^2}$$

$$= \sum [(\hat{y}_i - \bar{y})^2] / \sqrt{\sum (\hat{y}_i - \bar{y})^2} \sqrt{\sum (y_i - \bar{y})^2}$$

$$= \sqrt{\sum [(\hat{y}_i - \bar{y})^2]} / \sqrt{\sum (y_i - \bar{y})^2}$$

$$= \sqrt{SSE / SST}$$

$$= \sqrt{R^2}$$

$$\Rightarrow \rho_{xy}^2 = R^2 \text{ for single variate OLS}$$

$$\because \sum (\hat{y}_i - \bar{y})\hat{u}_i = 0$$

1.2

In general, $R^2 \neq \rho_{xy}^2$ in the multivariate case. Observe that in the proof for the single variable case, a transformation is made from X to \hat{Y} . However, in the multivariable case, the covariance terms would not reduce to 0 after making the transformation, and therefore the two values would not converge.

2 Problem 2

2.1

For a general linear model of the form $y_i = \sum_{j=0}^k \beta_j x_{ij} + u_i$

If MLR.1-MLR.6 are satisfied

$$\Rightarrow \hat{\beta}_j \sim \mathcal{N}(\beta_j, \sigma^2 / [SST_j(1 - R_j^2)])$$

For simplicity let $SST'_j = SST_j(1 - R_j^2)$

$$se[\hat{\beta}_j] = \sqrt{s^2 / SST'_j}$$

$$s^2 = \frac{1}{n-k-1} \sum \hat{u}_i^2$$

By MLR.6 $\hat{u} \sim \mathcal{N}(0, \sigma^2) \Rightarrow \hat{u}/\sigma \sim \mathcal{N}(0, 1) \Rightarrow \sum \hat{u}_i^2 \sim \sigma^2 \chi^2$

$$\Rightarrow s^2(n-k-1)/\sigma^2 \sim \chi^2_{n-k-1}$$

$$\Rightarrow (\hat{\beta}_j - \beta_j) / se(\hat{\beta}_j) = (\hat{\beta}_j - \beta_j) / \sqrt{s^2 / SST'_j}$$

$$= (\hat{\beta}_j - \beta_j) / \sqrt{\sigma^2 / SST'_j} / \sqrt{s^2 / \sigma^2}$$

\therefore multiplying by 1

$$= (\hat{\beta}_j - \beta_j) / \sqrt{\sigma^2 / SST'_j} / \sqrt{s^2(n-k-1) / [\sigma^2(n-k-1)]}$$

\therefore multiplying by 1

$$\sim \mathcal{N}(0, 1) / \sqrt{\chi^2_{n-k-1} / (n-k-1)}$$

$$\sim t_{n-k-1}$$

2.2

Based on results from previous problem, $s^2(df)/\sigma^2 \sim \chi^2_{df}$

Also in general, $SSR = (df)s^2 \sim \sigma^2 \chi^2_{df}$

Since unrestricted model has $n-k-1$ degrees of freedom \Rightarrow restricted model has $n-k-1-q$ degrees of freedom

This implied: $(SSR_{ur} - SSR_r)/q \sim \sigma^2(\chi^2_{n-k-1} - \chi^2_{n-k-1-q})/q$

$$\sim \sigma^2(\mathcal{X}_q^2/q)$$

And

$$SSR_{ur}/(n-k-1) \sim \sigma^2 \mathcal{X}_{n-k-q}^2$$

The σ^2 in top and bottom will cancel

$$\Rightarrow [(SSR_{ur} - SSR_r)/q]/[SSR_{ur}/(n-k-1)] \sim (\mathcal{X}_q^2/q)/(\mathcal{X}_{n-k-1}^2/(n-k-1))$$

$$\sim F_{q,n-k-1}$$

3 Problem 3

4 Problem 4

$$\log wage = \beta_0 + \beta_1 educ + \beta_2 exper + \beta_3 tenure + u$$

4.1

- i. Another year of general work experience has as much effect on $\log wage$ as another year of tenure with current employer

$$H_0 : \hat{\beta}_2 = \hat{\beta}_3$$

However, this can also be written as:

$$H_0 : \theta = \hat{\beta}_2 - \hat{\beta}_3 = 0$$

And the model equation will be transformed to:

$$\log(wage) = \beta_0 + \beta_1 educ + \theta exper + \beta_2(exper + tenure)$$

- ii. Calculating the t-statistic of the regression:

5 Problem 5

$$\log(psoda) = \beta_0 + \beta_1 prpbck + \beta_2 \log(income) + \beta_3 prppov + u$$

6 Problem 6

The second model is preferable since it has a higher \overline{R}^2 value than that of the other two models. Without